

Expected Value — Survival Function Method

For a continuous, non-negative X , the expected value of a function X , i.e. $g(X)$, can be rewritten as follows using integration by parts:

$$\begin{aligned}
 E[g(X)] &= \int_0^{\infty} g(x) \cdot f_X(x) \, dx \\
 &= [-g(x) \cdot S_X(x)] \Big|_0^{\infty} - \int_0^{\infty} g'(x) \cdot [-S_X(x)] \, dx \\
 &= \left[-g(\infty) \cdot \underbrace{S_X(\infty)}_0 + \underbrace{g(0)}_0 \cdot S_X(0) \right] + \int_0^{\infty} g'(x) \cdot S_X(x) \, dx \\
 &= \int_0^{\infty} g'(x) \cdot S_X(x) \, dx
 \end{aligned}$$

Note that in order to simplify to the final line, two requirements need to be fulfilled:

1. The upper bound of the integral needs to be the upper bound of X , resulting in $S_X(\infty) = 0$ in the equation above.
2. The function evaluated at the lower bound needs to be 0, i.e., $g(0) = 0$ in the equation above.

If the two requirements are not fulfilled, we need to adjust the formula accordingly.

It follows that for the scenario where $g(X) = X$,

$$\begin{aligned}
 \mathbb{E}[X] &= \int_0^{\infty} x \cdot f_X(x) \, dx \\
 &= [-x \cdot S_X(x)] \Big|_0^{\infty} - \int_0^{\infty} 1 \cdot [-S_X(x)] \, dx \\
 &= \int_0^{\infty} S_X(x) \, dx
 \end{aligned}$$

3rd Central Moment

Recall that $\mu'_k = \mathbb{E}[X^k]$. Thus,

$$\begin{aligned}
 \mu_3 &= \mathbb{E}[(X - \mu)^3] \\
 &= \mathbb{E}\left[\binom{3}{0} X^{3-0} \mu^0 - \binom{3}{1} X^{3-1} \mu^1 + \binom{3}{2} X^{3-2} \mu^2 - \binom{3}{3} X^{3-3} \mu^3\right] \\
 &= \mathbb{E}[X^3 - 3X^2\mu + 3X\mu^2 - \mu^3] \\
 &= \mathbb{E}[X^3] - 3\mu\mathbb{E}[X^2] + 3\mu^2\mathbb{E}[X] - \mu^3 \\
 &= \mu'_3 - 3\mu(\mu'_2) + 3\mu^2(\mu) - \mu^3 \\
 &= \mu'_3 - 3\mu(\mu'_2) + 3\mu^3 - \mu^3 \\
 &= \mu'_3 - 3\mu'_2\mu + 2\mu^3
 \end{aligned}$$

4th Central Moment

Recall that $\mu'_k = \mathbb{E}[X^k]$. Thus,

$$\begin{aligned}
\mu_4 &= \mathbb{E}[(X - \mu)^4] \\
&= \mathbb{E}\left[\binom{4}{0}X^{4-0}\mu^0 - \binom{4}{1}X^{4-1}\mu^1 + \binom{4}{2}X^{4-2}\mu^2 - \binom{4}{3}X^{4-3}\mu^3 + \binom{4}{4}X^{4-4}\mu^4\right] \\
&= \mathbb{E}[X^4 - 4X^3\mu + 6X^2\mu^2 - 4X\mu^3 + \mu^4] \\
&= \mathbb{E}[X^4] - 4\mu\mathbb{E}[X^3] + 6\mu^2\mathbb{E}[X^2] - 4\mu^3\mathbb{E}[X] + \mu^4 \\
&= \mu'_4 - 4\mu(\mu'_3) + 6\mu^2(\mu'_2) - 4\mu^3(\mu) + \mu^4 \\
&= \mu'_4 - 4\mu(\mu'_3) + 6\mu^2(\mu'_2) - 4\mu^4 + \mu^4 \\
&= \mu'_4 - 4\mu'_3\mu + 6\mu'_2\mu^2 - 3\mu^4
\end{aligned}$$

The Law of Total Variance

$$\text{Var}_X [X] = \mathbb{E}_X [X^2] - \mathbb{E}_X[X]^2$$

Apply the Law of Total Expectation.

$$\mathbb{E}_X [X] = \mathbb{E}_Y [\mathbb{E}_X [X | Y]]$$

$$\begin{aligned}
\mathbb{E}_X [X^2] &= \mathbb{E}_Y [\mathbb{E}_X [X^2 | Y]] \\
&= \mathbb{E}_Y [\text{Var}_X [X | Y] + \mathbb{E}_X [X | Y]^2] \\
&= \mathbb{E}_Y [\text{Var}_X [X | Y]] + \mathbb{E}_Y [\mathbb{E}_X [X | Y]^2]
\end{aligned}$$

Therefore,

$$\begin{aligned}
\text{Var}_X [X] &= \mathbb{E}_X [X^2] - \mathbb{E}_X[X]^2 \\
&= \mathbb{E}_Y [\text{Var}_X [X | Y]] + \mathbb{E}_Y [\mathbb{E}_X [X | Y]^2] - \mathbb{E}_Y [\mathbb{E}_X [X | Y]]^2
\end{aligned}$$

Recall that $\text{Var}[g(X)] = \mathbb{E}[g(X)^2] - \mathbb{E}[g(X)]^2$. Thus, we can translate the second half of the equation above to:

$$\mathbb{E}_Y \left[\mathbb{E}_X [X | Y]^2 \right] - \mathbb{E}_Y [\mathbb{E}_X [X | Y]]^2 = \text{Var}_Y [\mathbb{E}_X [X | Y]]$$

Thus,

$$\text{Var}_X [X] = \mathbb{E}_Y [\text{Var}_X [X | Y]] + \text{Var}_Y [\mathbb{E}_X [X | Y]]$$

Two Forms of the Biased Sample Variance

$$\begin{aligned} \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n} &= \frac{\sum_{i=1}^n (x_i^2 - 2\bar{x}x_i + \bar{x}^2)}{n} \\ &= \frac{\sum_{i=1}^n x_i^2 - 2\bar{x} \sum_{i=1}^n x_i + n \cdot \bar{x}^2}{n} \\ &= \frac{\sum_{i=1}^n x_i^2}{n} - 2\bar{x}^2 + \bar{x}^2 \\ &= \frac{\sum_{i=1}^n x_i^2}{n} - \bar{x}^2 \end{aligned}$$