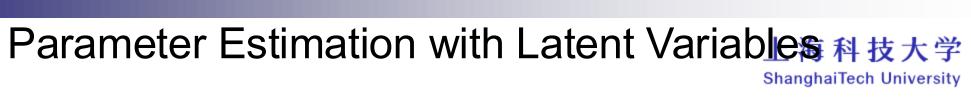


CS182: Introduction to Machine Learning –Expectation-Maximization (EM) algorithm and Gaussian Mixture Models (GMM)

Yujiao Shi SIST, ShanghaiTech Spring, 2025





- Consider a generative model with joint distr. $p(\mathbf{X}, \mathbf{Z}|\Theta) = \prod_{n=1}^{N} p(\mathbf{x}_n, \mathbf{z}_n)$
 - Observed data: $\mathbf{X} = \{x_n\}_{n=1}^N$
 - Latent variables: $\mathbf{Z} = \{z_n\}_{n=1}^N$. All the model parameters: Θ

Parameter Estimation with Latent Variables 科技大学 ShanghaiTech University



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 - Observed data: $\mathbf{X} = \{x_n\}_{n=1}^N$
 - Latent variables: $\mathbf{Z} = \{z_n\}_{n=1}^N$. All the model parameters: Θ
- Goal: Estimate the model parameters Θ via MLE (or MAP)

$$\hat{\Theta} = \arg\max_{\Theta} \log p(\mathbf{X}|\Theta) = \arg\max_{\Theta} \log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\Theta) \quad \text{(when } \mathbf{Z} \text{ is discrete)}$$

$$= \arg\max_{\Theta} \log \int_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\Theta) d\mathbf{Z} \quad \text{(when } \mathbf{Z} \text{ is continuous)}$$

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• Thus $\log p(\mathbf{X}|\Theta)$ is tightly lower-bounded by $\mathbb{E}[\log p(\mathbf{X}, \mathbf{Z}|\Theta)]$ which EM maximizes

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$$\Theta^{new} = \arg \max_{\Theta} \mathcal{Q}(\Theta, \Theta^{old}) \text{ (if doing MLE)}$$

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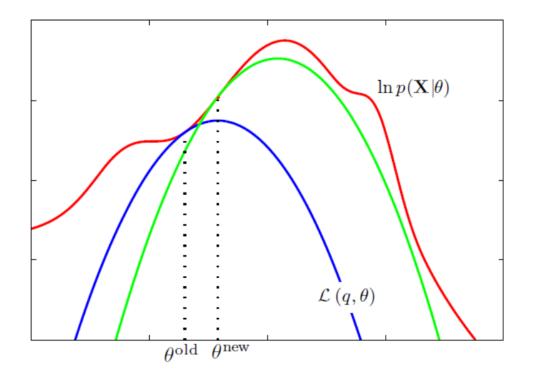
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EM: A View in the Parameter Space 上海科技大学

- 上海科技大学 ShanghaiTech University
- E-step: Update of q makes the $\mathcal{L}(q,\Theta)$ curve touch the $\log p(\mathbf{X}|\Theta)$ curve
- M-step gives the maxima Θ^{new} of $\mathcal{L}(q,\Theta)$
- Next E-step readjusts $\mathcal{L}(q,\Theta)$ curve (green) to meet $\log p(\mathbf{X}|\Theta)$ curve again
- This continues until a local maxima of $\log p(\mathbf{X}|\Theta)$ is reached





EM: Some Comments



- A general framework for parameter estimation in latent variable models
- Very widely used in problems with "missing data", e.g., missing features, or missing labels (semi-supervised learning)
 - \bullet "Missing" parts can be treated as latent variables z and estimated using EM
- More advanced probabilistic inference algorithms are based on similar ideas
 - E.g., variational Bayesian inference
- Very easy to extend to online learning setting and gives SGD like algorithms (will post a reading on "Online EM" on the class webpage)
- Note: The E and M steps may not always be possible to perform exactly (approximate inference methods may be needed in such cases)

Recap: GMM



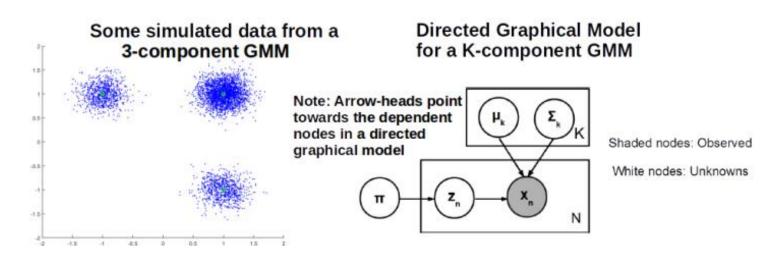
- The generative story for each x_n , n = 1, 2, ..., N
 - First choose one of the K mixture components as

$$z_n \sim \text{Multinomial}(z_n|\pi)$$
 (from the prior $p(z)$ over z)

• Suppose $z_n = k$. Now generate x_n from the k-th Gaussian as

$$\mathbf{x}_n \sim \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

(from the data distr. p(x|z))







We derive the Expectation-Maximization (EM) algorithm for GMM with K components:

- Observed data $X = \{x_1, ..., x_N\}$
- Latent variables $Z = \{z_1, ..., z_N\}$, where $z_i \in \{1, ..., K\}$.
- The parameters are $\Theta = {\{\pi_k, \mu_k, \Sigma_k\}_{k=1}^K}$.

1. Complete-Data Likelihood

The joint distribution of observed data and latent data is:

$$p(X,Z|\Theta) = \prod_{i=1}^{N} p(x_i, z_i) p(z_i|\Theta) = \prod_{i=1}^{N} \pi_{z_i} \mathcal{N}(x_i | \mu_{z_i}, \Sigma_{z_i})$$

The complete-data log-likelihood is:

$$\log p(X, Z|\Theta) = \sum_{i=1}^{N} \log \pi_{z_i} + \log \mathcal{N}\left(x_i \middle| \mu_{z_i}, \Sigma_{z_i}\right)$$



2. E-Step

 \square Compute the posterior responsibility $\gamma_{ik} = p(z_i = k | x_i, \Theta^{old})$:

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$$= \frac{p(x_i | z_i = k, \Theta^{old}) p(z_i = k | \Theta^{old})}{p(x_i | \Theta^{old})}$$

$$= \frac{\pi_k^{old} \mathcal{N}(x_i | \mu_k^{old}, \Sigma_k^{old})}{\sum_{j=1}^K \pi_j^{old} \mathcal{N}(x_i | \mu_j^{old}, \Sigma_j^{old})}$$

□ This is a soft assignment of x_i to cluster k.



- 3. M-step: Maximize $Q(\boldsymbol{\Theta}, \boldsymbol{\Theta}^{old})$
 - ☐ The *Q*-function is the expected complete-data log-likeligood

$$Q(\Theta, \Theta^{old}) = \mathbb{E}_{Z|X,\Theta^{old}}[\log p(X, Z|\Theta)] = \sum_{i=1}^{N} \sum_{k=1}^{K} \gamma_{ik}[\log \pi_k + \log \mathcal{N}(x_i|\mu_k, \Sigma_k)]$$

- \square Update Mixing coefficients π_k :
 - Maximize Q w.r.t. π_k under the constraint $\sum_k \pi_k = 1$:

$$\mathcal{L} = Q(\Theta, \Theta^{old}) + \lambda \left(1 - \sum_{k=1}^{K} \pi_k\right)$$

$$\frac{\partial \mathcal{L}}{\partial \pi_k} = \sum_{i=1}^N \frac{\gamma_{ik}}{\pi_k} - \lambda = 0 \implies \pi_k \propto \sum_{i=1}^N \gamma_{ik}$$



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 \Box Update Mixing coefficients π_k :

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• Enforce constraint $\sum_k \pi_k = 1$:

$$\lambda = \sum_{k=1}^{K} \sum_{i=1}^{N} \gamma_{ik} = N \quad \Rightarrow \quad \pi_k^{new} = \frac{1}{N} \sum_{i=1}^{N} \gamma_{ik}$$



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 \square The probability density function for a D -dimensional Gaussian:

$$\mathcal{N}(x \mid \mu_k, \Sigma_k) = \frac{1}{(2\pi)^{D/2} |\Sigma_k|^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \Sigma_k^{-1} (x - \mu_k)\right)$$

□ Update Means μ_k :

$$\frac{\partial Q}{\partial \mu_k} = \sum_{i=1}^N \gamma_{ik} \Sigma_k^{-1} (x_i - \mu_k) = 0 \implies \mu_k^{new} = \frac{\sum_{i=1}^N \gamma_{ik} x_i}{\sum_{i=1}^N \gamma_{ik}}$$



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$$\frac{\partial Q}{\partial \Sigma_k^{-1}} = \frac{1}{2} \sum_{i=1}^N \gamma_{ik} [\Sigma_k - (x_i - \mu_k)(x_i - \mu_k)^{\mathsf{T}}] = 0$$

$$\Sigma_k^{new} = \frac{\sum_{i=1}^N \gamma_{ik} (x_i - \mu_k^{new}) (x_i - \mu_k^{new})^{\mathsf{T}}}{\sum_{i=1}^N \gamma_{ik}}$$

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- 1. Initialize the Parameters $\Theta = \{\pi_k, \mu_k, \Sigma_k\}_{k=1}^K$ randomly, or using K-means
- 2. Iterate until convergence (e.g., when $\log p(X|\Theta)$ ceases to increase
 - a) E-step:

$$\gamma_{ik} = p(z_i = k | x_i, \Theta^{old}) = \frac{\pi_k^{old} \mathcal{N}(x_i | \mu_k^{old}, \Sigma_k^{old})}{\sum_{j=1}^K \pi_j^{old} \mathcal{N}(x_i | \mu_j^{old}, \Sigma_j^{old})}$$

M-step:

$$\pi_k^{new} = \frac{1}{N} \sum_{i=1}^N \gamma_{ik}$$

$$\mu_k^{new} = \frac{\sum_{i=1}^N \gamma_{ik} x_i}{\sum_{i=1}^N \gamma_{ik}}$$

$$\Sigma_k^{new} = \frac{\sum_{i=1}^N \gamma_{ik} (x_i - \mu_k^{new}) (x_i - \mu_k^{new})^\top}{\sum_{i=1}^N \gamma_{ik}}$$