

Multivariate Multiple Imputation

Utrecht University Winter School: Missing Data in R



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Outline

Flavors of MI

Fully Conditional Specification

Joint Modeling



Joint Modeling vs. Fully Conditional Specification

When imputing with *Joint Modeling* (JM) approaches, the missing data are replaced by samples from the joint posterior predictive distribution.

- To impute X , Y , and Z , we draw:

$$X, Y, Z \sim P(X, Y, Z | \theta)$$

With *Fully Conditional Specification* (FCS), the missing data are replaced with samples from the conditional posterior predictive distribution of each incomplete variable.

- To impute X , Y , and Z , we draw:

$$X \sim P(X | Y, Z, \theta_X)$$

$$Y \sim P(Y | X, Z, \theta_Y)$$

$$Z \sim P(Z | Y, X, \theta_Z)$$



Joint Modeling: Strengths

When correctly implemented, JM approaches are guaranteed to produce *Bayesianly proper* imputations.

- A sufficient condition for *properness* is that the imputations are randomly sampled from the correctly specified joint posterior predictive distribution of the missing data.
 - This is the defining characteristic of JM methods.

When using the correct distribution, imputations produced by JM methods will be the best possible imputations.

- Unbiased parameter estimates
- Well-calibrated sampling variability



Joint Modeling: Weaknesses

JM approaches don't scale well.

- The computational burden increases with the number of incomplete variables.

JM approaches are only applicable when the joint distribution of all incomplete variables follows a known form.

- Mixes of continuous and categorical variables are difficult to accommodate.



Software Implementations



Fully Conditional Specification: Strengths

FCS scales much better than JM.

- FCS only samples from a series of univariate distributions, not large joint distributions.

FCS approaches can create imputations for variables that don't have a sensible joint distribution.

- FCS can easily treat mixes of continuous and categorical variables.



Fully Conditional Specification: Weaknesses

FCS will usually be slower than JM.

- Each variable gets its own fully parameterized distribution, even if that granularity is unnecessary.

When the incomplete variables don't have a known joint distribution, FCS doesn't have theoretical support.

- There is, however, a large degree of empirical support for the tenability of the FCS approach.
- In practice, we usually choose FCS since real data rarely arise from a known joint distribution.



Software Implementations



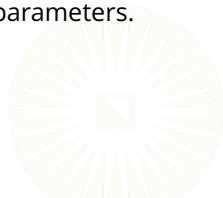
Aside: Gibbs Sampling

Up to this point, most of the models we've explored could be approximated by sampling directly from their posterior distributions.

- This won't be true with arbitrary, multivariate missing data.

To make inference regarding a multivariate distribution with multiple, interrelated, unknown parameters, we can use *Gibbs sampling*.

- Sample from the conditional distribution of each parameter, conditioning on the current best guesses of all other parameters.

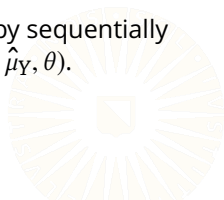


Aside: Gibbs Sampling

Suppose the following:

1. I want to make some inference about the tri-variate mean of $X, Y, Z = \mu_X, \mu_Y, \mu_Z \sim P(\mu|\theta)$
2. $P(\mu|\theta)$ is super hairy and difficult to sample
3. I can easily sample from the conditional distributions: $P(\mu_X|\hat{\mu}_Y, \hat{\mu}_Z, \theta)$, $P(\mu_Y|\hat{\mu}_X, \hat{\mu}_Z, \theta)$, and $P(\mu_Z|\hat{\mu}_X, \hat{\mu}_Y, \theta)$.

Then, I can approximate the full joint distribution $P(\mu|\theta)$ by sequentially sampling from $P(\mu_X|\hat{\mu}_Y, \hat{\mu}_Z, \theta)$, $P(\mu_Y|\hat{\mu}_X, \hat{\mu}_Z, \theta)$, and $P(\mu_Z|\hat{\mu}_X, \hat{\mu}_Y, \theta)$.



Aside: Gibbs Sampling

Starting with initial guesses of μ_Y , $\hat{\mu}_Y^{(0)}$, and μ_Z , $\hat{\mu}_Z^{(0)}$, and assuming θ is known, Gibbs sampling proceeds as follows:

$$\hat{\mu}_X^{(1)} \sim P(\mu_X | \hat{\mu}_Y^{(0)}, \hat{\mu}_Z^{(0)}, \theta)$$

$$\hat{\mu}_Y^{(1)} \sim P(\mu_Y | \hat{\mu}_X^{(1)}, \hat{\mu}_Z^{(0)}, \theta)$$

$$\hat{\mu}_Z^{(1)} \sim P(\mu_Z | \hat{\mu}_Y^{(1)}, \hat{\mu}_X^{(1)}, \theta)$$

$$\hat{\mu}_X^{(2)} \sim P(\mu_X | \hat{\mu}_Y^{(1)}, \hat{\mu}_Z^{(1)}, \theta)$$

$$\hat{\mu}_Y^{(2)} \sim P(\mu_Y | \hat{\mu}_X^{(2)}, \hat{\mu}_Z^{(1)}, \theta)$$

$$\hat{\mu}_Z^{(2)} \sim P(\mu_Z | \hat{\mu}_Y^{(2)}, \hat{\mu}_X^{(2)}, \theta)$$

\vdots



Why do we care?

Multivariate MI employs the same logic as Gibbs sampling.

- The imputations are created by conditioning on the current estimates of the imputation model parameters.
- The imputation model parameters are updated by conditioning on the most recent imputations.
- With FCS, each variable is imputed by conditioning on the most recent imputations of all other variables.

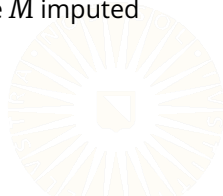


Fully Conditional Specification



Procedure: Fully Conditional Specification

1. Fill the missing data with reasonable guesses.
2. For each incomplete variable, do a single iteration of univariate Bayesian MI (e.g., as seen in the last set of slides).
 - After each variable on the data set is so treated, we've completed one iteration.
3. Repeat Step 2 many times.
4. After the imputation model parameters stabilize, save M imputed data sets.



Example: Fully Conditional Specification

```
## Simulate some data:
```

```
simData <-
```

```
  simCovData(nObs = 1000, sigma = 0.25, nVars = 4)
```

```
head(simData, 10)
```

	x1	x2	x3	x4
1	0.06313632	-1.4057704	-0.01709217	1.47929405
2	-1.31592547	1.2970920	-0.83500777	-0.44528158
3	-0.30997023	0.9782580	0.02731853	0.35507390
4	0.06927787	0.1836032	0.68794409	0.08049987
5	-0.99354894	-0.3038956	0.80918329	-1.72143555
6	0.36828016	-0.9423245	1.05155348	-0.11078496
7	1.33333163	2.3089780	1.47203000	0.85877495
8	-0.02759718	0.1714383	0.18927909	0.28627771
9	2.37929433	2.5080935	2.18344726	1.56980951
10	0.31841502	0.7886025	0.73658136	0.39445970

Example: Fully Conditional Specification

```
## Impose missing:
targets <- paste0("x", 1:3)
missData <- imposeMissData(data = simData,
                           targets = targets,
                           preds = "x4",
                           pm = 0.3,
                           types = c("low", "center", "high")
                           )
```

Example: Fully Conditional Specification

```
head(missData, 10)
```

	x1	x2	x3	x4
1	NA	-1.4057704	-0.01709217	1.47929405
2	-1.3159255	1.2970920	NA	-0.44528158
3	-0.3099702	NA	0.02731853	0.35507390
4	NA	0.1836032	0.68794409	0.08049987
5	-0.9935489	-0.3038956	0.80918329	-1.72143555
6	0.3682802	-0.9423245	NA	-0.11078496
7	1.3333316	2.3089780	NA	0.85877495
8	NA	NA	0.18927909	0.28627771
9	NA	NA	NA	1.56980951
10	NA	0.7886025	0.73658136	0.39445970

Example: Fully Conditional Specification

```
## Define iteration numbers:
nImps <- 100
nBurn <- 500
nSams <- nBurn + nImps

## Summarize missingness:
rMat <- !is.na(missData)
nObs <- colSums(rMat)
nMis <- colSums(!rMat)

## Fill the missingness with initial (bad) guesses:
mean0 <- colMeans(missData, na.rm = TRUE)
sigma0 <- cov(missData, use = "pairwise")
draws0 <- rmvnorm(nrow(missData), mean0, sigma0)

impData      <- missData
impData[!rMat] <- draws0[!rMat]
```

Example: Fully Conditional Specification

Define an elementary imputation function:

```
eif <- function(data, rVec, v) {  
  ## Get the expected betas:  
  fit <- lm(paste(v, "~ ."), data = data[rVec, ])  
  beta <- coef(fit)  
  
  ## Sample sigma:  
  sigScale <- (1 / fit$df) * crossprod(resid(fit))  
  sigmaSam <- rinvchisq(1, df = fit$df, scale = sigScale)  
  
  ## Sample beta:  
  betaVar <- sigmaSam * solve(crossprod(qr.X(fit$qr)))  
  betaSam <- rmvnorm(1, mean = beta, sigma = betaVar)  
  
  ## Return a randomly sampled imputation:  
  matrix(cbind(1, data[!rVec, ])) %*% betaSam +  
    rnorm(sum(!rVec), 0, sqrt(sigmaSam))  
}
```

Example: Fully Conditional Specification

Apply the elementary imputation function to each incomplete variable:

```
## Iterate through the FCS algorithm:
impList <- list()
for(s in 1 : nSams) {
  for(v in targets) {
    rVec      <- rMat[ , v]
    impData[!rVec, v] <- eif(impData, rVec, v)
  }

  ## If the chains are burnt-in, save imputed datasets:
  if(s > nBurn) impList[[s - nBurn]] <- impData
}

Error in matrix(cbind(1, data[!rVec, ])) %*% betaSam: requires
numeric/complex matrix/vector arguments
```

Example: Fully Conditional Specification

Analyze the multiply imputed datasets:

```
## First, our manual version:
fits1 <- lapply(impList,
               function(x) lm(x1 ~ x2 + x3, data = x)
               )
pool1 <- MIcombine(fits1)
```

Error in variances[[1]]: subscript out of bounds

```
## Do the same analysis with mice():
miceOut <- mice(data      = missData,
               m          = 100,
               method     = "norm",
               printFlag  = FALSE)
fits2 <- with(miceOut, lm(x1 ~ x2 + x3))
pool2 <- pool(fits2)
```

Example: Fully Conditional Specification

Compare approaches:

```
Error in coef(pool1): object 'pool1' not found
Error in vcov(pool1): object 'pool1' not found
Error in eval(expr, envir, enclos): object 'cf1' not found
Error in eval(expr, envir, enclos): object 'pool1' not found
Error in pt(t1, df1, lower.tail = FALSE): object 't1' not
found
Error in eval(expr, envir, enclos): object 'pool1' not found
Error in cbind(cf1, se1, t1, p1, fmi1): object 'cf1' not
found
Error in colnames(res1) <- c("Est", "SE", "t", "p", "FMI"):
object 'res1' not found
Error in eval(expr, envir, enclos): object 'res1' not found
Error in res1[flag, "p"] <- "<0.001": object 'res1' not
found
Error in rownames(res1): object 'res1' not found
Error in xtable(res1, caption = "Manual Version"): object
```

Joint Modeling



Aside: Definition of Regression Parameters

So far, we've been using the least-squares estimates of α , β , and σ^2 to parameterize our posterior distributions.

- We can also define the parameters in terms of sufficient statistics.

Given μ and Σ , we can define all of our regression moments as:

$$\begin{aligned}\beta &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y} \\ &= \text{Cov}(\mathbf{X})^{-1} \text{Cov}(\mathbf{X}, \mathbf{Y}) \\ \alpha &= \mu_Y - \beta^T \mu_X \\ \Sigma_\epsilon &= \Sigma_Y - \beta^T \Sigma_X \beta\end{aligned}$$

These definitions are crucial for JM approaches.

- Within the subset of data define by a given response pattern, the outcome variables will be entirely missing.



Multivariate Bayesian Regression

Previously, we saw examples of univariate Bayesian regression which used the following model:

$$\begin{aligned}\beta &\sim \text{MVN}\left(\hat{\beta}_{ls}, \sigma^2(\mathbf{X}^T\mathbf{X})^{-1}\right) \\ \sigma^2 &\sim \text{Inv-}\chi^2(N - P, \text{MSE})\end{aligned}$$

We can directly extend the above to the multivariate case:

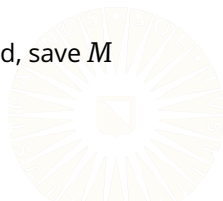
$$\begin{aligned}\Sigma^{(i)} &\sim \text{Inv-W}\left(N - 1, (N - 1)\Sigma^{(i-1)}\right) \\ \mu^{(i)} &\sim \text{MVN}\left(\mu^{(i-1)}, N^{-1}\Sigma^{(i)}\right)\end{aligned}$$

We get α , β , and Σ_ϵ via the calculations on the preceding slide



Procedure: Joint Modeling

1. Partition the incomplete data by response pattern.
 - Produce S subsets wherein each row shares the same response pattern.
2. Provide initial guesses for μ and Σ .
3. Within each subset, use the current guesses of μ and Σ to generate imputations via multivariate Bayesian regression.
4. Use the filled-in data matrix to updated the sufficient statistics.
5. Repeat Steps 3 and 4 many times.
6. After the imputation model parameters have stabilized, save M imputed data sets produced in Step 3.



Example: Joint Modeling

```
iStep <- function(data, pats, ind, p0, pars) {  
  ## Loop over non-trivial response patterns:  
  for(i in c(1 : nrow(pats))[-p0]) {  
    ## Define the current response pattern:  
    p1 <- pats[i, ]  
  
    ## Subset the data:  
    dat1 <- data[ind == i, ]  
  
    ## Replace missing data with imputations:  
    data[ind == i, !p1] <- getImps(data = dat1,  
                                   mu    = pars$mu,  
                                   sigma = pars$sigma,  
                                   p1    = p1)  
  }  
  
  ## Return the imputed data:  
  data  
}
```

Example: Joint Modeling

```
getImps <- function(data, mu, sigma, p1) {  
  ## Partition the parameter matrices:  
  mY <- matrix(mu[!p1])  
  mX <- matrix(mu[p1])  
  sY <- sigma[!p1, !p1]  
  sX <- sigma[p1, p1]  
  cXY <- sigma[p1, !p1]  
  
  ## Compute the imputation model parameters:  
  beta <- solve(sX) %*% cXY  
  alpha <- mY - t(beta) %*% mX  
  sE <- sY - t(beta) %*% sX %*% beta  
  
  ## Pull out predictors:  
  X <- as.matrix(data[, p1])  
  
  ## Generate and return the imputations:  
  n <- nrow(X)  
  matrix(1, n) %*% t(alpha) + X %*% beta + rmvnorm(n, sigma = sE)  
}
```

Example: Joint Modeling

```
pStep <- function(data) {  
  ## Update the complete-data sufficient statistics:  
  n <- nrow(data)  
  m <- colMeans(data)  
  s <- (n - 1) * cov(data)  
  
  ## Sample sigma and mu:  
  sigma <- riwish((n - 1), s)  
  mu <- rmvnorm(1, m, (sigma / n))  
  
  ## Return the updated parameters:  
  list(mu = mu, sigma = sigma)  
}
```

Example: Joint Modeling

Now that we've defined the necessary functions, do the imputation:

```
## Some preliminaries:
```

```
impData <- missData
```

```
nIter   <- 50
```

```
nImps   <- 100
```

```
## Summarize response patterns:
```

```
rMat <- !is.na(impData)
```

```
pats <- uniquecombs(rMat)
```

```
Error in uniquecombs(rMat): could not find function "uniquecombs"
```

```
ind <- attr(pats, "index")
```

```
Error in eval(expr, envir, enclos): object 'pats' not found
```

```
p0 <- which(apply(pats, 1, all))
```

```
Error in apply(pats, 1, all): object 'pats' not found
```

```
## Get starting values for the parameters:
```

```
pars <- list(mu = colMeans(impData, na.rm = TRUE),
```

```
sigma = cov(impData, use = "pairwise"))
```

Example: Joint Modeling

```
## Iterate over I- and P-Steps to generate imputations:
```

```
impList3 <- list()
for(m in 1 : nImps) {
  for(rp in 1 : nIter) {
    impData <- iStep(data = impData,
                     pats = pats,
                     ind = ind,
                     p0 = p0,
                     pars = pars)
    pars <- pStep(impData)

    if(rp == nIter) impList3[[m]] <- impData
  }
}
```

```
Error in iStep(data = impData, pats = pats, ind = ind, p0 = p0, pars =
pars): object 'pats' not found
```


Example: Joint Modeling

Do the same type of imputation with norm:

```
missData <- as.matrix(missData)
```

```
meta <- prelim.norm(missData)
```

```
Error in prelim.norm(missData): could not find function "prelim.norm"
```

```
theta0 <- em.norm(meta, showits = FALSE)
```

```
Error in em.norm(meta, showits = FALSE): could not find function "em.norm"
```

```
rngseed(235711)
```

```
Error in rngseed(235711): could not find function "rngseed"
```

```
impList4 <- list()
```

```
for(m in 1 : nImps) {
```

```
  theta1 <- da.norm(s      = meta,  
                   start = theta0,  
                   steps = nIter)
```

```
  impList4[[m]] <- imp.norm(s      = meta,  
                           theta = theta1,  
                           x      = missData)
```

Example: Joint Modeling

Analyze the multiply imputed data:

```
## Manual implementation:
fits3 <- lapply(impList3,
               function(x) lm(x1 ~ x2 + x3, data = x)
               )
pool3 <- MIcombine(fits3)
```

Error in variances[[1]]: subscript out of bounds

```
## Imputation using norm():
fits4 <- lapply(impList4,
               function(x) lm(x1 ~ x2 + x3,
                              data = as.data.frame(x)
                              )
               )
pool4 <- MIcombine(fits4)
```

Error in variances[[1]]: subscript out of bounds

Example: Joint Modeling

Compare approaches:

```
Error in coef(pool3): object 'pool3' not found
Error in vcov(pool3): object 'pool3' not found
Error in eval(expr, envir, enclos): object 'cf3' not found
Error in eval(expr, envir, enclos): object 'pool3' not found
Error in pt(t3, df3, lower.tail = FALSE): object 't3' not
found
Error in eval(expr, envir, enclos): object 'pool3' not found
Error in cbind(cf3, se3, t3, p3, fmi3): object 'cf3' not
found
Error in colnames(res3) <- c("Est", "SE", "t", "p", "FMI"):
object 'res3' not found
Error in eval(expr, envir, enclos): object 'res3' not found
Error in res3[flag, "p"] <- "<0.001": object 'res3' not
found
Error in coef(pool4): object 'pool4' not found
Error in vcov(pool4): object 'pool4' not found
```