# Bayesian polynomial regression

Kristian Wichmann

October 25, 2017

# 1 The setup

Given a training set  $\mathbf{x}$ ,  $\mathbf{t}$  of N points where the random t's are thought to depend of the non-random x's, we wish to find a set of weights  $\mathbf{w}$  that fits the model where the t's are normally distributed with precision  $\beta$  and mean:

$$y(x, \mathbf{w}) = \sum_{j=0}^{M} w_j x^j \tag{1.1}$$

More generally,  $x^{j}$  could be substituted for another set of basis functions  $f_{i}(x)$ .

### 2 Maximum likelihood estimation

The usual frequentist approach would be a maximum likelihood estimation of the parameters  $\beta$  and  $\mathbf{w}$ . The likelihood function for the model is:

$$L(\beta, \mathbf{w} | \mathbf{x}, \mathbf{t}) = \prod_{n=1}^{N} \mathcal{N}(y(x_n, \mathbf{w}), \beta^{-1}) = \prod_{n=1}^{N} \sqrt{\frac{\beta}{2\pi}} \exp\left[-\frac{1}{2}\beta(y(x_n, \mathbf{w}) - t_n)^2\right]$$
(2.1)

To turn the product into the sum, find the log-likelihood<sup>1</sup>

$$\ell(\beta, \mathbf{w}|\mathbf{x}, \mathbf{t}) = -\frac{N}{2}(\ln \beta - \ln(2\pi)) + \frac{\beta}{2} \sum_{n=1}^{N} (y(x_n, \mathbf{w}) - t_n)^2$$
 (2.2)

Now differentiate with respect to weights and precision to find the minimum:

$$\frac{\partial \ell}{\partial w_j} = \frac{\beta}{2} \sum_{n=1}^{N} 2(y(x_n, \mathbf{w}) - t_n) \frac{\partial y_n}{\partial w_j} = \beta \sum_{n=1}^{N} (y(x_n, \mathbf{w}) - t_n) x_n^j$$
 (2.3)

<sup>&</sup>lt;sup>1</sup>Here including a minus as well. Conventions might differ.

Setting equal to zero we get the usual least squares equations:

$$\sum_{n=1}^{N} (y(x_n, \mathbf{w}_{ML}) - t_n) x_n^j = 0$$
 (2.4)

For precision:

$$\frac{\partial \ell}{\partial \beta} = -\frac{N}{2\beta} + \frac{1}{2} \sum_{n=1}^{N} (y(x_n, \mathbf{w}) - t_n)^2$$
 (2.5)

Setting this equal to zero:

$$\frac{1}{\beta_{\text{ML}}} = \frac{1}{N} \sum_{n=1}^{N} (y(x_n, \mathbf{w}) - t_n)^2$$
 (2.6)

# 3 Bayesian treatment

In the Bayesian formulation, we need to specify a *prior distribution*. Initially, without any specific information, we have no idea whether a weight should be positive or negative, so let's choose a prior where each is centered at zero. The normal distribution is an obvious pick. Assuming independence of weights and them all having the same precision  $\alpha$ , this means our prior is:

$$p(\mathbf{w}|\alpha) = \mathcal{N}(\mathbf{0}, \alpha^{-1}\mathbf{I}_{M+1}) = \left(\frac{\alpha}{2\pi}\right)^{(M+1)/2} \exp\left\{-\frac{\alpha}{2}\mathbf{w}^t\mathbf{w}\right\}$$
 (3.1)

We can now use Bayes' theorem to obtain the posterior distribution as proportional to the likelihood function (2.1) times the prior (3.1):

$$p(\mathbf{w}|\alpha, \beta, \mathbf{x}, \mathbf{t}) \propto L(\beta, \mathbf{w}|\mathbf{x}, \mathbf{t})p(\mathbf{w}|\alpha)$$
 (3.2)

Ignoring factors, this means:

$$p(\mathbf{w}|\alpha, \beta, \mathbf{x}, \mathbf{t}) \propto \prod_{n=1}^{N} \exp\left[-\frac{1}{2}\beta(y(x_n, \mathbf{w}) - t_n)^2\right] \cdot \exp\left\{-\frac{\alpha}{2}\mathbf{w}^t\mathbf{w}\right\}$$
 (3.3)

### 3.1 Maximum posterity estimation

If we're just looking for a point estimate, we can now choose the parameters which maximizes equation 3.3, a process known as  $maximum\ posterior\ estimation$  or MAP for short. Since the exponential is strictly monotonous, we can maximize the exponent to get:

$$\frac{\beta}{2} sum_{i=1}^{N} (y(x_n, \mathbf{w}_{MAP}) - t_n)^2 + \frac{\alpha}{2} \mathbf{w}_{MAP}^t \mathbf{w}_{MAP} = 0$$
 (3.4)