# Missing Data Basics Utrecht University Winter School: Missing Data in R



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### Introduction

- 1. What's your name?
- 2. Where are you from/where do you work?
- 3. What type of research do you do?
- 4. What type of missing data problems do you encounter in your research?
- 5. What statistical software do you use/do you have programming experience?
- 6. What's your math background?

### Outline

Missing Data Descriptives

Missing Data Mechanisms

Missing Data Treatments



### What are Missing Data?

Missing data are empty cells in a dataset where there should be observed values.

 The missing cells correspond to true population values, but we haven't observed those values.



## What are Missing Data?

Missing data are empty cells in a dataset where there should be observed values.

 The missing cells correspond to true population values, but we haven't observed those values.

Not every empty cell is a missing datum.

- Quality-of-life ratings for dead patients in a mortality study
- Firm profitability after the company goes out of business
- Self-reported severity of menstrual cramping for men
- Empty blocks of data following "gateway" items

### A Little Notation

$$Y := An N \times P$$
 Matrix of Arbitrary Data

 $Y_{mis} := \text{The } missing \text{ part of } Y$ 

 $Y_{obs} := \text{The } observed \text{ part of } Y$ 

 $R := An N \times P$  response matrix

 $M := An N \times P$  missingness matrix

The R and M matrices are complementary.

- $r_{np} = 1$  means  $y_{np}$  is observed;  $m_{np} = 1$  means  $y_{np}$  is missing.
- $r_{np} = 0$  means  $y_{np}$  is missing;  $m_{np} = 0$  means  $y_{np}$  is observed.
- $M_p$  is the *missingness* of  $Y_p$ .

# MISSING DATA DESCRIPTIVES



### Missing Data Pattern

Missing data (or response) patterns represent unique combinations of observed and missing items.

• P items  $\Rightarrow 2^P$  possible patterns.

	Χ	Υ
1	Х	У
2	Χ	
3		У
4		

Patterns for 
$$P = 2$$

Χ	Υ	Z
Х	У	Z
Χ	У	
Χ		Z
	У	Z
Χ		
		Z
	У	
•		
	x x x	x y x y x

Patterns for P = 3

### Missing Data Pattern

The concept of a "missing data pattern" can also be used to classify the spatial arrangement of missing cells on a data set.

- Univariate
  - Missing data occur on only one variable
- Monotone
  - The proportion of complete elements, in both rows and columns, decreases when traversing the data set.
  - The observed cells can be arranged into a "staircase" pattern.
- Arbitrary
  - Missing values are "randomly" scattered throughout the data set.

### **Example Missing Data Patterns**

	Χ	Υ	Z		Χ	Υ	Z			Χ	Υ	Z
1	Х	У	Z	1	Х	У	Z		1	Х		Z
2	Χ	У	Z	2	Χ	У	Z		2	Χ	У	Z
3	Χ	У	Z	3	Χ	У	Z		3	Χ	У	Z
4	Χ	У	Z	4	Χ	У			4	Χ		Z
5	Χ	У	Z	5	Χ	У			5	Χ	У	Z
6	Χ		Z	6	Χ	У			6	Χ		Z
7	Χ		Z	7	Χ				7		У	Z
8	Χ		Z	8	Χ				8	Χ	У	Z
9	Χ		Z	9	Χ				9	Χ		
10	Χ	•	Z	10	•	•	•		10	Χ	У	•
Univa	riate	e Pat	ttern	Monotone Pattern			1	Arbit	Arbitrary Pattern			

### Nonresponse Rates

#### PROPORTION MISSING

- · The proportion of cells containing missing data
- Good early screening measure
- Should be computed for each variable, not for the entire dataset

#### ATTRITION RATE

 The proportion of participants that drop-out of a study at each measurement occasion

### Nonresponse Rates

#### PROPORTION OF COMPLETE CASES

- The proportion of observations with no missing data
- Often reported but nearly useless quantity

#### FRACTION OF MISSING INFORMATION

- Associated with an estimated parameter, not with an incomplete variable
- Like an  $R^2$  for the missing data
- Most important diagnostic value for missing data problems
- Can only be computed after treating the missing data

#### COVARIANCE COVERAGE

$$CC_{jk} = N^{-1} \sum_{n=1}^{N} r_{nj} r_{nk}$$

- The proportion of cases available to estimate a given pairwise relationship (e.g., a covariance between two variables)
- Very important to have adequate coverage of the parameters you want to estimate

#### INBOUND STATISTIC

$$I_{jk} = \frac{\sum_{n=1}^{N} (1 - r_{nj}) r_{nk}}{\sum_{n=1}^{N} (1 - r_{nj})}$$

• The proportion of missing cases in  $Y_i$  for which  $Y_k$  is observed

#### **OUTBOUND STATISTIC**

$$O_{jk} = \frac{\sum_{n=1}^{N} r_{nj} (1 - r_{nk})}{\sum_{n=1}^{N} r_{nj}}$$

• The proportion of observed cases in  $Y_i$  for which  $Y_k$  is missing

#### INflux Coefficient

$$I_{j} = \frac{\sum_{k=1}^{P} \sum_{n=1}^{N} (1 - r_{nj}) r_{nk}}{\sum_{k=1}^{P} \sum_{n=1}^{N} r_{nk}}$$

- The proportion of observed cells in Y that exists in cases for which  $Y_i$  is missing
- How well the missing values in  $Y_j$  connect to the observed values in  $Y_{-j}$

#### **OUTFLUX COEFFICIENT**

$$O_{j} = \frac{\sum_{k=1}^{P} \sum_{n=1}^{N} r_{nj} (1 - r_{nk})}{\sum_{k=1}^{P} \sum_{n=1}^{N} (1 - r_{nk})}$$

- The proportion of missing cells in Y that exists in cases for which Y<sub>i</sub> is observed
- How well the observed values in  $\mathbf{Y}_{j}$  connect to the missing values in  $\mathbf{Y}_{-j}$

### Examples

- 1. What is the coverage for cov(X, Y)?
- 2. What is the coverage for cov(W, Y)?
- **3**. What is the coverage for cov(X, Z)?
- 4. What is the outflux coefficient for *W*?
- **5.** What is the influx coefficient for *W*?

	W	Χ	Υ	Z	
1	W	Х	У		
2	W	Χ	У		
3	W	Χ	У		
4	W	Χ	У		
5	W	Χ	У		
6	W		У	Z	
7	W		У	Z	
8	W		У	Z	
9	W		У	Z	
10	W		У	Z	

# Examples

		T1	T2	Т3	T4
1. What is the percent missing at T2?	1	x1	x2	х3	x4
2 What is the attrition water at T22	2	x1	x2	х3	x4
2. What is the attrition rate at T3?	3	x1	x2	х3	x4
3. What is the inbound statistic $I_{32}$ ?	4	x1	x2	х3	
2-2	5	x1	x2	х3	
4. What is the outbound statistic $O_{42}$ ?	6	x1	x2		
[ What is the influx coefficient I 2	7	x1	x2	•	
5. What is the influx coefficient $I_3$ ?	8	x1		•	
6. What is the outflux coefficient $O_2$ ?	9	x1		•	
2.	10	x1			

# MISSING DATA MECHANISMS



### Missing Data Mechanisms

Missing Completely at Random (MCAR)

- $P(R|Y_{mis}, Y_{obs}) = P(R)$
- Missingness is unrelated to any study variables.

Missing at Random (MAR)

- $P(R|Y_{mis}, Y_{obs}) = P(R|Y_{obs})$
- Missingness is related to only the observed parts of study variables.

Missing not at Random (MNAR)

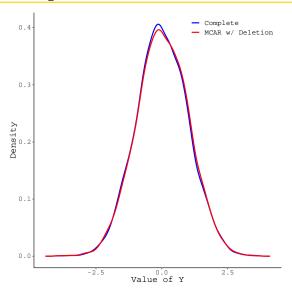
- $P(R|Y_{mis}, Y_{obs}) \neq P(R|Y_{obs})$
- Missingness is related to the unobserved parts of study variables.

### Simulate Some Toy Data

```
nObs <- 5000 # Sample Size
pm <- 0.3 # Proportion Missing
sigma \leftarrow matrix(c(1.0, 0.5, 0.3,
                   0.5, 1.0, 0.0,
                   0.3. 0.0. 1.0).
                 ncol = 3
tmp <- rmvnorm(n0bs, c(0, 0, 0), sigma)
x0 \leftarrow tmp[, 1]
y0 <- tmp[ , 2]
z0 \leftarrow tmp[, 3]
cor(y0, x0) # Check correlation between X and Y
[1] 0.5001822
```

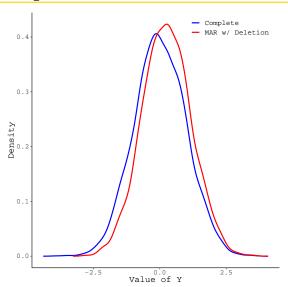
## MCAR Example

# MCAR Example



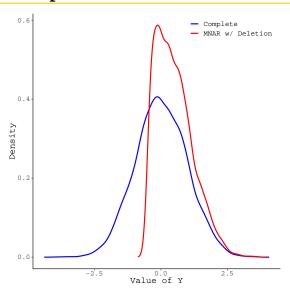
### MAR Example

# **MAR Example**



### MNAR Example

# MNAR Example



In our previous MAR example, ignoring the predictor of missingness actually produces *Indirect MNAR*.

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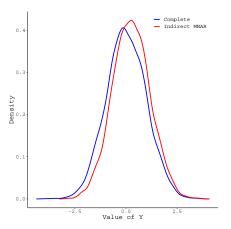
QUESTION: What happens if we ignore the predictor of missingness, but that predictor is independent of our study variables?

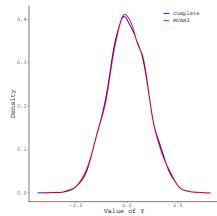
In our previous MAR example, ignoring the predictor of missingness actually produces *Indirect MNAR*.

QUESTION: What happens if we ignore the predictor of missingness, but that predictor is independent of our study variables?

ANSWER: We get back to MCAR:)

The missing data mechanisms are not simply characteristics of an incomplete dataset; we also need to account for the analysis.





### Testing the Missing Data Mechanism

We cannot fully test the MAR or MNAR assumptions.

- To do so would require knowing the values of the missing data.
- We can find observed predictors of missingness, but we can never know that we have them all.
- In practice, MAR and MNAR live on the ends of a continuum.
  - Our missing data problem exists at some unknown point along this continuum.
  - We can do a lot to nudge our problem towards the MAR side.

### Testing the Missing Data Mechanism

We can test the MCAR assumption.

- With MCAR, the missing data and the observed data should have the same distribution.
- We can test for MCAR by testing the distributions of *auxiliary* variables, **Z**.
  - Use a t-test to compare the subset of  $Z_p$  that corresponds to  $Y_{mis}$  to the subset corresponding to  $Y_{obs}$ .
  - The Little (1988) MCAR test is a multivariate version of this.

### Example

Create some toy datasets from the variables we generated above.

### T-Test Example

Test for dependence between X and  $M_Y$  in MCAR data.

```
mcarData %$% t.test(x ~ m) %>% wrap()
Welch Two Sample t-test
data: x by m
t = 0.68563, df = 2852.8, p-value = 0.493
alternative hypothesis: true difference in means between
group 0 and group 1 is not equal to 0
95 percent confidence interval:
-0.03921499 0.08138543
sample estimates:
mean in group 0 mean in group 1
    0.013908816 -0.007176408
```

### T-Test Example

Test for dependence between Z and  $M_{
m Y}$  in MCAR data.

```
mcarData %$% t.test(z ~ m) %>% wrap()
Welch Two Sample t-test
data: z by m
t = 0.38865, df = 2841.9, p-value = 0.6976
alternative hypothesis: true difference in means between
group 0 and group 1 is not equal to 0
95 percent confidence interval:
-0.04848298 0.07245421
sample estimates:
mean in group 0 mean in group 1
    0.009151786 -0.002833825
```

Test for dependence between X and  $M_Y$  in MAR data.

```
marData %$% t.test(x ~ m) %>% wrap()
Welch Two Sample t-test
data: x by m
t = 92.56, df = 3832.8, p-value < 2.2e-16
alternative hypothesis: true difference in means between
group 0 and group 1 is not equal to 0
95 percent confidence interval:
1.614203 1.684066
sample estimates:
mean in group 0 mean in group 1
      0.5023237 -1.1468112
```

Test for dependence between Z and  $M_Y$  in MAR data.

```
marData %$% t.test(z ~ m) %>% wrap()
Welch Two Sample t-test
data: z by m
t = 16.913, df = 2832.1, p-value < 2.2e-16
alternative hypothesis: true difference in means between
group 0 and group 1 is not equal to 0
95 percent confidence interval:
0.4491108 0.5669049
sample estimates:
mean in group 0 mean in group 1
      0.1579585 -0.3500494
```

Test for dependence between X and  $M_Y$  in MNAR data.

```
mnarData %$% t.test(x ~ m) %>% wrap()
Welch Two Sample t-test
data: x by m
t = 28.251, df = 2926.7, p-value < 2.2e-16
alternative hypothesis: true difference in means between
group 0 and group 1 is not equal to 0
95 percent confidence interval:
0.7439001 0.8548632
sample estimates:
mean in group 0 mean in group 1
      0.2473977 -0.5519839
```

Test for dependence between Z and  $M_{\mathrm{Y}}$  in MNAR data.

```
mnarData %$% t.test(z ~ m) %>% wrap()
Welch Two Sample t-test
data: z by m
t = -0.33313, df = 2778.5, p-value = 0.7391
alternative hypothesis: true difference in means between
group 0 and group 1 is not equal to 0
95 percent confidence interval:
-0.07145430 0.05070098
sample estimates:
mean in group 0 mean in group 1
    0.002443105 0.012819764
```

# Little (1988) MCAR Test Example

Use the Little (1988) MCAR test on MCAR data.



# Little (1988) MCAR Test Example

Use the Little (1988) MCAR test on MAR data.



## Little (1988) MCAR Test Example

Use the Little (1988) MCAR test on MNAR data.



## Logistic Regression Example

```
## Read in some data:
diabetes <- readRDS(pasteO(dataDir, "diabetes.rds"))</pre>
## Generate MAR missingness:
diabetes$m <- simLogisticMissingness0(data = diabetes,</pre>
                                      pm
                                          = 0.25,
                                      preds = c("bmi", "tc"),
                                      type = "high",
                                      stdData = TRUE)$r
## Predict the missingness using logistic regression:
fit <- diabetes %>%
    select(-glu) %>%
   glm(m ~ ., data = ., family = "binomial")
```

## Logistic Regression Example

```
partSummary(fit, 3)
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) -1.459e+01
                       4.031e+00 -3.619 0.000296
            1.205e-02 1.141e-02 1.056 0.290782
age
            2.269e-01
                       4.054e-02 5.596 2.19e-08
bmi
           -1.213e-02
                       1.147e-02 -1.057 0.290292
bp
t.c
            2.949e-02
                       2.897e-02 1.018 0.308696
1d1
            2.703e-03
                       2.625e-02 0.103 0.917986
hdl
           -5.961e-05
                       3.990e-02 -0.001 0.998808
                       2.889e-01
                                  -1.0940.274049
t.ch
           -3.160e-01
            5.588e-01
                       8.952e-01 0.624 0.532537
ltg
            2.501e-03
                       2.380e-03 1.051 0.293237
progress
sexmale
            4.336e-02
                       2.978e-01 0.146 0.884234
```

# MISSING DATA TREATMENTS



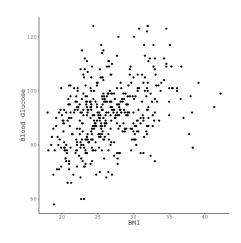
#### Listwise Deletion (Complete Case Analysis)

- Use only complete observations for the analysis
  - Very wasteful (can throw out lots of useful data)
  - Loss of statistical power

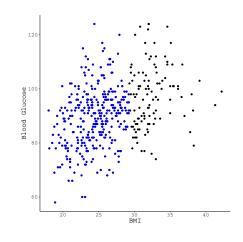
#### Pairwise Deletion (Available Case Analysis)

- Use only complete pairs of observations for analysis
  - Different samples sizes for different parameter estimates
  - Can cause computational issues

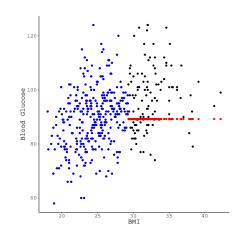
- Replace  $Y_{mis}$  with  $\bar{Y}_{obs}$ 
  - Negatively biases regression slopes and correlations
  - Attenuates measures of linear association



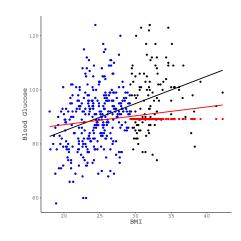
- Replace  $Y_{mis}$  with  $\bar{Y}_{obs}$ 
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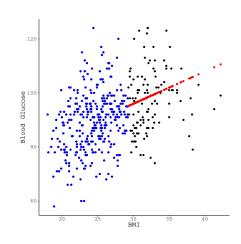
- Replace  $Y_{mis}$  with  $\bar{Y}_{obs}$ 
  - Negatively biases regression slopes and correlations
  - Attenuates measures of linear association



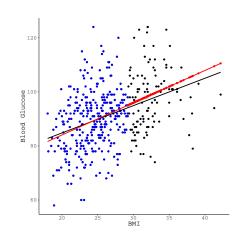
- Replace  $Y_{mis}$  with  $\bar{Y}_{obs}$ 
  - Negatively biases regression slopes and correlations
  - Attenuates measures of linear association



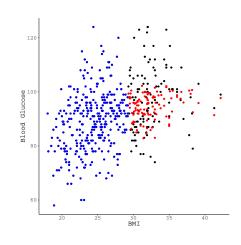
- Replace  $Y_{mis}$  with  $\widehat{Y}_{mis}$  from some regression equation
  - Positively biases regression slopes and correlations
  - Inflates measures of linear association



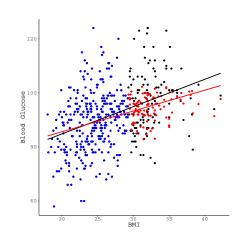
- Replace  $Y_{mis}$  with  $\widehat{Y}_{mis}$  from some regression equation
  - Positively biases regression slopes and correlations
  - Inflates measures of linear association



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  - Positively biases regression slopes and correlations
  - Inflates measures of linear association



## General Issues with Deletion-Based Methods

- Biased parameter estimates unless data are MCAR
- · Generalizability issues

## General Issues with Simple Single Imputation Methods

- Biased parameter estimates even when data are MCAR
- Attenuates variability in any treated variables

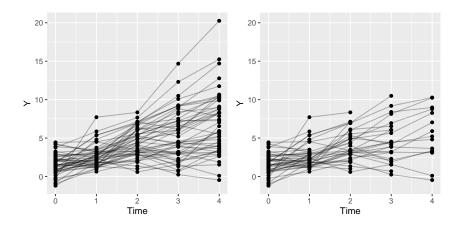
#### Averaging Available Items (Person-Mean Imputation)

- Compute aggregate scores using only available values
  - Missing data must be MCAR
  - Each item must contributes equally to the aggregate score

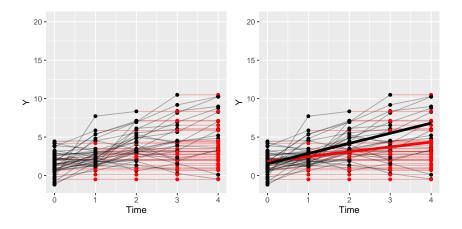
#### Last Observation Carried Forward (LOCF)

- Replace post-dropout values with the most recent observed value
  - Assume that dropouts would maintain their last known values
  - Attenuates estimates of growth/development

# LOCF



# LOCF



## OK Methods (These work in some situations)

## Stochastic Regression Imputation

- Fill  $Y_{mis}$  with  $\widehat{Y}_{mis}$  plus some random noise.
  - Produces unbiased parameter estimates and predictions
  - Computationally efficient
  - Attenuates standard errors

```
Error in '[.data.frame'(dat2, ,
c("bmi", "glu")): undefined
columns selected
Error in complete(miceS, 1):
object 'miceS' not found
Error in datS[!mVec, ] <- NA:</pre>
object 'datS' not found
Error in fortify(data): object
'datS' not found
Error: geom_point requires the
following missing aesthetics: x
and v
```

5 Pof Makes Claund prodiction

## OK Methods (These work in some situations)

```
Error in FUN(X[[i]], ...):
object 'bmi' not found
```

## Stochastic Regression Imputation

- Fill  $Y_{mis}$  with  $\widehat{Y}_{mis}$  plus some random noise.
  - Produces unbiased parameter estimates and predictions
  - Computationally efficient
  - Attenuates standard errors
  - Makes CIs and prediction intervals too narrow

## OK Methods (These work in some situations)

### Nonresponse Weighting

- Weight the observed cases to correct for nonresponse bias
  - Popular in survey research and official statistics
  - Only worth considering with Unit Nonresponse
  - Doesn't make any sense with Item Nonresponse



# **Expectation Maximization**



#### Multiple Imputation (MI)

- Replace the missing values with M plausible estimates
  - Essentially, a repeated application of stochastic regression imputation (with a particular type of regression model)
  - Produces unbiased parameter estimates and predictions
  - Produces "correct" standard errors, Cls, and prediction intervals
  - Very, very flexible
  - Computationally expensive



What happens when we apply MI to our previous MAR example?

```
## Estimate imputation model:
miceOut1 <- mice(data = data.frame(y3, x),
                m = 100,
                maxit = 1,
                method = c("norm", ""),
                printFlag = FALSE)
Error in data.frame(y3, x): object 'y3' not found
## Replace missing values with imputations:
impList1 <- list()</pre>
for(m in 1 : miceOut1$m)
   impList1[[m]] <- complete(miceOut1, m)</pre>
Error in eval(expr, envir, enclos): object 'miceOut1' not found
```

The MI-based parameter estimate looks good.

 MI produces unbiased estimates of the parameter when data are MAR.

```
Error in density.default(y): 'x' contains missing values
Error in impList1[[m]]: subscript out of bounds
Error in plot.window(...): need finite 'xlim' values
```

What about applying MI to our MNAR example?

```
## Estimate imputation model:
miceOut2 <- mice(data = data.frame(y4, x),
                m = 100,
                maxit = 1,
                method = c("norm", ""),
                printFlag = FALSE)
Error in data.frame(y4, x): object 'y4' not found
## Replace missing values with imputations:
impList2 <- list()</pre>
for(m in 1 : miceOut2$m)
   impList2[[m]] <- complete(miceOut2, m)</pre>
Error in eval(expr, envir, enclos): object 'miceOut2' not found
```

The MI-based parameter estimate is still biased.

 MI cannot correct bias in parameter estimates when data are MNAR.

```
Error in density(y4, na.rm = TRUE): object 'y4' not found
Error in impList2[[m]]: subscript out of bounds
Error in plot.window(...): need finite 'xlim' values
```

### **Bayesian Modeling**

- Treat missing values as just another parameter to be estimated
  - Models can be directly estimated in the presence of missing data
    - Essentially, runs MI behind-the-scenes during model estimation
  - The predictors of nonresponse must be included in the model, somehow
  - Computationally expensive



#### Full Information Maximum Likelihood (FIML)

- Adjust the objective function to only consider the observed parts of the data
  - Models are directly estimated in the presence of missing data
  - The predictors of nonresponse must be included in the model, somehow
  - Unless you write your own optimization program, FIML is only available for certain types of models
  - In linear regression models, FIML cannot treat missing data on predictors (if the predictors are taken as fixed)

## References

Little, R. J. A. (1988). Missing-data adjustments in large surveys. *Journal of Business & Economic Statistics*, 6(3), 287–296. doi: 10.1080/07350015.1988.10509663

