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## Simulation Methods

## 46-773

## Homework #1

```
import pandas as pd
import numpy as np
import scipy
import astplotlib.pyplot as plt
import statsmodels.api as sm
import statsmodels.atats.api as sms
import scipy.stats as stats
```

1) **Practice on random variable generation and plotting in Python and in studying the quality of various Normal generation methods.** Using Python Numpy and Matplotlib generate three different sets of observations (n = 100, n = 1,000 and n = 10,000) designed to have a standard normal distribution using the methods listed below, and present a Q-Q (normal) plot for each. Comment on the quality of the methods based on your normal plots

• (a) Use some version of the standard normal generator in Numpy

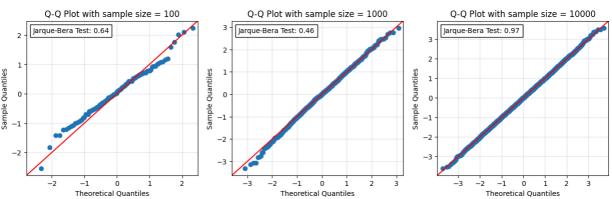
```
In | ]:
    from matplotlib.offsetbox import AnchoredText
    # Create a 0-0 plot
    fig, axes = pit.subplots(1, 3, figsize=(15, 4))
    n = 100
    measurements = np.random.normal(size = n)
    # Jarque-Bera text Chi<sup>2</sup> two-tail prob
    jbt = sms.jarque_bera(measurements) [1]
    sm.qaplot(measurements, line ='45', ax=axes[0])
    axes[0].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[0].sqt_id(alpha = 0.3)
    anchored_text = AnchoredText(ff')arque-Bera Test: {jbt:.2f}'', loc=2)
    axes[0].ada_artis(anchored_text)

    n = 1000
    measurements = np.random.normal(size = n)
    # Jarque-Bera test Chi<sup>2</sup> two-tail prob
    jbt = sms.jarque_Pera measurements[1])

sm.qaplot(measurements, line ='45', ax=axes[1])
    axes[1].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[1].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[1].set_itile(problem = 0.3)
    anchored_text = AnchoredText(ff')arque-Bera Test: {jbt:.2f}'', loc=2)
    axes[1].set_sine(problem = 0.3)
    axes[2].set_ctile(ff'0-0 plot with sample size = n)
    # Jarque-Bera test Chi<sup>2</sup> two-tail prob
    jbt = sms.jarque_Dera measurements[1]

sm.qaplot(measurements, line ='45', ax=axes[2])
    axes[2].set_ctile(ff'0-0 plot with sample size = {n}'')
    axes[2].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[2].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[2].grid(alpha = 0.3)
    anchored_text = AnchoredText(ff')arque-Bera Test: {jbt:.2f}'', loc=2)
    axes[2].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[2].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[2].set_itile(ff'0-0 plot with sample size = {n}'')
    axes[2].grid(alpha = 0.3)
    anchored_text = AnchoredText(ff')arque-Bera Test: {jbt:.2f}'', loc=2)
    axes[2].doc_artistilenchored_text)
```

Out[]: <matplotlib.offsetbox.AnchoredText at 0x1377e4a10>



```
In []: # To see the quality of method from the standard normal generator in Numpy, I run the whole experiment 1,000 time
# and see how much percentage the samples pass normality test
N = 1000
n1 = 100
n2 = 10000
n3 = 10000
jbts = []
for i in range(N):

measurements = np.random.normal(size = n1)
# Jarque-Bera test Chi^2 two-tail prob
jbt1 = sms.jarque_bera(measurements)[1]

measurements = np.random.normal(size = n2)
# Jarque-Bera test Chi^2 two-tail prob
jbt2 = sms.jarque_bera(measurements)[1]

measurements = np.random.normal(size = n3)
# Jarque-Bera test Chi^2 two-tail prob
jbt3 = sms.jarque_bera(measurements)[1]
jbts.append([jbt1, jbt2, jbt3])
```

```
In [ ]:
    table = pd.DataFrame(jbts, columns=[f"size = {n1}", f"size = {n2}", f"size = {n3}"])
    print('Percentage of accepting null hypothesis: pass normality test')
    (table > 0.05).mean()

Percentage of accepting null hypothesis: pass normality test
```

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• (b) The "poor man's" normal generator based on some version of the standard uniform generator in Numpy. The "poor man's" generator works as follows: generate 12 independent standard uniforms,  $\{U_i\}_{i=1}^{12}$  and compute  $\sum_{i=1}^{12} U_i - 6$ .

```
In []:

# Create a 0-0 plot
fig, axes = plt.subplots(1, 3, figsize=(15, 4))

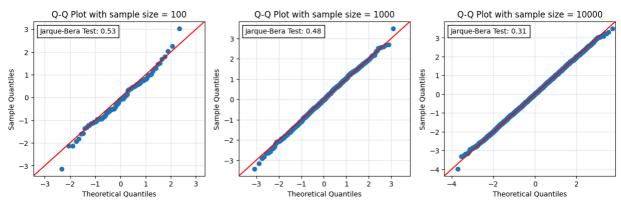
n = 100
measurements = np.random.uniform(size=(n, 12)).sum(axis = 1) - 6
# Jbrque-Bera test Chit? two-tail prob
jbt = sms.jarque_bera(measurements)[1]

sm.qqplot(measurements, line *45', ax-axes[0])
axes[0].set_title(""0-0 Plot with sample size = {n}")
axes[0].set_title(""0-0 Plot with sample size = {n}")
axes[0].set_title(""0-0 Plot with sample size = {n}")
axes[0].add_artist(anchored_text)

n = 1000
measurements = np.random.uniform(size=(n, 12)).sum(axis = 1) - 6
# Jarque-Bera test (ni? 2 bwo-tail prob
jbt = sms.jarque_bera(measurements)[1]

sm.qqplot(measurements, line *45', ax-axes[1])
axes[1].set_title(""0-0 Plot with sample size = {n}")
sxs.[1].set_title(""0-0 Plot with sample size = {n}")
axes[1].set_title(""0-0 Plot with sample size = {n}")
axes[2].set_title(""0-0 Plot with sample size = {n}")
```

Out[]: <matplotlib.offsetbox.AnchoredText at 0x137b6fc50>



```
In []:
# To see the quality of method from the (b), I run the whole experiment 1,000 time
# and see how much percentage the samples pass normality test

N = 1000
n1 = 1000
n2 = 10000
n3 = 10000
jbts = []
for i in range(N):

    measurements = np.random.uniform(size=(n1, 12)).sum(axis = 1) - 6
# Jarque-Bera test Chi^2 two-tail prob
jbt1 = sms.jarque_bera(measurements)[1]

    measurements = np.random.uniform(size=(n2, 12)).sum(axis = 1) - 6
# Jarque-Bera test Chi^2 two-tail prob
jbt2 = sms.jarque_bera(measurements)[1]

    measurements = np.random.uniform(size=(n3, 12)).sum(axis = 1) - 6
# Jarque-Bera test Chi^2 two-tail prob
jbt3 = sms.jarque_bera(measurements)[1]

    jbt3.append([jbt1, jbt2, jbt3])
```

```
In []:
    table = pd.DataFrame(jbts, columns=[f"size = {n1}", f"size = {n2}", f"size = {n3}"])
    print('Percentage of accepting null hypothesis: pass normality test')
    (table > 0.05).mean()
Percentage of accepting null hypothesis: pass normality test
```

Based on the (a) & (b) results, in small size samples method (b) is better than method (a). However, in large size sample, method (a) is better than method (b).

## 2) Practice on pricing a straddle

A straddle is constructed by owning both a call and a put option, each written on the same underlying with the same strike price (K) and expiration date (T). Suppose the rate of return and volatility of the underlying are μ and σ, and the riskless rate is r. Suppose the underlying price process is geometric Brownian motion and let S(0) = 100,K = 100,r = .05, μ = .10, σ = .10, T = 1. Using n = 10,000 sample paths price the straddle and give a standard error for your estimate. Use Numpy and vectorize.

```
In [ ]:
    def call_payoff(ST, strike):
        return np.maximum(ST - strike, 0)

def put_payoff(ST, strike):
        return np.maximum(strike - ST, 0)

def price_GBM_straddle_vec(S0, K, r, sigma, T, n, func1, func2):
        z = np.random.normal(size = (n,1))
        f1 = (r - 0.5 * sigma ** 2) * T
        f2 = sigma * np.sqrt(T)
        pv = np.exp(-r*T)
        ST = 50 * np.exp(f1+f2*z)
        C = pv * (func1(ST, K) + func2(ST, K))
        price = np.mean(C)
```

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```
stderr = np.std(C, ddof = 1)/np.sqrt(n)
return price, stderr
In [ ]: | S0 = 100
            S0 = 100

K = 100

r = 0.05

mu = 0.1

sigma = 0.1

T = 1

n = 10000
             \verb|price|, stderr = price_GBM\_straddle\_vec(S0, K, r, sigma, T, n, call\_payoff, put\_payoff)|
In [ ]: | print(f"straddle price = {price}, std = {stderr}")
```

straddle price = 8.637303176129077. std = 0.06850442786394907

## 3) Practice in generating a GBM path

Consider a GBM underlying under the risk neutral measure

$$dS_t = rS_t dt + \sigma S_t dW_t,$$

Simulate n = 10,000 paths of this process with  $S_0=100,r=.05,\,\sigma=.1$  and let T=1, each at N=52 time points, i.e. for a particular path, generate values of the price at the time points  $\{0,\Delta,2\Delta,\cdots,N\Delta\}$  where  $\Delta=T/N$ . The final price is given by  $S_T$  and an arithmetic Asian option will be based on the average price,  $A=\frac{1}{N}\sum_1^N S_{i\Delta}$ . Estimate the correlation between  $S_T$  and A and the correlation between  $S_T$  and (A-100)+. To get a standard error of your estimated correlation coefficient, you must replicate the experiment and derive the standard error from those replications. (Note, in lecture 2 or 3 we will learn the variance reduction technique called "control variables," and we will learn that  $S_T$  is a potential control variable for the Asian option. Its effectiveness will be determined by the magnitude of the correlation coefficient. If you are interested, you might change the strike to a number smaller than 100 or larger than 100 to see how it changes the correlation, but do not turn this in.)

```
In [ ]: 
    def gbm_path_z(S_0, a, b, z):
        zs = np.cumsum(a + b * z)
        S = S_0 * np.exp(zs)
        return S
                      def pay_off_asian_call(S_arr, strike, step):
    return np.maximum(np.mean(S_arr[::step]) - strike, 0)
```

```
In [ ]: | S0 = 100
             K = 100
r = 0.05
mu = 0.1
             sigma = 0.1
T = 1
n = 10000
             N = 52
             delta_t = T / N
             loc = (r - 0.5 * sigma**2) * delta_t
scale = sigma * np.sqrt(delta_t)
            def corr_st_with_sian_call_payoff_and_avg_st(S0, K, loc, scale, n, N):
    zs = np.random.normal(size=(n, N))
                   all_po_asian = np.zeros(n)
                  As = np.zeros(n)
ST = np.zeros(n)
for min range(n):
gbm_paths = gbm_path_z(S0, loc, scale, zs[m])
                        ST[m] = gbm_paths[-1]
all_po_asian[m] = pay_off_asian_call(gbm_paths, K, 1)
                  return np.corrcoef(ST, As)[0, 1], np.corrcoef(ST, all_po_asian)[0, 1]
              \textit{\# np.exp(-r * T) * all\_po\_asian.mean(), all\_po\_asian.std() / np.sqrt(n) }
```

```
In [ ]: | m = 1000
         corrs = []
for i in range(m):
              corrs.append(corr st with sian call pavoff and avg st(S0, K, loc, scale, n, N))
              # nrint(i)
```

In [ ]: pd.DataFrame(corrs, columns=['ST & A', 'ST & (A-100)+']).describe()

ST & A ST & (A-100)+ count 1000.000000 1000.000000 0.871627 0.815632 std 0.002481 0.003293 min 0.863715 0.804578 25% 0.870031 0.813446 0.871716 0.815626 50% 75% 0.873281 0.817937 0.878909 0.825649 max

# 4) Practice on the Probability Integral Transform

- a) Suppose X is a random variable with a Generalized Pareto distribution with shape parameter heta>0 having a c.d.f. given by

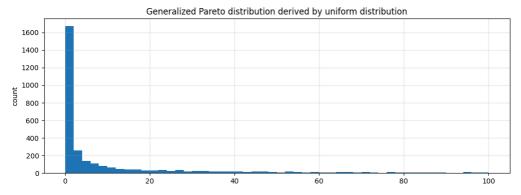
$$F(x) = egin{cases} 1 - (1 + heta x)^{-(1/ heta)} & ext{if } x \geq 0 \ 0 & ext{otherwise} \end{cases}$$

 Give the probability integral transform method for generating and observation from this distribution. (Note, this distribution arises in risk management characterizing the distribution of large losses. Here we assume  $\theta \geq 0$ ; however, when  $\theta = 0$ , this distribution corresponds to an exponential distribution and it can also be defined for  $\theta < 0$ . One can introduce scale and location parameters by transforming X to  $\sigma X + \mu$ ).

$$\begin{split} u &= 1 - (1 + \theta x)^{-(1/\theta)} \\ \Rightarrow 1 - u &= (1 + \theta x)^{-(1/\theta)} \\ \Rightarrow (1 - u)^{-\theta} &= 1 + \theta x \\ \Rightarrow x &= \frac{(1 - u)^{-\theta} - 1}{\theta} \end{split}$$

```
In [ ]: u = \text{np.random.uniform(0, 1, size} = 10000)
          theta = 20
x = ((1 - u) ** (-theta) - 1) / theta
```

```
In []: x.shape
Out[]: (10000,)
In [ ]: # descriptive statistcs
pd.Series(x).describe()
                               1.000000e+04
               count
                               1.939145e+76
1.939145e+78
9.257697e-06
2.030978e+01
               mean
std
               min
25%
                50%
75%
                                5.459377e+04
6.995200e+10
               max 1.939145e+80
dtype: float64
                fig = plt.figure(figsize = (12, 4))
plt.hist(x, bins=50, range=(0, 100))
plt.grid(alpha = 0.3)
plt.fitle('Generalized Pareto distribution derived by uniform distribution')
plt.ylabel('count')
Out[]: Text(0, 0.5, 'count')
```



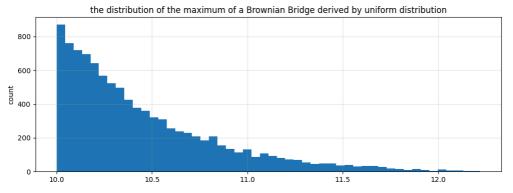
• b) (Note, the distribution will arise in the paper by Beaglehole, Dybvig and Zhou later in the course). Suppose X is a random variable having the distribution of the maximum of a Brownian Bridge having a c.d.f. given by

$$F(x) = egin{cases} 1 - e^{-2x(x-b)/h} & ext{if } max(0,b) \leq x \ 0 & ext{otherwise} \end{cases}$$

• where h and b are given constants. Give the probability integral transform method to generate this distribution. Be careful to ensure that the random variable X you generate satisfies  $X \ge max(0,b)$ .

$$\begin{split} u &= 1 - e^{-2x(x-b)/h} \\ &\Rightarrow 1 - u = e^{-2x(x-b)/h} \\ &\Rightarrow \log(1-u) = -2x(x-b)/h \\ &\Rightarrow -\frac{h}{2}log(1-u) = x(x-b) \\ &\Rightarrow \frac{b^2}{4} - \frac{h}{2}log(1-u) = (x-\frac{b}{2})^2 \\ &\Rightarrow x = \sqrt{\frac{b^2}{4} - \frac{h}{2}log(1-u)} + \frac{b}{2}, \quad \text{if } b > 0 \end{split}$$

```
In []:    h = 10
    b = 10
    u = np.random.uniform(0, 1, size = 10000)
    x = np.sqrt(b**2 / 4 - h / 2 * np.log(1 - u)) + b / 2
 In []: # descriptive statistcs
pd.Series(x).describe()
                                     10000.000000
Out[]:
                  count
                                           10.460075
0.431808
10.000039
10.141275
10.329667
                   mean
                  std
min
25%
50%
                   75%
                                           10.651598
                  max 13.599455
dtype: float64
                   \label{eq:fig} \begin{array}{lll} fig = plt.figure(figsize = (12, \ 4)) \\ plt.hist(x, \ bins=50, \ range=(b, \ b*h*** 0.05 + 1)) \\ plt.grid(alpha = 0.3) \\ plt.fitle('the distribution of the maximum of a Brownian Bridge derived by uniform distribution') \\ plt.ylabel('count') \end{array}
Out[]: Text(0, 0.5, 'count')
```



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ullet c) Let X be a random variable having a standard Cauchy distribution, that is X has p.d.f.  $f_X(x)$  given by

$$f_X(x) = rac{1}{\pi} rac{1}{1+x^2}, \quad ext{for} - \infty < x < \infty$$

Give the probability integral transform method to generate this random variable.

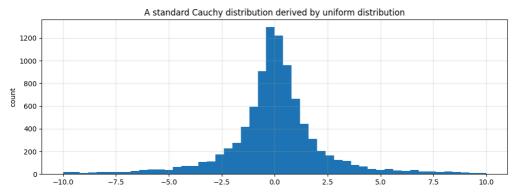
To get c.d.f, we have

$$egin{aligned} F_X(x) &= rac{1}{\pi} \int_{-\infty}^x rac{1}{1+t^2} dt \ &= rac{1}{\pi} arctan(x) + rac{1}{2}, \quad ext{for} -\infty < x < \infty \end{aligned}$$

Construction

$$u = \frac{1}{\pi} arctan(x) + \frac{1}{2}$$
 
$$\Rightarrow x = tan(\pi(u - \frac{1}{2}))$$

Out[]: Text(0, 0.5, 'count')



Appendix: timestamp

```
In []: from datetime import datetime print(f"Generated on {datetime.now()}")

Generated on 2024-05-06 20:39:05.430670
```