ME414 2S 2018

# ME414 - ESTATÍSTICA PARA EXPERIMENTALISTAS

PROVA II - 22/11/2018

## **FORMULÁRIO**

1. 
$$\bar{x} = \frac{\sum_{i=1}^{n} x_i}{n}$$
;

2. 
$$s^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} = \frac{1}{n-1} (\sum_{i=1}^n x_i^2 - n\bar{x}^2);$$

3. 
$$C.V = \frac{s}{\bar{r}}$$
.

## Probabilidade

1. 
$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$
;

2. 
$$P(A|B) = P(A \cap B)/P(B)$$
;

3. 
$$P(A) = \sum_{i=1}^{k} P(A|B_i)P(B_i);$$

4. 
$$P(B_i|A) = \frac{P(A|B_i)P(B_i)}{\sum_{i=1}^k P(A|B_i)P(B_i)}$$
.

## Distribuição de probabilidade

Seja X uma variável aleatória discreta. Então,

$$\mu = E(X) = \sum_{x} x P(X = x)$$
 e  $Var(X) = E(X - \mu)^2 = \sum_{x} (x - \mu)^2 P(X = x)$ .

1. Se 
$$X \sim \text{Uniforme Discreta} \quad \Rightarrow \quad P(X=x) = \left\{ \begin{array}{ll} 1/k, & x=1,2,\ldots,k; \\ 0, & \text{caso contrário.} \end{array} \right.$$

2. Se 
$$X \sim b(p)$$
  $\Rightarrow$   $P(X = x) = \begin{cases} p^x (1-p)^{1-x}, & x = 0, 1; \\ 0, & \text{caso contrário.} \end{cases}$ 

3. Se 
$$X \sim Bin(n,p)$$
  $\Rightarrow$   $P(X=x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x}, & x=0,1,2,\ldots,n; \\ 0, & \text{caso contrário,} \end{cases}$  onde  $\binom{n}{x} = \frac{n!}{x!(n-x)!}$ .

4. Se 
$$X \sim G(p) \implies P(X = x) = p(1 - p)^{x-1}, \quad x = 1, 2, \dots$$

5. Se 
$$X \sim Hip(N, n, r)$$
  $\Rightarrow$   $P(X = x) = \begin{cases} \frac{\binom{r}{x} \binom{N-r}{n-x}}{\binom{N}{n}}, & x = 0, \dots, min\{r, n\}; \\ \binom{N}{n}, & 0, & \text{caso contrário.} \end{cases}$ 

6. Se 
$$X \sim P(\lambda)$$
  $\Rightarrow$   $P(x = x) = \begin{cases} \frac{e^{-\lambda} \lambda^x}{x!}, & x = 0, 1, 2, \dots; \\ 0, & \text{caso contrário.} \end{cases}$ 

### Densidade de probabilidade

Seja X uma variável aleatória contínua. Então,

1. Se 
$$X \sim \text{Uniforme(a,b)} \quad \Rightarrow \quad f(x) = \begin{cases} \frac{1}{(b-a)}, & a \le x \le b; \\ 0, & \text{caso contrário.} \end{cases}$$

ME414 2S 2018

2. Se 
$$X \sim \text{Exp}(\lambda) \implies f(x) = \lambda e^{-\lambda x}, \quad x \ge 0.$$

3. Se 
$$X \sim \text{Normal}(\mu, \sigma^2)$$
  $\Rightarrow$   $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{(x-\mu)^2}{2\sigma^2}\}, -\infty \le x \le \infty.$ 

### Inferência

1. Se 
$$X \sim \text{Normal}(\mu, \sigma^2)$$
, então  $Z = \frac{X - \mu}{\sigma} \sim \text{Normal}(0, 1)$ ;

2. 
$$P(-z_{\alpha/2} < Z <_{\alpha/2}) = 1 - \alpha$$
,  $P(Z > z_{\alpha}) = \alpha$  e  $P(Z < -z_{\alpha}) = \alpha$ ;

3. 
$$\hat{p} \sim \text{Normal}(p, \frac{p(1-p)}{n}) \Rightarrow \frac{\hat{p}-p}{\sqrt{p(1-p)/n}} \sim \text{Normal}(0, 1);$$

4. 
$$\left[\hat{p}-z_{\alpha/2}\sqrt{\frac{\hat{p}(1-\hat{p})}{n}};\hat{p}+z_{\alpha/2}\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}\right];$$

5. 
$$\left[\hat{p} - z_{\alpha/2}\sqrt{\frac{1}{4n}}; \hat{p} + z_{\alpha/2}\sqrt{\frac{1}{4n}}\right];$$

6. 
$$n = \left(\frac{z_{\alpha/2}}{2m}\right)^2$$
;

7. 
$$\overline{X} \sim \text{Normal}(\mu, \frac{\sigma^2}{n})$$
 então  $\frac{\overline{X} - \mu}{\sigma/\sqrt{n}} \sim \text{Normal}(0, 1), \frac{\overline{X} - \mu}{s/\sqrt{n}} \sim t_{n-1};$ 

8. 
$$\left[ \overline{x} - z_{\alpha/2} \sqrt{\frac{\sigma^2}{n}}; \overline{x} + z_{\alpha/2} \sqrt{\frac{\sigma^2}{n}} \right];$$

9. 
$$\left[\overline{x} - t_{n-1,\alpha/2}\sqrt{\frac{s^2}{n}}; \overline{x} + t_{n-1,\alpha/2}\sqrt{\frac{s^2}{n}}\right];$$

10. 
$$n = \left(\frac{z_{\alpha/2}}{m}\right)^2 \sigma^2$$
;

11. Sob 
$$H_0: \mu_x = \mu_y$$
,

$$\frac{\overline{X} - \overline{Y}}{\sqrt{\sigma_x^2/n + \sigma_y^2/m}} \sim \text{Normal}(0, 1), \quad \frac{\overline{X} - \overline{Y}}{\sqrt{s_p^2(1/n + 1/m)}} \sim t_{n+m-2},$$

$$s_p^2 = \frac{(n-1)s_x^2 + (m-1)s_y^2}{n+m-2};$$

12. Sob 
$$H_0: p_1 = p_2$$
,

$$\frac{\hat{p_1} - \hat{p_2}}{\sqrt{p_0(1 - p_0)(1/n + 1/m)}} \sim \text{Normal}(0, 1),$$

$$p_0 = \frac{n\hat{p_1} + m\hat{p_2}}{n+m}.$$