

## Assignment 8

### Partial solutions for selected problems

1. (a) This follows from  $\{u < U \leq V \leq v\} = \{u < X \leq v\} \cap \{u < Y \leq v\}$ .  
 (b)

$$F_{U,V}(u, v) = \begin{cases} F(u)(2F(v) - F(u)) & \text{if } u \leq v, \\ (F(v))^2 & \text{if } v < u. \end{cases}$$

It follows that  $F_U(u) = F(u)(2 - F(u))$  and  $F_V(v) = (F(v))^2$ .

- (c)  $f_{U,V}(u, v) = 2f(u)f(v)1_{u < v}$ . [This *does* integrate to 1.]  $f_U(u) = 2f(u)(1 - F(u))$  and  $f_V(v) = 2f(v)F(v)$ .
2. First, the possible values of  $(X_1, X_2)$  are the whole real plane. Next, compute  $\frac{dx dy}{dr d\theta} = r = \sqrt{x_1^2 + x_2^2}$ . Thus,

$$f_{X_1, X_2}(x_1, x_2) = \frac{1}{r} f_{R, \theta}(r, \theta) = \dots = \frac{e^{-\sqrt{x_1^2 + x_2^2}/2}}{4\pi\sqrt{x_1^2 + x_2^2}}.$$

Clearly,  $X_1$  and  $X_2$  are not independent.

[This actually was not the problem I intended. Instead it should have been  $(X_1, X_2) = (\sqrt{R}\sin(\Theta), \sqrt{R}\cos(\Theta))$  which are independent standard normal rvs.]

3. (a) Use what is known about gamma random variables:  $E(S) = 2$ ,  $\text{Var}(S) = 1$ .  
 (b) (Recall that the conditional distribution for  $R$ , given  $S = s$ , is uniform(0, s).)  $E(R|S) = \frac{1}{2}S$  and  $E(R^2|S) = \frac{1}{3}S^2$ . Thus,  $\text{Var}(R|S) = \frac{1}{12}S^2$ . Additionally,

$$E(R) = E(E(R|S)) = E(\frac{1}{2}S) = 1, \quad \text{Var}(R) = \text{Var}(\frac{1}{2}S) + E(\frac{1}{12}S^2) = \frac{2}{3}.$$

(c,d)

$$E(RS) = E(SE(R|S)) = E(\frac{1}{2}S^2) = \frac{5}{2},$$

and thus  $\text{Cov}(R, S) = \frac{1}{2}$ .

4. First, using the given property for  $W$ , find

$$E(W^2) = \lambda + \lambda^2 \quad \text{and} \quad E(W^4) = \lambda + 7\lambda^2 + 6\lambda^3 + \lambda^4.$$

Then  $E(X) = E(E(X|W)) = E(W^2) = \lambda + \lambda^2$ . Also,

$$E(X^2) = E(E(X^2|W)) = E(2(W^2)^2) = 2\lambda + 14\lambda^2 + 12\lambda^3 + 2\lambda^4.$$

Then  $\text{Var}(X) = E(X^2) - (E(X))^2 = 2\lambda + 13\lambda^2 + 10\lambda^3 + \lambda^4$ . Alternatively, use the variance partition formula.

5. (a)  $f_{Y|X}(y|x) = \frac{x+2y}{x+1}$ ,  $g(x) = E(Y|X = x) = \frac{3x+4}{6(x+1)}$ .  $f_{X|Y}(x|y) = \frac{x+2y}{1+2y}$ ,  $h(y) = E(X|Y = y) = \frac{4(2+3y)}{3(1+2y)}$ .  $g(x)$  and  $h(y)$  are not inverses of each other. [Nor could they be – they never are unless  $X$  and  $Y$  have a perfect 1-1 relationship.]  
 (c,d)  $E(X) = \frac{7}{6}$ ,  $E(Y) = \frac{7}{12}$ ,  $E(XY) = \frac{2}{3}$ ,  $\text{Cov}(X, Y) = -\frac{1}{72}$ .

6.  $\text{Var}(S) = \text{Var}(T) = \frac{7}{4}$ ,  $\text{Cov}(S, T) = -\frac{1}{4}$ ,  $\text{Corr}(S, T) = -\frac{1}{7}$ .

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7. (a)  $X$  and  $Y$  are not independent. They have the same pmf  $f(x) = \frac{5-2|x|}{13}$  for  $x \in \{-2, -1, 0, 1, 2\}$ .  
(b)  $\text{Cov}(X, Y) = 0$ .
8. (a) Apply the theorem to the independent random variables  $e^{tX_1}, \dots, e^{tX_1}$ .  
(b) Use the binomial( $n, p$ ) mgf  $(1 - p + pe^t)^n$ , the gamma( $\alpha, \gamma$ ) mgf  $(1 - \gamma t)^{-\alpha}$  and the Poisson( $\lambda$ ) mgf  $e^{\lambda(e^t - 1)}$ .