# Interest Rate Model - Vasicek and Hull White

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# Ordinary differential equation

ODE type		solution	application
1st order linear <i>ODE</i>	f'(x) = p(x) + q(x)f(x)	integration factor	Hull White bond price
1st order linear ODE no const	f'(x) = q(x)f(x)	separation of variables	Hull White bond price
1st order quadratic ODE	$f'(x) = p(x) + q(x)f(x) + r(x)f^{2}(x)$	Riccati equation	Heston

Please refer to  ${\it Heston.doc}$  for  ${\it Riccati}$  equation.

#### 1. Short rate models

Instantaneous rate (or short rate) is defined as interest rate accrued for infinitesimal period of time from t to t+dt as of time t.

$dr_t$	=	$a(b-r_t)dt + \sigma dz_t$	Vasicek model	const parameter : a, b and $\sigma$
$dr_t$	=	$(\theta_t - ar_t)dt + \sigma dz_t$	Hull White model	const parameter : a, $\sigma$ and $ heta_{t}$
$dr_t$	=	$(\theta_t - a_t r_t) dt + \sigma dz_t$	Hull White extended Vasicek model	const parameter : $\sigma$ , $a_t$ and $\theta_t$
$dr_t$	=	$\theta_t dt + \sigma dz_t$	Ho Lee model	

All above short rate models are *normal short rate* model, which implies *lognormal bond price*.

- Hull White model can be degenerated to Vasicek by setting  $\theta_t = ab$
- Hull White model can be degenerated to Ho Lee by setting a = 0
- *Vasicek* fits market data of yield curve and volatility matrix with only 3 constant parameters a, b and  $\sigma$
- Vasicek does not have enough degree of freedom to achieve a perfect match with market quotes of yield curve
- Hull White introduces time dependent  $\theta_t$  which can achieve a perfect match with market quotes of yield curve
- Hull White has extra parameters a and  $\sigma$  for fitting the volatility matrix
- both Hull White and Ho Lee guarantee consistency with the yield curve
- given short rate model, we can derive bond price formula (which also offers discount factor) as shown in the following ...

#### Vasciek bond price

$$P_{t}(T) = A(t,T)e^{-r_{t}B(t,T)}$$

$$\Rightarrow B(t,T) = \frac{1}{a}(1-e^{-a(T-t)})$$

$$\Rightarrow A(t,T) = \exp\left((b-\frac{\sigma^{2}}{2a^{2}})(B(t,T)-(T-t)) - \frac{\sigma^{2}}{4a}B^{2}(t,T)\right)$$

$$\Rightarrow B(t,T) = A(t,T)e^{-r_{t}B(t,T)}$$

$$\Rightarrow B(t,T) = \frac{1}{a}(1-e^{-a(T-t)})$$

$$\Rightarrow B(t,T) = \exp\left(-\int_{t}^{T} \theta_{s}B(s,T)ds - \frac{\sigma^{2}}{2a^{2}}(B(t,T)-(T-t)) - \frac{\sigma^{2}}{4a}B^{2}(t,T)\right)$$

$$\Rightarrow A(t,T) = \exp\left(-\int_{t}^{T} \theta_{s}B(s,T)ds - \frac{\sigma^{2}}{2a^{2}}(B(t,T)-(T-t)) - \frac{\sigma^{2}}{4a}B^{2}(t,T)\right)$$

$$\Rightarrow A(t,T) = \exp\left(-\int_{t}^{T} \theta_{s}B(s,T)ds - \frac{\sigma^{2}}{2a^{2}}(B(t,T)-(T-t)) - \frac{\sigma^{2}}{4a}B^{2}(t,T)\right)$$

$$\Rightarrow eq(2)$$

eq(2)

Lets check if we can get eq(1) by putting  $\theta_t = ab$  into eq(2). We consider this part for simplicity.

$$\begin{split} &-\int_t^T \theta_s B(s,T) ds - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= -\int_t^T ab B(s,T) ds - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= -ab \frac{1}{a} \int_t^T (1 - e^{-a(T-s)}) ds - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= -b \left[ \int_t^T 1 ds - \int_t^T e^{-a(T-s)} ds \right] - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= -b \left[ (T-t) - \frac{1}{a} (e^{-a(T-T)} - e^{-a(T-t)}) \right] - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= -b \left[ (T-t) - \frac{1}{a} (1 - e^{-a(T-t)}) \right] - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= b(B(t,T) - (T-t)) - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) \\ &= (b - \frac{\sigma^2}{2a^2}) (B(t,T) - (T-t)) \end{split}$$

## 2. Risk neutral expectation method (RNE) to bond price

Consider current time t, and time s is a time point in the future, such that s > t:

•	step 1 : solve for short rate	$r_S$	this is random
•	step 2 : solve for discount path	$\int_{t}^{T} r_{s} ds$	this is random
•	step 3 : expectation and variance of short rate	$E[r_s \mid I_t]$ and $V[r_s \mid I_t]$	deterministic given $r_t$
•	step 4 : expectation and variance of discount path	$E[\int_t^T r_s ds \mid I_t]$ and $V[\int_t^T r_s ds \mid I_t]$	deterministic given $r_t$
•	step 5 : bond price in form of	$P_t(T) = A(t,T)e^{-r_t B(t,T)}$	deterministic given r <sub>t</sub>

This method is applicable to both *Vasicek* and *Hull White*. Lets try *RNE* with *Vasicek*  $dr_t = a(b - r_t)dt + \sigma dz_t$ , the results are:

$$\begin{array}{lll} r_{s} & = & r_{t}e^{-a(s-t)} + b(1-e^{-a(s-t)}) + \sigma\int_{t}^{s}e^{-a(s-u)}dz_{u} & besides \ r_{s} \ is \ normal \\ E[r_{s} \mid I_{t}] & = & r_{t}e^{-a(s-t)} + b(1-e^{-a(s-t)}) \\ V[r_{s} \mid I_{t}] & = & \frac{\sigma^{2}}{2a}(1-e^{-2as}) \\ \int_{t}^{T}r_{s}ds & = & b(T-t) + \frac{1}{a}(r_{t}-b)(1-e^{-a(T-t)}) + \sigma\int_{t}^{T}\int_{u}^{T}e^{-a(s-u)}dsdz_{u} & besides \int_{t}^{T}r_{s}ds \ is \ normal \\ E[\int_{t}^{T}r_{s}ds\mid I_{t}] & = & b(T-t) + \frac{1}{a}(r_{t}-b)(1-e^{-a(T-t)}) \\ V[\int_{t}^{T}r_{s}ds\mid I_{t}] & = & \frac{\sigma^{2}}{a^{2}}\Big((T-t) - \frac{2}{a}(1-e^{-a(T-t)}) + \frac{1}{2a}(1-e^{-2a(T-t)})\Big) \end{array}$$

# step 1 : solve for interest rate

Lets remove  $r_t$  from RHS by considering:

$$d(r_t e^{at}) = r_t de^{at} + e^{at} dr_t$$

$$= r_t a e^{at} dt + e^{at} (a(b - r_t) dt + \sigma dz_t)$$

$$= r_t a e^{at} dt + (abe^{at} - r_t a e^{at}) dt + \sigma e^{at} dz_t$$

$$= abe^{at} dt + \sigma e^{at} dz_t$$

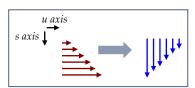
Integrate both sides from t to s:

$$\begin{split} \int_t^s d(r_u e^{au}) &= \int_t^s abe^{au} du + \int_t^s \sigma e^{au} dz_u \\ r_s e^{as} - r_t e^{at} &= \int_{u=t}^{u=s} be^{au} dau + \sigma \int_t^s e^{au} dz_u \\ r_s e^{as} - r_t e^{at} &= b(e^{as} - e^{at}) + \sigma \int_t^s e^{au} dz_u \\ r_s e^{as} &= r_t e^{at} + b(e^{as} - e^{at}) + \sigma \int_t^s e^{au} dz_u \\ r_s &= r_t e^{a(s-t)} + b(1 - e^{-a(s-t)}) + \sigma \int_t^s e^{-a(s-u)} dz_u \\ r_s &= r_t w_{t,s} + b(1 - w_{t,s}) + \sigma \int_t^s w_{u,s} dz_u \\ &= \varepsilon (E[r_s], V[r_s]) \end{split}$$
where  $w_{t,s} = e^{-a(s-t)}$  is DF

## step 2 : solve for discounting path

$$\begin{split} \int_{t}^{T} r_{s} ds &= \int_{t}^{T} (r_{t} e^{-a(s-t)} + b(1 - e^{-a(s-t)})) ds + \sigma \int_{t}^{T} (\int_{t}^{s} e^{-a(s-u)} dz_{u}) ds \\ &= \int_{t}^{T} (b + (r_{t} - b) e^{-a(s-t)}) ds + \sigma \int_{t}^{T} \int_{u}^{T} e^{-a(s-u)} ds dz_{u} \\ &= b(T - t) + (r_{t} - b) \int_{s=t}^{s=T} e^{-a(s-t)} d(s - t) + \sigma \int_{t}^{T} \int_{u}^{T} e^{-a(s-u)} ds dz_{u} \\ &= b(T - t) - \frac{1}{a} (r_{t} - b) (e^{-a(T - t)} - e^{-a(t-t)}) + \sigma \int_{t}^{T} \int_{u}^{T} e^{-a(s-u)} ds dz_{u} \\ &= b(T - t) + \frac{1}{a} (r_{t} - b) (1 - e^{-a(T - t)}) + \sigma \int_{t}^{T} \int_{u}^{T} e^{-a(s-u)} ds dz_{u} \\ &= \varepsilon (E[\int_{t}^{T} r_{s} ds], V[\int_{t}^{T} r_{s} ds]) \end{split}$$

swap integration sequence



#### step 3: find the expectation and volatility of interest rate

$$\begin{split} E[r_s \mid I_t] &= r_t e^{-a(s-t)} + b(1-e^{-a(s-t)}) & expectation of Itos is zero \\ V[r_s \mid I_t] &= E[(\sigma)_t^s e^{-a(s-u)} dz_u)^2] \\ &= \sigma^2 \int_t^s E[e^{-2a(s-u)}] du & variance of Itos becomes time integral \\ &= \sigma^2 \int_t^s e^{-2a(s-u)} du & \\ &= \frac{\sigma^2}{2a} [e^{-2a(s-u)}]_0^s & \\ &= \frac{\sigma^2}{2a} (1-e^{-2as}) \end{split}$$

# step 4: find the expectation and volatility of discounting path

$$\begin{split} E[\int_t^T r_s ds \, | \, I_t \, ] &= b(T-t) + \frac{1}{a} (r_t - b)(1 - e^{-a(T-t)}) & expectation of Itos is zero \\ V[\int_t^T r_s ds \, | \, I_t \, ] &= E[(\sigma)_t^T \int_u^T e^{-a(s-u)} ds dz_u)^2 \, ] \\ &= \sigma^2 \int_t^T E[(\int_u^T e^{-a(s-u)} ds)^2 \, ] du \\ &= \sigma^2 \int_t^T \left( \int_u^T e^{-a(s-u)} ds \right)^2 du \\ &= \sigma^2 \int_t^T \left( \frac{1}{-a} [e^{-a(s-u)}]_u^T \right)^2 du \\ &= \frac{\sigma^2}{a^2} \int_t^T (1 - e^{-a(T-u)})^2 du \\ &= \frac{\sigma^2}{a^2} \int_t^T (1 - 2e^{-a(T-u)} + e^{-2a(T-u)}) du \\ &= \frac{\sigma^2}{a^2} \left( (T-t) - \frac{2}{a} (e^{-a(T-T)} - e^{-a(T-t)}) + \frac{1}{2a} (e^{-2a(T-T)} - e^{-2a(T-t)}) \right) \\ &= \frac{\sigma^2}{a^2} \left( (T-t) - \frac{2}{a} (1 - e^{-a(T-t)}) + \frac{1}{2a} (1 - e^{-2a(T-t)}) \right) \end{split}$$

#### step 5: find spot and forward bond price

$$P_{t}(T) = E\left[ s1 \times e^{-\int_{t}^{T} r_{t} ds} \mid r_{t} \right]$$
 FTAP 
$$= E\left[ e^{s(-E[\int_{t}^{T} r_{t} ds], V[\int_{t}^{T} r_{t} ds]}) \mid r_{t} \right]$$
 
$$= \exp\left( -E[\int_{t}^{T} r_{s} ds] + \frac{1}{2}V[\int_{t}^{T} r_{s} ds] \right)$$
 
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$$= \exp\left( -E[\int_{t}^{T} r_{s} ds] + \frac{1}{2}V[\int_{t}^{T} r_{s} ds]$$

then 
$$A(t,T) = P_t(T)e^{+B(t,T)r_t} \qquad since \ P_t(T) = A(t,T)e^{-B(t,T)r_t}$$

$$= \exp\left(-\left(b(T-t) + \frac{1}{a}(r_t - b)(1 - e^{-a(T-t)})\right) + \frac{1}{2}\frac{\sigma^2}{a^2}\left((T-t) - \frac{2}{a}(1 - e^{-a(T-t)}) + \frac{1}{2a}(1 - e^{-2a(T-t)})\right) + r_t B(t,T)\right)$$

$$= \exp\left(-b(T-t) - (r_t - b)B(t,T) + \frac{\sigma^2}{2a^2}\left((T-t) - 2B(t,T) + B(t,T) - \frac{1}{2}aB^2(t,T)\right) + r_t B(t,T)\right)$$

$$= \exp\left(-b(T-t) + bB(t,T) + \frac{\sigma^2}{2a^2}\left((T-t) - B(t,T) - \frac{1}{2}aB^2(t,T)\right)\right)$$

$$= \exp\left(b(B(t,T) - (T-t)) + \frac{\sigma^2}{2a^2}((T-t) - B(t,T)) - \frac{\sigma^2}{4a}B^2(t,T)\right)$$

$$= \exp\left((b - \frac{\sigma^2}{2a^2})(B(t,T) - (T-t)) - \frac{\sigma^2}{4a}B^2(t,T)\right)$$

$$= eq(4)$$

## Ito's integral and Ito's isometry

Itos integral  $\int_0^t f_s dz_s$  is an integral in  $z_t$  space with deterministic or stochastic integrand  $f_t$ . Itos integral is itself stochastic.

	property	deterministic integrad	stochastic integrand	
(1)	martingale	yes	yes	
(2)	$E[\int_t^T f_s dz_s \mid I_t]$	0	0	
(3)	$E[(\int_t^T f_s dz_s)^2 \mid I_t]$	$\int_{t}^{T} f_{s}^{2} ds$	$\int_{t}^{T} E[f_{s}^{2} \mid I_{t}] ds$	also called Ito's isometry
(4)	$E[(\int_t^T f_s dz_s)(\int_t^T g_s dz_s)   I_t]$	$\int_{t}^{T} f_{s} g_{s} ds$	$\int_{t}^{T} E[f_{s}g_{s} \mid I_{t}]ds$	also called Ito's isometry
(5)	distribution	$\varepsilon(0, \int_t^T f_s^2 ds)$	unknown	

Lets prove. We put t=0 for simplicity.

$$(4) \quad E[(\int_{0}^{t} f_{s} dz_{s})(\int_{0}^{t} g_{s} dz_{s}) | I_{0}] = E[\lim_{N \to \infty} (\sum_{n=0}^{N-1} f_{n\Delta t}(z_{(n+1)\Delta t} - z_{n\Delta t})) \times (\sum_{m=0}^{N-1} g_{m\Delta t}(z_{(m+1)\Delta t} - z_{m\Delta t})) | I_{0}]$$

$$= E[\lim_{N \to \infty} \sum_{n=0}^{N-1} \sum_{m=0}^{N-1} f_{n\Delta t} g_{m\Delta t}(z_{(n+1)\Delta t} - z_{n\Delta t})(z_{(m+1)\Delta t} - z_{m\Delta t}) | I_{0}]$$

$$= \lim_{N \to \infty} \sum_{n=0}^{N-1} \sum_{m=0}^{N-1} E[f_{n\Delta t} g_{m\Delta t}(z_{(n+1)\Delta t} - z_{n\Delta t})(z_{(m+1)\Delta t} - z_{m\Delta t}) | I_{0}]$$

$$= \lim_{N \to \infty} \sum_{n=0}^{N-1} E[f_{n\Delta t} g_{n\Delta t}(z_{(n+1)\Delta t} - z_{n\Delta t})(z_{(m+1)\Delta t} - z_{m\Delta t}) | I_{0}]$$

$$= \lim_{N \to \infty} \sum_{n=0}^{N-1} E[f_{n\Delta t} g_{n\Delta t}(z_{(n+1)\Delta t} - z_{n\Delta t})(z_{(n+1)\Delta t} - z_{m\Delta t}) | I_{0}]$$

$$= \lim_{N \to \infty} \sum_{n=0}^{N-1} E[f_{n\Delta t} g_{n\Delta t}(z_{(n+1)\Delta t} - z_{n\Delta t})(z_{(n+1)\Delta t} - z_{m\Delta t}) | I_{0}]$$

$$= \lim_{N \to \infty} \sum_{n=0}^{N-1} E[f_{n\Delta t} g_{n\Delta t} | I_{0}] | I_{0}]$$

$$= E[f_{t} g_{t} E[(z_{t'} - z_{t})^{2} | I_{t}] | I_{0}]$$

$$= E[f_{t} g_{t} E[(z_{t'} - z_{t'})^{2} | I_{t}] | I_{0}]$$

$$= E[f_{t} g_{t}(z_{t'} - z_{t'})(z_{s'} - z_{s}) | I_{0}] | E[f_{t} g_{s}(z_{s'} - z_{s}) | I_{0}] | E[f_{t} g_{s}(z_{s'} - z_{s}) | I_{0}]$$

$$= E[f_{t} g_{s}(z_{s'} - z_{s}) | I_{0}] | E[f_{t} g_{s}(z_$$

If  $f_t$  is deterministic, then Ito's integral must be normal (because the sum of normal is also normal).

$$(5) \quad \int_0^t f_s dz_s = \lim_{N \to \infty} \sum_{n=0}^{N-1} f_{n\Delta t} (z_{(n+1)\Delta t} - z_{n\Delta t})$$

$$= \lim_{N \to \infty} \sum_{n=0}^{N-1} f_{n\Delta t} (\varepsilon \Delta t)$$

$$= \varepsilon(\mu = 0, \sigma^2 = \int_0^t f_s^2 ds)$$

# 3. Partial differential equation method (PDE) to bond price

Let's try PDE approach with Hull White, which does not involve solving short rate and discounting path:

- step 1 : given SDE of underlying short rate
- step 2: derive PDE of contingent claim using Dr Yan's fastlane
- step 3 : break into two ODEs
- step 4 : solve first order linear ODE by integrating factor
- step 5 : solve first order linear ODE by separation of variables

# step 2 : fastlane to PDE

With Dr Yan's fastlane, we can write the *PDE* of any contingent claim  $P_t(T)$  with underlyings  $x_t$  and  $y_t$  in single step.

$$\begin{array}{rcl} dx_t & = & \alpha_{xt}dt + \beta_{xt}dz_{1t} \\ dy_t & = & \alpha_{yt}dt + \beta_{yt}dz_{2t} \\ dz_{1t}dz_{2t} & = & \rho dt \\ \\ r_tP_t(T) & = & \partial_tP_t + \alpha_{xt}(\partial_xP_t) + \alpha_{yt}(\partial_yP_t) + \frac{1}{2}\beta_{xt}^2(\partial_{xx}P_t) + \rho\beta_{xt}\beta_{yt}(\partial_{xy}P_t) + \frac{1}{2}\beta_{yt}^2(\partial_{yy}P_t) + \dots & \textit{for 2 risk factors} \\ & = & \partial_tP_t + \alpha_{xt}(\partial_xP_t) + \frac{1}{2}\beta_{xt}^2(\partial_{xx}P_t) & \textit{for 1 risk factor} \\ & = & \partial_tP_t + (\theta_t - ar_t)(\partial_rP_t) + \frac{\sigma^2}{2}(\partial_{rr}P_t) & x_t = r_t, \ \alpha_t = \theta_t - ar_t, \ \beta_t = \sigma t \\ P_T(T) & = & 1 & \textit{boundary condition} \end{array}$$

## step 3: breakdown into two ODEs

We put an ansatz (educated guess) into the PDE:

$$P_t(T) = A(t,T)e^{-r_tB(t,T)}$$

$$\begin{split} r_t A(t,T) e^{-r_t B(t,T)} &= & \hat{\sigma}_t (A(t,T) e^{-r_t B(t,T)}) + (\theta_t - a r_t) \hat{\sigma}_r (A(t,T) e^{-r_t B(t,T)}) + \frac{\sigma^2}{2} \hat{\sigma}_{rr} (A(t,T) e^{-r_t B(t,T)}) \\ &= & (\hat{\sigma}_t A(t,T)) e^{-r_t B(t,T)} - r_t (\hat{\sigma}_t B(t,T)) A(t,T) e^{-r_t B(t,T)} - (\theta_t - a r_t) A(t,T) B(t,T) e^{-r_t B(t,T)} + \frac{\sigma^2}{2} A(t,T) B^2(t,T) e^{-r_t B(t,T)} \\ r_t A(t,T) &= & \hat{\sigma}_t A(t,T) - r_t (\hat{\sigma}_t B(t,T)) A(t,T) - (\theta_t - a r_t) A(t,T) B(t,T) + \frac{\sigma^2}{2} A(t,T) B^2(t,T) \end{split}$$

Move all instantaneous rate  $r_t$  terms to *LHS*:

$$r_t A(t,T) + r_t A(t,T) \partial_t B(t,T) - a r_t A(t,T) B(t,T) = \partial_t A(t,T) - \theta_t A(t,T) B(t,T) + \frac{\sigma^2}{2} A(t,T) B^2(t,T)$$

$$r_t A(t,T) (1 + \partial_t B(t,T) - a B(t,T)) = \partial_t A(t,T) - \theta_t A(t,T) B(t,T) + \frac{\sigma^2}{2} A(t,T) B^2(t,T)$$

Since equality holds for all short rate  $r_t$ , we can breakdown the *ODE* into two linear *ODE*:

$$eq(5) \qquad \partial_t B(t,T) - aB(t,T) + 1 \qquad = \qquad 0 \qquad and \left[ B(T,T) = 0 \right] \qquad \qquad linear \ ODE \ solved \ by \ integrating \ factor$$

$$eq(6) \qquad \partial_t A(t,T) + \left( -\theta_t B(t,T) + \frac{\sigma^2}{2} B^2(t,T) \right) A(t,T) \quad = \qquad 0 \qquad and \left[ A(T,T) = 1 \right] \qquad \qquad linear \ ODE \ solved \ by \ separation \ of \ variables$$

Why boundary condition can be broken down in this way?

$$P_T(T) = 1$$

$$\Rightarrow A(T,T)e^{-r_TB(T,T)} = 1e^{-r_T0}$$

$$\Rightarrow \begin{bmatrix} A(T,T) \\ B(T,T) \end{bmatrix} = \begin{bmatrix} 1 \\ 0 \end{bmatrix}$$

The following *first order linear ODE* can be solved by scaling both sides with an integrating factor  $e^{-\int q(x)dx}$ .

$$f'(x) = p(x) + q(x)f(x)$$

$$(\frac{df(x)}{dx} - q(x)f(x))e^{-\int q(x)dx} = p(x)e^{-\int q(x)dx} \qquad no fin RHS$$

$$\frac{d}{dx}(f(x)e^{-\int q(x)dx}) = p(x)e^{-\int q(x)dx} \qquad no fin both sides, only f' exists$$

$$f(x)e^{-\int q(x)dx} = \int (p(x)e^{-\int q(x)dx})dx$$

$$f(x) = e^{\int q(x)dx} \int (p(x)e^{-\int q(x)dx})dx$$

$$f(x) = Q(x)\int \frac{p(x)}{Q(x)}dx \qquad where Q(x) = e^{\int q(x)dx}$$

Now solve *eq*(5) with integrating factor technique.

$$\frac{dB(t,T)}{dt} - aB(t,T) = -1 \qquad such that \ B(T,T) = 0 \quad recap \ eq(5)$$

$$(\frac{dB(t,T)}{dt} - aB(t,T))e^{-at} = -e^{-at}$$

$$\frac{d}{dt}(B(t,T)e^{-at}) = -e^{-at}$$

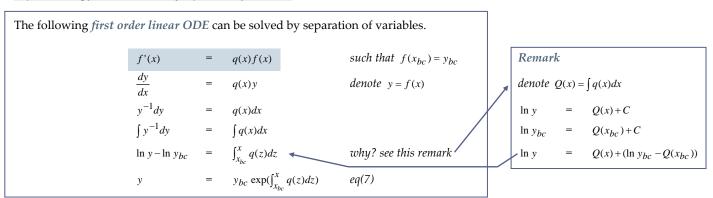
$$B(t,T)e^{-at} = -\int e^{-at} dt = \frac{1}{a}e^{-at} + C$$

$$B(t,T) = \frac{1}{a} + Ce^{at}$$

$$C = [B(t,T) - \frac{1}{a}]e^{-at} = [\underline{B(T,T)} - \frac{1}{a}]e^{-aT} = -\frac{1}{a}e^{-aT}$$

$$B(t,T) = \frac{1}{a} - \frac{1}{a}e^{-aT}e^{at} = \frac{1}{a}(1 - e^{-a(T-t)}) \qquad exactly the same as Vasicek$$

# step 5 : solving first order ODE by separation of variables



Now solve eq(6) with separation of variables technique.

$$\frac{dA(t,T)}{dt} + (-\theta_t B(t,T) + \frac{\sigma^2}{2} B^2(t,T))A(t,T) = 0 \quad \text{such that } A(T,T) = 1$$

$$\frac{dA(t,T)}{dt} = q(t)A(t,T)$$

$$A(t,T) = A(T,T) \exp\left(-\int_t^T q(s)ds\right) \quad \text{using eq(7)}$$

$$= \exp\left(-\int_t^T \theta_s B(s,T)ds + \frac{\sigma^2}{2} \int_t^T B^2(s,T)ds\right) \quad \text{using } A(T,T) = 1$$

$$= \exp\left(-\int_t^T \theta_s B(s,T)ds - \frac{\sigma^2}{2} \left(\frac{1}{a^2} (B(t,T) - (T-t)) + \frac{1}{2a} B^2(t,T)\right)\right)$$

$$= \exp\left(-\int_t^T \theta_s B(s,T)ds - \frac{\sigma^2}{2a^2} (B(t,T) - (T-t)) - \frac{\sigma^2}{4a} B^2(t,T)\right)$$

### B(t,T) related formulae

We recap eq(3) and eq(3') and derive eq(3'') below, as they are useful in simplifying the messy A(t,T) expression.

$$\frac{1}{a}(1-e^{-a(T-t)}) = B(t,T) \qquad eq(3)$$

$$\frac{1}{a}(1-e^{-2a(T-t)}) = 2B(t,T)-aB^{2}(t,T) \qquad eq(3')$$

$$\int_{t}^{T} B^{2}(s,T)ds = \frac{1}{a^{2}} \int_{t}^{T} (1-e^{-a(T-s)})^{2} ds$$

$$= \frac{1}{a^{2}} \left[ \int_{t}^{T} ds - 2 \int_{t}^{T} e^{-a(T-s)} ds + \int_{t}^{T} e^{-2a(T-s)} ds \right]$$

$$= \frac{1}{a^{2}} \left[ (T-t) - \frac{2}{a} [e^{-a(T-s)}]_{t}^{T} + \frac{1}{2a} [e^{-2a(T-s)}]_{t}^{T} \right]$$

$$= \frac{1}{a^{2}} \left[ (T-t) - \frac{2}{a} [1-e^{-a(T-t)}] + \frac{1}{2a} [1-e^{-2a(T-t)}] \right]$$

$$= \frac{1}{a^{2}} \left[ (T-t) - 2B(t,T) + \frac{1}{2} (2B(t,T) - aB^{2}(t,T)) \right]$$

$$= -\frac{1}{a^{2}} (B(t,T) - (T-t)) - \frac{1}{2a} B^{2}(t,T) \qquad eq(3'')$$

Various derivatives of B(t,T) for the sake of *part4*.

$$\begin{array}{llll} \partial_t B(t,T) & = & -e^{-a(T-t)} & = & aB(t,T)-1 & \Rightarrow & aB(t,T)-\partial_t B(t,T)=1 & eq(3a) \\ \partial_{tt} B(t,T) & = & a\partial_t B(t,T) & = & a(aB(t,T)-1) & \Rightarrow & a\partial_t B(t,T)-\partial_{tt} B(t,T)=0 & eq(3b) \\ \partial_T B(t,T) & = & e^{-a(T-t)} & = & 1-aB(t,T) & \Rightarrow & aB(t,T)+\partial_T B(t,T)=1 & eq(3c) \\ \partial_{TT} B(t,T) & = & -a\partial_T B(t,T) & = & a(aB(t,T)-1) & \Rightarrow & a\partial_T B(t,T)+\partial_{TT} B(t,T)=0 & eq(3d) \end{array}$$

# 4. Perfect match with yield curve

Calibrate Hull White with term structure

In this section, we are going to prove that Hull White parameters can be decoupled into two groups:

- $\theta_t$  for perfect match with yield curve
- a and  $\sigma$  for matching volatility matrix

The proof makes use of forward short rate  $r_t(s)$ . We recap its definition in eq(8) of chap4.doc.

$$P_t(T) = A(t,T)e^{-r_tB(t,T)}$$

$$P_t(T) = e^{-\int_t^T r_t(s)ds}$$

eq(8) of chap4.doc

using eq(3c)

and

We have not derived relation between  $r_t$  and  $r_t(s)$  in *chap4.doc*, we are going to do it now. Comparing 2 equations above :

$$e^{-\int_{t}^{T} r_{t}(s)ds} = A(t,T)e^{-r_{t}B(t,T)}$$

$$\int_{t}^{T} r_{t}(s)ds = -\ln A(t,T) + r_{t}B(t,T)$$

$$= -\ln \exp\left(-\int_{t}^{T} \theta_{s}B(s,T)ds - \frac{\sigma^{2}}{2a^{2}}(B(t,T) - (T-t)) - \frac{\sigma^{2}}{4a}B^{2}(t,T)\right) + r_{t}B(t,T)$$

$$\int_{t}^{T} r_{t}(s)ds = \int_{t}^{T} \theta_{s}B(s,T)ds + \frac{\sigma^{2}}{2a^{2}}(B(t,T) - (T-t)) + \frac{\sigma^{2}}{4a}B^{2}(t,T) + r_{t}B(t,T)$$

$$eq(8)$$

#### step2

We differentiate both *LHS* and *RHS* of *eq*(8) *wrt* to maturity *T* using Leibniz rule. We have *eq*(9).

$$\bullet \qquad \partial_T \int_t^T r_t(s) ds \qquad = \qquad r_t(T)$$

$$\partial_T \int_t^T r_t(s) ds = r_t(T)$$

$$\partial_T \int_t^T \theta_s B(s, T) ds = \int_t^T \partial_T (\theta_s B(s, T)) ds + \theta_T \underbrace{B(T, T) \partial_T T}_{0} - \theta_t B(t, T) \underbrace{\partial_T t}_{0}$$

$$= \int_{t}^{T} \theta_{s} \partial_{T} B(s, T) ds$$

$$= \int_{t}^{T} \theta_{s} \hat{\sigma}_{T} B(s,T) ds$$

$$\bullet \quad \hat{\sigma}_{T} (B(t,T) - (T-t)) = \hat{\sigma}_{T} B(t,T) - 1$$

$$= 1 - aB(t,T) - 1$$

$$= -aB(t,T)$$

$$= -aB(t,T)$$

$$\bullet \quad \partial_T \left[ \frac{\sigma^2}{4a} B^2 + r_t B \right] = \frac{\sigma^2}{2a} B(t,T) \partial_T B(t,T) + r_t \partial_T B(t,T)$$

$$\Rightarrow r_t(T) = \int_t^T \theta_s \partial_T B(s,T) ds - \frac{\sigma^2}{2a} B(t,T) + \frac{\sigma^2}{2a} B(t,T) \partial_T B(t,T) + r_t \partial_T B(t,T)$$
 eq(9)

We differentiate *again* both *LHS* and *RHS* of *eq*(9) *wrt* maturity *T* using Leibniz rule. We have *eq*(10).

$$\begin{array}{lll} \bullet & \partial_T \int_t^T \theta_s \partial_T B(s,T) ds & = & \int_t^T \partial_T (\theta_s \partial_T B(s,T)) ds + \theta_T \underbrace{\partial_T B(T,T)}_{1-aB(T,T)=1} \underbrace{\partial_T T}_{1} - \theta_t \partial_T B(t,T) \underbrace{\partial_T t}_{0} \\ & = & \int_t^T \theta_s \partial_{TT} B(s,T) ds + \theta_T \\ \\ \Rightarrow & \partial_T r_t(T) & = & \int_t^T \theta_s \partial_{TT} B(s,T) ds + \theta_T - \frac{\sigma^2}{2a} \partial_T B(t,T) + \frac{\sigma^2}{2a} [(\partial_T B(t,T))^2 + B(t,T) \partial_{TT} B(t,T)] + r_t \partial_{TT} B(t,T) \end{array} \end{aligned}$$

### step4

Since derivative of *B* gives a linear function of *B*, see eq(3a-d) just like exponential function, so we repeatedly differentiate eq(8) so as to obtain eq(9,10), whose sum does help to remove many terms by applying eq(3c,d):

$$\begin{split} ar_t(T) + \partial_T r_t(T) &= \begin{bmatrix} \int_t^T \theta_s(a\partial_T B(s,T)) ds - \frac{\sigma^2}{2a} aB(t,T) + \frac{\sigma^2}{2a} B(t,T) (a\partial_T B(t,T)) + r_t(a\partial_T B(t,T)) + \\ \int_t^T \theta_s \partial_{TT} B(s,T) ds + \theta_T - \frac{\sigma^2}{2a} \partial_T B(t,T) + \frac{\sigma^2}{2a} [(\partial_T B(t,T))^2 + B(t,T) \partial_{TT} B(t,T)] + r_t \partial_{TT} B(t,T) \end{bmatrix} \\ &= \theta_T - \frac{\sigma^2}{2a} \underbrace{(aB(t,T) + \partial_T B(t,T))}_{1} + \frac{\sigma^2}{2a} \underbrace{(\partial_T B(t,T))}_{e^{-a(T-t)}}^2 \\ &= \theta_T - \frac{\sigma^2}{2a} (1 - e^{-2a(T-t)}) \\ \theta_t &= ar_t(T) + \partial_T r_t(T) + \frac{\sigma^2}{2a} (1 - e^{-2at}) \qquad putting \ t = 0 \ and \ T = t \end{aligned}$$

We write  $\theta_t$  in terms of forward rate term structure and the other two Hull White parameters a and  $\sigma$ .

- as opposed to Vasciek, Hull White has enough freedom to fit market yield curve perfectly
- in Quantlib, we construct term structure using rate-helpers and then ...
- Hull White model can ask for forward short rate from term structure to calculate  $\theta_t$
- Hull White model parameters a and  $\sigma$  are then calibrated to fit volatility matrix

# Recall Leibniz rule of differentiation

General and specialized versions

$$\begin{array}{lcl} \partial_T \int_{a(T)}^{b(T)} f(s,T) ds & = & \int_{a(T)}^{b(T)} \partial_T f(s,T) ds + f(b(T),T) \partial_T b(T) - f(a(T),T) \partial_T a(T) \\ \partial_T \int_{a(T)}^{b(T)} f(s) ds & = & f(b(T)) \partial_T b(T) - f(a(T)) \partial_T a(T) \\ \partial_T \int_{c}^{T} f(s) ds & = & f(T) \end{array}$$

#### 5. Hull White tree

and

Construction of *Hull White tree* (*from market data*) is done in two steps:

- (1) construction of a tree for symmetric process  $x_t$
- (2) apply time dependent *shift process yt* on the tree
- step 1 finds branching probabilities which are uniquely determined by Hull White parameters  $(a, \sigma)$
- step 2 finds shifts by pricing bonds with the tree and compared to market quoted yield curve without knowing  $\theta_1$
- *Hull White* parameter  $\theta_t$  is not needed in both steps

#### Step 1: Construction of symmetric process tree

Markov process  $x_t$  (*NOT* the short rate process) is defined as a process that fluctuates around zero symmetrically:

$$dx_t = -ax_t dt + \sigma dz_t$$
$$= -ax_t dt + \sigma \varepsilon \sqrt{dt}$$
$$x_0 = 0$$

boundary condition

Discretization in time domain (with constant time resolution  $\Delta t$ ):

Discretization in price domain (constant price resolution  $\Delta x$ ):

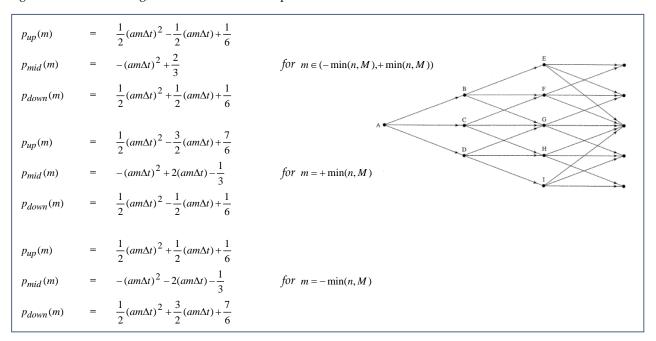
$$x_{n\Delta t} = m\Delta x$$
 denoted as state  $(n,m)$  where  $m$  is price index  $m \in [-\min(n,M),+\min(n,M)]$  where  $M$  is price limit hence  $x_{(n+1)\Delta t} = x_{n\Delta t} + \varepsilon_m \Delta x$  where  $\varepsilon_m$  is RNG  $= m\Delta x + \varepsilon_m \Delta x$   $= (m + \varepsilon_m)\Delta x$  eq(13)

Given current state (n,m), transition probability matrix is:

- each edge is a feasible state transition
- each node in timestep n connects to 3 nodes in timestep n+1, with probabilities  $p_u(m)$ ,  $p_m(m)$  and  $p_d(m)$
- topology for nodes in the middle is different from those at both ends in price domain
- with the constraint that sum of transition probabilites is *one*, degree of freedom is *two* only
- probabilities are found by matching 1st and 2nd moments of eq(12) and eq(13)
- probabilities are *price dependent* but *time independent*, thus they bear index *m*, as there is no *t* index in moment matching:

moments of eq(13) 
$$eq(14a) \quad p_{up}(m) + p_{mid}(m) + p_{down}(m) = 1$$
 
$$eq(14b) \quad p_{up}(m)(+\Delta x) + p_{mid}(m)(0) + p_{down}(m)(-\Delta x) = -ax_{n\Delta t}\Delta t$$
 
$$eq(14c) \quad p_{up}(m)(+\Delta x)^2 + p_{mid}(m)(0)^2 + p_{down}(m)(-\Delta x)^2 = (-a(m\Delta x)\Delta t)^2 + \sigma^2\Delta t$$
 
$$E[(\varepsilon_m \Delta x)^2] \qquad E^2[\varepsilon_m \Delta x] + V[\varepsilon_m \Delta x]$$

By solving the moment matching, we have answers for all probabilities.



Let's verify the answer for case  $m \in (-\min(n, M), +\min(n, M))$ .

$$\begin{aligned} p_{up}(m) + p_{mid}(m) + p_{down}(m) \\ &= \frac{1}{2}(am\Delta t)^2 - \frac{1}{2}(am\Delta t) + \frac{1}{6} - (am\Delta t)^2 + \frac{2}{3} + \frac{1}{2}(am\Delta t)^2 + \frac{1}{2}(am\Delta t) + \frac{1}{6} \\ &= 1 \\ p_{up}(m)(+\Delta x) + p_{mid}(m)(0) + p_{down}(m)(-\Delta x) \\ &= \left(\frac{1}{2}(am\Delta t)^2 - \frac{1}{2}(am\Delta t) + \frac{1}{6}\right)(+\Delta x) + \left(\frac{1}{2}(am\Delta t)^2 + \frac{1}{2}(am\Delta t) + \frac{1}{6}\right)(-\Delta x) \\ &= -am\Delta x\Delta t \\ p_{up}(m)(+\Delta x)^2 + p_{mid}(m)(0)^2 + p_{down}(m)(-\Delta x)^2 \\ &= \left(\frac{1}{2}(am\Delta t)^2 - \frac{1}{2}(am\Delta t) + \frac{1}{6}\right)(+\Delta x)^2 + \left(\frac{1}{2}(am\Delta t)^2 + \frac{1}{2}(am\Delta t) + \frac{1}{6}\right)(-\Delta x)^2 \\ &= ((am\Delta t)^2 + \frac{1}{3})(\Delta x)^2 \\ &= (am\Delta x\Delta t)^2 + \frac{1}{3}(\Delta x)^2 \\ &= (am\Delta x\Delta t)^2 + \sigma^2 \Delta t \end{aligned} \qquad \text{by putting we set } (\Delta x)^2 = 3\sigma^2 \Delta t \\ &\Rightarrow verified \ eq(14c) \end{aligned}$$

Therefore, with the second order moment matching, we effectively determine the resolution of price domain:

$$\Delta x = \sigma \sqrt{3\Delta t}$$

Besides, this price resolution gives  $\frac{kurtosis}{2} (please prove it yourself)$ , which is consistent with the  $\frac{kurtosis}{2} (please prove it yourself)$ , which is consistent with the  $\frac{kurtosis}{2} (please prove it yourself)$ .

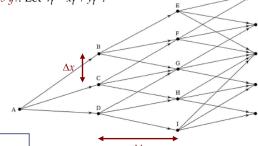
### Step 2: Deformation of tree

Deformation is done by adding time dependent but price independent shift to each node in the original tree. Shift process is:

$$dy_t = (\theta_t - ay_t)dt$$

*Hull White* short rate process is the sum of *symmetric process*  $x_t$  and *shift process*  $y_t$ . Let  $r_t = x_t + y_t$ :

$$\begin{array}{lll} dr_t & = & d(x_t + y_t) \\ & = & dx_t + dy_t \\ & = & -ax_t dt + \sigma dz_t + (\theta_t dt - ay_t) dt \\ & = & (\theta_t - a(x_t + y_t)) dt + \sigma dz_t \\ & = & (\theta_t - ar_t) dt + \sigma dz_t \end{array}$$



There are *two* steps in deforming the *Hull White* tree :

- (1) define *Q* and derive a recursive formula for *Q*, update it by *forward induction*
- (2) match with market quoted yield curve
- i.e. calibrating the tree (or adjusting  $y_t$ ) so that it describes the market data

## step1: update Q value recursively

$$Q_{n,m} = PV \text{ at time 0 of a security that pays $1 \text{ on reaching state } (n,m) \text{ and nothing otherwise, by FTAP} \dots$$

$$= E \left[ I(r_{n\Delta t} = m\Delta t + y_{n\Delta t}) \times \exp(-\int_{0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\int_{0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\int_{\infty}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right]$$

$$= E \left[ E \left[ I(x_{n\Delta t} = m\Delta t) \times \exp(-\sum_{k=0}^{n\Delta t} r_t dt) \mid I_0 \right] \right]$$

$$= E \left[ \frac{+I(x_{(n-1)\Delta t} = m\Delta t)}{+I(x_{(n-1)\Delta t} = (m+1)\Delta t)} \times \frac{I(x_{m-1} = t)}{+I(x_{(n-1)\Delta t} = m\Delta t)} \times \frac{I(x_{m-1} = t)}{+I(x_{(n-1)\Delta t} = (m+1)\Delta t)} \times \frac{I(x_{m-1} = t)}{+I(x_{m-1} = t)} \times \frac{I(x_{m-$$

where  $Q_{0,0} = 1$  boundary condition

#### Useful formulae involving delta function

In order to derive recursive formula for Q, we need a recursive formula for its payoff:

$$\begin{split} 1(x_{n\Delta t} = m\Delta x) &\equiv \begin{cases} 1(x_{(n-1)\Delta t} = m\Delta x) \times 1(\varepsilon_m = 0) + \\ 1(x_{(n-1)\Delta t} = (m-1)\Delta x) \times 1(\varepsilon_{m-1} = +1) + \\ 1(x_{(n-1)\Delta t} = (m+1)\Delta x) \times 1(\varepsilon_{m+1} = -1) \end{cases} \\ eq(16) \end{split}$$

Suppose x is a random variable with  $x_n$  as possible realization.

$$E[1(x = x_0) f(x)] = \int 1(x = x_0) f(x) p(x) dx$$

$$= f(x_0) p(x_0)$$

$$= f(x_0) \int 1(x = x_0) p(x) dx$$

$$= f(x_0) E[1(x = x_0)]$$
eq(17)

Conversion from random x (a tree layer at a specific time) to realized value  $x_n$  (a specific layer node)

$$f(x) = \sum_{n} f(x_n) 1(x = x_n)$$
 eq(18)

### step2: calibration to market yield curve

Consider bond price (*not Q contract*) with maturity  $(n+1)\Delta t$ :

$$P_{0}((n+1)\Delta t) = E[\exp(-\int_{0}^{(n+1)\Delta t} r_{t}dt) | I_{0}]$$

$$= E[\exp(-\sum_{k=0}^{n-1} r_{k\Delta t}\Delta t) \exp(-r_{n\Delta t}\Delta t) | I_{0}] \qquad note \ r_{n\Delta t} \ is \ random \ w$$

$$= E[\exp(-\sum_{k=0}^{n-1} r_{k\Delta t}\Delta t) [\sum_{m} \exp(-(m\Delta x + y_{n\Delta t})\Delta t) | I_{n\Delta t} = m\Delta x + y_{n\Delta t}] | I_{0}] \qquad by \ eq(18)$$

$$= \sum_{m} \underbrace{E[I(r_{n\Delta t} = m\Delta x + y_{n\Delta t}) \times \exp(-\sum_{k=0}^{n-1} r_{k\Delta t}\Delta t) | I_{0}] \times \exp(-(m\Delta x + y_{n\Delta t})\Delta t)}_{Q_{n,m}} \qquad rearrange \ terms, \ m \ is \ deterministic$$

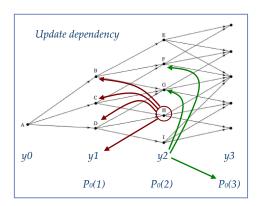
$$= \sum_{m} Q_{n,m} e^{-(m\Delta x + y_{n\Delta t})\Delta t}$$

$$= (\sum_{m} Q_{n,m} e^{-m\Delta x\Delta t}) \times e^{-y_{n\Delta t}\Delta t}$$

fctb 
$$y_{n\Delta t} = -\frac{1}{\Delta t} \ln \left[ \frac{P_0((n+1)\Delta t)}{\sum_m Q_{n,m} e^{-m\Delta x \Delta t}} \right]$$

Here is the algorithm, forward induction of Q and y should be done simultaneously.

```
Q[0,0] = 1;
y[0] = yieldcurve.zero_rate(dt); // y[0] = short rate at t=0
for (n=1:N)
{
    // step1 - Q[n,:] depends on Q[n-1,:] and y[n-1]
    for(m=-n:+n) Q[n,m] = fctA(Q[n-1,:], y[n-1]);
    // step2 - y[n] depends on Q[n,:] and bond_price[n+1]
    y[n] = fctB(Q[n,:], bond_price[n+1]);
}
```



#### Pricing with the tree

Pricing of IRD is done by marking possible coupons on appropriate nodes of the tree and discount:

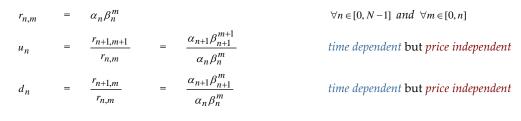
$$f(n,m) = \sum_{n,m} coupon_{n,m} Q_{n,m}$$
 or 
$$f(n,m) = coupon_{n,m} + (p_{up}(m)f(n+1,m+1) + p_{mid}(m)f(n+1,m) + p_{down}(m)f(n+1,m-1)) \times e^{-(m\Delta x + y_{n\Delta t})\Delta t}$$

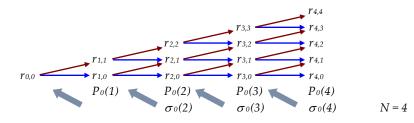
fctA/fctB/fctC are quite intuitive, we can jump to the answer directly, like a fastlane.

# 6. Black Derman Toy (BDT) tree

BDT tree is a binary recombining tree modelling short rate:

- up-prob equals to down-prob (which is ½)
- up-scale and down-scale are time dependent but price independent
- root vertex in layer 0 representing current time 0
- n+1 vertices in layer n representing forward time  $n\Delta t$
- vertex (*N*, :) in the last layer do not bear any short rate
- vertex (n,m) denotes short rate  $r_{n,m}$  within  $[n\Delta t, (n+1)\Delta t]$





bond price bond price volatility

#### Calibration

There are 2N-1 parameters  $(\alpha_0)(\alpha_1\beta_1)(\alpha_2\beta_2)...(\alpha_{N-1}\beta_{N-1})$  calibration needs 2N-1 data:

- *N* data from bond price term structure :  $P_0(1)$   $P_0(2)$  ...  $P_0(N)$
- *N*-1 data from bond price volatility term structure :  $\sigma_0(2) \sigma_0(3) \dots \sigma_0(N)$
- calibration is done by forward propagation, which repeatedly calls Newton Raphson
- the *nth* Newton Raphson solves two nonlinear equations for parameters  $\alpha_{n-1}$  and  $\beta_{n-1}$
- one equation is constructed from *bond price* with maturity  $n\Delta t$  and
- one equation is constructed from *bond price volatility* with maturity  $n\Delta t$
- for n=1,  $\beta_0$  is redundant, *bond price volatility* with maturity  $1\Delta t$  is not needed
- lets derive the system of nonlinear equation ...

# How to read bond price from a tree?

Fill the payoff with all one, then propagate the value using the similar (but simplier) equation as in HW tree :

$$f(n,m) = coupon_{n,m} + (p_{up}(m)f(n+1,m+1) + p_{mid}(m)f(n+1,m) + p_{down}(m)f(n+1,m-1)) \times e^{-(m\Delta x + y_{n\Delta t})\Delta t}$$
 from HW 
$$P_{n,m} = \left(\frac{1}{2}P_{n+1,m+1} + \frac{1}{2}P_{n+1,m}\right) \times \exp(-\alpha_n \beta_n^m \Delta t)$$
 for BDT

How to read bond price volatility from a tree?

$$\begin{split} dP_t(T) &= \mu_t(T)P_t(T)dt + \sigma_t(T)P_t(T)dz_t \\ d\ln P_t(T) &= (\mu_t(T) - \sigma_t^2(T)/2)dt + \sigma_t(T)dz_t \\ \sigma_t^2(T)\Delta t &= V[\Delta \ln P_t(T)] \\ &= V[\ln P_{t+\Delta t}(T) - \ln P_t(T)] \\ &= V\left[\ln \frac{P_{t+\Delta t}(T)}{P_t(T)}\right] \end{split}$$

Now we have reference day t = 0, hence we have bond price volatility in terms of *current bond price* and *forward bond price*:

$$\begin{split} \sigma_0(T) &= \sqrt{\frac{1}{\Delta t}} V \bigg[ \ln \frac{P_{\Delta t}(T)}{P_0(T)} \bigg] \\ \sigma_0(n\Delta t) &= \sqrt{\frac{1}{\Delta t}} V \bigg[ \ln \frac{P_{\Delta t}(n\Delta t)}{P_0(n\Delta t)} \bigg]^2 - E^2 \bigg[ \ln \frac{P_{\Delta t}(n\Delta t)}{P_0(n\Delta t)} \bigg] \\ &= \frac{1}{\sqrt{\Delta t}} \sqrt{E \bigg[ \ln \frac{P_{\Delta t}(n\Delta t)}{P_0(n\Delta t)} \bigg]^2 - E^2 \bigg[ \ln \frac{P_{\Delta t}(n\Delta t)}{P_0(n\Delta t)} \bigg]} \\ &= \frac{1}{\sqrt{\Delta t}} \sqrt{\frac{1}{2}} (\ln \frac{P_{1,1}}{P_0})^2 + \frac{1}{2} (\ln \frac{P_{1,0}}{P_0})^2 - \left( \frac{1}{2} \ln \frac{P_{1,1}}{P_0} + \frac{1}{2} \ln \frac{P_{1,0}}{P_0} \right)^2 \\ &= \frac{1}{\sqrt{\Delta t}} \sqrt{\frac{1}{2}} (\ln \frac{P_{1,1}}{P_0})^2 + \frac{1}{2} (\ln \frac{P_{1,0}}{P_0})^2 - \frac{1}{4} (\ln \frac{P_{1,0}}{P_0})^2 - \frac{1}{2} (\ln \frac{P_{1,1}}{P_0}) (\ln \frac{P_{1,0}}{P_0}) \\ &= \frac{1}{2\sqrt{\Delta t}} \sqrt{(\ln \frac{P_{1,1}}{P_0})^2 + (\ln \frac{P_{1,0}}{P_0})^2 - 2(\ln \frac{P_{1,1}}{P_0}) (\ln \frac{P_{1,0}}{P_0})} \\ &= \frac{1}{2\sqrt{\Delta t}} \sqrt{(\ln \frac{P_{1,1}}{P_0} - \ln \frac{P_{1,0}}{P_0})^2} \\ &= \frac{1}{2\sqrt{\Delta t}} \ln \frac{P_{1,1}}{P_{1,0}} \end{split}$$

# <u>Algorithm</u>

Suppose both is in the middle of forward propagation, the following are known:

```
bdt.a[0] ... bdt.a[n-2]
bdt.b[0] ... bdt.b[n-2]
```

Now given  $P_0(T)$  and  $\sigma_0(T)$  such that :

=  $market bond price of maturity n\Delta t$  $P_0(T)$ 

market bond price volatility of maturity  $n\Delta t$ 

we setup the system of two nonlinear equation solve for bdt.a[n-1] and bdt.b[n-1]:

```
P_0(T)
             bdt.node(0)
\sigma_0(T)
              ln(bdt.node(1,1)/ bdt.node(1,0))/(2*sqrt(bdt.dt))
```

#### Short summary

Comparison of common trees

binomial CRR tree 1

2 binomial 1/2 prob tree

trinomial HW tree 3

BDT tree 4

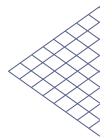
two degree of freedom: u and p two degree of freedom: u and d two degree of freedom:  $p_u$  and  $p_d$ 

two degree of freedom:

constraining d = 1/uconstraining p = 1/2 $constraining \ p_m = 1 - p_u - p_d$ 

multiplication topology multiplication topology addition topology multiplication topology

multiplication topology



addition topology

