How Markov Chain Monte Carlo (MCMC) Works

Markov Chain Monte Carlo (MCMC) is a class of algorithms used to **sample from a probability distribution** for which direct sampling is difficult or impossible. This is often the case for complex, high-dimensional distributions, especially the **posterior distribution** in Bayesian inference.

The core idea behind MCMC is to construct a **Markov chain** – a sequence of random variables where the probability of transitioning to the next state depends only on the current state. ¹² This Markov chain is designed such that its **stationary distribution** (the distribution it converges to after a large number of steps) is the target probability distribution we want to sample from. ³

Here's a breakdown of the key steps involved:

- 1. **Define the Target Distribution:** This is the probability distribution we want to sample from (e.g., the posterior distribution in Bayesian inference). We usually know a function that is proportional to this distribution, even if we don't know the normalization constant. Let's call this unnormalized target distribution f(x). Our goal is to get samples that look like they came from P(x)=Zf(x), where $Z=\int f(x)dx$ (the normalization constant) might be unknown.
- 2. **Initialize:** Start at an arbitrary initial state x0 in the sample space of the target distribution.⁵
- 3. **Propose a New State:** From the current state xt, propose a new candidate state x* according to a **proposal distribution** Q(x*|xt). The choice of the proposal distribution is crucial for the efficiency of the MCMC algorithm. Common choices include:
 - **Random Walk:** Proposing a new state by taking a random step from the current state (e.g., adding a random value from a normal distribution).
 - **Independent Proposal:** Proposing a new state independently of the current state (though this requires some knowledge of the target distribution to be efficient).
- 4. **Accept or Reject the Proposed State:** Decide whether to move to the proposed state x* or stay at the current state $xt.^8$ This decision is made based on an **acceptance probability** $\alpha(x*|xt).^9$ The acceptance probability is designed to ensure that the Markov chain converges to the target distribution. A common acceptance criterion is the **Metropolis-Hastings algorithm:**
 - $\alpha(x*|xt)=\min(1,f(xt)Q(x*|xt)f(x*)Q(xt|x*))$
 - The ratio f(xt)f(x*) compares the "likelihood" of the proposed state to the current state under the target distribution.
 - The ratio Q(x*|xt)Q(xt|x*) accounts for the asymmetry in the proposal distribution (how likely it is to propose x* from x*). If the proposal distribution is symmetric (i.e., Q(x*|xt)=Q(xt|x*)), the acceptance probability simplifies to: $\alpha(x*|xt)=\min(1,f(xt)f(x*))$
- 5. **Update the State:**
 - \circ Generate a random number u from a uniform distribution between 0 and 1 (U(0,1)). 10
 - If $u \le \alpha(x*|xt)$, then **accept** the proposed state and set xt+1=x*.
 - If $u > \alpha(x * | xt)$, then **reject** the proposed state and stay at the current state, setting xt+1=xt.
- **6**. **Repeat:** Continue steps 3-5 for a large number of iterations. ¹¹
- 7. **Collect Samples:** After a "burn-in" period (initial iterations discarded to allow the chain to converge to the stationary distribution), the sequence of accepted states (xburn-in+1,xburn-in+2,...,xT) forms a sample from the target distribution.

Why does this work?

The acceptance probability is carefully constructed to satisfy the **detailed balance condition** for the Markov chain with respect to the target distribution. Detailed balance ensures that the rate of transitioning from state x to state y is

equal to the rate of transitioning from state y to state x under the target distribution. If detailed balance holds, then the target distribution is the stationary distribution of the Markov chain. ¹²

Example: Sampling from a Beta Distribution using Metropolis-Hastings

Let's say we want to sample from a Beta distribution with parameters α =2 and β =5. The probability density function (PDF) of a Beta distribution is:

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P(x|\alpha,\beta)=B(\alpha,\beta)x\alpha-1(1-x)\beta-1, for 0 \le x \le 1
```

where $B(\alpha,\beta)$ is the Beta function (the normalization constant). Let $f(x)=x\alpha-1(1-x)\beta-1$ be the unnormalized target distribution.

We will use a simple **random walk proposal** where we propose a new state x* by taking a step from the current state x*:

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x = xt + \epsilon
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where ϵ is drawn from a symmetric distribution centered at zero, such as a uniform distribution [-0.1,0.1]. Since the proposal distribution is symmetric (Q(x*|xt)=Q(xt|x*)), the acceptance probability simplifies to:

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\alpha(x*|xt) = \min(1, f(xt)f(x*)) = \min(1, (xt)2 - 1(1-xt)5 - 1(x*)2 - 1(1-x*)5 - 1) = \min(1, xt(1-xt)4x*(1-x*)4) = \min(1, xt(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt)4x*(1-xt
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Steps:

- 1. **Initialize:** Let's start with x0=0.5.
- **2**. **Propose:** Draw ϵ from U(-0.1,0.1). Let's say ϵ =0.05. Then x*=0.5+0.05=0.55.
- 3. Calculate Acceptance Probability: $f(xt)=0.5(1-0.5)4=0.5\times0.0625=0.03125$ $f(x*)=0.55(1-0.55)4=0.55\times(0.45)4=0.55\times0.04100625\approx0.02255$ $\alpha(0.55|0.5)=\min(1,0.031250.02255)=\min(1,0.7216)=0.7216$
- **4. Accept or Reject:** Generate a random number u from U(0,1). Let's say u=0.3. Since u=0.3 \le 0.7216, we **accept** the proposed state. So, x1=0.55.
- **5**. **Repeat:** We continue this process for many iterations. The sequence of accepted x values will eventually approximate samples from the Beta(2, 5) distribution.

Important Considerations:

- **Burn-in Period:** The initial samples of the Markov chain might be far from the stationary distribution, so they are typically discarded.¹³
- **Mixing:** The rate at which the Markov chain converges to the stationary distribution is crucial. A chain that mixes slowly will require many more iterations to produce representative samples.
- Proposal Distribution: The choice of the proposal distribution significantly affects the efficiency of the MCMC algorithm.¹⁴ A good proposal distribution will lead to a reasonable acceptance rate and good mixing.¹⁵
- **Autocorrelation:** Successive samples in an MCMC chain are often correlated. ¹⁶ Techniques like thinning (keeping only every n-th sample) can be used to reduce autocorrelation. ¹⁷

MCMC algorithms are powerful tools for sampling from complex distributions and are widely used in Bayesian statistics, computational physics, and other fields where direct sampling is intractable. ¹⁸ Common MCMC algorithms include Metropolis-Hastings, Gibbs sampling, and Hamiltonian Monte Carlo.