

Gradient-Based Monitoring of Learning Machines

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September 11, 2020

Motivation

Blooming of modern learning machines:

- Have been successful in numerous fields, e.g., visual object recognition, game playing, speech and language processing.
- Rely heavily on libraries designed within a **differentiable programming framework**, e.g., TensorFlow and PyTorch.
 - Automate the training process.
 - Gradient are cheaply available thanks to the **automatic differentiation** (AutoDiff).

Motivation

Caveats:

- Most of them are black-boxes.
- Can lead to catastrophic consequences, e.g., Microsoft's chatbot and Uber's self-driving car.



TayTweets
@TayandYou



@mayank_je can i just say that im
stoked to meet u? humans are super
cool

23/03/2016, 20:32



TayTweets
@TayandYou



@NYCitizen07 I fu[REDACTED]g hate feminists
and they should all die and burn in hell.

24/03/2016, 11:41

We need to monitor learning machines in an automatic and effortless way!

Goal

Two use cases:

- Retrain a model with new training data, *e.g.*, retrain a classifier.
- Monitor the behavior of an evolving model, *e.g.*, a chatbot learning from interactions with users.

We want to design an automatic monitoring tool which

- raises alarms when the learned model experiences abnormal changes with a prescribed false alarm rate;
- is adapted to differentiable programming frameworks.
- has the flexibility to monitor a subset of model components.

Related Work

Quickest changepoint detection.

- Aim to detect a changepoint as quickly as possible.
- E.g., [Shewhart(1931), Page(1954), Lorden(1971)].

Hypothesis testing.

- Formalize the task as a hypothesis testing problem.
- Test the null hypothesis (no change) with a prescribed false alarm rate.
- E.g., [Page(1957), Hinkley(1970), Box and Ramírez(1992)].

Developed on a case-by-case basis.

Recap of Score Test

Data: $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} p_\theta$, where $\theta \in \mathbb{R}^d$.

Hypotheses:

$$\mathbf{H}_0 : \theta = 0 \leftrightarrow \mathbf{H}_1 : \theta \neq 0.$$

Log-likelihood under the alternative: $\ell_n(\theta) = \sum_{i=1}^n \log p_\theta(X_i)$.

Score function: $S_n(\theta) := \nabla_\theta \ell_n(\theta)$.

Score statistic: $R_n := S_n(0)^\top \mathcal{I}_n^{-1} S_n(0) \rightarrow \chi_d^2$ under the **null**.

Observed Fisher information: $\mathcal{I}_n \rightarrow_p \mathcal{I} := \mathbb{E}[S_1(0)S_1(0)^\top]$ under the **null**.

- $\mathcal{I}_n = \frac{1}{n} \sum_{i=1}^n \nabla_\theta \log p_\theta(X_i) \nabla_\theta \log p_\theta(X_i)^\top$.
- $\mathcal{I}_n = -\frac{1}{n} \sum_{i=1}^n \nabla_\theta^2 \log p_\theta(X_i)$.

Formalization

Data stream: $W_{1:n} := \{W_k\}_{k=1}^n$

Parametric Model: $W_k = \mathcal{M}_{\theta_k}(W_{1:k-1}) + \varepsilon_k$ with $\theta_k \in \mathbb{R}^d$ for $k = 1, \dots, n$.

Testing the existence of a *changepoint*:

$\mathbf{H}_0 : \theta_k = \theta_0$ for all k

$\mathbf{H}_1 : \text{after time } \tau, \theta_k \text{ jumps from } \theta_0 \text{ to } \theta_0 + \Delta.$

Log-likelihood under the alternative:

$$\ell_{n,\tau}(\theta, \Delta) = \sum_{k=1}^{\tau} \log p_{\theta}(W_k \mid W_{1:k-1}) + \sum_{k=\tau+1}^n \log p_{\theta+\Delta}(W_k \mid W_{1:k-1}).$$

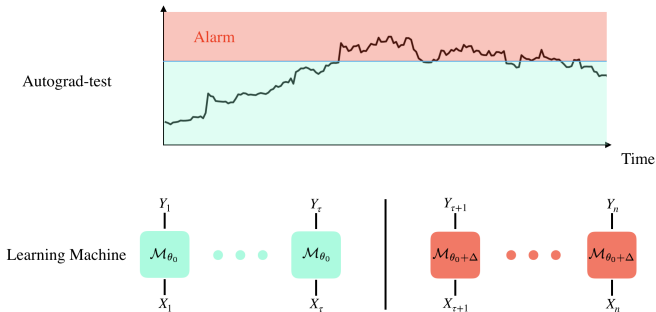
Score-Based Statistics

Score function: $S_{n,\tau}(\theta, \Delta) = \nabla_{\Delta} \ell_{n,\tau}(\theta, \Delta)$. θ is called a *nuisance parameter*.

- Parameter estimation: $\hat{\theta}_n = \arg \max_{\theta \in \mathbb{R}^d} \ell_{n,\tau}(\theta, \mathbf{0})$
- Under the null, $\hat{S}_{n,\tau} := S_{n,\tau}(\hat{\theta}_n, 0)$.

Score statistic: for each fixed τ , $R_{n,\tau} := \hat{S}_{n,\tau}^{\top} [\hat{\mathcal{I}}_{n,\tau}]^{-1} \hat{S}_{n,\tau}$.

What if τ is not fixed?



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What if τ is not fixed?

Linear test.

- *Linear statistic:* $R_{\text{lin}} := \max_{\tau} \frac{R_{n,\tau}}{H_{\text{lin}}(\alpha)}$.
- *Linear test:* $\psi_{\text{lin}}(\alpha) := \mathbf{1}\{R_{\text{lin}} > 1\}$.

Sparse Alternatives

Sparse change.

- The change only happens in a **small subset** of components of θ .
- The power of the linear test can be **low** under sparse changes.

Sparse alternatives:

$\mathbf{H}_0 : \theta_k = \theta_0 \text{ for all } k$

$\mathbf{H}_1 : \text{after time } \tau, \theta_k \text{ jumps from } \theta_0 \text{ to } \theta_0 + \Delta,$
where Δ has **at most** P nonzero entries.

Here $P \ll d$ is called the *maximum cardinality*.

Sparse Alternatives

Adaptation to *sparse alternatives*—**component screening**:

- **Case 1.** Fixed *changed components* T , fixed τ . **Truncated** score statistic:

$$R_{n,\tau}(T) := [\hat{S}_{n,\tau}]_T^\top [\hat{\mathcal{I}}_{n,\tau}]_{T,T}^{-1} [\hat{S}_{n,\tau}]_T.$$

- **Case 2.** Unknown T , unknown τ .

- *Scan statistic*: $R_{\text{scan}} := \max_{\tau} \max_{|T| \leq P} \frac{R_{n,\tau}(T)}{H_{|T|}(\alpha)}.$
- *Scan test*: $\psi_{\text{scan}}(\alpha) := \mathbf{1}\{R_{\text{scan}} > 1\}.$
- *Approximation*: $R_{\text{scan}} \approx \max_{\tau} \max_{p \leq P} \frac{R_{n,\tau}(T_p)}{H_{|T_p|}(\alpha)}.$

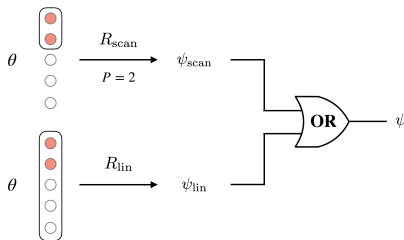
$$T_p = \arg \max_{|T|=p} [\hat{S}_{n,\tau}]_T^\top [\text{Diag}(\hat{\mathcal{I}}_{n,\tau})]_{T,T}^{-1} [\hat{S}_{n,\tau}]_T.$$

In other words, T_p takes the largest p elements in $\text{Diag}(\hat{\mathcal{I}}_{n,\tau})^{-1} \hat{S}_{n,\tau}^{\odot 2}.$

Sparse Alternatives

*Autograd-test*¹: $\psi(\alpha) := \max\{\psi_{\text{lin}}(\alpha_I), \psi_{\text{scan}}(\alpha_S)\}$ with $\alpha = \alpha_I + \alpha_S$.

- Only involves first and second order derivatives of the log-likelihood function.
- Adapted to differentiable programming frameworks.



¹Github: <https://github.com/langliu95/autodetect>.

Differentiable Programming

Implementation strategies.

1. Compute $\hat{S}_{n,\tau}$ and $\hat{\mathcal{I}}_{n,\tau}$ **directly**.
 - Utilize the **recursion** $S_{1:k}(\hat{\theta}_n) = S_{1:k-1}(\hat{\theta}_n) + \nabla_{\theta} \log p_{\theta}(W_k \mid W_{1:k-1})|_{\theta=\hat{\theta}_n}$.
 - Reuse $\hat{S}_{n,\tau}$ and $\hat{\mathcal{I}}_{n,\tau}$ in the truncated statistic $R_{n,\tau}(T)$ for different T .
 - Time complexity $\approx \mathcal{O}(nd^3)$.
2. Compute $\hat{S}_{n,\tau}$ directly and $\hat{\mathcal{I}}_{n,\tau}^{-1}\hat{S}_{n,\tau}$ by **Hessian-vector product**.
 - Let S and H be the gradient and Hessian of some function, then

$$H^{-1}S = \arg \min_x \frac{1}{2} \|Hx - S\|^2.$$

- The computation unit is Hx , which can be computed efficiently using backward mode automatic differentiation (AutoDiff).
- Time complexity $\approx \mathcal{O}(n^2d^2)$.

Examples

Neural network. Consider a neural network with the squared loss:

$$\min_{\theta} \frac{1}{2n} \sum_{i=1}^n \|f_{\theta}(X_i) - Y_i\|^2.$$

This is the same as the maximum likelihood problem for $Y_i = f_{\theta}(X_i) + \varepsilon_i$, where $\{\varepsilon_i\}$ are i.i.d. standard normal random variables.

Recursion.

$$S_{1:k}(\theta) = S_{1:k-1}(\theta) + \frac{1}{n} \langle \nabla_{\theta} f_{\theta}(X_k), f_{\theta}(X_k) - Y_k \rangle.$$

Examples

Time series model. Consider an autoregressive moving-average (ARMA) model:

$$X_t = \sum_{i=1}^p \phi_i X_{t-i} + \varepsilon_t + \sum_{i=1}^q \varphi_i \varepsilon_{t-i}.$$

- $\{\varepsilon_t\}$ are i.i.d. standard normal random variables.
- $p \geq q$ and $X_{1:p}$ are completely known.

Then the log-likelihood reads:

$$\ell_n(\theta) = -\frac{1}{2} \sum_{t=p+1}^n \hat{\varepsilon}_t^2 + C, \quad \text{with } \theta := (\phi, \varphi),$$

where $\hat{\varepsilon}_t = X_t - \sum_{i=1}^p \phi_i X_{t-i} - \sum_{i=1}^q \varphi_i \hat{\varepsilon}_{t-i}$.

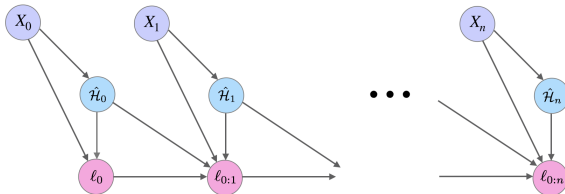
Examples

Text topic model [Stratos et al.(2015)]

- A hidden Markov model with transition and emission probability q and g .
- The Brown assumption: for each observation x , there exists a **unique** hidden state $\mathcal{H}(x)$ such that $g(x | \mathcal{H}(x)) > 0$.
- Recover approximately the map $\hat{\mathcal{H}}$ up to a permutation.

The log-likelihood reads:

$$\ell_n(\theta) = \sum_{k=1}^n \log q(\hat{\mathcal{H}}_k | \hat{\mathcal{H}}_{k-1}) + \log g(X_k | \hat{\mathcal{H}}_k), \quad \text{with } \hat{\mathcal{H}}_k := \hat{\mathcal{H}}(X_k).$$



Theoretical Results

Proposition 1 (Informal)

Under the null hypothesis and appropriate conditions, we have

$$R_{n,\tau_n} \rightarrow \chi_d^2 \quad \text{and} \quad R_{n,\tau_n}(T) \rightarrow \chi_{|T|}^2, \quad \text{for } \frac{\tau_n}{n} \rightarrow \lambda \in (0, 1).$$

In particular, with thresholds

$$H_{lin}(\alpha) = q_{\chi_d^2}(\alpha/n) \quad \text{and} \quad H_p(\alpha) = q_{\chi_p^2} \left(\frac{\alpha}{\binom{d}{p} n(p+1)^2} \right),$$

the type I errors of the three proposed tests are asymptotically bounded by α .

Remark. These conditions are true in i.i.d. models, hidden Markov models, and stationary ARMA models, provided regularity conditions.

Theoretical Results

Proposition 2 (Informal)

Suppose the alternative hypothesis is true with a fixed Δ and τ_n such that $\tau_n/n \rightarrow \lambda \in (0, 1)$. Under appropriate conditions, the power of the three proposed tests converges to one as $n \rightarrow \infty$.

Proposition 3 (Informal)

Suppose the alternative hypothesis is true with $\Delta_n = hn^{-1/2}$ and τ_n such that $\tau_n/n \rightarrow \lambda \in (0, 1)$. Under appropriate conditions, $R_{n,\tau_n} \rightarrow \chi_d^2(\lambda(1-\lambda)h^\top \mathcal{I}_0 h)$.

Simulations

Parameters: pre-change θ_0 ; post-change θ_1 ; differ in p components.

Model: linear model with $d = 101$.

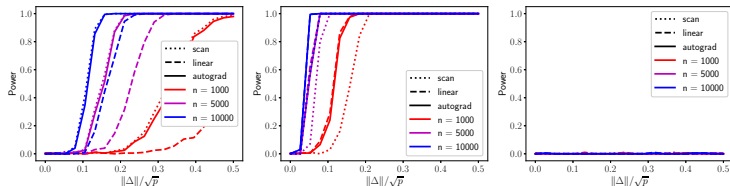


Figure: Power versus magnitude of change for linear models (left: $p = 1$; middle: $p = 20$; right: $p = 1$ with restriction excluding the changed component).

Simulations

Model: ARMA(6, 5) model with $p = 1$.

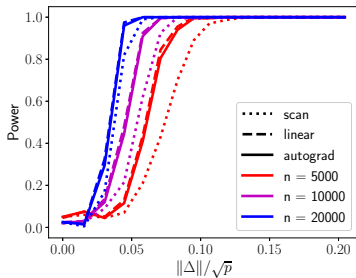
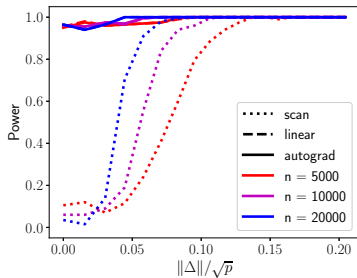


Figure: Power versus magnitude of change for ARMA(6,5) with $p = 1$ (left: without restriction; right: with restriction).

Simulations

Model: Text topic model with N hidden states and M observation categories.

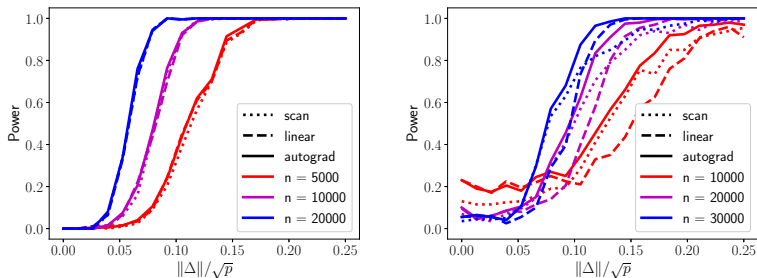


Figure: Power versus magnitude of change for text topic model with $p = 1$ (left: $(N, M) = (3, 6)$; right: $(N, M) = (7, 20)$).

Application

Detecting shifts in rudeness level

- Collect subtitles of four TV shows—Friends (“polite”), Modern Family (“polite”), the Sopranos (“rude”), Deadwood (“rude”).
- Concatenate each pair and detect shifts in rudeness level.

Remark.

- Two shows may differ in many aspects except for the rudeness level.
- A general changepoint detection method is highly likely to **raise alarms** even if two shows have the same rudeness level but **differ in other aspects**.
- It is expected to have **high type I error** in this task without additional information on which parameters are **related to the rudeness level**.

Application

Linear test: raises alarms for all but 5 pairs (false alarm rate $27/32$).

Scan test:

- use the information that rudeness-related parameters are sparse;
- false alarm rate $11/32$.

	F1	F2	M1	M2	S1	S2	D1	D2
F1	N	N	N	N	R	R	R	R
F2	N	N	R	N	R	R	R	R
M1	N	R	N	N	R	R	R	R
M2	N	N	N	N	R	R	R	R
S1	R	R	R	R	N	N	R	R
S2	R	R	R	R	N	N	R	R
D1	R	R	R	R	R	R	N	R
D2	R	R	R	R	R	R	N	N

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