## MSE of ML estimator

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Given a data-point  $\mathbf{x}_i$ , we have:

$$y_i = \mathbf{w}^T \mathbf{x}_i + \epsilon_i$$

Here,  $\epsilon_i \sim \mathcal{N} \big(0, \sigma^2\big)$ . Also, we assume that  $\epsilon_i$  and  $\epsilon_j$  to be independent, hence  $\text{cov}(\epsilon_i, \epsilon_j) = 0$ . Now, we "treat"  $\mathbf{x}_i$  as fixed and  $y_i$  as a random variable that is governed by the following conditional distribution:

$$y_i \mid \mathbf{x}_i \sim \mathcal{N} ig( \mathbf{w}^T \mathbf{x}_i, \sigma^2 ig)$$

Here,  $\mathbf{w}$  is also fixed. But the difference between  $\mathbf{w}$  and  $\mathbf{x}_i$  is that  $\mathbf{x}_i$  is known and  $\mathbf{w}$  is unknown. We can add all the  $y_i$ s into a random vector  $\mathbf{y} = \begin{bmatrix} y_1 & \cdots & y_n \end{bmatrix}^T$ . The conditional distribution of this random vector given the data-matrix is:

$$\mathbf{y} \mid \mathbf{X} \sim \mathcal{N} ig( \mathbf{X}^T \mathbf{w}, \sigma^2 \mathbf{I} ig)$$

We wish to estimate  ${\bf w}$ . The ML (maximum likelihood) estimator of  ${\bf w}$  is  $\widehat{{\bf w}}$  and given by:

$$\widehat{\mathbf{w}} = \left(\mathbf{X}\mathbf{X}^T\right)^{-1}\mathbf{X}\mathbf{y}$$

Here, we are deriving the result for the special case of  $\mathbf{X}\mathbf{X}^T$  being invertible. This happens when the rows of  $\mathbf{X}$  are linearly independent, that is, when there is no linear dependence among the features. The estimator  $\widehat{\mathbf{w}}$  is also a random vector since it is a function of the random vector  $\mathbf{y}$ . The estimator will turn into an estimate when we replace the random vector  $\mathbf{y}$  with its realization. Let us first compute some useful quantities. Using the linearity of expectation:

$$E[\widehat{\mathbf{w}}] = \left[ \left( \mathbf{X} \mathbf{X}^T \right)^{-1} \mathbf{X} \right] E[\mathbf{y}]$$
$$= \left( \mathbf{X} \mathbf{X}^T \right)^{-1} \mathbf{X} \mathbf{X}^T \mathbf{w}$$
$$= \mathbf{w}$$

Since  $E[\widehat{\mathbf{w}}] = \mathbf{w}$ , we have an unbiased estimator. Since the bias is zero, the MSE actually captures the variance in the estimator:

$$\begin{split} E \big[ ||\widehat{\mathbf{w}} - \mathbf{w}||^2 \big] &= E \big[ ||\widehat{\mathbf{w}} - E[\widehat{\mathbf{w}}]||^2 \big] \\ &= \mathrm{trace} \big( \mathrm{cov}(\widehat{\mathbf{w}}) \big) \end{split}$$

The trace of the covariance matrix is the sum of the variances of the d components of the random vector  $\widehat{\mathbf{w}}$ . Let  $\mathbf{A} = \left(\mathbf{X}\mathbf{X}^T\right)^{-1}\mathbf{X}$ , then  $\widehat{\mathbf{w}} = \mathbf{A}\mathbf{y}$ . We note the following facts:

• 
$$\mathbf{A}\mathbf{X}^T = \mathbf{I}$$

• 
$$XA^T = I$$

• 
$$\widehat{\mathbf{w}}\widehat{\mathbf{w}}^T = \mathbf{A}\mathbf{y}\mathbf{y}^T\mathbf{A}^T$$

Now, let us compute the covariance matrix  $cov(\widehat{\mathbf{w}})$ :

$$\mathrm{cov}(\widehat{\mathbf{w}}) = E \Big[\widehat{\mathbf{w}}\widehat{\mathbf{w}}^T\Big] - E[\widehat{\mathbf{w}}]E[\widehat{\mathbf{w}}]^T$$

We will again use the linearity of expectation at several places. We will also use the following fact:

$$\begin{aligned} \mathsf{cov}(\mathbf{y}) &= E \big[ \mathbf{y} \mathbf{y}^T \big] - E[\mathbf{y}] E[\mathbf{y}]^T \\ \sigma^2 \mathbf{I} &= E \big[ \mathbf{y} \mathbf{y}^T \big] - \mathbf{X}^T \mathbf{w} \mathbf{w}^T \mathbf{X} \end{aligned}$$

Now, we continue to expand the RHS of the covariance matrix for  $\widehat{\mathbf{w}}$ :

$$\begin{split} E \Big[ \widehat{\mathbf{w}} \widehat{\mathbf{w}}^T \Big] - E [\widehat{\mathbf{w}}] E [\widehat{\mathbf{w}}]^T &= E \Big[ \mathbf{A} \mathbf{y} \mathbf{y}^T \mathbf{A}^T \Big] - \mathbf{w} \mathbf{w}^T \\ &= \mathbf{A} E \Big[ \mathbf{y} \mathbf{y}^T \Big] \mathbf{A}^T - \mathbf{w} \mathbf{w}^T \\ &= \mathbf{A} \Big[ \mathbf{cov}(\mathbf{y}) + E [\mathbf{y}] E [\mathbf{y}]^T \Big] \mathbf{A}^T - \mathbf{w} \mathbf{w}^T \\ &= \mathbf{A} \Big[ \sigma^2 \mathbf{I} + \mathbf{X}^T \mathbf{w} \mathbf{w}^T \mathbf{X} \Big] \mathbf{A}^T - \mathbf{w} \mathbf{w}^T \\ &= \sigma^2 \cdot \mathbf{A} \mathbf{A}^T + \mathbf{A} \mathbf{X}^T (\mathbf{w} \mathbf{w}^T) \mathbf{X} \mathbf{A}^T - \mathbf{w} \mathbf{w}^T \\ &= \sigma^2 \cdot \mathbf{A} \mathbf{A}^T + \mathbf{w} \mathbf{w}^T - \mathbf{w} \mathbf{w}^T \\ &= \sigma^2 \cdot \mathbf{A} \mathbf{A}^T \\ &= \sigma^2 \Big[ (\mathbf{X} \mathbf{X}^T)^{-1} \mathbf{X} \Big] \Big[ \mathbf{X}^T (\mathbf{X} \mathbf{X}^T)^{-1} \Big] \\ &= \sigma^2 (\mathbf{X} \mathbf{X}^T)^{-1} \end{split}$$

We can now compute the MSE as:

$$\begin{split} E \Big[ ||\widehat{\mathbf{w}} - \mathbf{w}||^2 \Big] &= E \Big[ ||\widehat{\mathbf{w}} - E[\widehat{\mathbf{w}}]||^2 \Big] \\ &= \mathrm{trace} \Big( \mathrm{cov}(\widehat{\mathbf{w}}) \Big) \\ &= \sigma^2 \mathrm{trace} \Big[ \left( \mathbf{X} \mathbf{X}^T \right)^{-1} \Big] \end{split}$$