

ECON 3510: Poverty and Economic Development

Lecture 10: Matching

Instructor: Weizheng Lai

Bowdoin College

Fall 2025

Conditional Independence Assumption (CIA)

- ▶ We have been using the following regression

$$Y_i = \beta D_i + \mathbf{X}_i' \boldsymbol{\gamma} + \varepsilon_i.$$

- ▶ To interpret β as a causal effect, we need to assume:
 - (i) $D_i \perp\!\!\!\perp \varepsilon_i \mid \mathbf{X}_i$;
 - (ii) the linear relationship between Y_i and \mathbf{X}_i is true.
- ▶ (ii) is often implicitly taken as granted. But it might be a strong assumption.
 - How do we know the correct form of \mathbf{X}_i ? Linear, quadratic, cubic, log?
 - How much should we trust that \mathbf{X}_i has constant effect $\boldsymbol{\gamma}$ across i ?
- ▶ Ideally, we want to exploit the CIA to estimate causal effects without a strong functional assumption.
- ▶ **Fix: matching.** Basic idea: find comparable controls ($D_i = 0$) for the treated ($D_i = 1$) based upon \mathbf{X}_i .

Matching Basics

► Step 1: Decide covariates \mathbf{X} .

- Guided by economic theory.

► Step 2: Match treated and control observations with similar values of \mathbf{X} .

• How similar?

- *Exact matching*: If \mathbf{X} is binary or discrete, then it is possible to match observations with the same values.
- *Nearest matching*: If \mathbf{X} is continuous (so not possible to match exactly), can match the treated to a control with the closest value of \mathbf{X} .
- *Radius matching*: can also match treated i to control j , as long as $|\mathbf{X}_i - \mathbf{X}_j| < r$, where radius/caliper r is chosen by the researcher.

• With replacement or without replacement?

- *With replacement*: After one time of matching, the control observation goes back to the pool for the next time of matching. Thus, it's possible for a control observation to be matched for multiple treated observations.
- Either way is fine. With replacement may be preferred in small samples.

- Thus, matching can be one-to-one or one-to-many. It's also likely that we can't find matches for some treated.

► Step 3: Estimate Causal Effects.

- $\hat{\beta}_{\text{matching}} = \frac{1}{N_p} \sum_p (Y_p^T - Y_p^C)$, where p indexes matched pairs, and N_p is the number of matched pairs.
- Run linear regression using the matched sample: $Y_i = \alpha + \beta D_i + \varepsilon_i$.
 - Variants: (i) control for pair FEs; (ii) control for covariates.

Remarks

- ▶ Basic matching can be done in Stata by `calipmatch` and other commands.
- ▶ Matching deals with selection on observables (i.e., CIA is assumed). It can't address selection on unobservables.
- ▶ In fact, we can never be sure which covariates are correct ones to match on.
- ▶ The state-of-the-art implementation of matching is to use it for selecting comparable controls, and implement some quasi-experimental methods in the matched sample.
 - E.g., the parallel trends assumption for the DiD might be more plausible in a matched sample.
- ▶ Important to check the validity of matching. Are covariates you match on indeed balanced between treatment and control groups? What about untargeted covariates?

Propensity Score Matching

- ▶ *Curse of Dimensionality*: One problem with the basic matching method is that if there are many variables in \mathbf{X}_i , it is difficult to find matches for all treated, yielding a matches sample that is too small to be useful.
 - But for a plausible CIA argument, we do want more variables in \mathbf{X}_i .
- ▶ One solution: **Propensity Score Matching (PSM)**.
- ▶ Rather than matching on \mathbf{X}_i , it's enough to match on the scalar **propensity score**

$$p(\mathbf{X}_i) = \Pr(D_i = 1 \mid \mathbf{X}_i).$$

- **Theorem**: $D_i \perp\!\!\!\perp \varepsilon_i \mid \mathbf{X}_i$ implies $D_i \perp\!\!\!\perp \varepsilon_i \mid p(\mathbf{X}_i)$.
- **Key Condition (“Overlap”)**: $0 < p(\mathbf{X}_i) < 1$.

Procedures

► Step 1: Decide covariates \mathbf{X} .

► Step 2: Estimate propensity score $p(\mathbf{X}_i)$.

- Run a Probit regression

$$\Pr(D_i = 1 \mid \mathbf{X}_i) = \Phi(\mathbf{X}_i' \boldsymbol{\delta}),$$

where $\Phi(\cdot)$ is the cdf of $N(0, 1)$.

- Obtain estimated propensity score

$$\widehat{p(\mathbf{X}_i)} = \Phi(\mathbf{X}_i' \hat{\boldsymbol{\delta}}).$$

► Step 3: Match treated and control observations based upon $\widehat{p(\mathbf{X}_i)}$.

- Can do nearest matching or radius matching.
- *Blocking*: block $\widehat{p(\mathbf{X}_i)}$ into several bins; treated and control observations in the same block are matched together.

► Step 4: Estimation.

- Mean difference: $\hat{\beta}_{\text{matching}} = \frac{1}{N_p} \sum_p (\bar{Y}_p^T - \bar{Y}_p^C)$.
- Regression with the matched sample: $Y_i = \alpha + \beta D_i + \varepsilon_i$.
 - Variants: (i) control for pair/block FEs; (ii) control for covariates.

Remarks

- ▶ PSM can be done in Stata by `psmatch2`.
- ▶ Again, PSM only addresses selection on observables, not selection on unobservables.
- ▶ Pros of matching:
 - Easy to tell what comparisons are used;
 - Does not rely on strong functional form assumptions.
- ▶ Cons of matching:
 - Low statistical power: samples are smaller;
 - Data greedy.