1) The GENDER variable is configurated with 2 levels It is encoded with 1 dummy voichle, i.e.

The RACE variable is categorical with 3 Revels.

It is encoded with 2 (= 3-1, To evoid collinearity) dummy variables, Specifically, from the output we see that we have parametres associated with the levels 2 and 3 (RACE2, RACE3), hence RACE=1 is the baseline.

Thus

RACEL:= $\begin{cases} 1 & \text{if } RACE_i = 2 \text{ (individual } i \text{ is hispanic}) \\ 0 & \text{otherwise} \end{cases}$

RACE3; =
$$\begin{cases} 1 & \text{if } RACEi = 3 \text{ (individual it is white)} \\ 0 & \text{otherwise} \end{cases}$$

2) Gaussian linear model

denoting with
$$gi$$
 the wage of individual i

$$Y_{i} = \beta_{1} + \beta_{2} \text{ 5DU}_{i} + \beta_{3} \text{ SOUTH1}_{i} + \beta_{4} \text{ GENDER1}_{i} + \beta_{5} \text{ 5XPER}_{i} +$$

+ \$\beta_{\infty} union1; + \$\beta_{\infty} AQF; + \$\beta_{8} RACE1; + \$\beta_{9} RACE3; + \$\beta_{10} KARR1; + &:

Ei ~ N(0, 52) Li.d.

The model's assumptions one: (i) normality, homosædashaity, independence \Rightarrow E: $N(0,6^2)$ iid for i=1,...,534

(ii) einearity writ. \$1,..., B20

(iii) the covoristes are linearly independent

3) It-value of
$$\beta_{c}$$
 (union1)

$$\tau^{obs} = \frac{\hat{\beta}_{c} - o}{\sqrt{\hat{\phi}_{c}(\hat{\beta}_{c})}} = \frac{\hat{\beta}_{c}}{\hat{s}_{c}(\hat{\beta}_{c})} = \frac{1.4336}{0.5087} = 2.819$$

p-value of the test { Ho: $\beta_6 = 0$ } p. value = PHo (|T| > |tobs|) = 2. PHo (T> |tobs|) = 2. PHo (T> 2.818) = 2. (1-0.9975) = 0

where
$$T = \frac{\hat{\beta}_{6}(\underline{Y})}{\hat{se}(\hat{\beta}_{6}(\underline{Y}))}$$

The t_{n-p} = t₅₂₄
 $\hat{se}(\hat{\beta}_{6}(\underline{Y}))$

Estimate of β_{3} (RACE3)

 $t^{obs}: \frac{\hat{\beta}_3}{\hat{s}_2(\hat{\beta}_3)} \longrightarrow \hat{\beta}_3 = t^{obs}: \hat{s}_2(\hat{\beta}_3) = 1.66 \cdot 0.5860 = 0.9727$

4) EDUCATION is numeric

Hence β_3 represents the (adolitie) change in the expected wase for an adolitional year of education, keeping the other covariates fixed In other words, for every additional year of education, the mean wase increases of $1.26 \, \beta$, with all other variables held constart.

RACE is categorical

I consider two individuals j and k such that RACE; = 1 and RACEK = 2

while all other covariates one equal (i.e.,
$$EDU_j = EDU_k$$
, $SOUTH_j = SOUTH_k$,...)

 $\mu_j = \beta_1 + \beta_2 = EDU_j + ... + \beta_7 AGE_j + \beta_8 RACE2_j + \beta_9 RACE3_j + \beta_{10} HARR1_j$

consider un-mi = E[Yk] - E[Yi]

$$\mu_{K}$$
 - μ_{j} = β_{1} + β_{2} (Boundary) + ... + β_{2} (AGE, AGE;) + β_{3} (RACE2, RACE2;) + β_{3} (RACE3, RACE3;) + β_{10} (MARRIX MARRI) = β_{8}

Hence (Bg) represents the adolitive change in the mean hourey wage if I consider an individual in the hispanic population composed to an individual in the "other" population (keeping other covariates equal). parameter associated with RACEL

= 2. (1-0.9975)= 0.005

Following a similar reasoning, By (associated with RACE3) represents the additive change in the mean hourey wage

if I consider on individual in the white population composed to on individual in the "other" population (keeping other covortates equal).

If I consider two individuals, identical for all correlates but the monital status, the married one has a mean hourly wage of 0.4563\$ higher than the unmarried one.

If we want a text with a 5% significan a level

HARRI is binory (cotteporical with 2 cottepories)

5)
$$\begin{cases} H_0: \beta_2 = \beta_3 = ... = \beta_{10} = 0 \\ H_1: \text{ at least one } \beta_j \text{ is } \neq 0 \end{cases}$$
 ($j = 2, ..., 10$)

The observed value of the text stanship is $f^{obs} = \frac{0.2753}{40.2753} \cdot \frac{524}{3} = 22.117$

I use the test statistic $f = \frac{R^2}{4R^2} \cdot \frac{RP}{P-1} = \frac{R^2}{4-R^2} \cdot \frac{524}{9}$ such that $f \approx F_{3,524}$

the reject region is
$$R = (f_{9,524}; 0.35; +\infty) = (1.8977, +\infty)$$

F_{9,524}; 0.35 Since $f^{obs} \in R$, I reject the

6) The model is

) The model is
$$Y_{i} = Y_{1} + Y_{2} \in \text{EXPER}_{i} + Y_{3} \in \text{GENDER1}_{i} + E_{i}$$

$$E_{i} \sim N(O_{i} e^{2}) \text{ i.i.d.}$$

This model is nested to model A, hence I can compone them through a text Sto: B2 = B3 = B6 = B7 = B8 = B9 = B10 = 0

 $F = \frac{85F_B - 85F_A}{85F_A} \cdot \frac{n - P_A}{P_A - P_B} = \frac{85F_B - 85F_A}{55F_A} \cdot \frac{534 - 40}{40 - 3}$ $F \approx \frac{10 - 3}{10 - 3}$

To compute the observed value, we first need to obtain SSEB and SSEA.
SSE =
$$\sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

SSEA = $(residual s.e.)^2$. $(n-10) = 4.412^2$. S24 = 10.200.05

SSEB = (residual s.c.)2. (n-3) = 5.0112.531 = 13.333.47

$$\rho$$
 = $\frac{13 \, 333 - 10 \, 200}{10 \, 100}$. $\frac{524}{7}$ = 22.99
I reject to for all usual significance certal. I prefer model A.

4) No, because the R2 always increase (or stays the some) when I add ovariates. I should up the adjusted R2.

8) The inclusion of the interaction allows studying if the effect on the mean wage

of an additional year of experience is different for men and women.

The model is

Y: = \(\xi_1 + \xi_2 \) EXPER: 4 \(\xi_3 \) GENDER1; + \(\xi_4 \) EXPER: GENDER1; + \(\xi_4 \) Ei $\sim N(0, \sigma^2)$ Li.d.

If I consider a mon, the expected wage is IE[Yi] = \$1 + \$2 EXPER:

a women.

E[Yi] = \$1 + \$2 EXPER: + \$3 + \$4 EXPER: $= (\xi_1 + \xi_3) + (\xi_2 + \xi_4)$ EXPERI

Hence \$4 is the change in the effect of an adolitional year of experience on the moon wage due to boing a woman (componed to being a mon) In other terms: an additional year of experience leads to an increase of 0.809 \$ in the mean wase for a man, while it cooks to an increase of (0.0809 - 0.0798) = 0.011\$ for