

Statistics for High Dimensional Data (and CompStat Lab)

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Lesson 3



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Spatial model with latent variable

$$y(\mathbf{s}) = \mathbf{x}(\mathbf{s})' \boldsymbol{\beta} + \alpha w(\mathbf{s}) + \varepsilon(\mathbf{s}) \quad (2)$$

- $w(\mathbf{s}) \sim GP(0, \rho(\|\mathbf{s} - \mathbf{s}'\|; \boldsymbol{\theta}))$
- $\rho(\|\mathbf{s} - \mathbf{s}'\|; \boldsymbol{\theta}) = \text{corr}(w(\mathbf{s}), w(\mathbf{s}'))$
- $\varepsilon(\mathbf{s}) \sim N(0, \sigma_\varepsilon^2)$
- The unknown parameter set is $\Psi = \{\boldsymbol{\beta}, \alpha, \sigma_\varepsilon^2, \boldsymbol{\theta}\}$
- How to estimate Ψ from data?

Maximum likelihood estimate

- We rely on MLE to estimate the model parameter vector
- Because MLE has good properties and...
- Because we can use the EM algorithm (if needed)
- Likelihood function of (2) is:

$$L(\Psi; \mathbf{y}, \mathbf{w}, \mathbf{X}) = L(\Psi; \mathbf{y}|\mathbf{w}, \mathbf{X})L(\Psi; \mathbf{w})$$

- $\mathbf{y} = (y_1, \dots, y_n)'$ is the vector of observations at n spatial locations
- $\mathbf{w} = (w_1, \dots, w_n)'$ is the vector of latent variables at n spatial locations
- \mathbf{X} is the $n \times p$ design matrix

Likelihood decomposition

$$\begin{aligned} L(\Psi; \mathbf{y}, \mathbf{w}, \mathbf{X}) &= L(\Psi; \mathbf{y}|\mathbf{w}, \mathbf{X})L(\Psi; \mathbf{w}) \\ &= L(\boldsymbol{\beta}, \alpha, \sigma_{\varepsilon}^2; \mathbf{y}|\mathbf{w}, \mathbf{X})L(\boldsymbol{\theta}; \mathbf{w}) \end{aligned}$$

- Each likelihood term depends on a subset of Ψ .
- $L(\Psi; \mathbf{y}, \mathbf{w}, \mathbf{X})$ is the complete-data likelihood (which assumes \mathbf{w} to be known)
- $L(\boldsymbol{\beta}, \alpha, \sigma_{\varepsilon}^2; \mathbf{y}|\mathbf{w}, \mathbf{X})$ and $L(\boldsymbol{\theta}; \mathbf{w})$ are densities of n –variate normal distributions

Log-likelihood function

- As usual we prefer to work with $\log(L_\Psi)$
- $-2\log L_\Psi$ is given by:

$$\log |\Sigma_\varepsilon| + \mathbf{e}' \Sigma_\varepsilon^{-1} \mathbf{e} + \log |\Sigma_w| + \mathbf{w}' \Sigma_w^{-1} \mathbf{w}$$

where

- $\mathbf{e} = \mathbf{y} - \mathbf{X}\boldsymbol{\beta} - \alpha \mathbf{w}$
- $\Sigma_\varepsilon = \sigma_\varepsilon^2 \mathbf{I}_n$, with \mathbf{I}_n the identity matrix of dimension n
- Σ_w is the $n \times n$ correlation matrix (e.g., $\exp(-D/\theta)$, with D the distance matrix)



ML estimate

- MLE is given by

$$\hat{\Psi} = \underset{\beta, \alpha, \sigma_{\varepsilon}^2, \theta}{\operatorname{argmin}} \log |\Sigma_{\varepsilon}| + \mathbf{e}' \Sigma_{\varepsilon}^{-1} \mathbf{e} + \log |\Sigma_w| + \mathbf{w}' \Sigma_w^{-1} \mathbf{w}$$

- Argmin because we are considering $-2\log(L_{\Psi})$
- Unfortunately minimizing $-2\log(L_{\Psi})$ is not feasible, plus \mathbf{w} is latent and not observed
- We must rely on the EM algorithm

EM algorithm

- The EM is an iterative algorithm for MLE
- First iteration starts with initial values $\hat{\Psi}^{(0)}$ (usually given by OLS and method of moments)

- E-step

$$Q(\Psi, \hat{\Psi}^{(m)}) = E_{\hat{\Psi}^{(m)}}(-2\log L(\Psi; \mathbf{y}, \mathbf{w}, \mathbf{X}) | \mathbf{y})$$

- M-step

$$\hat{\Psi}^{(m+1)} = \operatorname{argmax}_{\Psi} Q(\Psi, \hat{\Psi}^{(m)})$$

EM algorithm, E-step

- E-step

$$\begin{aligned} & E_{\hat{\Psi}^{(m)}}(-2\log L(\Psi; \mathbf{y}, \mathbf{w}, \mathbf{X}) | \mathbf{y}) \\ &= \text{tr} \left[\Sigma_{\varepsilon}^{-1} \left(E(\mathbf{e} | \mathbf{y}) E(\mathbf{e} | \mathbf{y})' + \text{Var}(\mathbf{e} | \mathbf{y}) \right) \right] \\ &+ \text{tr} \left[\Sigma_{\mathbf{w}}^{-1} \left(E(\mathbf{w} | \mathbf{y}) E(\mathbf{w} | \mathbf{y})' + \text{Var}(\mathbf{w} | \mathbf{y}) \right) \right] \end{aligned}$$

- $E(\mathbf{e} | \mathbf{y}) = \mathbf{y} - \mathbf{X}\boldsymbol{\beta} - \alpha E(\mathbf{w} | \mathbf{y})$
- $\text{Var}(\mathbf{e} | \mathbf{y}) = \text{Var}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta} - \alpha \mathbf{w} | \mathbf{y}) = \alpha^2 \text{Var}(\mathbf{w} | \mathbf{y})$

EM algorithm, E-step

- $E(\mathbf{w}|\mathbf{y}) = Cov(\mathbf{w}, \mathbf{y})Var(\mathbf{y})^{-1}[\mathbf{y} - \mathbf{X}\boldsymbol{\beta}]$ (see multivariate normal)
- $Var(\mathbf{w}|\mathbf{y}) = \Sigma_{\mathbf{w}} - Cov(\mathbf{w}, \mathbf{y})Var(\mathbf{y})^{-1}Cov(\mathbf{w}, \mathbf{y})'$
- $Var(\mathbf{y}) = Var(\mathbf{X}\boldsymbol{\beta} + \alpha\mathbf{w} + \boldsymbol{\varepsilon}) = Var(\alpha\mathbf{w} + \boldsymbol{\varepsilon}) = \alpha^2 Var(\mathbf{w}) + Var(\boldsymbol{\varepsilon}) + 2Cov(\mathbf{w}, \boldsymbol{\varepsilon})$
- $Var(\mathbf{w}) = \Sigma_{\mathbf{w}}$
- $Var(\boldsymbol{\varepsilon}) = \Sigma_{\boldsymbol{\varepsilon}} = \sigma_{\boldsymbol{\varepsilon}}^2 \mathbf{I}_n$
- $2Cov(\mathbf{w}, \boldsymbol{\varepsilon}) = \mathbf{0}$ (from model assumptions)
- $Cov(\mathbf{w}, \mathbf{y}) = Cov(\mathbf{w}, \mathbf{X}\boldsymbol{\beta} + \alpha\mathbf{w} + \boldsymbol{\varepsilon}) = Cov(\mathbf{w}, \alpha\mathbf{w}) = \alpha Cov(\mathbf{w}, \mathbf{w}) = \alpha Var(\mathbf{w}) = \alpha \Sigma_{\mathbf{w}}$

EM algorithm, M-step

$$\hat{\Psi}^{\langle m+1 \rangle} = \operatorname{argmax}_{\Psi} Q(\Psi, \hat{\Psi}^{\langle m \rangle})$$

$$\frac{dQ(\Psi, \hat{\Psi}^{\langle m \rangle})}{d\Psi} = 0$$

$$\alpha^{\langle m+1 \rangle} = \frac{\operatorname{tr}[(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}^{\langle m \rangle})E(\mathbf{w}|\mathbf{y})']}{\operatorname{tr}[E(\mathbf{w}|\mathbf{y})E(\mathbf{w}|\mathbf{y})' + \operatorname{Var}(\mathbf{w}|\mathbf{y})]}$$

$$\boldsymbol{\beta}^{\langle m+1 \rangle} = (\mathbf{X}'\mathbf{X})^{-1} \left[\mathbf{X}' \left(\mathbf{y} - \alpha^{\langle m+1 \rangle} E(\mathbf{w}|\mathbf{y}) \right) \right]$$

EM algorithm, M-step

$$\sigma_{\varepsilon}^{2\langle m+1 \rangle} = \frac{1}{n} \text{tr}[E(\mathbf{e}|\mathbf{y})E(\mathbf{e}|\mathbf{y})' + \text{Var}(\mathbf{e}|\mathbf{y})]$$

$$\boldsymbol{\theta}^{\langle m+1 \rangle} = \underset{\boldsymbol{\theta}}{\text{argmin}} \log |\Sigma_{\mathbf{w}}^{-1}(\boldsymbol{\theta})| + \text{tr}[\Sigma_{\mathbf{w}}^{-1}(\boldsymbol{\theta})(\hat{\mathbf{w}}\hat{\mathbf{w}}')]$$

Where $\hat{\mathbf{w}} = E_{\boldsymbol{\theta}^{\langle m \rangle}}(\mathbf{w}|\mathbf{y}) = E(\mathbf{w}|\mathbf{y})$