

Self-Validated Computations for the
Probabilities of the Central
Bivariate Chi-Square Distribution

Kevin Wright
Iowa State University
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Ph.D. advisor: William Kennedy

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Outline

- Introduction to Interval Analysis
- Bivariate Chi-Square Distributions
- Interval Bracket-Secant/Bisection
- Numerical results

Motivation

- Error analysis
- Cancellation
- Rounding errors

Solution? Interval analysis

Goal: Guaranteed error bounds

Interval Analysis (Moore, 1966, 1979)

A real interval \mathbf{x} is defined $\mathbf{x} = [\underline{x}, \overline{x}]$ where $\underline{x} \in \mathbb{R}, \overline{x} \in \mathbb{R}, \underline{x} \leq \overline{x}$.

Let $\mathbf{x} = [\underline{x}, \overline{x}]$ and $\mathbf{y} = [\underline{y}, \overline{y}]$.

Arithmetic operations for intervals are defined:

$$\mathbf{x} * \mathbf{y} = \{x * y : x \in \mathbf{x}, y \in \mathbf{y}\} \text{ for } * \in \{+, -, \times, \div\}$$

Equivalently:

$$\mathbf{x} + \mathbf{y} = [\underline{x} + \underline{y}, \overline{x} + \overline{y}]$$

$$\mathbf{x} - \mathbf{y} = [\underline{x} - \overline{y}, \overline{x} - \underline{y}]$$

$$\mathbf{x} \cdot \mathbf{y} = [\min(\underline{x}\underline{y}, \underline{x}\overline{y}, \overline{x}\underline{y}, \overline{x}\overline{y}), \max(\underline{x}\underline{y}, \underline{x}\overline{y}, \overline{x}\underline{y}, \overline{x}\overline{y})]$$

$$1/\mathbf{y} = [1/\overline{y}, 1/\underline{y}], \quad 0 \notin \mathbf{y}$$

$$\mathbf{x}/\mathbf{y} = \mathbf{x} * (1/\mathbf{y}), \quad \mathbf{y} \neq \mathbf{0}$$

Examples:

$$[1, 1] + [-2, 5] = [-1, 6]$$

$$[-2, 3] * [1, 4] = [-8, 12]$$

Note: Subtraction and division are not the inverse of addition and multiplication.

$$[0, 1] - [0, 1] = [-1, 1]$$

$$[1, 2]/[1, 2] = [0.5, 2]$$

Theorem: Rational interval functions are inclusion monotone, i.e. $f(\mathbf{x}) \subseteq f(\mathbf{y})$ whenever $\mathbf{x} \subseteq \mathbf{y}$.

- IEEE Floating Point Specifications

Round to zero

Round to nearest

Round to $+\infty$

Round to $-\infty$

Compute $[1, 1]/[3, 3]$:

$$[\nabla(1/3), \Delta(1/3)] = [0.333, 0.334] = [0.33\frac{4}{3}]$$

- Ensures guaranteed enclosure for rational interval functions.

- C++ BIAS/PROFIL (O. Knuppel, 1993)

Operator overloading : $\mathbf{x} + \mathbf{y}$

Rounding mode control

Dependency difficulties

Reduce the number of occurrences of a given interval in an expression.

- Ex: Suppose $x = [-1, 2]$.

$$x^2 = x \cdot x = [-1, 2] \cdot [-1, 2] = [-2, 4]$$

Fix this with an appropriate definition of x^2

$$x^2 = \{x^2 : x \in x\}$$

$$x^2 = [0, 4]$$

- Ex: Two extensions of $f(x) = x^2 - x$ are:

$$f_1(x) = x^2 - x$$

$$f_2(x) = (x - \frac{1}{2})^2 - \frac{1}{4}$$

$$f_3(x) = x(x - 1)$$

$$f_1([0, 2]) = [-2, 4]$$

$$f_2([0, 2]) = [-\frac{1}{4}, 2]$$

$$f_3([0, 2]) = [-2, 2]$$

The range of f over $[0, 2]$ is $[-\frac{1}{4}, 2]$.

Replacing each occurrence of x by x is called the *natural interval extension* of $f(x)$.

Bivariate Chi-Square Applications

- Simultaneous inferences for variances
- Simultaneous tests in ANOVA
- Simultaneous tests for goodness of fit
- Distribution of larger of correlated χ^2 -variates
- Density of linear combination of independent χ^2 -variates

See Gunst & Webster, 1973; Jensen & Howe, 1968

Bivariate Chi-square distribution : Case I
Gunst, 1973

$(Z_{1i}, Z_{2i}), i = 1, \dots, m$ independent

$Z_{ij} \sim N(0, 1)$

$Corr(Z_{1i}, Z_{2i}) = \rho, i = 1, \dots, m$

$Y_i = \sum_{j=1}^m Z_{ij}^2, i = 1, 2$

$(Y_1, Y_2) \sim Biv\chi^2(m, m, m)$

Density

$$f(y_1, y_2) = (1 - \rho^2)^{m/2} \sum_{j=0}^{\infty} \frac{\Gamma(\frac{m}{2} + j) \rho^{2j}}{j! \Gamma(\frac{m}{2})} \times$$

$$\frac{(y_1 y_2)^{(m/2)+j-1} \exp[-(y_1 + y_2)/2(1 - \rho^2)]}{[2^{(m/2)+j} \Gamma(\frac{m}{2} + j) (1 - \rho^2)^{(m/2+j)/2}]^2}$$

Bivariate Chi-square distribution : Case I

Distribution

$$P[Y_1 \leq d_1, Y_2 \leq d_2] = (1 - \rho^2)^{m/2} \times \sum_{j=0}^{\infty} \frac{\Gamma(\frac{m}{2} + j)}{j! \Gamma(\frac{m}{2})} \rho^{2j} \gamma(\frac{m}{2} + i, d_1^*) \gamma(\frac{m}{2} + i, d_2^*)$$

$$\gamma(\alpha, d) = \int_0^d \frac{1}{\Gamma(\alpha)} x^{\alpha-1} e^{-x} dx$$

$$d_j^* = d_j / (1 - \rho_{12}^2).$$

Let P_t be the sum truncated at t

Truncation error (Gunst, 1973)

$$0 \leq R_t \leq 1 - (1 - \rho^2)^{m/2} \sum_{j=0}^t \frac{\Gamma(\frac{m}{2} + j)}{j! \Gamma(\frac{m}{2})} \rho^{2j} \quad (1)$$

$\mathbf{p}_t, \mathbf{r}_t$ interval extensions of P_t, R_t

$$P[Y_1 \leq d_1, Y_2 \leq d_2] = P_t + R_t \text{ for all } t$$

$$\in [\underline{\mathbf{p}}_t, \overline{\mathbf{p}}_t + \overline{\mathbf{r}}_t] \text{ for all } t$$

- t determined by machine/software precision
- Stopping rules

$$\mathbf{r}_{t-1} = \mathbf{r}_t$$

$$width(\mathbf{r}_t) < \epsilon$$

Bivariate Chi-square distribution : Case II

$(Z_{1i}, Z_{2i}), i = 1, \dots, m$ independent

$Z_{2i}, i = m + 1, \dots, m + n$

$Z_{ij} \sim N(0, 1)$

$Corr(Z_{1i}, Z_{2i}) = \rho, i = 1, \dots, m$

$Y_1 = \sum_{j=1}^m Z_{1j}^2, Y_2 = \sum_{j=1}^{m+n} Z_{2j}^2$

$(Y_1, Y_2) \sim Biv\chi^2(m, m + n, \rho)$

Distribution

$$P[Y_1 \leq d_1, Y_2 \leq d_2,] = \\ (1 - \rho^2)^{(m+n)/2} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \frac{\Gamma(\frac{m}{2} + j)}{j! \Gamma(\frac{m}{2})} \frac{\Gamma(\frac{n}{2} + k)}{k! \Gamma(\frac{n}{2})} \times \\ \rho^{2(j+k)} \gamma(\frac{m}{2} + j, d_1^*) \gamma(\frac{n}{2} + k, d_2^*)$$

Truncation error

$$R_{t_1, t_2} \leq 1 - (1 - \rho^2)^{m/2+n/2} \times \sum_{j=0}^{t_1} \sum_{k=0}^{t_2} \frac{\Gamma(\frac{m}{2} + j)}{j! \Gamma(\frac{m}{2})} \frac{\Gamma(\frac{n}{2} + k)}{k! \Gamma(\frac{n}{2})} \rho^{2(j+k)}$$

$P[Y_1 \leq d_1, Y_2 \leq d_2] \in [\underline{\mathbf{p}}_{t_1, t_2}, \bar{\mathbf{p}}_{t_1, t_2} + \bar{\mathbf{r}}_{t_1, t_2}]$ for all pairs (t_1, t_2) .

Bivariate Chi-square distribution : Case III

$(Z_{1i}, Z_{2i}), i = 1, \dots, m$ independent

$Z_{1i}, i = m + 1, \dots, m + n$

$Z_{2i}, i = m + 1, \dots, m + p$

$Z_{ij} \sim N(0, 1)$

$Corr(Z_{1i}, Z_{1i}) = \rho, 1 = 1, \dots, m$

$Y_1 = \sum_{j=1}^{m+n} Z_{1j}^2, Y_2 = \sum_{j=1}^{m+p} Z_{2j}^2$

$(Y_1, Y_2) \sim Biv\chi^2(m + n, m + p, m)$

Distribution

$$P[Y_1 \leq d_1, Y_2 \leq d_2,] = (1 - \rho^2)^{(m+n+p)/2} \times \\ \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \sum_{l=0}^{\infty} \frac{\Gamma(\frac{m}{2} + j)}{j! \Gamma(\frac{m}{2})} \frac{\Gamma(\frac{n}{2} + k)}{k! \Gamma(\frac{n}{2})} \frac{\Gamma(\frac{p}{2} + l)}{l! \Gamma(\frac{p}{2})} \times \\ \rho^{2(j+k+l)} \gamma(\frac{m}{2} + \frac{n}{2} + k + j, d_1^*) \gamma(\frac{m}{2} + \frac{p}{2} + j + l, d_2^*)$$

Truncation error

$$R_t \leq 1 - (1 - \rho^2)^{m/2+n/2+p/2} \times \sum_{j=0}^{t_1} \sum_{k=0}^{t_2} \sum_{l=0}^{t_3} \frac{\Gamma(\frac{m}{2} + j)}{j! \Gamma(\frac{m}{2})} \frac{\Gamma(\frac{n}{2} + k)}{k! \Gamma(\frac{n}{2})} \frac{\Gamma(\frac{p}{2} + l)}{l! \Gamma(\frac{p}{2})} \rho^{2(j+k+l)}$$

$P[Y_1 \leq d_1, Y_2 \leq d_2] \in [\underline{\mathbf{p}}_{t_1, t_2, t_3}, \overline{\mathbf{p}}_{t_1, t_2, t_3} + \bar{\mathbf{r}}_{t_1, t_2, t_3}]$
for all triples (t_1, t_2, t_3)

Critical points of distributions

$$P(Y_1 \leq x, Y_2 \leq x) - \alpha = 0$$

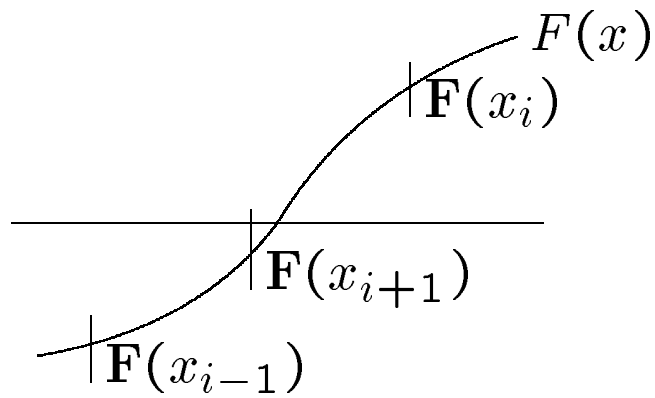
- Intervalized Newton methods
- Interval extension of F, f

Alternative: Derivative free method

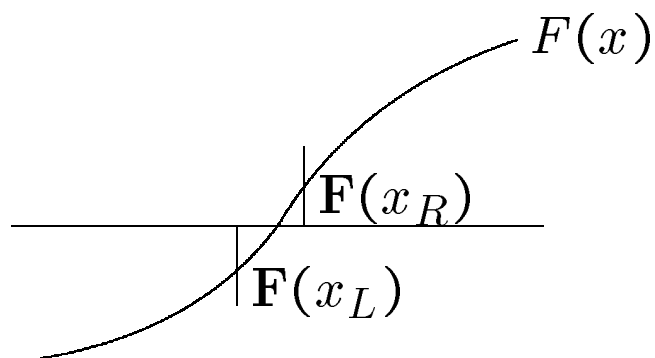
- Bracket-Secant / Illinois Modification

Bracket-Secant stops too soon

- Bisection



Termination of the bracket-secant



Termination of the bisection portion

Upper 0.05 Percentile points of the Bivariate
Chi-Square Distribution : Case I

$$(Y_1, Y_2) \sim Biv\chi^2(m, m, m)$$

ρ	m = 2	m = 12
0.1	7.348735242636 ₆₂ ⁹⁴	23.291675614644 ₃ ⁹
0.2	7.337736654468 ₅₂ ⁷³	23.279893907495 ₀ ⁵
0.3	7.318116097295 ₀₀ ³³	23.257752618706 ₃ ⁸
0.4	7.28777721964 ₁₉₇ ²³¹	23.22124105410 ₁₃ ²⁰
0.5	7.243389878426 ₀₄ ³⁵	23.16405309924 ₇₁ ⁸⁰
0.6	7.179739084402 ₀₁ ⁴⁷	23.07634122520 ₀₇ ¹⁸
0.7	7.088168635581 ₂₃ ⁷³	22.94173918779 ₁₀ ²⁶
0.8	6.95217862545 ₄₆₂ ⁵⁶¹	22.72905146810 ₀₃ ²⁸
0.9	6.73002568707 ₄₉₇ ⁶⁹⁵	22.35957467992 ₀ ⁶

Upper 0.05 Percentile points of the Bivariate
Chi-Square Distribution : Case II
 $(Y_1, Y_2) \sim Biv\chi^2(m, m + n, m)$

m	$m + n$	$\rho = 0.2$	$\rho = 0.4$
8	10	19.25562949840 ₅₇ ⁸³	19.2145575272 ₀₅₁ ¹¹⁸
8	12	21.43987191019 ₀₁ ³¹	21.4161617406 ₂₄₄ ³¹³
8	14	23.8527709314 ₀₉₃ ¹²⁵	23.8412410690 ₀₈₆ ¹⁵⁴
8	16	26.3613220302 ₆₆₅ ⁷⁰⁰	26.356264614 ₇₉₄₂ ⁸⁰²⁸
8	18	28.89382780550 ₅₇ ⁹³	28.8917353286 ₆₅₇ ⁷⁶⁸

m	$m + n$	$\rho = 0.6$	$\rho = 0.8$
8	10	19.1173225955 ₂₆₈ ⁴⁶⁴	18.90189596 ₂₉₅₀₃ ³⁰³⁴¹
8	12	21.3604107620 ₄₆₀ ⁶²⁵	21.243609643 ₈₃₈₈ ⁹¹⁶³
8	14	23.8140331030 ₆₅₄ ⁸⁸⁶	23.759702671 ₅₂₃₆ ⁶⁴⁵⁹
8	16	26.3442149104 ₄₁₉ ⁷³⁵	26.321129600 ₂₃₃₈ ³⁴⁰⁰
8	18	28.8866926661 ₄₈₀ ⁷³²	28.877401302 ₄₆₄₇ ⁶²⁴¹

Upper 0.05 Percentile points of the Bivariate
Chi-Square Distribution: Case III
 $(Y_1, Y_2) \sim Biv\chi^2(m + n, m + p, m)$

m	n	p	$\rho = 0.4$	$\rho = 0.6$
7	1	11	28.892120984 ⁷²³⁵ ₆₄₈₇	28.887920316 ⁷⁴⁶¹ ₃₄₅₀
6	2	12	28.8924903043 ⁹²⁸ ₁₉₀	28.88907530 ⁴²⁹¹⁴ ₃₉₀₀₅
5	3	13	28.892843220 ⁴²³³ ₃₅₀₆	28.890155051 ⁴⁶⁴¹ ₁₁₀₅
4	4	14	28.893179676 ⁸⁵⁶² ₇₈₅₂	28.891157355 ⁶¹⁹⁶ ₁₂₇₉
3	5	15	28.893499628 ⁹¹⁶¹ ₈₄₉₈	28.89208036 ⁹³²⁶⁹ ₈₈₆₃₂
2	6	16	28.893803042 ⁸⁷⁶⁷ ₇₉₆₁	28.892922583 ³⁸⁸⁶ ₀₀₂₃
1	7	17	28.894089895 ⁹⁰⁴⁵ ₈₀₅₅	28.89368281 ¹⁴⁰³⁸ ₀₉₉₂₆

Conclusions

- Theoretical error analysis
- Rounding errors
- Cancellation

- Interval analysis
- Guaranteed error bound
- Improved Bivariate Chi-Square values
- Corrections to previous tables

- Further research in statistics

- Univariate Distributions : Wang & Kennedy, 1994