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Description Provides functions which perform Bayesian estimations of a covariance matrix for multivariate normal data. Assumes that the covariance matrix is sparse or band matrix and positive-definite. This software has been developed using funding supported by Basic Science Research Program through the National Research Foundation of Korea (NRF) funded by the Ministry of Education (RS-2023-00211979, NRF-2022R1A5A7033499, NRF-2020R1A4A1018207 and NRF-2020R1C1C1A01013338).
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Bayesian Estimation of a Banded Covariance Matrix

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Description

Provides a post-processed posterior for Bayesian inference of a banded covariance matrix.

Usage

```
bandPPP(X, k, eps, prior = list(), nsample = 2000)
```

Arguments

Χ	a $n \times p$ data matrix with column mean zero.
k	a scalar value (natural number) specifying the bandwidth of covariance matrix.
eps	a small positive number decreasing to 0 with default value $(log(k))^2*(k+log(p))/n.$
prior	a list giving the prior information. The list includes the following parameters (with default values in parentheses): A (I) giving the positive definite scale matrix for the inverse-Wishart prior, nu (p + k) giving the degree of freedom of the inverse-Wishar prior.
nsample	a scalar value giving the number of the post-processed posterior samples.

Details

Lee, Lee, and Lee (2023+) proposed a two-step procedure generating samples from the post-processed posterior for Bayesian inference of a banded covariance matrix:

• Initial posterior computing step: Generate random samples from the following initial posterior obtained by using the inverse-Wishart prior $IW_p(B_0, \nu_0)$

$$\Sigma \mid X_1, \dots, X_n \sim IW_p(B_0 + nS_n, \nu_0 + n),$$

where $S_n = n^{-1} \sum_{i=1}^{n} X_i X_i^{\top}$.

• Post-processing step: Post-process the samples generated from the initial samples

$$\Sigma_{(i)} := \left\{ \begin{array}{ll} B_k(\Sigma^{(i)}) + \left[\epsilon_n - \lambda_{\min}\{B_k(\Sigma^{(i)})\}\right] I_p, & \text{ if } \lambda_{\min}\{B_k(\Sigma^{(i)})\} < \epsilon_n, \\ B_k(\Sigma^{(i)}), & \text{ otherwise }, \end{array} \right.$$

where $\Sigma^{(1)},\ldots,\Sigma^{(N)}$ are the initial posterior samples, ϵ_n is a small positive number decreasing to 0 as $n\to\infty$, and $B_k(B)$ denotes the k-band operation given as

$$B_k(B) = \{b_{ij}I(|i-j| \le k)\} \text{ for any } B = (b_{ij}) \in R^{p \times p}.$$

For more details, see Lee, Lee and Lee (2023+).

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Value

Sigma a nsample \times p(p+1)/2 matrix including lower triangular elements of covariance

matrix.

p dimension of covariance matrix.

Author(s)

Kwangmin Lee

References

Lee, K., Lee, K., and Lee, J. (2023+), "Post-processes posteriors for banded covariances", *Bayesian Analysis*, DOI: 10.1214/22-BA1333.

See Also

cv.bandPPP estimate

Examples

```
## Not run:
n <- 25
p <- 50
Sigma0 <- diag(1, p)
X <- MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = Sigma0)
res <- bspcov::bandPPP(X,2,0.01,nsample=100)
## End(Not run)</pre>
```

bmspcov

Bayesian Sparse Covariance Estimation

Description

Provides a Bayesian sparse and positive definite estimate of a covariance matrix via the beta-mixture shrinkage prior.

Usage

```
bmspcov(X, Sigma, prior = list(), nsample = list())
```

Arguments

X a $n \times p$ data matrix with column mean zero.

Sigma an initial guess for Sigma.

prior a list giving the prior information. The list includes the following parameters

(with default values in parentheses): a (1/2) and b (1/2) giving the shape parameters for beta distribution, lambda (1) giving the hyperparameter for the diagonal elements, tau1sq $(10000/(n*p^4))$ giving the hyperparameter for the

shrinkage prior of covariance.

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nsample

a list giving the MCMC parameters. The list includes the following integers (with default values in parentheses): burnin (1000) giving the number of MCMC samples in transition period, nmc (1000) giving the number of MCMC samples for analysis.

Details

Lee, Jo and Lee (2022) proposed the beta-mixture shrinkage prior for estimating a sparse and positive definite covariance matrix. The beta-mixture shrinkage prior for $\Sigma = (\sigma_{jk})$ is defined as

$$\pi(\Sigma) = \frac{\pi^u(\Sigma)I(\Sigma \in C_p)}{\pi^u(\Sigma \in C_p)}, \ C_p = \{ \ \text{all} \ p \times p \ \text{positive definite matrices} \ \},$$

where $\pi^u(\cdot)$ is the unconstrained prior given by

$$\pi^{u}(\sigma_{jk} \mid \rho_{jk}) = N\left(\sigma_{jk} \mid 0, \frac{\rho_{jk}}{1 - \rho_{jk}} \tau_{1}^{2}\right)$$
$$\pi^{u}(\rho_{jk}) = Beta(\rho_{jk} \mid a, b), \ \rho_{jk} = 1 - 1/(1 + \phi_{jk})$$
$$\pi^{u}(\sigma_{jj}) = Exp(\sigma_{jj} \mid \lambda).$$

For more details, see Lee, Jo and Lee (2022).

Value

Sigma a nmc \times p(p+1)/2 matrix including lower triangular elements of covariance matrix.

Phi a $\text{nmc} \times \text{p(p+1)/2}$ matrix including shrinkage parameters corresponding to lower

triangular elements of covariance matrix.

p dimension of covariance matrix.

Author(s)

Kyoungjae Lee and Seongil Jo

References

Lee, K., Jo, S., and Lee, J. (2022), "The beta-mixture shrinkage prior for sparse covariances with near-minimax posterior convergence rate", *Journal of Multivariate Analysis*, 192, 105067.

See Also

sbmspcov

```
## Not run:
set.seed(1)
n <- 100
p <- 20

# generate a sparse covariance matrix:
True.Sigma <- matrix(0, nrow = p, ncol = p)
diag(True.Sigma) <- 1
Values <- -runif(n = p*(p-1)/2, min = 0.2, max = 0.8)</pre>
```

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```
nonzeroIND <- which(rbinom(n=p*(p-1)/2,1,prob=1/p)==1)
zeroIND = (1:(p*(p-1)/2))[-nonzeroIND]
Values[zeroIND] <- 0</pre>
True.Sigma[lower.tri(True.Sigma)] <- Values</pre>
True.Sigma[upper.tri(True.Sigma)] <- t(True.Sigma)[upper.tri(True.Sigma)]</pre>
if(min(eigen(True.Sigma)$values) <= 0){</pre>
  delta <- -min(eigen(True.Sigma)$values) + 1.0e-5</pre>
  True.Sigma <- True.Sigma + delta*diag(p)</pre>
}
# generate a data
X \leftarrow MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = True.Sigma)
# compute sparse, positive covariance estimator:
fout <- bspcov::bmspcov(X = X, Sigma = diag(diag(cov(X))))</pre>
post.est.m <- bspcov::estimate(fout)</pre>
sqrt(mean((post.est.m - True.Sigma)^2))
sqrt(mean((cov(X) - True.Sigma)^2))
## End(Not run)
```

cv.bandPPP

CV for Bayesian Estimation of a Banded Covariance Matrix

Description

Performs leave-one-out cross-validation (LOOCV) to calculate the predictive log-likelihood for a post-processed posterior of a banded covariance matrix and selects the optimal parameters.

Usage

```
cv.bandPPP(X, kvec, epsvec, prior = list(), nsample = 2000, ncores = 2)
```

Arguments

X a $n \times p$ data matrix with column mean zero.

kvec a vector of natural numbers specifying the bandwidth of covariance matrix.

epsvec a vector of small positive numbers decreasing to 0.

prior a list giving the prior information. The list includes the following parameters

(with default values in parentheses): A (I) giving the positive definite scale matrix for the inverse-Wishart prior, nu (p + k) giving the degree of freedom of

the inverse-Wishar prior.

nsample a scalar value giving the number of the post-processed posterior samples.

ncores a scalar value giving the number of CPU cores.

Details

The predictive log-likelihood for each k and ϵ_n is estimated as follows:

$$\sum_{i=1}^{n} \log S^{-1} \sum_{s=1}^{S} p(X_i \mid B_k^{(\epsilon_n)}(\Sigma_{i,s})),$$

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where X_i is the ith observation, $\Sigma_{i,s}$ is the sth posterior sample based on $(X_1, \ldots, X_{i-1}, X_{i+1}, \ldots, X_n)$, and $B_k^{(\epsilon_n)}$ represents the banding post-processing function. For more details, see (3) in Lee, Lee and Lee (2023+).

Value

elpd

a $M \times 3$ dataframe having the expected log predictive density (ELPD) for each k and eps, where M = length(kvec) * length(epsvec).

Author(s)

Kwangmin Lee

References

Lee, K., Lee, K., and Lee, J. (2023+), "Post-processes posteriors for banded covariances", *Bayesian Analysis*, DOI: 10.1214/22-BA1333.

Gelman, A., Hwang, J., and Vehtari, A. (2014). "Understanding predictive information criteria for Bayesian models." *Statistics and computing*, 24(6), 997-1016.

See Also

bandPPP

Examples

```
## Not run:
Sigma0 <- diag(1,50)
X <- mvtnorm::rmvnorm(25,sigma = Sigma0)
kvec <- 1:2
epsvec <- c(0.01,0.05)
res <- bspcov::cv.bandPPP(X,kvec,epsvec,nsample=10,ncores=4)
plot(res)
## End(Not run)</pre>
```

cv.thresPPP

CV for Bayesian Estimation of a Sparse Covariance Matrix

Description

Performs cross-validation to estimate spectral norm error for a post-processed posterior of a sparse covariance matrix.

Usage

```
cv.thresPPP(
   X,
   thresvec,
   epsvec,
   hyperparam = NULL,
   thresfun = "hard",
```

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```
nsample = 2000,
ncores = 2
)
```

Arguments

X a $n \times p$ data matrix with column mean zero

thresvec a vector of real numbers specifying the parameter of the threshold function.

epsvec a vector of small positive numbers decreasing to 0.

thresfun a string to specify the type of threshold function. fun ('hard') giving the

thresholding function ('hard' or 'soft') for the thresholding PPP procedure.

nsample a scalar value giving the number of the post-processed posterior samples.

ncores a scalar value giving the number of CPU cores.

prior a list giving the prior information. The list includes the following parameters

(with default values in parentheses): A (I) giving the positive definite scale matrix for the inverse-Wishart prior, nu (p + 1) giving the degree of freedom of

the inverse-Wishar prior.

Details

Given a set of train data and validation data, the spectral norm error for each γ and ϵ_n is estimated as follows:

 $||\hat{\Sigma}(\gamma, \epsilon_n)^{(train)} - S^{(val)}||_2$

where $\hat{\Sigma}(\gamma, \epsilon_n)^{(train)}$ is the estimate for the covariance based on the train data and $S^{(val)}$ is the sample covariance matrix based on the validation data. The spectral norm error is estimated by the 10-fold cross-validation. For more details, see the first paragraph on page 9 in Lee and Lee (2023).

Value

 ${\sf CVdf}$ a ${\sf M} \times {\sf 3}$ dataframe having the estimated spectral norm error for each thres and

eps, where M = length(thresvec) * length(epsvec)

Author(s)

Kwangmin Lee

References

Lee, K. and Lee, J. (2023), "Post-processes posteriors for sparse covariances", *Journal of Econometrics*, 236(3), 105475.

See Also

thresPPP

```
## Not run:
Sigma0 <- diag(1,50)
X <- mvtnorm::rmvnorm(25,sigma = Sigma0)
thresvec <- c(0.01,0.1)
epsvec <- c(0.01,0.1)</pre>
```

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```
res <- bspcov::cv.thresPPP(X,thresvec,epsvec,nsample=100)
plot(res)
## End(Not run)</pre>
```

estimate

Point-estimate of posterior distribution

Description

Compute the point estimate (mean) to describe posterior distribution.

Usage

```
estimate(object, ...)
## S3 method for class 'bspcov'
estimate(object, ...)
```

Arguments

object

an object from bandPPP, bmspcov, sbmspcov, and thresPPP.

Author(s)

Seongil Jo

See Also

plot.postmean.bspcov

```
## Not run:
n <- 25
p <- 50
Sigma0 <- diag(1, p)
X <- MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = Sigma0)
res <- bspcov::bandPPP(X,2,0.01,nsample=100)
est <- bspcov::estimate(res)
## End(Not run)</pre>
```

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plot.bspcov

Plot Diagnostics of Posterior Samples and Cross-Validation

Description

Provides a trace plot of posterior samples and a plot of a learning curve for cross-validation

Usage

```
## S3 method for class 'bspcov'
plot(object, cols, rows, ...)
```

Arguments

object an object from bmspcov, sbmspcov, cv.bandPPP, and cv.thresPPP.

cols a scalar or a vector including specific column indices for the trace plot.

rows a scalar or a vector including specific row indices greater than or equal to columns indices for the trace plot.

Author(s)

Seongil Jo

```
## Not run:
set.seed(1)
n <- 100
p <- 20
# generate a sparse covariance matrix:
True.Sigma <- matrix(0, nrow = p, ncol = p)</pre>
diag(True.Sigma) <- 1</pre>
Values <- -runif(n = p*(p-1)/2, min = 0.2, max = 0.8)
nonzeroIND <- which(rbinom(n=p*(p-1)/2,1,prob=1/p)==1)
zeroIND = (1:(p*(p-1)/2))[-nonzeroIND]
Values[zeroIND] <- 0</pre>
True.Sigma[lower.tri(True.Sigma)] <- Values</pre>
True.Sigma[upper.tri(True.Sigma)] <- t(True.Sigma)[upper.tri(True.Sigma)]</pre>
if(min(eigen(True.Sigma)$values) <= 0){</pre>
  delta <- -min(eigen(True.Sigma)$values) + 1.0e-5</pre>
  True.Sigma <- True.Sigma + delta*diag(p)</pre>
# generate a data
X \leftarrow MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = True.Sigma)
# compute sparse, positive covariance estimator:
fout <- bspcov::sbmspcov(X = X, Sigma = diag(diag(cov(X))))</pre>
plot(fout, cols = c(1, 3, 4), rows = c(1, 3, 4))
\#plot(fout, cols = 1, rows = 1:3)
# Cross-Validation for Banded Structure
```

plot.postmean.bspcov

```
Sigma0 <- diag(1,50)
X <- mvtnorm::rmvnorm(25,sigma = Sigma0)
kvec <- 1:2
epsvec <- c(0.01,0.05)
res <- bspcov::cv.bandPPP(X,kvec,epsvec,nsample=10,ncores=4)
plot(res)
## End(Not run)</pre>
```

plot.postmean.bspcov Draw a Heat Map for Point Estimate of Covariance Matrix

Description

Provides a heat map for posterior mean estimate of sparse covariance matrix

Usage

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```
## S3 method for class 'postmean.bspcov'
plot(object, ...)
```

Arguments

object

an object from estimate.

Author(s)

Seongil Jo

See Also

estimate

```
## Not run:
n <- 25
p <- 50
Sigma0 <- diag(1, p)
X <- MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = Sigma0)
res <- bspcov::thresPPP(X, eps=0.01, thres=list(value=0.5,fun='hard'), nsample=100)
est <- bspcov::estimate(res)
plot(est)
## End(Not run)</pre>
```

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sbmspcov

Bayesian Sparse Covariance Estimation using Sure Screening

Description

Provides a Bayesian sparse and positive definite estimate of a covariance matrix via screened betamixture prior.

Usage

```
sbmspcov(X, Sigma, cutoff = list(), prior = list(), nsample = list())
```

Arguments

X a $n \times p$ data matrix with column mean zero.

Sigma an initial guess for Sigma.

cutoff a list giving the information for the threshold. The list includes the follow-

ing parameters (with default values in parentheses): method ('FNR') giving the method for determining the threshold value (method='FNR' uses the false negative rate (FNR)-based approach, method='corr' chooses the threshold value by sample correlations), rho a lower bound of meaningfully large correlations whose the defaults values are 0.25 and 0.2 for method = 'FNR' and method = 'corr', respectively. Note. If method='corr', rho is used as the threshold value. FNR (0.05) giving the prespecified FNR level when method = 'FNR'. nsimdata (1000) giving the number of simulated datasets for calculating Jeffreys' default

Bayes factors when method = 'FNR'.

prior a list giving the prior information. The list includes the following parameters

(with default values in parentheses): a (1/2) and b (1/2) giving the shape parameters for beta distribution, lambda (1) giving the hyperparameter for the diagonal elements, tau1sq $(\log(p)/(p^2*n))$ giving the hyperparameter for

the shrinkage prior of covariance.

nsample a list giving the MCMC parameters. The list includes the following integers

(with default values in parentheses): burnin (1000) giving the number of MCMC samples in transition period, nmc (1000) giving the number of MCMC samples

for analysis.

Details

Lee, Jo, Lee, and Lee (2022+) proposed the screened beta-mixture shrinkage prior for estimating a sparse and positive definite covariance matrix. The screened beta-mixture shrinkage prior for $\Sigma = (\sigma_{jk})$ is defined as

$$\pi(\Sigma) = \frac{\pi^u(\Sigma)I(\Sigma \in C_p)}{\pi^u(\Sigma \in C_p)}, \ C_p = \{ \ \text{all} \ p \times p \ \text{positive definite matrices} \ \},$$

where $\pi^u(\cdot)$ is the unconstrained prior given by

$$\pi^{u}(\sigma_{jk} \mid \psi_{jk}) = N\left(\sigma_{jk} \mid 0, \frac{\psi_{jk}}{1 - \psi_{jk}} \tau_{1}^{2}\right), \ \psi_{jk} = 1 - 1/(1 + \phi_{jk})$$
$$\pi^{u}(\psi_{jk}) = Beta(\psi_{jk} \mid a, b) \text{ for } (j, k) \in S_{r}(\hat{R})$$

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$$\pi^u(\sigma_{jj}) = Exp(\sigma_{jj} \mid \lambda),$$

where $S_r(\hat{R}) = \{(j,k): 1 < j < k \le p, |\hat{\rho}_{jk}| > r\}$, $\hat{\rho}_{jk}$ are sample correlations, and r is a threshold given by user.

For more details, see Lee, Jo, Lee and Lee (2022+).

Value

Sigma a nmc \times p(p+1)/2 matrix including lower triangular elements of covariance matrix. p dimension of covariance matrix. Phi a nmc \times p(p+1)/2 matrix including shrinkage parameters corresponding to lower triangular elements of covariance matrix.

INDzero a list including indices of off-diagonal elements screened by sure screening.

cutoff the cutoff value specified by FNR-approach.

Author(s)

Kyoungjae Lee and Seongil Jo

References

Lee, K., Jo, S., Lee, K., and Lee, J. (2022+), "Scalable and optimal Bayesian inference for sparse covariance matrices via screened beta-mixture prior", arXiv:2206.12773.

See Also

bmspcov

```
## Not run:
set.seed(1)
n <- 100
p <- 20
# generate a sparse covariance matrix:
True.Sigma <- matrix(0, nrow = p, ncol = p)</pre>
diag(True.Sigma) <- 1</pre>
Values <- -runif(n = p*(p-1)/2, min = 0.2, max = 0.8)
nonzeroIND <- which(rbinom(n=p*(p-1)/2,1,prob=1/p)==1)</pre>
zeroIND = (1:(p*(p-1)/2))[-nonzeroIND]
Values[zeroIND] <- 0</pre>
True.Sigma[lower.tri(True.Sigma)] <- Values</pre>
True.Sigma[upper.tri(True.Sigma)] <- t(True.Sigma)[upper.tri(True.Sigma)]</pre>
if(min(eigen(True.Sigma)$values) <= 0){</pre>
  delta <- -min(eigen(True.Sigma)$values) + 1.0e-5</pre>
  True.Sigma <- True.Sigma + delta*diag(p)</pre>
}
# generate a data
X \leftarrow MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = True.Sigma)
# compute sparse, positive covariance estimator:
```

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```
fout <- bspcov::sbmspcov(X = X, Sigma = diag(diag(cov(X))))
post.est.m <- bspcov::estimate(fout)
sqrt(mean((post.est.m - True.Sigma)^2))
sqrt(mean((cov(X) - True.Sigma)^2))
## End(Not run)</pre>
```

summary.bspcov

Summary of Posterior Distribution

Description

Provides the summary statistics for posterior samples of covariance matrix.

Usage

```
## S3 method for class 'bspcov'
summary(object, cols, rows, ...)
```

Arguments

object an object from bandPPP, bmspcov, sbmspcov, and thresPPP.

cols a scalar or a vector including specific column indices.

rows a scalar or a vector including specific row indices greater than or equal to columns

indices.

Note

If both cols and rows are vectors, they must have the same length.

Author(s)

Seongil Jo

```
## Not run:
set.seed(1)
n <- 100
p <- 20
# generate a sparse covariance matrix:
True.Sigma <- matrix(0, nrow = p, ncol = p)</pre>
diag(True.Sigma) <- 1</pre>
Values <- -runif(n = p*(p-1)/2, min = 0.2, max = 0.8)
nonzeroIND <- which(rbinom(n=p*(p-1)/2,1,prob=1/p)==1)
zeroIND = (1:(p*(p-1)/2))[-nonzeroIND]
Values[zeroIND] <- 0
True.Sigma[lower.tri(True.Sigma)] <- Values</pre>
True.Sigma[upper.tri(True.Sigma)] <- t(True.Sigma)[upper.tri(True.Sigma)]</pre>
if(min(eigen(True.Sigma)$values) <= 0){</pre>
  delta <- -min(eigen(True.Sigma)$values) + 1.0e-5</pre>
  True.Sigma <- True.Sigma + delta*diag(p)</pre>
```

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```
# generate a data
X <- MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = True.Sigma)
# compute sparse, positive covariance estimator:
fout <- bspcov::sbmspcov(X = X, Sigma = diag(diag(cov(X))))
summary(fout, cols = c(1, 3, 4), rows = c(1, 3, 4))
#summary(fout, cols = 1, rows = 1:p)
## End(Not run)</pre>
```

thresPPP

Bayesian Estimation of a Sparse Covariance Matrix

Description

Provides a post-processed posterior (PPP) for Bayesian inference of a sparse covariance matrix.

Usage

```
thresPPP(X, eps, thres = list(), prior = list(), nsample = 2000)
```

Arguments

X a $n \times p$ data matrix with column mean zero.

eps a small positive number decreasing to 0.

thres a list giving the information for thresholding PPP procedure. The list includes the following parameters (with default values in parentheses): value (0.1) giv-

ing the positive real number for the thresholding PPP procedure, fun ('hard') giving the thresholding function ('hard' or 'soft') for the thresholding PPP pro-

cedure.

prior a list giving the prior information. The list includes the following parameters

(with default values in parentheses): A (I) giving the positive definite scale matrix for the inverse-Wishart prior, nu (p + 1) giving the degree of freedom of

the inverse-Wishar prior.

nsample a scalar value giving the number of the post-processed posterior samples.

Details

Lee and Lee (2023) proposed a two-step procedure generating samples from the post-processed posterior for Bayesian inference of a sparse covariance matrix:

• Initial posterior computing step: Generate random samples from the following initial posterior obtained by using the inverse-Wishart prior $IW_p(B_0, \nu_0)$

$$\Sigma \mid X_1, \dots, X_n \sim IW_p(B_0 + nS_n, \nu_0 + n),$$

where $S_n = n^{-1} \sum_{i=1}^{n} X_i X_i^{\top}$.

• Post-processing step: Post-process the samples generated from the initial samples

$$\Sigma_{(i)} := \left\{ \begin{array}{ll} H_{\gamma_n}(\Sigma^{(i)}) + \left[\epsilon_n - \lambda_{\min}\{H_{\gamma_n}(\Sigma^{(i)})\}\right] I_p, & \text{if } \lambda_{\min}\{H_{\gamma_n}(\Sigma^{(i)})\} < \epsilon_n, \\ H_{\gamma_n}(\Sigma^{(i)}), & \text{otherwise} \end{array} \right.$$

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where $\Sigma^{(1)},\ldots,\Sigma^{(N)}$ are the initial posterior samples, ϵ_n is a positive real number, and $H_{\gamma_n}(\Sigma)$ denotes the generalized threshodling operator given as

$$(H_{\gamma_n}(\Sigma))_{ij} = \left\{ \begin{array}{ll} \sigma_{ij}, & \text{if } i = j, \\ h_{\gamma_n}(\sigma_{ij}), & \text{if } i \neq j, \end{array} \right.$$

where σ_{ij} is the (i,j) element of Σ and $h_{\gamma_n}(\cdot)$ is a generalized thresholding function. For more details, see Lee and Lee (2023).

Value

Sigma a nsample \times p(p+1)/2 matrix including lower triangular elements of covariance

matrix.

p dimension of covariance matrix.

Author(s)

Kwangmin Lee

References

Lee, K. and Lee, J. (2023), "Post-processes posteriors for sparse covariances", *Journal of Econometrics*.

See Also

cv.thresPPP

```
## Not run:
n <- 25
p <- 50
Sigma0 <- diag(1, p)
X <- MASS::mvrnorm(n = n, mu = rep(0, p), Sigma = Sigma0)
res <- bspcov::thresPPP(X, eps=0.01, thres=list(value=0.5,fun='hard'), nsample=100)
est <- bspcov::estimate(res)
## End(Not run)</pre>
```

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