

# Project2

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## Importing required libraries

```
library(tidyverse)      # ggplot(), %>%, mutate(), and friends
library(scales)         # Format numbers with functions like comma(), percent(), and dollar()
library(broom)          # Convert models to data frames
library(modelsummary)   # Side-by-side regression tables
library(foreign)        # for importing stata data
library(readr)
library(haven)
library(modelsummary)
library(stargazer)
library(tidyverse)
```

## Importing dataset “INJURY”

```
dataset <- read_dta("C:/Users/utkrsh/Desktop/EC402_Assignment_2_DID/INJURY.DTA")
```

## Cleaning the data so that it only includes rows from (ky == 1)

```
injury <- dataset %>% filter(ky==1)
```

## Renaming the columns

```
injury <- injury %>% rename(duration = durat, log_duration = ldurat, after_1980 = afchnge)
```

Viewing the dataset

```
injury
```

```
## # A tibble: 5,626 x 30
##   duration after_1980 highearn male married hosp indust injtype age prewage
##   <dbl>      <dbl>    <dbl> <dbl>    <dbl> <dbl>  <dbl>  <dbl> <dbl>
## 1         1         1         1     1      0      1      3      1    26    405.
## 2         1         1         1     1      1      0      3      1    31    644.
## 3        84         1         1     1      1      1      3      1    37    398.
## 4         4         1         1     1      1      1      3      1    31    528.
## 5         1         1         1     1      1      0      3      1    23    529.
## 6         1         1         1     1      1      0      3      1    34    614.
## 7         7         1         1     1      1      0      3      1    35    546.
## 8         2         1         1     1      1      1      3      1    45    660.
## 9        175         1         1     1      1      1      3      1    41    479.
## 10        60         1         1     1      1      1      3      1    33    481.
## # i 5,616 more rows
## # i 20 more variables: totmed <dbl>, injdes <dbl>, benefit <dbl>, ky <dbl>,
## #   mi <dbl>, log_duration <dbl>, afhigh <dbl>, lprewage <dbl>, lage <dbl>,
## #   ltotmed <dbl>, head <dbl>, neck <dbl>, upextr <dbl>, trunk <dbl>,
## #   lowback <dbl>, lowextr <dbl>, occdis <dbl>, manuf <dbl>, construc <dbl>,
## #   highlpre <dbl>
```

## Converting “industry” and “injury type” to categories/factors

```
df <- injury %>% mutate(indust = as.factor(indust), injtype = as.factor(injtype))
```

1. Calculating the policy effect on duration, without running any regression, here we see the mean duration of the weeks for both treated and control group before and after 1980 (treatement)

```
difr <- df %>% group_by(after_1980, highearn) %>% summarize(mean.duration = mean(duration))
```

```
## 'summarise()' has grouped output by 'after_1980'. You can override using the
## '.groups' argument.
```

```
print(difr)
```

```
## # A tibble: 4 x 3
## # Groups:   after_1980 [2]
##   after_1980 highearn mean.duration
##   <dbl>      <dbl>    <dbl>
## 1         0         0         6.27
## 2         0         1        11.2
## 3         1         0         7.04
## 4         1         1        12.9
```

[after\_1980(0);highearn(0)]: pre-treatment control group mean\_duration : 6.47 weeks

[after\_1980(0);highearn(1)]: pre-treatment treatment group mean\_duration : 11.76 weeks

[after\_1980(1);highearn(0)]: post-treatment control group mean\_duration : 7.03 weeks

[after\_1980(1);highearn(1)]: post-treatment treatment group mean\_duration : 12.89 weeks

policy\_effect = [(avg\_duration(post-treatment treated group)-avg\_duration(pre-treatment treated group)]-[avg\_duration(post-treatment control group)-avg\_duration(pre-treatment control group)]

policy\_effect = [avg\_duration of treated group (POST-PRE)]-[avg\_duration of control group(POST-PRE)]

```
pre_treatment_treated_group <- difr %>%  
  filter(after_1980 == 0, highearn == 1) %>%  
  pull(mean.duration)
```

```
pre_treatment_control_group<- difr %>%  
  filter(after_1980 == 0, highearn == 0) %>%  
  pull(mean.duration)
```

```
post_treatment_treated_group <- difr %>%  
  filter(after_1980 == 1, highearn == 1) %>%  
  pull(mean.duration)
```

```
post_treatment_control_group <- difr %>%  
  filter(after_1980 == 1, highearn == 0) %>%  
  pull(mean.duration)
```

```
treatment_group_before_after <- post_treatment_treated_group - pre_treatment_treated_group  
control_group_before_after <- post_treatment_control_group - pre_treatment_control_group
```

## Policy Effect(DiD Estimate:)

```
policy_effect_nive <- treatment_group_before_after - control_group_before_after  
print(policy_effect_nive)
```

```
## [1] 0.9512506
```

2. Calculating the policy effect on duration, without running any regression, here we see the mean log\_duration in weeks for both treated and control group before and after 1980(treatemnt)

```
difr_log <- df %>% group_by(after_1980,highearn) %>% summarize(mean.log_duration = mean(log_duration))
```

## 'summarise()' has grouped output by 'after\_1980'. You can override using the  
## '.groups' argument.

```
head(difr_log)
```

```
## # A tibble: 4 x 3
## # Groups:   after_1980 [2]
##   after_1980 highearn mean.log_duration
##         <dbl>     <dbl>           <dbl>
## 1         0         0             1.13
## 2         0         1             1.38
## 3         1         0             1.13
## 4         1         1             1.58
```

[after\_1980(0);highearn(0)]: pre-treatment control group mean\_log\_duration : 1.12 weeks

[after\_1980(0);highearn(1)]: pre-treatment treated group mean\_log\_duration : 1.38 weeks

[after\_1980(1);highearn(0)]: post-treatment control group mean\_log\_duration : 1.13 weeks

[after\_1980(1);highearn(1)]: post-treatment treated group mean\_log\_duration : 1.58 weeks

policy\_effect = [(log\_avg\_duration(post-treatment treated group)-log\_avg\_duration(pre-treatment treated group))-log\_avg\_duration(post-treatment control group)-log\_avg\_duration(pre-treatment control group)]

policy\_effect = [log\_avg\_duration of treated group (POST-PRE)]-[log\_avg\_duration of control group(POST-PRE)]

```
pre_treatment_treated_group_log <- difr_log %>%
  filter(after_1980 == 0, highearn == 1) %>%
  pull(mean.log_duration)
```

```
pre_treatment_control_group_log<- difr_log %>%
  filter(after_1980 == 0, highearn == 0) %>%
  pull(mean.log_duration)
```

```
post_treatment_treated_group_log <- difr_log %>%
  filter(after_1980 == 1, highearn == 1) %>%
  pull(mean.log_duration)
```

```
post_treatment_control_group_log <- difr_log %>%
  filter(after_1980 == 1, highearn == 0) %>%
  pull(mean.log_duration)
```

```
log_treatment_group_before_after <- post_treatment_treated_group_log -
  pre_treatment_treated_group_log
log_control_group_before_after <- post_treatment_control_group_log -
  pre_treatment_control_group_log
```

## Policy Effect log(DiD log Estimate:)

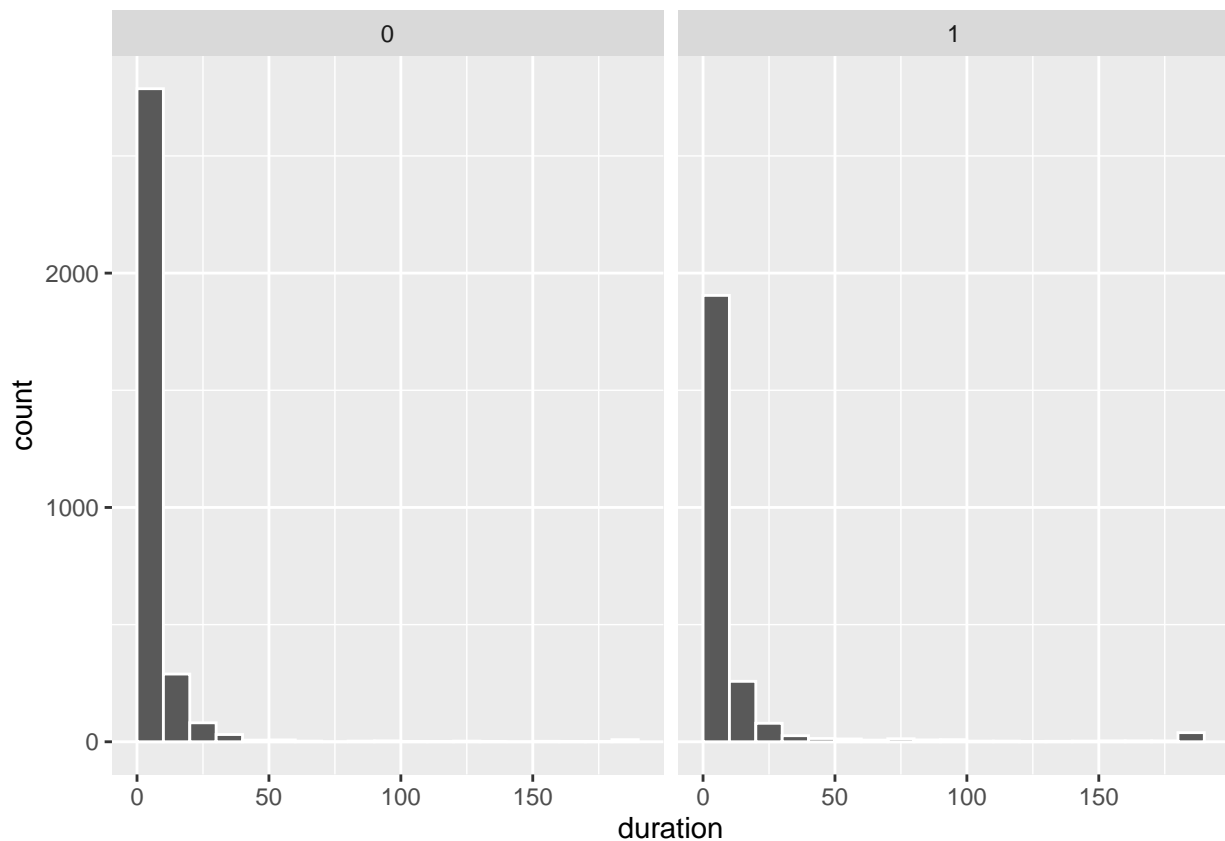
```
policy_effect_log <- log_treatment_group_before_after - log_control_group_before_after  
print(policy_effect_log)
```

```
## [1] 0.1906012
```

## Plotting

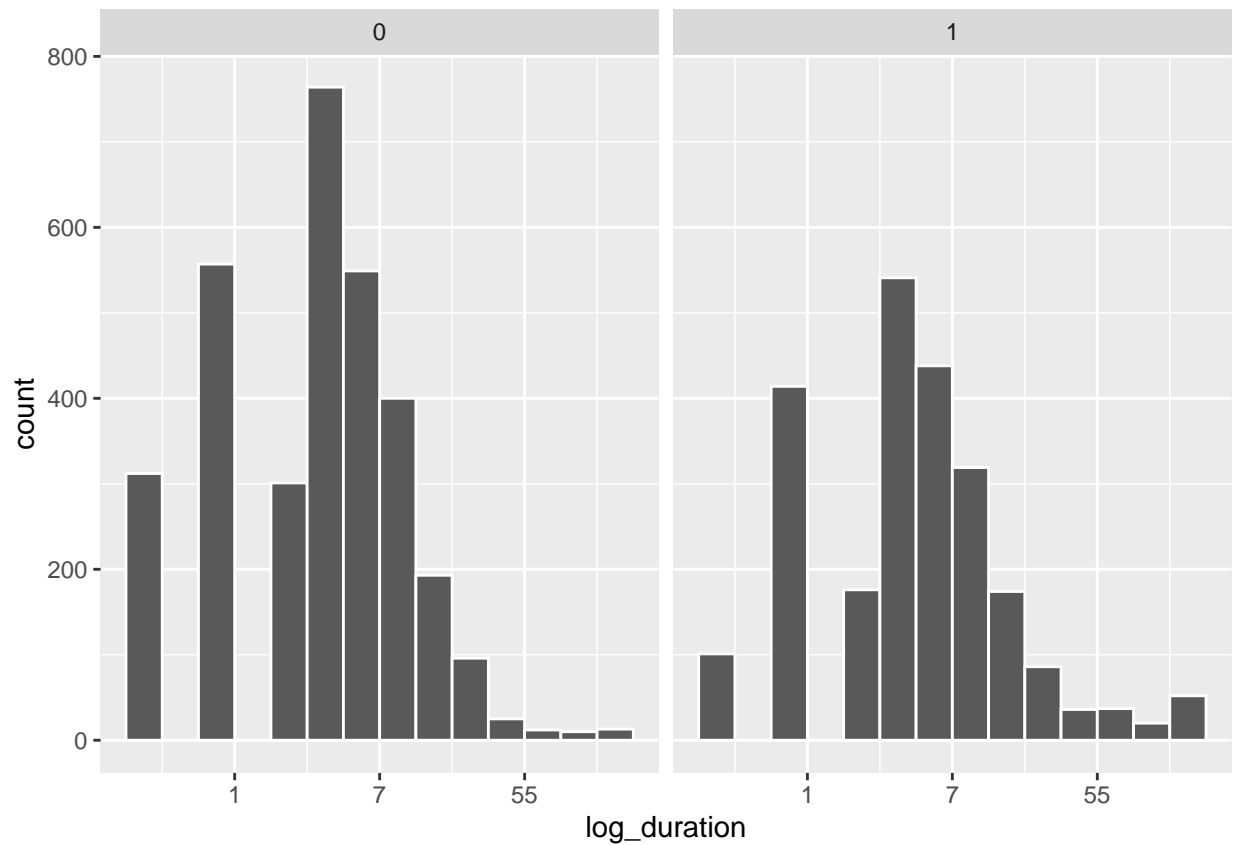
### Distribution of Duration by Category

```
ggplot(data = df, aes(x = duration)) +  
  geom_histogram(binwidth = 10, color = "white", boundary = 0) +  
  facet_wrap(~ highearn)
```



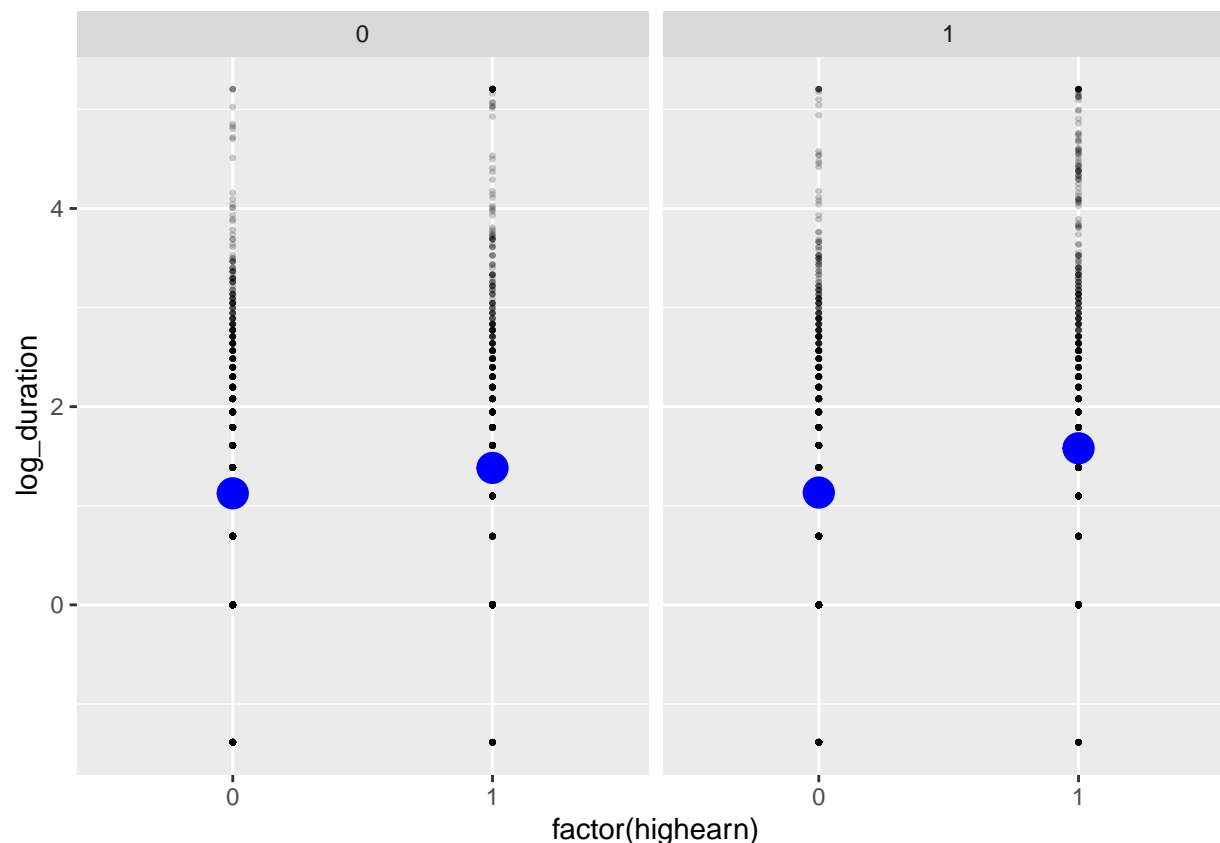
### Plotting log duration based on category

```
ggplot(data = df, mapping = aes(x = log_duration)) +  
  geom_histogram(binwidth = 0.5, color = "white", boundary = 0) +  
  scale_x_continuous(labels = trans_format("exp", format = round)) +  
  facet_wrap(~ highearn)
```



Here we just calculate Mean

```
ggplot(df, aes(x = factor(highearn), y = log_duration)) +  
  geom_point(size = 0.5, alpha = 0.2) +  
  stat_summary(geom = "point", fun = "mean", size = 5, color = "blue") +  
  facet_wrap(vars(after_1980))
```



### 3. Basic Regression analysis to calculate the estimates without any control

#### Basic Regression model

$$duration = \beta_0 + \beta_1 after_{1980} + \beta_2 highearn + \delta afterchange \cdot highearn + u$$

```
basic_model_1 <- lm(duration ~ after_1980 + highearn + after_1980*highearn, data = df)
tidy(basic_model_1)
```

```
## # A tibble: 4 x 5
##   term                estimate std.error statistic  p.value
##   <chr>              <dbl>    <dbl>    <dbl>    <dbl>
## 1 (Intercept)        6.27      0.523     12.0 9.51e-33
## 2 after_1980         0.766     0.761      1.01 3.14e- 1
## 3 highearn           4.91      0.807      6.08 1.30e- 9
## 4 after_1980:highearn 0.951     1.17      0.816 4.14e- 1
```

The notation  $\delta_1 = 0.9513$ , indicates that the policy change might have increased the duration of benefits for high-income workers by about 0.9513 weeks more than for low-income workers. However, this effect is not statistically significant ( $p=0.414$ ). The coefficient on `after_1980` is small 0.7658 and statistically insignificant which means the increase in the earnings cap has no effect on duration for low-income workers.

$$\log(duration) = \beta_0 + \beta_1 after_{1980} + \beta_2 highearn + \delta_1 afterchange \cdot highearn + u$$

```
basic_model_2 <- lm(log(duration) ~ after_1980 + highearn + after_1980*highearn, data = df)
tidy(basic_model_2)
```

```
## # A tibble: 4 x 5
##   term                estimate std.error statistic    p.value
##   <chr>              <dbl>    <dbl>    <dbl>    <dbl>
## 1 (Intercept)        1.13      0.0307    36.6  1.62e-263
## 2 after_1980         0.00766   0.0447     0.171 8.64e- 1
## 3 highearn           0.256     0.0474     5.41 6.72e- 8
## 4 after_1980:highearn 0.191     0.0685     2.78 5.42e- 3
```

The notation  $\delta$  signifies that the average duration of workers' compensation among high earners rose approximately by 19.06% due to the increased earnings cap. The coefficient on after\_1980 is small 0.007 and statistically insignificant which means the increase in the earnings cap has no effect on duration for low-income workers.

#### 4. regression adjustment procedure to evaluate the impact of policy change on the duration, and log\_duration.

##### Regression Adjustment Model 1

$$\begin{aligned} \text{duration} = & \beta_0 + \beta_1 \text{after1980} + \beta_2 \text{highearn} + \delta_1 \text{afterchange} \cdot \text{highearn} \\ & + \gamma_1 \text{male} + \gamma_2 \text{married} + \gamma_3 \text{age}^2 + \gamma_4 \text{hosp} + \gamma_5 \text{indust} \\ & + \gamma_6 \text{injtype} + \gamma_7 \text{lprewage} + u \end{aligned}$$

### Creating a new column age squared

```
df <- df %>% mutate(age_squared = age^2)
```

```
adv_model_1 <- lm(duration ~ after_1980 + highearn +
after_1980*highearn + male + married + age_squared +
hosp + indust + injtype + lprewage, data =df)
print(summary(adv_model_1))
```

```
##
## Call:
## lm(formula = duration ~ after_1980 + highearn + after_1980 *
##   highearn + male + married + age_squared + hosp + indust +
##   injtype + lprewage, data = df)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -27.603  -6.471  -2.251   1.503  181.142
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  -1.863e+01  7.171e+00  -2.598  0.009390 **
## after_1980     1.101e+00  7.031e-01   1.566  0.117441
## highearn     -1.100e+00  1.517e+00  -0.726  0.468082
## male         -9.995e-01  7.165e-01  -1.395  0.163113
```



```
## married          1.581e+00  6.286e-01  2.515 0.011938 *
## age_squared      3.412e-04  2.820e-04  1.210 0.226412
## hosp            1.346e+01  6.296e-01 21.379 < 2e-16 ***
## indust2         7.637e-01  9.210e-01  0.829 0.407036
## indust3         2.781e+00  6.440e-01  4.318 1.6e-05 ***
## injtype2         5.904e+00  2.446e+00  2.414 0.015813 *
## injtype3         1.104e+00  1.454e+00  0.760 0.447436
## injtype4        -5.895e-01  1.580e+00 -0.373 0.709034
## injtype5         3.461e+00  1.454e+00  2.381 0.017308 *
## injtype6         3.725e-01  1.469e+00  0.254 0.799757
## injtype7         1.071e+01  3.241e+00  3.303 0.000962 ***
## injtype8        -1.464e-02  2.023e+00 -0.007 0.994227
## lprewage         3.454e+00  1.362e+00  2.536 0.011226 *
## after_1980:highearn 1.790e+00  1.088e+00  1.644 0.100192
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 19.56 on 5329 degrees of freedom
## (279 observations deleted due to missingness)
## Multiple R-squared:  0.1102, Adjusted R-squared:  0.1074
## F-statistic: 38.83 on 17 and 5329 DF, p-value: < 2.2e-16
```

After controlling for all the X's (male ,married , age\_squared ,hosp,indust,injtype,lprewage), the coefficient of interaction term comes out to be 1.79 which is insignificant.

## Regression Adjustment Model 2

$$\begin{aligned} \log(\text{duration}) = & \beta_0 + \beta_1 \text{after1980} + \beta_2 \text{highearn} + \delta_1 \text{afterchange} \cdot \text{highearn} \\ & + \gamma_1 \text{male} + \gamma_2 \text{married} + \gamma_3 \text{age}^2 + \gamma_4 \text{hosp} + \gamma_5 \text{indust} \\ & + \gamma_6 \text{injtype} + \gamma_7 \text{lprewage} + u \end{aligned}$$

After controlling for all the X's (male ,married , age\_squared ,hosp,indust, injtype,lprewage), the coefficient of interaction term comes out to be 1.69 which is significant , which means that the he average duration of workers' compensation among high earners rose approximately by 16.07% due to the increased earnings cap.

```
adv_model_2 <- lm(log_duration ~ after_1980 + highearn +
after_1980*highearn + male + married + age_squared + hosp
+ indust + injtype + lprewage, data =df)
print(summary(adv_model_2))
```

```
##
## Call:
## lm(formula = log_duration ~ after_1980 + highearn + after_1980 *
##     highearn + male + married + age_squared + hosp + indust +
##     injtype + lprewage, data = df)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.0550 -0.7754  0.0930  0.7318  4.3620
##
## Coefficients:
```

```
##               Estimate Std. Error t value Pr(>|t|)
## (Intercept)    -1.460e+00  4.216e-01  -3.463 0.000538 ***
## after_1980      4.937e-02  4.134e-02   1.194 0.232437
## highearn       -1.532e-01  8.916e-02  -1.718 0.085832 .
## male           -9.264e-02  4.213e-02  -2.199 0.027931 *
## married         6.771e-02  3.696e-02   1.832 0.067005 .
## age_squared     7.357e-05  1.658e-05   4.437 9.32e-06 ***
## hosp           1.131e+00  3.702e-02  30.546 < 2e-16 ***
## indust2         1.831e-01  5.415e-02   3.380 0.000729 ***
## indust3         1.629e-01  3.787e-02   4.301 1.73e-05 ***
## injtype2        9.364e-01  1.438e-01   6.512 8.09e-11 ***
## injtype3        6.362e-01  8.547e-02   7.443 1.14e-13 ***
## injtype4        5.571e-01  9.288e-02   5.998 2.13e-09 ***
## injtype5        6.446e-01  8.547e-02   7.542 5.42e-14 ***
## injtype6        6.163e-01  8.634e-02   7.137 1.08e-12 ***
## injtype7        9.966e-01  1.906e-01   5.229 1.77e-07 ***
## injtype8        4.361e-01  1.190e-01   3.666 0.000248 ***
## lprewage        2.943e-01  8.005e-02   3.676 0.000239 ***
## after_1980:highearn 1.697e-01  6.400e-02   2.651 0.008043 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.15 on 5329 degrees of freedom
## (279 observations deleted due to missingness)
## Multiple R-squared:  0.1893, Adjusted R-squared:  0.1867
## F-statistic: 73.17 on 17 and 5329 DF, p-value: < 2.2e-16
```

Combining the data set to contain (ky==0) We here combine the data set including the state of michigan state sample .

```
dataset_combined <- dataset %>% rename(duration = durat,
log_duration = ldurat, after_1980 = afchnge)%>%
mutate(indust = as.factor(indust), injtype = as.factor(injtype))%>%
mutate(age_squared = age^2)
```

## 5. Robustness check

$$\begin{aligned}\log(\text{duration}) = & \beta_0 + \beta_1 \text{after1980} + \beta_2 \text{highearn} + \delta_1 \text{afterchange} \cdot \text{highearn} \\ & + \gamma_1 \text{male} + \gamma_2 \text{married} + \gamma_3 \text{age}^2 + \gamma_4 \text{hosp} + \gamma_5 \text{indust} \\ & + \gamma_6 \text{injtype} + \gamma_7 \text{lprewage} + \delta_2 \text{after1980} \cdot \text{ky} + \delta_3 \text{ky} + u\end{aligned}$$

```
robust_model1 <- lm(log_duration ~ after_1980 + highearn +
after_1980*highearn + male + married + age_squared + hosp
+ indust + injtype + lprewage + ky +ky:after_1980, data =dataset_combined)
print(summary(robust_model1))
```

```
##
## Call:
## lm(formula = log_duration ~ after_1980 + highearn + after_1980 *
##     highearn + male + married + age_squared + hosp + indust +
```

```
##      injtype + lprewage + ky + ky:after_1980, data = dataset_combined)
##
## Residuals:
##      Min        1Q    Median        3Q        Max
## -4.7673 -0.7489  0.0699  0.7357  4.2709
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   -1.175e+00  4.514e-01  -2.603  0.009270 **
## after_1980      8.059e-02  6.422e-02   1.255  0.209597
## highearn     -1.326e-01  8.690e-02  -1.526  0.126941
## male          -1.438e-01  3.852e-02  -3.732  0.000192 ***
## married        4.594e-02  3.324e-02   1.382  0.167018
## age_squared    8.785e-05  1.501e-05   5.853  5.04e-09 ***
## hosp          1.101e+00  3.346e-02  32.910 < 2e-16 ***
## indust2        2.640e-01  4.724e-02   5.588  2.38e-08 ***
## indust3        1.551e-01  3.370e-02   4.603  4.24e-06 ***
## injtype2        8.898e-01  1.333e-01   6.674  2.68e-11 ***
## injtype3        6.550e-01  7.936e-02   8.254 < 2e-16 ***
## injtype4        5.969e-01  8.558e-02   6.975  3.35e-12 ***
## injtype5        6.315e-01  7.956e-02   7.938  2.39e-15 ***
## injtype6        6.068e-01  8.025e-02   7.562  4.49e-14 ***
## injtype7        1.115e+00  1.612e-01   6.918  5.00e-12 ***
## injtype8        5.605e-01  1.080e-01   5.190  2.16e-07 ***
## lprewage        2.758e-01  7.919e-02   3.483  0.000499 ***
## ky             -1.628e-01  5.932e-02  -2.745  0.006064 **
## after_1980:highearn 1.674e-01  5.887e-02   2.843  0.004476 **
## after_1980:ky     -2.950e-02  6.997e-02  -0.422  0.673382
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.179 on 6802 degrees of freedom
## (328 observations deleted due to missingness)
## Multiple R-squared:  0.1832, Adjusted R-squared:  0.1809
## F-statistic: 80.27 on 19 and 6802 DF, p-value: < 2.2e-16
```

## DDD

$$\log(\text{duration}) = \beta_0 + \beta_1 * ky + \beta_2 * after1980_t + \beta_3 * highEarn_i + \gamma_1(ky_s * after1980_t) + \gamma_2(ky * highearn) + \gamma_2(after1980 * highearn) + \delta_1(ky * after1980 * highearn) + \psi * X + u$$

Considering the entire dataset including Michigan, we need to account for state-specific and income-specific time trends that might influence compensation benefits. To do this, we use Triple Differences, which involves comparing the differences between Kentucky and Michigan data to eliminate these confounding factors.

```
DDD= lm( duration ~ ky + after_1980 + highearn + ky*after_1980 + ky*highearn
        + after_1980*highearn + ky*after_1980*highearn + male + married
        + age_squared + hosp
+ indust + injtype + lprewage, data=dataset_combined)
summary(DDD)
```

```
##
```

```
## Call:
## lm(formula = duration ~ ky + after_1980 + highearn + ky * after_1980 +
##     ky * highearn + after_1980 * highearn + ky * after_1980 *
##     highearn + male + married + age_squared + hosp + indust +
##     injtype + lprewage, data = dataset_combined)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -41.970  -7.650  -2.916   1.287  179.648
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    -1.483e+01  8.685e+00  -1.708  0.087693 .
## ky             -3.296e+00  1.295e+00  -2.545  0.010941 *
## after_1980      2.708e+00  1.421e+00   1.906  0.056712 .
## highearn        5.330e-01  2.303e+00   0.231  0.816988
## male           -1.865e+00  7.407e-01  -2.518  0.011813 *
## married         8.539e-01  6.387e-01   1.337  0.181296
## age_squared     8.690e-04  2.888e-04   3.009  0.002634 **
## hosp            1.380e+01  6.430e-01  21.468 < 2e-16 ***
## indust2         2.311e+00  9.102e-01   2.539  0.011141 *
## indust3         2.485e+00  6.476e-01   3.838  0.000125 ***
## injtype2        6.412e+00  2.562e+00   2.503  0.012331 *
## injtype3        1.889e+00  1.525e+00   1.239  0.215275
## injtype4        5.805e-01  1.644e+00   0.353  0.723992
## injtype5        4.376e+00  1.528e+00   2.863  0.004204 **
## injtype6        8.398e-01  1.542e+00   0.545  0.585941
## injtype7        1.297e+01  3.101e+00   4.181  2.94e-05 ***
## injtype8        2.665e+00  2.075e+00   1.284  0.199142
## lprewage        3.250e+00  1.522e+00   2.136  0.032710 *
## ky:after_1980   -1.636e+00  1.638e+00  -0.999  0.317866
## ky:highearn     -1.269e+00  1.975e+00  -0.643  0.520521
## after_1980:highearn  5.057e-01  2.576e+00   0.196  0.844378
## ky:after_1980:highearn 1.361e+00  2.868e+00   0.474  0.635199
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 22.64 on 6800 degrees of freedom
## (328 observations deleted due to missingness)
## Multiple R-squared:  0.09577,    Adjusted R-squared:  0.09298
## F-statistic: 34.3 on 21 and 6800 DF,  p-value: < 2.2e-16
```

here we see that the triple interaction term comes out to be 1.36 positive but insignificant, when controlling for all the other possible X's.

## 6. Generating Graph

```
graph_data <- dataset_combined %>%
  mutate(highearn = factor(highearn, labels = c("Low earner", "High earner")),
         after_1980 = factor(after_1980, labels = c("Before 1980", "After 1980"))) %>%
  group_by(highearn, after_1980) %>%
```

```

summarize(mean_duration = mean(log_duration),
          se_duration = sd(log_duration) / sqrt(n()),
          upper = mean_duration + (1.96 * se_duration),
          lower = mean_duration - (1.96 * se_duration))

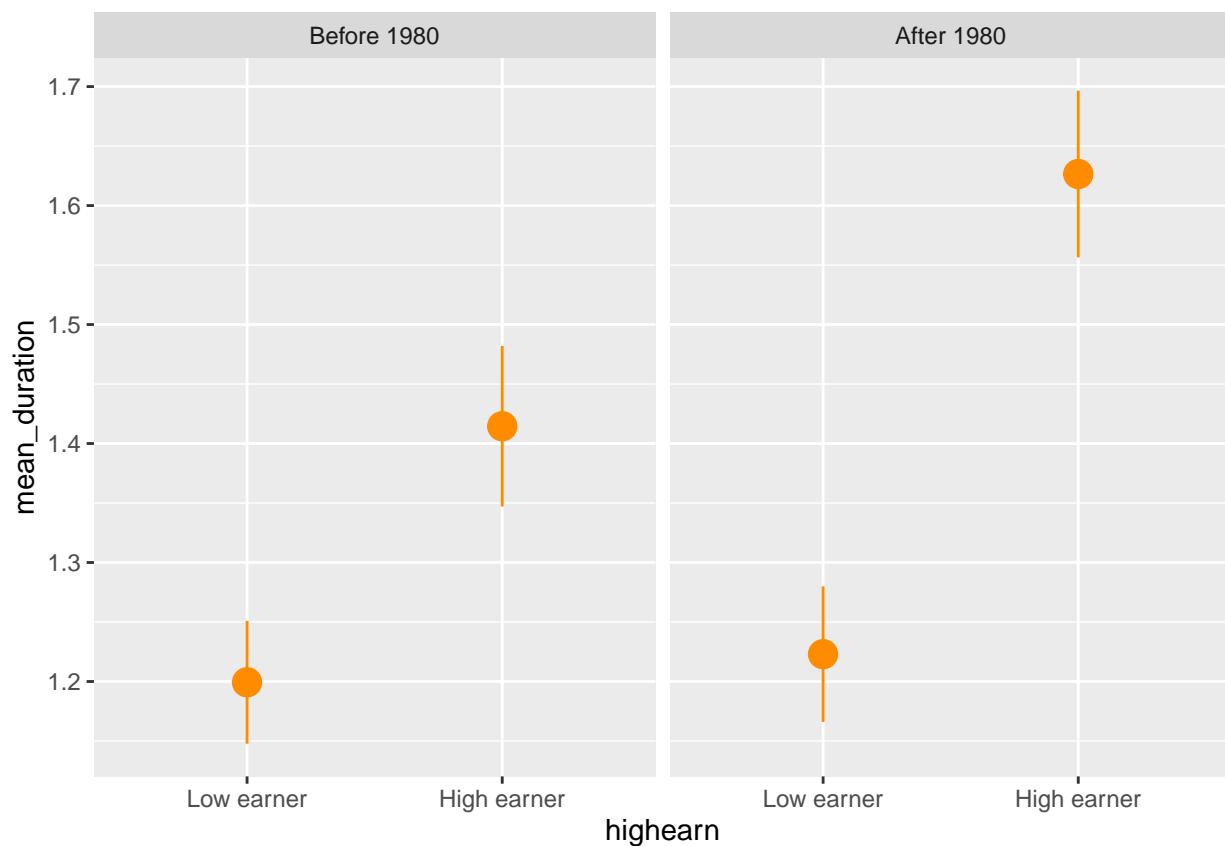
```

## 'summarise()' has grouped output by 'highearn'. You can override using the  
## '.groups' argument.

```

ggplot(graph_data, aes(x = highearn, y = mean_duration)) +
  geom_pointrange(aes(ymin = lower, ymax = upper),
                 color = "darkorange", size = 1) +
  facet_wrap(vars(after_1980))

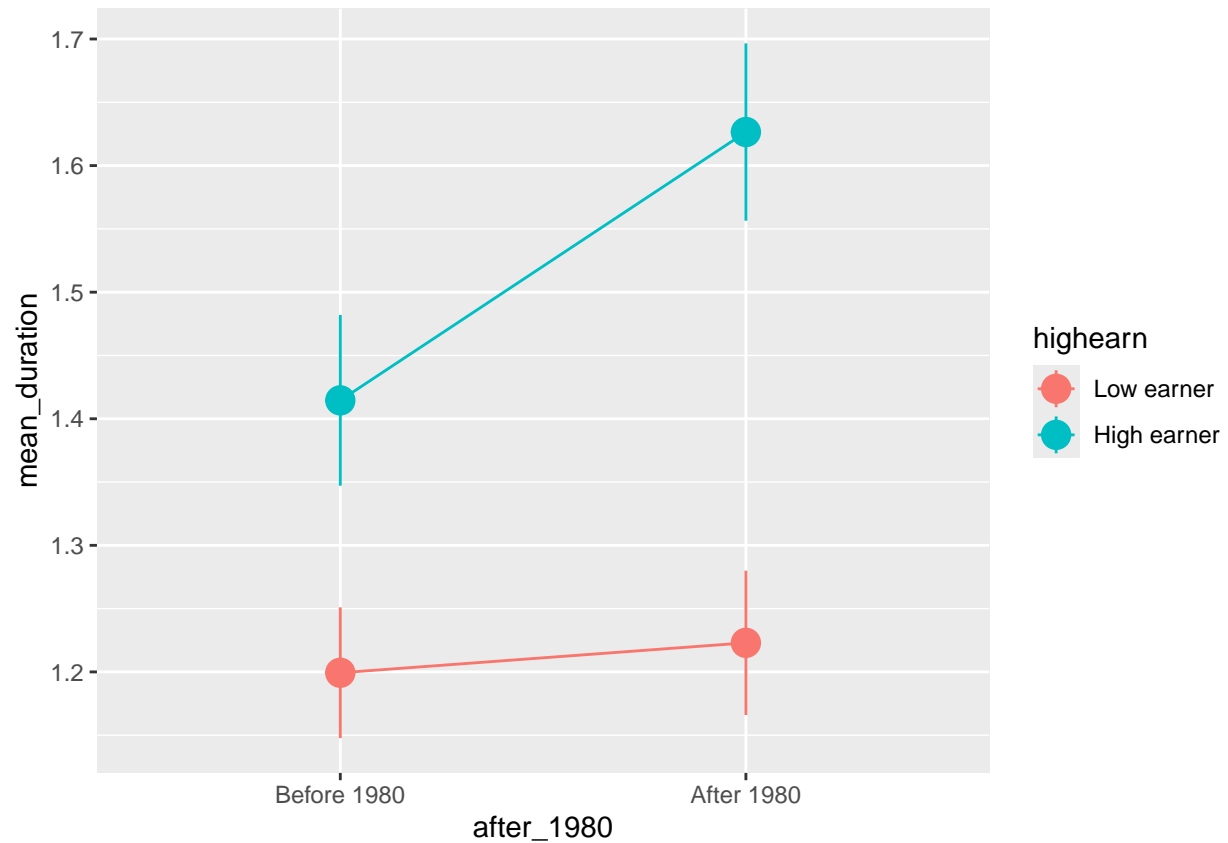
```



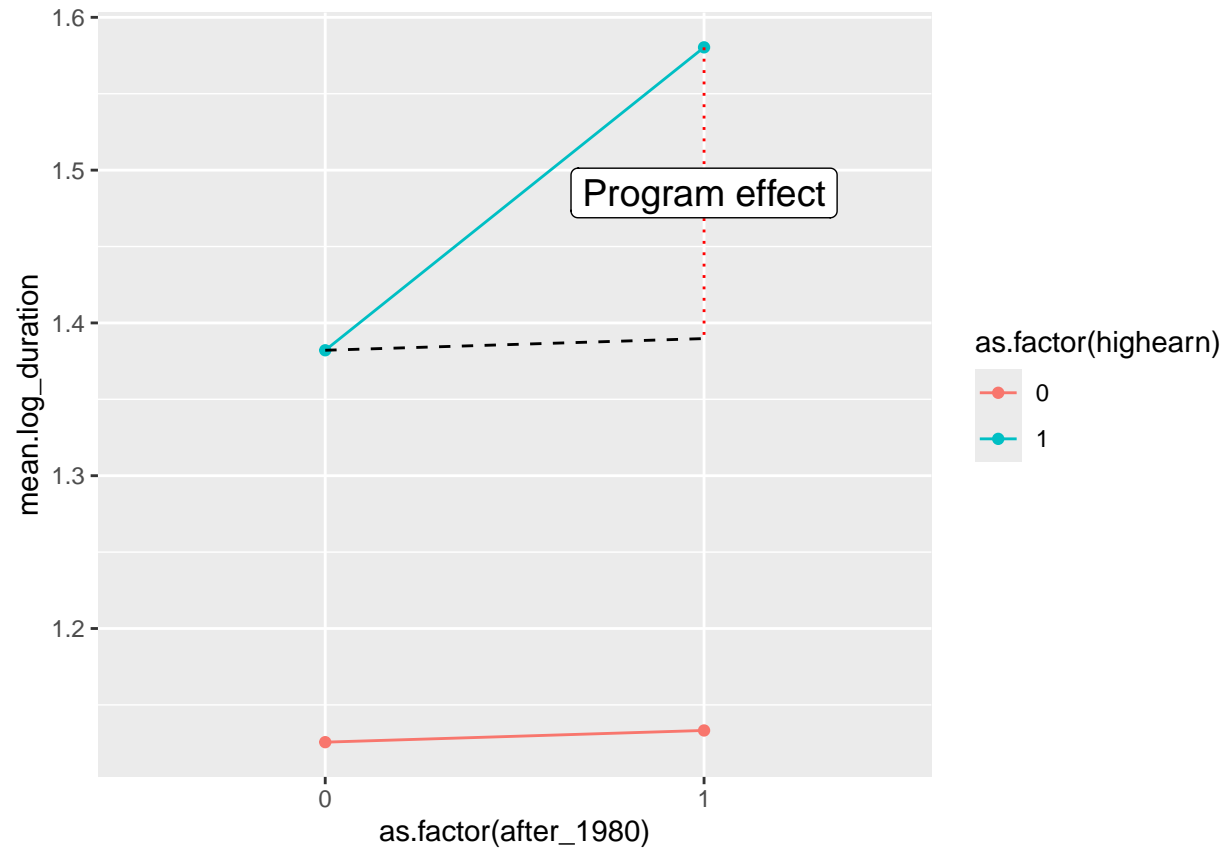
```

ggplot(graph_data, aes(x = after_1980, y = mean_duration, color = highearn)) +
  geom_pointrange(aes(ymin = lower, ymax = upper), size = 1) +
  # The group = highearn here makes it so the lines go across categories
  geom_line(aes(group = highearn))

```



```
ggplot(difr_log, aes(x = as.factor(after_1980),
  y = mean.log_duration,
  color = as.factor(highearn)))+
  geom_point() +
  geom_line(aes(group = as.factor(highearn))) +
  annotate(geom = "segment", x = "0", xend = "1",
    y = pre_treatment_treated_group_log, yend =
      post_treatment_treated_group_log - policy_effect_log,
    linetype = "dashed", color = "black") +
  annotate(geom = "segment", x = "1", xend = "1",
    y = post_treatment_treated_group_log, yend =
      post_treatment_treated_group_log - policy_effect_log,
    linetype = "dotted", color = "red") +
  annotate(geom = "label", x = "1", y = post_treatment_treated_group_log
    - (policy_effect_log / 2),
    label = "Program effect", size = 5)
```



```
modelsummary(list(basic_model_2,adv_model_2,DDD,robust_model1),output='latex')
```

	(1)	(2)	(3)	(4)
(Intercept)	1.126 (0.031)	-1.460 (0.422)	-14.834 (8.685)	-1.175 (0.451)
after_1980	0.008 (0.045)	0.049 (0.041)	2.708 (1.421)	0.081 (0.064)
highearn	0.256 (0.047)	-0.153 (0.089)	0.533 (2.303)	-0.133 (0.087)
after_1980 $\times$ highearn	0.191 (0.069)	0.170 (0.064)	0.506 (2.576)	0.167 (0.059)
male		-0.093 (0.042)	-1.865 (0.741)	-0.144 (0.039)
married		0.068 (0.037)	0.854 (0.639)	0.046 (0.033)
age_squared		0.000 (0.000)	0.001 (0.000)	0.000 (0.000)
hosp		1.131 (0.037)	13.804 (0.643)	1.101 (0.033)
indust2		0.183 (0.054)	2.311 (0.910)	0.264 (0.047)
indust3		0.163 (0.038)	2.485 (0.648)	0.155 (0.034)
injtype2		0.936 (0.144)	6.412 (2.562)	0.890 (0.133)
injtype3		0.636 (0.085)	1.889 (1.525)	0.655 (0.079)
injtype4		0.557 (0.093)	0.581 (1.644)	0.597 (0.086)
injtype5		0.645 (0.085)	4.376 (1.528)	0.631 (0.080)
injtype6		0.616 (0.086)	0.840 (1.542)	0.607 (0.080)
injtype7		0.997 (0.191)	12.966 (3.101)	1.115 (0.161)
injtype8		0.436 (0.119)	2.665 (2.075)	0.561 (0.108)
lprewage		0.294 (0.080)	3.250 (1.522)	0.276 (0.079)
ky			-3.296 (1.295)	-0.163 (0.059)
ky $\times$ after_1980			-1.636 (1.638)	
ky $\times$ highearn			-1.269 (1.975)	
ky $\times$ after_1980 $\times$ highearn			1.361 (2.868)	