Complete Pass-Through in Levels

Kunal Sangani

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Disclaimer

This presentation contains my own analyses calculated (or derived) based in part on data from Nielsen Consumer LLC and marketing databases provided through the NielsenIQ Datasets at the Kilts Center for Marketing Data Center at The University of Chicago Booth School of Business. The conclusions drawn from the NielsenIQ data are those of the author and do not reflect the views of NielsenIQ. NielsenIQ is not responsible for, had no role in, and was not involved in analyzing and preparing the results reported herein.

Pass-Through in Logs and Levels

- Incomplete long-run pass-through of input cost changes.
 - E.g., Peltzman (2000), Kim and Cotterill (2008), Nakamura and Zerom (2010), Hong and Li (2017).
 - When costs increase 10%, firms raise prices <10%.
 - Incomplete even after accounting for input cost share and at long horizons.
 - Prevailing explanation: "Superelasticity" of demand leads to markup adjustment.

Pass-Through in Logs and Levels

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 - When costs increase 10%, firms raise prices <10%.
 - Incomplete even after accounting for input cost share and at long horizons.
 - Prevailing explanation: "Superelasticity" of demand leads to markup adjustment.
- Today: Measure pass-through on a dollars-and-cents basis.
- Result: Firms exhibit complete pass-through in levels.
 - \$1/unit increase in cost leads firms to increase prices \$1/unit. Looks incomplete in logs.
 - Pass-through in levels explains variation in "log pass-through" across products/firms.
 - Pattern spans many markets (gas stations, food products, manufacturing industries).

Explaining Pass-Through in Levels

- ullet In workhorse macro models, pass-through in levels should equal the markup, $\mu>1$.
 - For CES demand. (E.g., Dixit and Stiglitz 1977, Melitz 2003.)
 - For richer **homothetic** demand systems, for cost shocks that affect all firms in industry. (E.g., Kimball 1995; Atkeson and Burstein 2008; HSA from Matsuyama and Ushchev 2022.)

Explaining Pass-Through in Levels

- In workhorse macro models, pass-through in levels should equal the markup, $\mu >$ 1.
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 - For richer **homothetic** demand systems, for cost shocks that affect all firms in industry. (E.g., Kimball 1995; Atkeson and Burstein 2008; HSA from Matsuyama and Ushchev 2022.)
- Discipline departure from homotheticity using discrete choice framework for demand.
 - Parameter η dictates intensive margin elasticity of demand.
 - Demand within industry is homothetic (e.g., CES) iff $\eta = 1$.
 - Result: Pass-through in levels of aggregate cost shocks whenever $\eta = 0$.
 - Identifies class of models that generate pass-through in levels.
 E.g., Heterogeneous coefficient logit (Berry, Levinsohn, and Pakes 1995); probit demand; spatial competition (Hotelling 1929; Salop 1979; Barro 2024).

Macro Implications

- Asymmetry, convexity, and heterogeneity in "log pass-through."
- Oifferential pass-through of unit labor costs relative to materials costs (Okun 1981).
- Response of industry gross margins, profits, and entry to upstream costs.
- Low volatility of consumer inflation relative to commodity costs.
- Fluctuations in inflation inequality over commodity cost cycle.
 - Explains "cheapflation" during recent inflation without price gouging / excess demand.
 - Absent this channel, inflation inequality from 2020–2023 would have been 1/3 as large.

Selected Related Literature

Theoretical and empirical determinants of pass-through:

- E.g., Bulow and Pfleiderer (1983); Nakamura and Zerom (2010); Weyl and Fabinger (2013); Hong and Li (2017); Minton and Wheaton (2022); Mongey and Waugh (2023); (Exchange rates) Campa and Goldberg (2005); Burstein et al. (2006); Burstein and Gopinath (2014); Fitzgerald and Haller (2014); Amiti et al. (2019).
- Abstract from (1) asymmetry in speed of adjustment (Borenstein et al. 1997; Peltzman 2000; Benzarti et al. 2020) and (2) firm-specific shocks (e.g., Amiti et al. 2019).

Studies that measure pass-through in levels (not exhaustive):

- Retail Gasoline: (Pass-through asymmetry) Karrenbrock (1991), Borenstein et al. (1997), Lewis (2011) (Cycles) Wang (2009), Noel (2009, 2015), Lewis and Noel (2011), Atkinson et al. (2014), Byrne and de Roos (2017, 2019).
- Food: (Coffee) Bettendorf and Verboven (2000), Leibtag et al. (2007), Nakamura and Zerom (2010), Bonnet et al. (2013), (Cheese) Kim and Cotterill (2008), (Spirits) Conlon and Rao (2020), (Cigarettes, Beer, Milk) Butters et al. (2022).

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Differential pass-through of labor costs
Industry profits, margins, and entry
Stability of consumer price index
Inflation inequality

| Cost per unit | Baseline |
|-----------------------------------|----------|
| Commodity | \$1 |
| Other components of marginal cost | \$1 |
| Total marginal cost | \$2 |
| Price | \$4 |

| Cost per unit | Baseline | |
|-----------------------------------|----------|---------|
| Commodity | \$1 | +\$0.20 |
| Other components of marginal cost | \$1 | |
| Total marginal cost | \$2 | +\$0.20 |
| Price | \$4 | |

| Baseline | | New | |
|----------|-------------------|-----------------------------------|--|
| \$1 | +\$0.20 | \$1.20 | |
| \$1 | | \$1.00 | |
| \$2 | +\$0.20 | \$2.20 | |
| \$4 | ? | ? | |
| | \$1 \$1 \$2 | \$1 +\$0.20 \$1 \$2 +\$0.20 | \$1 +\$0.20 \$1.20 \$1 \$1.00 \$2 +\$0.20 \$2.20 |

| Cost per unit | Baseline | | New | % Change |
|-----------------------------------|----------|---------|--------|----------|
| Commodity | \$1 | +\$0.20 | \$1.20 | +20% |
| Other components of marginal cost | \$1 | | \$1.00 | |
| Total marginal cost | \$2 | +\$0.20 | \$2.20 | +10% |
| Price | \$4 | ? | ? | |

• Leontief production in commodity (\$1/unit) and other variable costs (\$1/unit).

| Cost per unit | Baseline | | New | % Change |
|-----------------------------------|----------|---------|--------|----------|
| Commodity | \$1 | +\$0.20 | \$1.20 | +20% |
| Other components of marginal cost | \$1 | | \$1.00 | |
| Total marginal cost | \$2 | +\$0.20 | \$2.20 | +10% |
| Price | \$4 | +\$0.40 | \$4.40 | +10% |

• Complete pass-through in logs: $p = \mu(c + w) \Rightarrow \Delta p = \mu \cdot \Delta c$.

| Cost per unit | Baseline | | New | % Change |
|-----------------------------------|----------|---------|--------|----------|
| Commodity | \$1 | +\$0.20 | \$1.20 | +20% |
| Other components of marginal cost | \$1 | | \$1.00 | |
| Total marginal cost | \$2 | +\$0.20 | \$2.20 | +10% |
| Price | \$4 | +\$0.20 | \$4.20 | +5% |

- Complete pass-through in logs: $p = \mu(c + w) \Rightarrow \Delta p = \mu \cdot \Delta c$.
- Complete pass-through in levels: $p = (c + w) + m \Rightarrow \Delta p = \Delta c$. Appears incomplete in logs, even relative to cost share (0.25 < 0.5).

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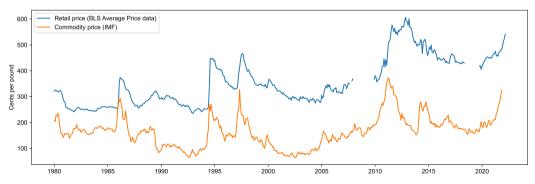
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Pass-through for six staple food products

- BLS tracks price levels for staple food products (Average Price Data program).
- Identify six products where we can match units and prices of commodity inputs.
 - Coffee, sugar, beef, rice, flour, frozen orange juice concentrate.

Figure: Arabica coffee commodity costs (IMF) and retail ground coffee prices (BLS).



Canonical approach to measure pass-through of cost changes

- Specification à la Campa and Goldberg (2005), Nakamura and Zerom (2010), etc.
- Price change at time *t* in market *m* due to commodity cost changes in last *K* periods:

$$\Delta
ho_{m,t} = a_m + \sum_{k=0}^K b_k \Delta c_{m,t-k} + \varepsilon_{m,t}.$$

Long-run pass-through is $\sum_{k=0}^{K} b_k$.

- Details:
 - For large *K*, long-run pass-through estimates gross markup if *c* is unit root.
 - All commodity series used appear unit root (autocorrelation coefficients 0.967–0.997).
 - Ensure one way Granger causality from Δc to Δp .
 - Use K = 8 weeks for gasoline, K = 12 months for all other products.

Example: Pass-through of coffee commodity costs to CPI

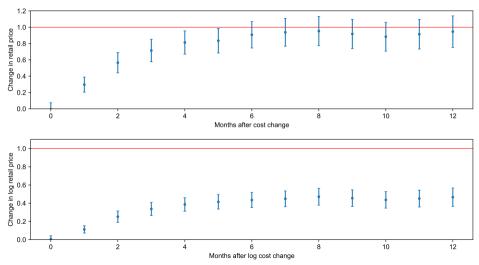


Figure: Passthrough in levels (top) and in logs (bottom).

Pass-through for six food products matched to commodities

| Commodity (IMF) | Final Good (BLS) |
|-------------------------------------|-----------------------------------|
| Arabica coffee price, per lb. | Coffee, 100%, ground roast |
| Sugar, No. 16, per lb. | Sugar, white, per lb. |
| Beef, global price, per lb. | Ground beef, 100% beef |
| Rice, Thailand, per metric ton | Rice, white, long grain, uncooked |
| Wheat, global price, per metric ton | Flour, white, all purpose |
| Frozen orange juice solids, per lb. | Orange juice, frozen concentrate |

- Monthly commodity prices (IMF), retail prices (BLS Average Price Data), 1990–2020.
- Match units (e.g., lbs flour per bushel of wheat) using USDA conversion tables.

Pass-through for six food products matched to commodities

| | | Р | ass-throug | gh (12 m | os.) |
|-------------------------------------|-----------------------------------|-------|------------|----------|---------|
| Commodity (IMF) | Final Good (BLS) | L | ogs | Le | vels |
| Arabica coffee price, per lb. | Coffee, 100%, ground roast | 0.466 | (0.051) | 0.946 | (0.099) |
| Sugar, No. 16, per lb. | Sugar, white, per lb. | 0.370 | (0.035) | 0.691 | (0.072) |
| Beef, global price, per lb. | Ground beef, 100% beef | 0.410 | (0.068) | 0.899 | (0.126) |
| Rice, Thailand, per metric ton | Rice, white, long grain, uncooked | 0.307 | (0.049) | 0.882 | (0.169) |
| Wheat, global price, per metric ton | Flour, white, all purpose | 0.240 | (0.048) | 0.865 | (0.160) |
| Frozen orange juice solids, per lb. | Orange juice, frozen concentrate | 0.327 | (0.040) | 0.974 | (0.111) |

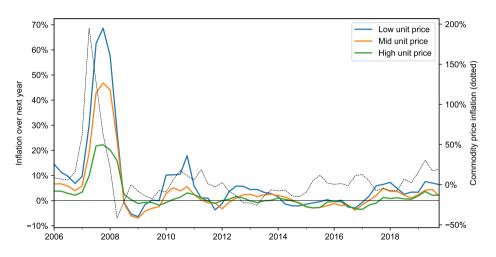
- Monthly commodity prices (IMF), retail prices (BLS Average Price Data), 1990–2020.
- Match units (e.g., lbs flour per bushel of wheat) using USDA conversion tables.
- Cannot reject complete pass-through in levels for 5 of 6. (Reject in logs for all.)

Further exploiting cross-sectional variation

- We see complete pass-through in levels across seemingly different products.
- Exploit UPC-level data from NielsenIQ for three food products (flour, rice, coffee).
 - Low-priced products have systematically lower "log pass-through."
 - Explained by similar pass-through in levels across all products.
- Exploit variation across retailers for broader set of products:
 - Retailers selling same UPC at lower price systematically have lower "log pass-through."
 - Explained by similar pass-through in levels across all retailers.
 - Holds for universe of food products in NielsenIQ data (>\$2T in sales, 235 retail chains).

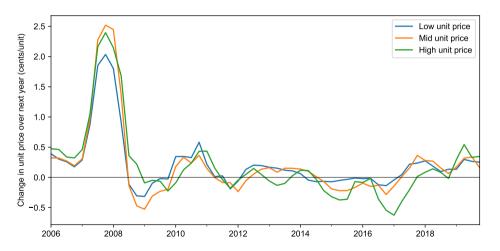
High-price products within category lower "log pass-through"

Figure: Inflation of Rice products in Nielsen data, split by tercile of unit price.



Differences in pass-through disappear in absolute terms (levels)

Figure: Change in unit price of Rice products in Nielsen data, split by tercile of unit price.



All categories: High-price items have systematically lower "log pass-through"

$$\Delta \log p_{it} = lpha_i + eta_1 \Delta \log c_t + \sum_{g=2}^3 eta_g (\mathbb{1}\{G(i,t) = g\} imes \Delta \log c_t) + arepsilon_{it}.$$

| Panel A: In percentages | | | |
|--|----------|-----------------|----------|
| | Ret | ail price infla | tion |
| | Rice | Flour | Coffee |
| Commodity Inflation × Mid Unit Price | -0.075** | -0.007 | -0.064** |
| | (0.014) | (0.009) | (0.015) |
| Commodity Inflation \times High Unit Price | -0.150** | -0.045** | -0.091** |
| | (0.022) | (0.009) | (0.017) |
| UPC FEs | Yes | Yes | Yes |
| N (thousands) | 399.4 | 101.4 | 1570.0 |
| R ² | 0.15 | 0.05 | 0.14 |

All categories: No systematic differences in pass-through in levels

$$\Delta p_{it} = \alpha_i + \beta_1 \Delta c_t + \sum_{g=2}^3 \beta_g \left(1\{G(i,t) = g\} \times \Delta c_t \right) + \varepsilon_{it}.$$

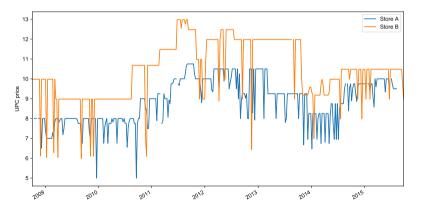
| Panel B: In levels | | | |
|--|---------|--------------|---------|
| | | ∆ Retail pri | ce |
| | Rice | Flour | Coffee |
| Δ Commodity Price $	imes$ Mid Unit Price | 0.059 | 0.027 | -0.069 |
| | (0.052) | (0.040) | (0.046) |
| Δ Commodity Price $	imes$ High Unit Price | 0.042 | -0.067 | -0.099* |
| | (0.100) | (0.044) | (0.058) |
| UPC FEs | Yes | Yes | Yes |
| N (thousands) | 399.4 | 101.4 | 1570.0 |
| R^2 | 0.07 | 0.05 | 0.14 |

• No systematic difference in pass-through in levels across unit price groups.

Exploiting variation in markups across retailers

- Consider two retailers selling a UPC at low vs. high markup.
- Test: When UPC cost rises, high-markup retailer should increase price more in levels.

Figure: Prices of identical coffee UPC in two stores in same 3-digit ZIP code in Philadelphia, PA.



Exploiting variation in markups across retailers: Predictions

$$\Delta p_{ikt} = \beta^{\text{level}} (\Delta \bar{p}_{kt} \times \text{RelativePrice}_{ikt}) + \delta \text{RelativePrice}_{ikt} + \alpha_{kt} + \varepsilon_{ikt}, \tag{1}$$

$$\Delta \log p_{ikt} = \beta^{\log} \left(\Delta \log \bar{p}_{kt} \times \text{RelativePrice}_{ikt} \right) + \tilde{\delta} \text{RelativePrice}_{ikt} + \tilde{\alpha}_{kt} + \varepsilon_{ikt}, \tag{2}$$

- Δp_{ikt} is change in price of UPC k at retailer i from t to t+4.
- $\Delta \bar{p}_{kt}$ is average change in UPC k's price across all retailers.
- RelativePrice_{ikt} = $log(p_{ikt}/\bar{p}_{kt})$ is retailer's price relative to others.

| | Fixed percentage markup $p_i = \mu_i c$ | Fixed unit margin $p_i = c + m_i$ |
|--|---|-----------------------------------|
| Price change relative to average: Levels $(dp_i/d\bar{p})$ Logs $(d\log p_i/d\log \bar{p})$ | $pprox 1 + \log(p_i/ar{p})$ | $1 pprox 1 - \log(p_i/ar{p})$ |
| Predicted interaction coefficient: Levels specification (1), β^{level} Logs specification (2), β^{log} | 1 0 | 0 -1 |

Exploiting variation in markups across retailers

| | | Δ UPC Pi | rice (Δp_{irt}) | |
|---|---------|----------|---------------------------|---------|
| | Rice | Flour | Coffee | All |
| | (1) | (2) | (3) | (4) |
| Δ Price _{it} \times Markup _{irt} | -0.022 | 0.058 | 0.192 | -0.001 |
| , | (0.141) | (0.209) | (0.170) | (0.839) |
| | | | | |
| UPC-Quarter FEs | Yes | Yes | Yes | Yes |
| N (millions) | 0.399 | 0.101 | 1.570 | 100.4 |
| R^2 | 0.46 | 0.48 | 0.51 | 0.59 |

• Result: $\beta^{\text{level}} \approx 0 \Rightarrow$ retailers change UPC price by same amount in *levels*.

Exploiting variation in markups across retailers

| | Δ UPC Price (Δp_{irt}) | | | | Δ Log UPC Price ($\Delta \log p_{irt}$) | | | | |
|--|---|---------|---------|---------|--|----------|----------|----------|--|
| | Rice | Flour | Coffee | All | Rice | Flour | Coffee | All | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | |
| Δ Price _{it} $	imes$ Markup _{irt} | -0.022 | 0.058 | 0.192 | -0.001 | | | | | |
| | (0.141) | (0.209) | (0.170) | (0.839) | | | | | |
| Δ Log Price _{it} $	imes$ Markup _{irt} | | | | | -1.004** | -1.014** | -1.261** | -1.043** | |
| | | | | | (0.125) | (0.184) | (0.084) | (0.073) | |
| UPC-Quarter FEs | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | |
| N (millions) | 0.399 | 0.101 | 1.570 | 100.4 | 0.399 | 0.101 | 1.570 | 100.4 | |
| R^2 | 0.46 | 0.48 | 0.51 | 0.59 | 0.62 | 0.61 | 0.58 | 0.61 | |

- Result: $\beta^{\text{level}} \approx 0 \Rightarrow$ retailers change UPC price by same amount in *levels*.
- Makes "log pass-through" fall with markup. ($\beta^{\log} \approx -1$.)
- Holds for broader sample of nearly 1M UPCs across 235 retail chains!

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Manufacturing Industries

- NBER-CES Manufacturing Data: 459 four-digit SIC industries from 1958–2018.
 - Sales (shipments) and input costs: materials, energy, and production labor.
 - Price indices for output, materials, energy, labor from BLS PPI, BEA make-use tables.
- Challenge: Price indices, not price levels! But starting from pass-through in levels,

$$ho^{ ext{level}} \equiv rac{\Delta
ho}{\Delta c} \qquad \Rightarrow \qquad rac{\Delta
ho}{
ho} =
ho^{ ext{level}} \left(rac{c ext{y}}{
ho ext{y}}
ight) rac{\Delta c}{c}.$$

So we can estimate ρ^{level} using

$$\Delta \log p_{it} = \beta \left(\Delta \log c_{it} \times (\text{InputCosts/Sales})_{it-1}\right) + \varepsilon_{it}.$$

ullet Multiplicative markups imply $eta=\mu>$ 1, while pass-through in levels implies etapprox 1.

Manufacturing Industries: Pass-Through

| | Δ Log Output Price $_t$ | | | | | | | | |
|--|--------------------------------|---------|----------|---------|--------------------|---------|--|--|--|
| Inputs: | Materials | | + Energy | | + Production Labor | | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | | | |
| Δ Log Input Price $_t$ | 0.690** | 0.079 | 0.704** | 0.005 | 0.796** | 0.052 | | | |
| | (0.072) | (0.132) | (0.073) | (0.134) | (0.083) | (0.232) | | | |
| $(InputCost/Sales)_{t-1}$ | | 0.004 | | 0.008 | | 0.023** | | | |
| | | (0.011) | | (0.011) | | (0.011) | | | |
| Δ Log Input Price _t × (InputCost/Sales) _{t-1} | | 0.947** | | 1.041** | | 0.984** | | | |
| | | (0.203) | | (0.201) | | (0.286) | | | |
| Industry FEs | Yes | Yes | Yes | Yes | Yes | Yes | | | |
| Year FEs | Yes | Yes | Yes | Yes | Yes | Yes | | | |
| N | 27381 | 27381 | 27381 | 27381 | 27381 | 27381 | | | |
| R^2 | 0.40 | 0.42 | 0.40 | 0.42 | 0.41 | 0.42 | | | |

- Estimate $\beta \approx 1$ (complete pass-through in levels). Variation in InputCosts/Sales leads to incomplete log pass-through.
- Similar results with commodity price IV for input costs, horizons 1,...,5 years.

Taking Stock of the Empirical Evidence

- Retail gasoline. [in paper]
 - Pass-through in levels explains variation in "log pass-through" across stations.
- Food products.
 - Pass-through in levels across nearly all markets.
 - Explains variation in "log pass-through" across products within category.
 - Explains variation in "log pass-through" across retailers.
- Manufacturing industries.
 - Pass-through in levels across broad set of industries.
- Previous literature. [in paper]
 - Survey of papers that estimate pass-through in levels of aggregate cost shocks.
 - Weight of evidence points toward complete pass-through in levels of commodity/input/shipping costs, excise taxes.

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Homothetic demand systems violate complete pass-through in levels

Workhorse models in macroeconomics/trade use homothetic demand systems.

E.g., CES, Kimball (1995), HSA (Matsuyama and Ushchev 2022), Atkeson and Burstein (2008).

This class of models does not deliver pass-through in levels of aggregate cost shocks.

Homothetic demand systems violate complete pass-through in levels

- Workhorse models in macroeconomics/trade use homothetic demand systems.
 E.g., CES, Kimball (1995), HSA (Matsuyama and Ushchev 2022), Atkeson and Burstein (2008).
- This class of models does not deliver pass-through in levels of aggregate cost shocks.
- Firms $j \in \{1,...,J\}$, each with constant marginal cost c_i , set prices p_i to max profits.
- ullet Rep. household maximizes utility $U(q_1,...,q_J)$ s.t. budget constraint $\sum_j p_j q_j = I$.

Proposition (Pass-through of aggregate shock with homothetic preferences)

Suppose U is homothetic. If all firms' marginal costs increase by $d \log c_j = d \log c$, then $\forall j$,

$$d \log p_j = d \log c, \quad \Rightarrow \quad dp_j = \mu_j dc_j.$$

That is, "log pass-through" is complete and pass-through in levels $\mu_i > 1$.

Departing from homothetic demand

- Depart from homotheticity using discrete choice model of industry demand.
 Anderson, Palma, and Thisse (1992).
- Parameter η will allow us to discriminate between demand systems.
 - Industry demand is homothetic (e.g., CES) iff $\eta = 1$.
 - Models with unit demand (e.g., heterogeneous coefficient logit, Hotelling) have $\eta=0$.
- Equilibrium exhibits complete pass-through in levels whenever $\eta = 0$.

Discrete Choice Model of Industry Demand

- Mass M of consumers indexed by $i \in [0,1]$. Each consumer has income Y.
- Firms $j \in \{1,...,J\}$. Consumers choose one firm to purchase from.
- Quasilinear preferences over q_{ij} and consumption of numeraire good q_{i0} ,

$$U_i = \max_{j,q_{ij}} \{u_{ij}\}, \qquad ext{where} \qquad u_{ij} = rac{oldsymbol{\delta}_{ij}}{oldsymbol{q}_{ij}} + q_{i0},$$

subject to budget constraint, $p_j q_{ij} + p_0 q_{i0} \leq Y$.

- Tastes δ_{ij} drawn from arbitrary distribution without mass points.
- Impose $q_{ij} = \bar{q}(Y/p_j)^{\eta}$.
 - η is the intensive margin elasticity of consumption with respect to price.

Special Cases of Demand

- Fixed budget ($\eta = 1$):
 - Constant elasticity of substitution (e.g., Dixit and Stiglitz 1977):

$$\delta_{ij}=\beta_{j}\varepsilon_{ij},$$

where β_j are taste shifters and ε_{ij} are i.i.d. Frechet shocks.

- Unit demand ($\eta = 0$):
 - Heterogeneous coefficient logit demand (e.g., BLP 1995; Nevo 2001):

$$\delta_{ij} = \xi_{ij} + \varepsilon_{ij},$$

where ξ_{ii} are systematic tastes and ε_{ii} are i.i.d. Type 1 Extreme Value shocks.

• Hotelling (1929) model of spatial differentiation (e.g., Salop 1979, Barro 2024):

$$\delta_{ij}=t|i-j|.$$

Firm Pricing

- Firm j produces with constant marginal cost c_i .
- Demand curve for firm j is

$$D_j(p) = \int_0^1 1\{u_{ij} \geq u_{ik}, \forall k \neq j\} q_{ij} M di.$$

ullet Firms choose p_j to maximize variable profits, taking tastes and other prices as given,

$$\max_{p_j} (p_j - c_j) D_j(p).$$

Pass-Through in Logs and Levels

Proposition (Pass-Through in Logs)

Suppose all marginal costs increase $d \log c_j = d \log c$. If $\eta = 1$, then pass-through in logs

$$d\log p_j = d\log p = d\log c,$$

is an equilibrium solution.

- Intuition: With $\eta = 1$, consumer choice depends on comparison of δ_{ij}/p_j across all j.
- If we multiply all p_j by same proportion, no change to consumer choices.
- No change to elasticity of demand \Rightarrow no change to desired markup.

Pass-Through in Logs and Levels

Proposition (Pass-Through in Levels)

Suppose all marginal costs increase $dc_j = dc$. If $\eta = 0$, then pass-through in levels

$$dp_j = dp = dc,$$

is an equilibrium solution.

- Intuition: With $\eta=0$, consumer choice depends on comparison of $\delta_{ij}-p_j$ across all j.
- If we increase all p_j by dp, no change to consumer choices and desired margin.
- ullet Any model with $\eta=0$ generates pass-through in levels of aggregate cost shocks.

Example: Symmetric Firms

Proposition (Prices with symmetric firms)

If firms are symmetric ($c_j = c$ for all j and δ_{ij} are i.i.d. draws from F), prices solve

$$(p-c)\left((1-\eta)\frac{\phi_0}{\rho_0}+\eta\frac{\phi_1}{\rho}\right)=1,$$

where ϕ_0, ϕ_1 are constants that depend on the number of firms J and taste distribution F. If $\eta=1$, price equals cost c times a percentage markup,

$$p=\frac{\phi_1}{\phi_1-1}c.$$

If $\eta=0$, price equals cost c plus a margin (priced relative to the numeraire p_0),

$$p=c+rac{
ho_0}{\phi_0}.$$

Simple Example

Empirical Evidence
Food products
Manufacturing industries

Explaining Pass-Through in Levels

Implications

Asymmetry, convexity, and heterogeneity in "log pass-through"
Differential pass-through of labor costs
Industry profits, margins, and entry
Stability of consumer price index
Inflation inequality

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Misspecification in measuring "log pass-through"

 Previous work finds "log pass-through" is asymmetric, convex, and falls with firm size / product quality.

(Asymmetry: Peltzman 2000; Size: Berman et al. 2012; Amiti et al. 2019; Gupta 2020; Quality: Chen and Juvenal 2016; Auer et al. 2018).

• If $\eta=0$ and firms set prices $p=c+p_0/\phi_0$, to a second order

$$\hat{
ho}^{\log} = rac{\Delta \log
ho}{\Delta \log c} pprox rac{c}{
ho} \left(1 + rac{
ho_0/\phi_0}{
ho} \Delta \log c
ight).$$

- **Asymmetry**: Positive shocks appear to have higher pass-through: $\hat{\rho}^{\log}(\Delta \log c) > \hat{\rho}^{\log}(-\Delta \log c)$.
- Convexity: Large positive shocks appear to have higher pass-through than small shocks: $\hat{\rho}^{\log}(\Delta \log c_H) > \hat{\rho}^{\log}(\Delta \log c_L)$ if $\Delta \log c_H > \Delta \log c_L$.
- **Heterogeneity**: If markups increase with firm size / quality, $\hat{\rho}^{\log}$ falls with both.

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Okun (1981) on materials costs

- Okun (1981) speculates a "special role for materials costs."
 - "Some views of marking up direct costs distinguish increases in the costs of purchased
 materials from increases in standard unit labor costs, implying that the former are likely to
 be passed through to customers essentially on a dollars-and-cents basis, while the latter
 are passed through with a percentage markup."
- Okun (1981) was right! Labor costs appear to be passed-through at higher rate.
- But Okun (1981) was wrong.
 - Not material costs that are special, but labor that is special!
 - In GE, wages can determine numeraire price p_0 and thus additive margin.
 - Correct for this, and we restore equal pass-through in levels of labor vs. other costs.

Okun's "special role": Empirical evidence

| | Δ Log Output Price $_t$ | | | |
|---|--------------------------------|---------------------|---------------------|--|
| | (1) Cost Shares | (2) Sales Shares | (3) Sales Shares | |
| Δ Log Material Price $_t 	imes$ Material Share $_{t-1}$ | 0.798** (0.085) | 1.046** (0.108) | | |
| Δ Log Energy Price $_t$ × Energy Share $_{t-1}$ | 0.715** (0.271) | 0.859** (0.378) | | |
| Δ Log Production Wage $_t 	imes$ Labor Share $_{t-1}$ | 1.070** (0.217) | 2.095** (0.281) | | |
| Δ Log Production Wage $_{t}$ $	imes$ (Labor $+$ Margin) Share $_{t-1}$ | | | | |
| Industry FEs | Yes | Yes | | |
| Year FEs | Yes | Yes | | |
| N R ² | 27374 0.49 | 27374 0.50 | | |

ullet Labor costs appear to be passed through completely in percentages, > 1:1 in levels.

Okun's "special role": Empirical evidence

| | Δ Log Output Price $_t$ | | | |
|---|--------------------------------|-----------------------------|-----------------------------|--|
| | (1) Cost Shares | (2) Sales Shares | (3) Sales Shares | |
| Δ Log Material Price $_t$ $	imes$ Material Share $_{t-1}$ | 0.798** (0.085) | 1.046** (0.108) | 1.049** (0.109) | |
| Δ Log Energy Price $_{t}$ $	imes$ Energy Share $_{t-1}$ | 0.715** (0.271) | 0.859** (0.378) | 0.997** (0.362) | |
| Δ Log Production Wage $_t 	imes$ Labor Share $_{t-1}$ | 1.070** (0.217) | 2.095** (0.281) | | |
| Δ Log Production Wage $_t$ $	imes$ (Labor $+$ Margin) Share $_{t-1}$ | | | 0.951** (0.209) | |
| Industry FEs Year FEs N R ² | Yes Yes 27374 0.49 | Yes Yes 27374 0.50 | Yes Yes 27374 0.49 | |

- Labor costs appear to be passed through completely in percentages, > 1:1 in levels.
- Corrected when we assume that additive margin is priced relative to wage.

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Response of industry profits and entry to cost changes

- With percentage markups, variable profits per unit increase with upstream costs.
- If aggregate industry demand is relatively inelastic, either operating profits must increase or entry must increase.

• Puzzle: Neither appears in the data.

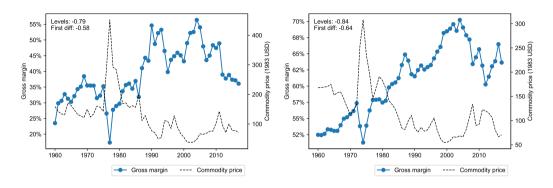
| Dep var: | Δ Log Operating Margin (1) | Δ Log Number of Establishments (2) |
|---------------------------------------|-----------------------------------|---|
| Δ Log Input Price _t | -0.122 (0.079) | 0.007 (0.013) |
| Industry FEs | Yes | Yes |
| Year FEs | Yes | Yes |
| N | 27305 | 18201 |
| R ² | 0.02 | 0.22 |

With pass-through in levels, gross margin instead adjusts

- Embed our demand system in standard monopolistic competition setup.
 - Mass N of symmetric firms, marginal cost c. Fixed entry cost f_e . Overhead cost f_o .
 - Aggregate industry demand inelastic, $Q = p^{-\theta}$, with $\theta < 1$.
 - Gross and operating margins = variable / operating profits as % of sales.
 - ζ is elasticity of entry to economic profits ($\zeta = 0$: fixed mass, $\zeta \to \infty$: free entry).

| Respor | se to cost inc | erease, dc > 0 | Gross margins dm ^{gross} | Operating margins dm ^{op} | Mass of firms $d \log N$ |
|-----------|---------------------------|----------------|--------------------------------------|---------------------------------------|--------------------------|
| $\eta=1$ | $\zeta=0$ | (Fixed mass) | 0 | > 0 | 0 |
| $\eta=$ 1 | $\zeta \in (0, \infty)$ | | 0 | > 0 | > 0 |
| $\eta=$ 1 | $\zeta ightarrow \infty$ | (Free entry) | 0 | 0 | > 0 |
| $\eta=0$ | $\zeta=0$ | (Fixed mass) | < 0 | \leq 0 | 0 |
| $\eta=$ 0 | $\zeta \in (0, \infty)$ | | < 0 | \leq 0 | ≤ 0 |
| $\eta=0$ | $\zeta ightarrow \infty$ | (Free entry) | < 0 | 0 | ≤ 0 |

Gross margins exhibit strong negative correlation with input costs



(b) Bread, cake & related products mfg vs. wheat.

(a) Roasted coffee mfg vs. coffee.

- ullet In contrast to $\eta=$ 1, gross margins for NBER-CES industries are not constant.
- Exhibit strong negative correlation with commodity input prices.

In data, adjustment along gross margins

| Panel A: Retail Gasoline Source: | Δ Log Gro ARTS (1) | oss Margin IRS (2) | Δ Log Op ARTS (3) | per. Margin IRS (4) | Δ <i>Log No</i> BDS (5) | um. Estabs SUSB (6) |
|--|--------------------------|--------------------------|--------------------------|---------------------------|-------------------------------|---------------------------|
| Δ Log Wholesale Price $_t$ | -0.263** | -0.291** | 0.218 | -0.167 | -0.002 | 0.001 |
| | (0.045) | (0.061) | (0.377) | (0.331) | (0.006) | (0.007) |
| N | 39 | 26 | 15 | 26 | 39 | 24 |
| R ² | 0.54 | 0.49 | 0.04 | 0.01 | 0.00 | 0.00 |
| Panel B: Manufacturing Industries | Δ Log Gro | oss Margin (2) | Δ Log O _l (3) | per. Margin (4) | Δ <i>Log No</i> (5) | um. Estabs (6) |
| Δ Log Input Price _t | -0.188** | 0.154 | -0.122 | 0.095 | 0.007 | -0.028 |
| | (0.039) | (0.103) | (0.079) | (0.211) | (0.013) | (0.044) |
| Δ Log Input Price $_t \times$ Inputs/Sales $_{t-1}$ | | -0.504** (0.188) | | -0.285 (0.376) | | 0.049 (0.063) |
| Industry FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Year FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| N | 27381 | 27381 | 27305 | 27305 | 18201 | 18201 |
| R ² | 0.05 | 0.11 | 0.02 | 0.06 | 0.22 | 0.23 |

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Stability of consumer price index

Inflation inequality

Low volatility of consumer inflation

ullet Volatility of PCE / CPI is too high with $\eta=1$ compared to the data.

| | Std. deviation of inflation, 1982–201 | Cost-weighted 8 avg. markup |
|--|---------------------------------------|--------------------------------|
| PCE Chain-Type Price Index (BEA) Consumer Price Index (BLS) | 1.1% 1.3% | |
| Model: $\eta=1$, Variable costs $=$ Materials | Flexible Calvo 6.0% 2.2% | 2.19 |

Simulation using BEA 402-industry I-O table, taking PPI for upstream commodities ("Stage 1 Intermediates") as exogenous.

Calvo model uses industry-level frequency of price adjustment from Pasten, Schoenle, and Weber (2020).

Low volatility of consumer inflation

- Volatility of PCE / CPI is too high with $\eta=1$ compared to the data.
- Reduce volatility by assuming more costs variable. But need low markups.

| | Std. deviation of inflation, 1982–2018 | | Cost-weighted avg. markup |
|--|--|-------|---------------------------|
| PCE Chain-Type Price Index (BEA) | 1.1% | | |
| Consumer Price Index (BLS) | 1.3% | | |
| Model: | Flexible | Calvo | |
| $\eta=$ 1, Variable costs $=$ Materials | 6.0% | 2.2% | 2.19 |
| $\eta=$ 1, Variable costs $=$ Materials $+$ wages | 2.9% | 1.6% | 1.31 |
| $\eta=$ 1, Variable costs $=$ Materials $+$ wages $+$ consumption of capital | 2.4% | 1.5% | 1.17 |

Simulation using BEA 402-industry I-O table, taking PPI for upstream commodities ("Stage 1 Intermediates") as exogenous.

Calvo model uses industry-level frequency of price adjustment from Pasten, Schoenle, and Weber (2020).

Low volatility of consumer inflation

- Volatility of PCE / CPI is too high with $\eta=1$ compared to the data.
- Reduce volatility by assuming more costs variable. But need low markups.
- Additive margins ($\eta = 0$) reconciles low volatility of agg. inflation with large markups.

| | Std. deviation of inflation, 1982–2018 | | Cost-weighted avg. markup |
|--|--|-------|---------------------------|
| PCE Chain-Type Price Index (BEA) | 1.1% | | |
| Consumer Price Index (BLS) | 1.3% | | |
| Model: | Flexible | Calvo | |
| $\eta=$ 1, Variable costs $=$ Materials | 6.0% | 2.2% | 2.19 |
| $\eta=$ 1, Variable costs $=$ Materials $+$ wages | 2.9% | 1.6% | 1.31 |
| $\eta=$ 1, Variable costs $=$ Materials $+$ wages $+$ consumption of capital | 2.4% | 1.5% | 1.17 |
| $\eta=$ 0, Additive margins | 1.9% | 1.3% | 1.0-2.2 |

Simulation using BEA 402-industry I-O table, taking PPI for upstream commodities ("Stage 1 Intermediates") as exogenous.

Calvo model uses industry-level frequency of price adjustment from Pasten, Schoenle, and Weber (2020).

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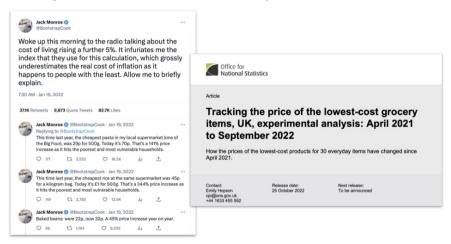
Explaining Pass-Through in Levels

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Asymmetry, convexity, and heterogeneity in "log pass-through"
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Inflation inequality

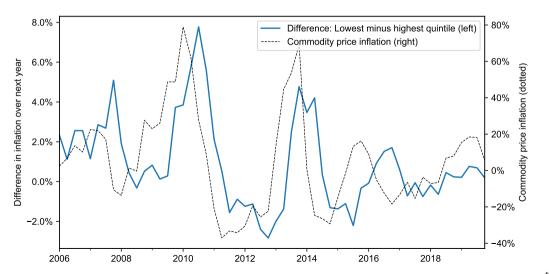
Attention to heightened inflation of low-end products in 2021



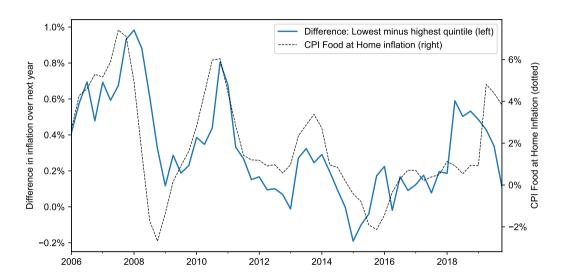
- "Supermarkets recouping margins on budget products" / elevated demand at low-end.
- Simpler explanation: Pass-through in levels → rising costs increase inflation inequality.

Inflation inequality moves with commodity costs within category (e.g. coffee)

• **Coffee**: Same Δp across products \rightarrow higher % inflation for low-price products.

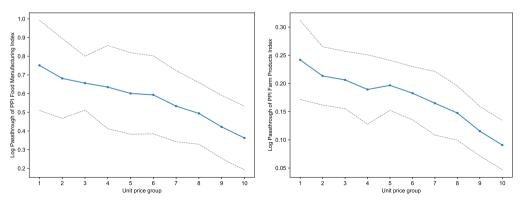


Cycles in inflation inequality over entire food-at-home bundle



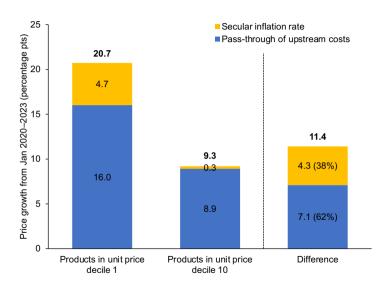
Can pass-through in levels explain recent "cheapflation"?

- Approach: split each product category into 10 groups by relative price.
- Fit pass-through of upstream costs to prices over 2006–2019, then extrapolate using 2020–2023 increase in costs.
- (a) Log pass-through of Food Manufacturing PPI.
- (b) Log pass-through of Farm Products PPI.



Predictions for food-at-home inflation, 2020–2023

- 20pp (11pp) price growth for high- (low-) price products.
- 3x higher than implied by secular inflation diffs. (Jaravel 2019).
- Magnitude similar to recent estimates.
 Cavallo and Kryvtsov (2024),
 Chen et al. (2024).



Conclusion

- Empirical evidence: Pass-through of costs tends to be complete in *levels*.
- Not consistent with homothetic models of demand used in workhorse models.
- Class of demand systems with $\eta = 0$ can explain pass-through behavior.
- Helps us understand dynamics of prices and industry outcomes:
 - Incomplete log pass-through + asymmetry, convexity, heterogeneity.
 - Differential pass-through of unit labor costs.
 - Dynamics of industry profits, margins, and entry.
 - Low volatility of aggregate inflation, despite substantial markups.
 - "Cheapflation" & unequal incidence of commodity inflation across income distribution.

Extra slides
Retail gasoline

Okun mode

Food-at-home inflation

Price change announcements

Inflation Expectations

Misspecification in "log pass-through": Positive squared $\Delta \log(c)$ term

| | Unleaded petrol (ULP) | | Premium (PULP) | |
|------------------------------|-----------------------|---------|----------------|---------|
| $\Delta \log(Price)_{it}$ | (1) (2) | | (3) | (4) |
| $\Delta \log(\text{Cost})_t$ | 0.870** | 0.889** | 0.865** | 0.881** |
| | (0.031) | (0.024) | (0.032) | (0.025) |
| $(\Delta \log(Cost)_t)^2$ | | 0.155** | | 0.147 |
| | | (0.068) | | (0.097) |
| N | 312215 | 312215 | 259437 | 259437 |
| R ² | 0.88 | 0.89 | 0.87 | 0.87 |

• Since cost share varies with commodity cost, misspecification in log regression.

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Inflation Expectations

Station-level panel dataset of gas prices in Perth, Australia

- 2.3M price observations (2001–present) for 875 stations in Perth metropolitan area.
- Spot price sold to retailers (Terminal Gate Price) available daily.
 - Pass-through is complete in levels.
 - No apparent heterogeneity in pass-through in levels.
 - Pass-through in levels + margins explains incomplete "log pass-through" and cross-sectional heterogeneity.

Pass-through of terminal gas price to station gas prices: Unleaded

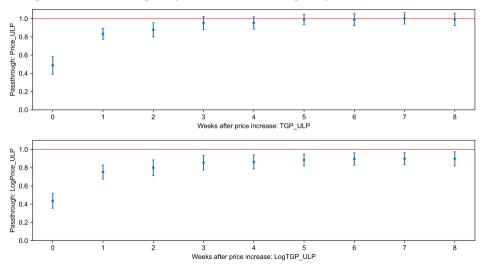


Figure: Passthrough in levels (top) and in logs (bottom). SEs two-way clustered by postcode \times year.

Summary of retail gasoline pass-through estimates

| | Pass-through (8 weeks) | | | |
|---|------------------------|--------------------|----------------|--------------------|
| Description | L | ogs | Le | vels |
| Australia, station-level, 2001–2022 Terminal to retail, Unleaded Terminal to retail, Premium Unleaded | 0.899 0.887 | (0.043) (0.041) | 0.991 0.985 | (0.038) (0.036) |
| Canada, city-level, 2007–2022 Crude to wholesale Wholesale to retail (excl. taxes) | 0.553 0.859 | (0.098) (0.016) | 0.927 1.008 | (0.100) (0.022) |
| South Korea, station-level, 2008–2022 Refinery to retail, Unleaded | 0.926 | (0.044) | 0.997 | (0.052) |
| United States, national, 1990–2022 NY Harbor spot price to retail | 0.570 | (0.051) | 0.954 | (0.053) |

• Cannot reject complete pass-through in levels. (Reject in logs for all.)

- Low markups, hard to differentiate pass-through in levels of 1 from 1.02–1.05.
- Test: Pass-through in levels should be higher for stations with 5% vs. 2% markup.

$$\Delta p_{it} = \alpha + \beta_1 \Delta c_{it} + \delta \text{AvgMarkup}_{it} + \beta_2 (\Delta c_{it} \times \text{AvgMarkup}_{it}) + \varepsilon_{it},$$

- where $\Delta p_{i,t}$, $\Delta c_{i,t}$ are change in station retail price and wholesale cost over 16 weeks.
- Exploit cross-sectional / time series variation in AvgMarkup_{it}, with IVs to isolate markups.
- Prediction: If constant multiplicative markup, $\beta_2 > 0$.

| Δ Price _{it} | (1) (OLS) | (2) (OLS) | (3) (IV1) | (4) (OLS) | (5) (IV2) |
|---|--------------|--------------|--------------|--------------|--------------|
| | . , | | | | |
| $\Delta Cost_t$ | 0.950** | | | | |
| | (0.021) | | | | |
| $\Delta Cost_t \times Avg$. Station Markup _i (N | let %) | | | | |
| | | | | | |
| $\Delta Cost_t \times Avg$. Quarter $Markup_t$ (N | Net %) | | | | |
| | | | | | |
| N | 312215 | | | | |
| | 012210 | | | | |

| (4) (OLS) | (5) (IV2) |
|--------------|--------------|
| | |
| | |
| | |
| | |
| | |
| | |

• Stations with higher markups do not have higher pass-through in levels ($\beta_2 \approx 0$).

| Δ Price $_{it}$ | (1) (OLS) | (2) (OLS) | (3) (IV1) | (4) (OLS) | (5) (IV2) |
|---|--------------------|--------------------|--------------------|--------------|--------------|
| $\Delta Cost_t$ | 0.950** (0.021) | 0.989** (0.037) | 0.952** (0.044) | | |
| $\Delta Cost_t \times Avg$. Station $Markup_i$ (Net %) | , , | -0.005 (0.003) | -0.000 (0.005) | | |
| $\Delta Cost_t \times Avg$. Quarter $Markup_t$ (Net %) | | . , | , , | | |
| N R ² | 312215 0.89 | 312215 0.89 | 312215 0.89 | | |

• Stations with higher markups do not have higher pass-through in levels ($\beta_2 \approx 0$).

| Δ Price _{it} | (1) | (2) | (3) | (4) | (5) |
|---|---------|-------------------|-------------------|-------------------|-------------------|
| | (OLS) | (OLS) | (IV1) | (OLS) | (IV2) |
| $\Delta Cost_t$ | 0.950** | 0.989** | 0.952** | 0.987** | 0.971** |
| | (0.021) | (0.037) | (0.044) | (0.034) | (0.043) |
| $\Delta Cost_t \times Avg$. Station $Markup_i$ (Net %) | , , | -0.005 (0.003) | -0.000 (0.005) | , , | , |
| $\Delta Cost_t \times Avg$. Quarter $Markup_t$ (Net %) | | | | -0.003 (0.003) | -0.002 (0.004) |
| N | 312215 | 312215 | 312215 | 312215 | 312215 |
| R ² | 0.89 | 0.89 | 0.89 | 0.89 | 0.89 |

• Stations with higher markups do not have higher pass-through in levels ($\beta_2 \approx 0$).

Pass-through in levels explains extent & variation of "log pass-through"

| $\Delta \log(Price)_{it}$ | (1) (OLS) | (2) (OLS) | (3) (IV1) | (4) (OLS) | (5) (IV2) |
|---|--------------------|--------------|--------------|--------------|--------------|
| $\Delta \log(Cost)_t$ $\Delta \log(Cost)_t 	imes Avg. Station Markup_i (Net \%)$ | 0.870** (0.031) | | | | |
| $\Delta \log(Cost)_t \times Avg$. Quarter $Markup_t$ (Net %) | | | | | |
| N | 312215 | | | | |
| R^2 | 0.88 | | | | |

Pass-through in levels explains extent & variation of "log pass-through"

| $\Delta \log(Price)_{it}$ | (1) (OLS) | (2) (OLS) | (3) (IV1) | (4) (OLS) | (5) (IV2) |
|--|--------------|---------------------|---------------------|--------------|--------------|
| $\Delta \log(Cost)_t$ | 0.870** | 0.998** | 0.968** | 0.977** | 0.967** |
| $\Delta \log(Cost)_t \times Avg$. Station Markup _i (Net %) | (0.031) | (0.035) -0.015** | (0.041) -0.011** | (0.026) | (0.033) |
| $\Delta \log(Cost)_t \times Avg$. Quarter Markup _t (Net %) | | (0.003) | (0.004) | -0.010** | -0.010** |
| $\Delta \log(\text{Oost})_t \wedge \text{Avg.}$ Quarter Markup, (Net 76) | | | | (0.002) | (0.003) |
| N | 312215 | 312215 | 312215 | 312215 | 312215 |
| R ² | 0.88 | 0.89 | 0.89 | 0.89 | 0.89 |

[•] As a result, stations with high margins appear to have "incomplete" pass-through.

Pass-through in levels explains extent & variation of "log pass-through"

| $\Delta \log(Price)_{it}$ | (1) (OLS) | (2) (OLS) | (3) (IV1) | (4) (OLS) | (5) (IV2) |
|--|--------------|--------------|--------------|--------------|--------------|
| $\Delta \log(Cost)_t$ | 0.870** | 0.998** | 0.968** | 0.977** | 0.967** |
| | (0.031) | (0.035) | (0.041) | (0.026) | (0.033) |
| $\Delta \log(\text{Cost})_t \times \text{Avg. Station Markup}_i$ (Net %) | | -0.015** | -0.011** | | |
| | | (0.003) | (0.004) | | |
| $\Delta \log(\text{Cost})_t \times \text{Avg. Quarter Markup}_t$ (Net %) | | | | -0.010** | -0.010** |
| | | | | (0.002) | (0.003) |
| N | 312215 | 312215 | 312215 | 312215 | 312215 |
| R^2 | 0.88 | 0.89 | 0.89 | 0.89 | 0.89 |

- As a result, stations with high margins appear to have "incomplete" pass-through.
- Intercept: Pass-through is complete as Net Markup_{i,t} \rightarrow 0.

Retail Gasoline: Taking Stock

- Pass-through complete in levels.
- Pass-through incomplete in logs, even accounting for cost share of gasoline.
- No apparent heterogeneity in pass-through in levels.
- Differences in margins rationalize cross-sectional heterogeneity in log pass-through.

- In paper: Similar results from other geographies (Canada, South Korea, U.S.).
- Similar results using K\u00e4nzig (2021) OPEC announcement IV for upstream costs.

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Explaining Okun (1981): Setup

- Continuum of products indexed by ω .
- Households choose which varieties $j(\omega)$ to buy and total mass of products Ω ,

$$Y = \max_{\Omega, j(\omega)} \int_0^{\Omega} \delta_{ij(\omega)} q_{ij(\omega)} d\omega, \quad \text{s.t.} \quad \int_0^{\Omega} p_{j(\omega)} q_{ij(\omega)} = I.$$

• Production of each product ω is Leontief in labor (wage w) and commodity (cost c),

$$mc(\omega) = \beta(\omega)w + (1 - \beta(\omega))c.$$

- We will assume for simplicity:
 - Identical firms selling each variety (\Rightarrow symmetric prices $p_{j(\omega)} = p_{\omega}$).
 - Each $\delta_{ij(\omega)}$ drawn i.i.d. from same distribution across all i, j, ω .

Numeraire determined endogenously

• The change in welfare from purchasing some product ω from firm $j(\omega)$ is

$$dYpprox q_{j(\omega)}d\omega \left[d\delta_{ij(\omega)}-\left[\etarac{\delta_{ij(\omega)}}{
ho_{j(\omega)}}+\left(1-\eta
ight)rac{\delta_{\Omega}}{
ho_{\Omega}}
ight]d
ho_{j(\omega)}
ight].$$

Note: "Numeraire" price determined by price/taste of marginal product Ω .

• When $\eta = 0$, each product's price is given by

$$p(\omega) = \operatorname{mc}(\omega) + \underbrace{(\mu(\Omega) - 1)\operatorname{mc}(\Omega)}_{\substack{\text{Additive margin} \\ \text{depends on }\operatorname{mc}(\Omega), \mu(\Omega)}}.$$

Pass-through of commodity vs. wage changes

Pass-through of change to unit commodity costs and wage costs:

$$\begin{split} \frac{\Delta p}{\left(1-\beta\left(\omega\right)\right)\Delta c} &= 1+\left(\mu\left(\Omega\right)-1\right)\frac{1-\beta\left(\Omega\right)}{1-\beta\left(\omega\right)},\\ \frac{\Delta p}{\beta\left(\omega\right)\Delta w} &= \mu\left(\omega\right)+\left(\mu\left(\Omega\right)-1\right)\left[\frac{\beta\left(\Omega\right)-\beta\left(\omega\right)}{\beta\left(\omega\right)}\right]\frac{c}{\text{mc}\left(\omega\right)}. \end{split}$$

- If marginal product Ω does not use commodity, i.e. $\beta(\Omega)=$ 1, we get
 - Complete pass-through in levels for commodity input.
 - (Slightly more than) complete "log pass-through" for wages.
- Okun's (1981) "special role" actually is a special role of labor costs!

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Extra slides
Retail gasoline

Okun mode

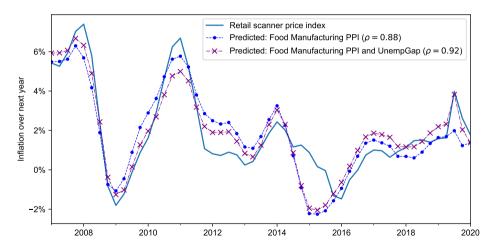
Food-at-home inflation

Price change announcements

Inflation Expectations

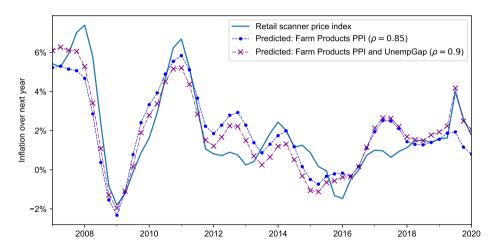
Backup: Upstream PPI explains food-at-home inflation well

Figure: Predicted Retail Scanner index inflation using Food Manufacturing PPI and unemp. gap.



Backup: Upstream PPI explains food-at-home inflation well

Figure: Predicted Retail Scanner index inflation using Farm Products PPI and unemp. gap.



Backup: Match rate of consumer expenditures to retail scanner infl. data

Table: Percent of expenditures matched to retail scanner and inflation data, by income group.

| Income | Matched to UPC | | Matched to | retailer-UPC |
|----------|----------------|------------|------------|--------------|
| quintile | Total | With infl. | Total | With infl. |
| 1 | 60.2 | 52.7 | 22.5 | 18.5 |
| 2 | 59.9 | 52.6 | 23.1 | 19.0 |
| 3 | 60.2 | 53.5 | 24.0 | 20.1 |
| 4 | 60.7 | 54.5 | 25.7 | 21.7 |
| 5 | 59.7 | 52.6 | 27.2 | 22.7 |

Backup: R^2 upstream PPI changes on inflation by income

Table: R^2 from long-term log pass-through regression.

| | Food Manufacturing PPI | | Farm | Products PPI |
|----------------------|------------------------|--------------|------|--------------|
| Index | UPC | Retailer-UPC | UPC | Retailer-UPC |
| Food-at-home CPI | 0.59 | 0.59 | 0.42 | 0.42 |
| Retail scanner index | 0.58 | 0.62 | 0.50 | 0.51 |
| All income groups | 0.42 | 0.45 | 0.31 | 0.27 |
| 1st quintile | 0.43 | 0.43 | 0.30 | 0.21 |
| 2 | 0.43 | 0.44 | 0.30 | 0.25 |
| 3 | 0.43 | 0.44 | 0.30 | 0.27 |
| 4 | 0.42 | 0.45 | 0.30 | 0.27 |
| 5th quintile | 0.43 | 0.44 | 0.30 | 0.30 |

Backup: R² upstream PPI changes on inflation by unit price group

Table: R^2 from long-term log pass-through regression.

| Index | Food Ma | nufacturing PPI | Farm I | Products PPI |
|----------------------|---------|-----------------|--------|--------------|
| | UPC | Retailer-UPC | UPC | Retailer-UPC |
| Food-at-home CPI | 0.59 | 0.59 | 0.42 | 0.42 |
| Retail scanner index | 0.58 | 0.62 | 0.50 | 0.51 |
| 1st decile | 0.61 | 0.58 | 0.55 | 0.42 |
| | 0.41 | 0.49 | 0.34 | 0.34 |
| 3 | 0.34 | 0.64 | 0.27 | 0.51 |
| 4 | 0.41 | 0.54 | 0.33 | 0.46 |
| 5 | 0.46 | 0.50 | 0.42 | 0.49 |
| 6 | 0.42 | 0.55 | 0.40 | 0.42 |
| 7 | 0.47 | 0.52 | 0.47 | 0.39 |
| 8 | 0.40 | 0.52 | 0.23 | 0.43 |
| 9 | 0.38 | 0.46 | 0.38 | 0.43 |
| 10th decile | 0.39 | 0.46 | 0.33 | 0.43 |

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Pass-through quotes

- "P&G, of Cincinatti, said Friday that it raised the list price of Folgers ground coffee by 28 cents per 11.5-ounce equivalent [...] to \$2.56 from \$2.28."
- "In futures trading Friday, arabica coffee beans for March delivery closed down 1.30 cents at \$1.3445 a pound on the New York Board of Trade. A Folgers spokeswoman said the last time the company changed retail coffee prices, in December, the contract was trading around 96 cents a pound."
- $0.28 \times (16/11.5 \text{ ounces}) = 0.38 \text{ per pound}$. Wall Street Journal.

Pass-through quotes

- "A weighted average price increase of approximately 8 percent across the company's
 instant consumable, multi-pack, packaged candy and grocery lines is effective today.
 These changes will help offset part of the significant increases in Hershey's input
 costs, including raw materials, packaging, fuel, utilities and transportation, which the
 company expects to incur in the future." Source.
- "Kruger Products L.P. ("KPLP"), a company in which KP Tissue Inc. (TSX:KPT) holds
 a limited partnership interest, announced that it will increase the price of its consumer
 branded tissue products sold in Canada effective Sunday, October 14, 2018. The
 increase is required to offset unprecedented and sustained cost increases on input
 materials and freight with pulp costs being up over 23% since our last pricing
 announcement in July 2017." Source. Article.
- "Mars Chocolate North America, the maker of M&M's and Snickers, said on Wednesday that it will raise its prices by an average of 7 percent "to offset rising costs," its first increase in three years." Source.
- "Michael Bronner, president of California-based Dr. Bronner's, said the natural-products company is increasing the price of soaps it sells to Whole Foods, Target Corp., Costco Corp., Walmart Inc. and other retailers by 3%. Prices for the

Kahneman et al. (1986) surveys on occasions for price changes

- Price increases to offset cost increases are perceived as fair.
 - "Suppose that, due to a transportation mixup, there is a local shortage of lettuce and the wholesale price has
 increased. A local grocer has bought the usual quantity of Itetuce at a price that is 30 cents per head higher
 than normal. The grocer raises the price of lettuce to customers by 30 cents per head."
 Acceptable 79%, Unfair 21%.
 - "A landlord owns and rents out a single small house to a tenant who is living on a fixed income. A higher rent would mean the tenant would have to move. Other small rental houses are available. The landlord's costs have increased substantially over the past year and the landlord raises rent to cover the cost increases when the tenant's lease is due for renewal."

Acceptable 75%, Unfair 25%.

With inventories, fairness considerations delay pass-through.

"A grocery store has several months supply of peanut butter in stock which it has on the shelves and in the storeroom. The owner hears that the wholesale price of peanut butter has increased and immediately raises the price on the current stock of peanut butter."

Acceptable 21%. Unfair 79%.

Price increases in response to demand shocks deemed unfair.

"A severe shortage of Red Delicious apples has developed in a community and none of the grocery stores or produce markets have any of this type of apple on their shelves. Other varieties of apples are plentiful in all of the other stores. One grocer receives a single shipment of Red Delicious apples at the regular wholesale cost and raises the retail price of these Red Delicious apples by 25% over the regular price."

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Food inflation exp. of low-income households more sensitive to inflation rate

| Data | Food 1-yr Inflation Expectations NY Fed, 2013–2023 | | | | | |
|---|---|---------------|---------------|--|--|--|
| Aggregation | Median (1) | Indiv. (2) | Indiv. (3) | | | |
| Mid-Income | -0.371** | -1.657** | - | | | |
| | (0.160) | (0.078) | - | | | |
| High-Income | -0.799** | -2.361** | - | | | |
| | (0.223) | (0.093) | - | | | |
| Food-at-home infl. over past year | 0.430** | 0.462** | 0.215** | | | |
| | (0.034) | (0.061) | (0.078) | | | |
| Mid -Income \times Food-at-home infl. | -0.255** | -0.163** | -0.075** | | | |
| | (0.038) | (0.024) | (0.037) | | | |
| High-Income \times Food-at-home infl. | -0.293** | -0.250** | -0.129** | | | |
| | (0.041) | (0.026) | (0.043) | | | |
| Respondent FEs N | No 360 | No 133780 | Yes 133780 | | | |

• Food infl. exp. of low-income higher, but also more sensitive to past food-at-home infl.

Overall inflation exp. of low-income more sensitive to food-at-home infl.

| _ | 1-yr Inflation Expectations | | | | |
|--|--|----------|---------|----------------------------|----------|
| Data | NY Fed, 2013-2023 | | | Michigan Survey, 1978–2024 | |
| Aggregation | Median | Indiv. | Indiv. | Mean | Median |
| | (1) | (2) | (3) | (4) | (5) |
| Mid-Income | -0.885** | -1.647** | - | -0.779** | -0.775** |
| | (0.216) | (0.289) | - | (0.092) | (0.071) |
| High-Income | -0.984** | -3.130** | - | -1.430** | -1.244** |
| | (0.263) | (0.380) | - | (0.136) | (0.091) |
| Food-at-home infl. | 0.258** | 0.243** | 0.129* | 0.139** | 0.104** |
| | (0.040) | (0.071) | (0.074) | (0.026) | (0.023) |
| $\label{eq:mid-Income} \mbox{Mid-Income} \times \mbox{Food-at-home infl.}$ | -0.142** | -0.055 | -0.042 | -0.027 | -0.023* |
| | (0.036) | (0.042) | (0.039) | (0.024) | (0.013) |
| $\label{eq:high-Income} \mbox{High-Income} \times \mbox{Food-at-home infl.}$ | -0.148** | -0.150** | -0.079 | -0.060* | -0.044** |
| | (0.043) | (0.064) | (0.063) | (0.031) | (0.019) |
| Controls | CPI less food & energy infl., gas infl., income interactions | | | | |
| Respondent FEs | - | No | Yes | - | - |
| N | 360 | 153051 | 153051 | 1599 | 1599 |

• Overall infl. exp. more sensitive to food-at-home infl (not true for broader CPI or gas).

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