

STATISTICAL MODELLING: Theory and practice

Project 1: Wind power data



GOALS: ASSIGNMENT 1

Descriptive statistics

Simple models



- 1. Fit different probability density models to wind power, wind speed and wind direction data.
- 2. Conclude on the most appropriate model for each variable.
- 3. Report parameters including assessment of their uncertainty.



Descriptive statistics

| r.day | month | day | pow.obs | ws30 | wd30 |
|-------|-------|-----|--------------|----------|-----------|
| 1 | 1 | 1 | 243.0277778 | 6.723611 | 4.0343405 |
| 2 | 1 | 2 | 2780.0136986 | 4.272603 | 2.1365208 |
| 3 | 1 | 3 | 2118.6164384 | 4.272603 | 1.6240318 |
| 4 | 1 | 4 | 1660.8767123 | 6.541096 | 0.2269022 |
| 5 | 1 | 5 | 1872.7945205 | 9.713699 | 5.3161852 |
| 6 | 1 | 6 | 3212.2602740 | 8.161644 | 0.9522963 |

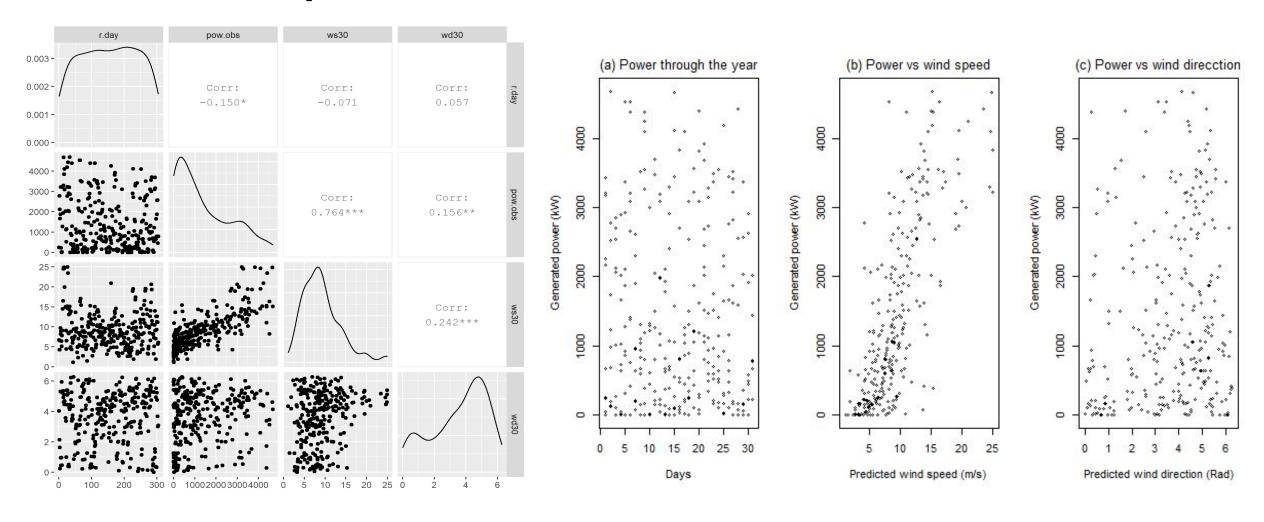
| Variable | Meaning | Unit |
|----------|--|------|
| r.day: | Days since 1/1 2003 | days |
| month: | Month in year | |
| day: | Day in month | |
| pow.obs: | Average daily wind power production | kW |
| ws30: | Predicted wind speed 30 meters above ground level | m/s |
| wd30: | Predicted wind direction (0 north, $\pi/2$ east) 30 meters | |
| | above ground level | rad |

288 x 6

| pow.obs | ws30 | wd30 |
|------------------|----------------|------------------|
| Min. : 0.123 | Min. : 1.139 | Min. :0.000095 |
| 1st Qu.: 254.158 | 1st Qu.: 5.779 | 1st Qu.:2.474999 |
| Median : 964.123 | Median : 8.498 | Median :4.079297 |
| Mean :1381.196 | Mean : 9.112 | Mean :3.602390 |
| 3rd Qu.:2196.579 | 3rd Qu.:11.202 | 3rd Qu.:4.945443 |
| Max. :4681.062 | Max. :24.950 | Max. :6.274642 |



Descriptive statistics





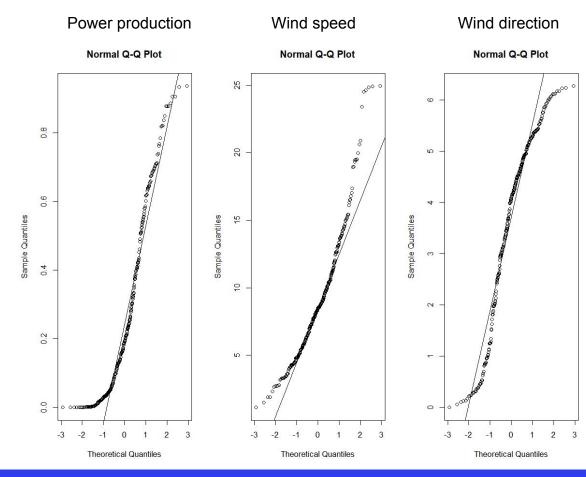
Normalization of power production:

$$normpow = \frac{pow.obs - min(pow.obs)}{max(pow.obs) - min(pow.obs)}$$

The normalization is based on the installed capacity, which maximum is 5000 kW.

Values that from 0 to 1 for power production.

Checking for normality of the data:

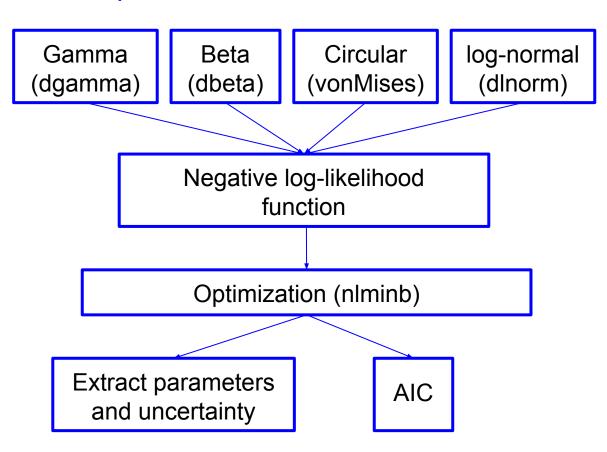


Project 1: Wind power data

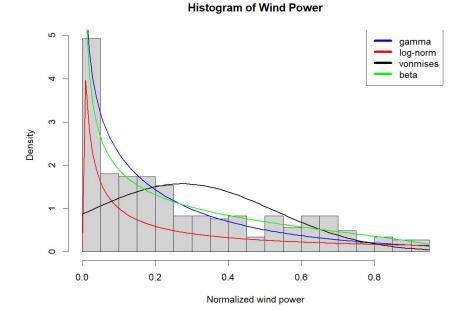


| Model | Parameter 1 | Cl Parameter 1 | Parameter 2 | CI Parameter 2 |
|------------|-------------|-------------------------|-------------|----------------------|
| Gamma | 0.6926402 | [0.6246109, 0.7606695] | 2.5073938 | [2.159441, 2.855347] |
| Beta | 0.5571045 | [0.4952471, 0.6189619] | 1.4918277 | [1.286340, 1.697315] |
| VonMises | 0.2738893 | [0.2443170, 0.3034616] | 15.7607520 | [13.23072, 18.29079] |
| Log-normal | 0.0000000 | [-0.3335482, 0.3335482] | 2.888063 | [2.652209, 3.123918] |

Power production: Non-normal models



| Model | Parameters | AIC |
|------------|-------------------------|-----------|
| Gamma | [0.6926402, 2.5073938] | -190.7635 |
| Beta | [0.5571045, 1.4918277] | -239.3236 |
| VonMises | [0.2738893, 15.7607520] | 36.75738 |
| Log-normal | [0.000000, 2.888063] | 187.8728 |





Power production: Normal models

Box-Cox Transformation

Transformation 1

Transformation 2

$$y^{(\lambda)} = \begin{cases} \frac{y^{\lambda} - 1}{\lambda} & \lambda \neq 0 \\ \log(y_i) & \lambda = 0 \end{cases}$$

$$y^{(\lambda)} = \frac{1}{\lambda} \log \left(\frac{y^{\lambda}}{1 - y^{\lambda}} \right); \quad \lambda > 0$$

$$y^{(\lambda)} = \frac{1}{\lambda} \log \left(\frac{y^{\lambda}}{1 - y^{\lambda}} \right); \quad \lambda > 0 \qquad y^{(\lambda)} = 2 \log \left(\frac{y^{\lambda}}{(1 - y)^{1 - \lambda}} \right); \quad \lambda \in (0, 1)$$

$$\lambda = 0.3467256$$

$$\lambda = 0.2620668$$

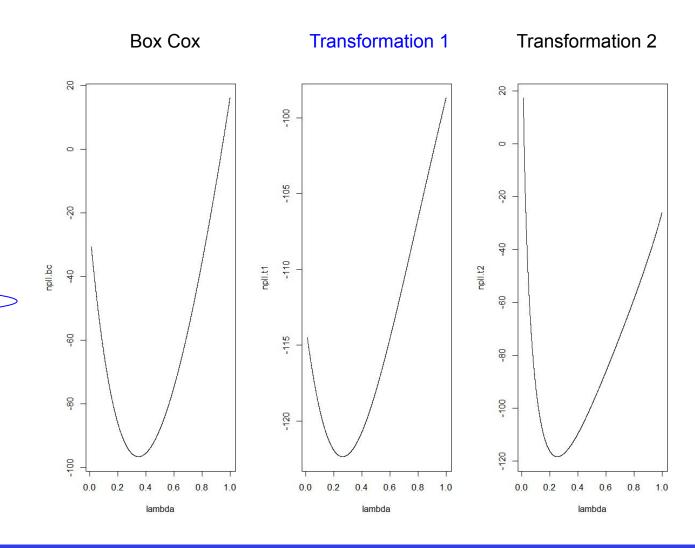
$$\lambda = 0.2523665$$

15 December 2020 Project 1: Wind power data **DTU Compute**



Power production: Normal models

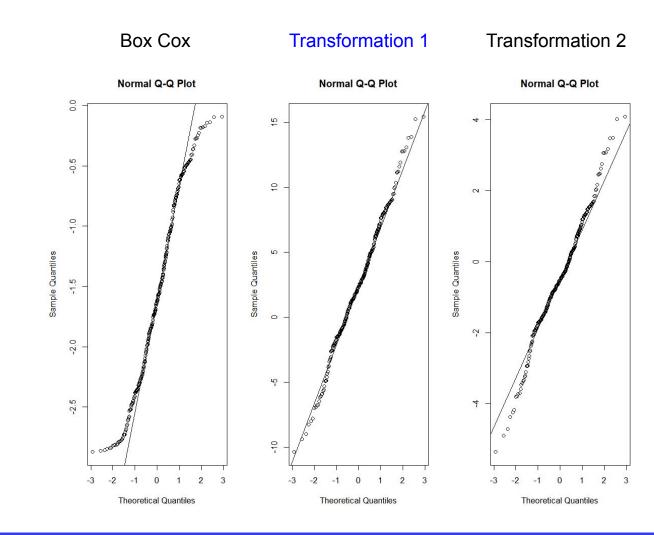
| Transformation | λ | min(npll) |
|----------------|-----------|-----------|
| Box Cox | 0.3467256 | -96.61813 |
| 1 | 0.2620668 | -122.3459 |
| 2 | 0.2523665 | -118.4156 |





Power production: Normal models

Transformation 1 is more suitable for the variable *Power production.*





Power production:

Transformation 1 to *Power production*

NON- NORMAL DISTRIBUTION

NORMAL DISTRIBUTION

FIT NON-NORMAL MODELS

NORMAL MODEL

BETA REGRESSION fits better to the variable *Power production*.

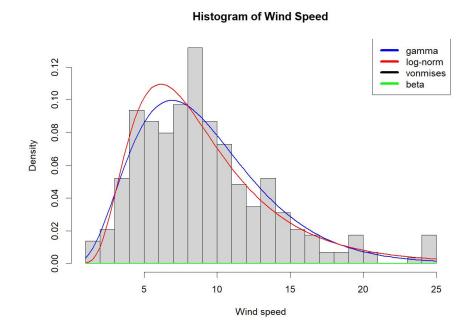
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Wind speed: non-normal models

log-normal Gamma Beta Circular (vonMises) (dgamma) (dbeta) (dlnorm) Negative log-likelihood function Optimization (nlminb) Extract parameters AIC and uncertainty

| Log norm model | | |
|----------------|------------------------|--|
| μ | 2.0844793 | |
| CI μ | [2.024706, 2.144253] | |
| σ | 0.5175574 | |
| CΙ σ | [0.4752910, 0.5598238] | |

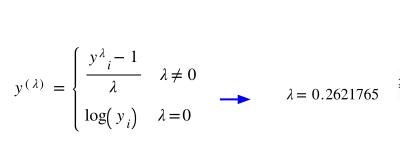


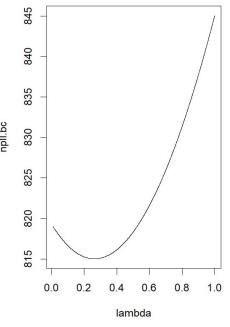
Lognorm is the most appropriate model for wind speed

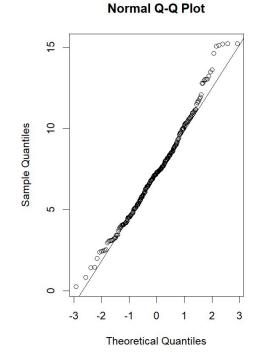


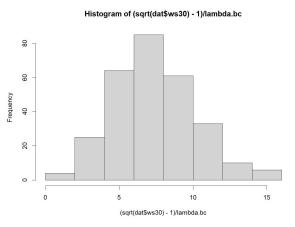
Wind speed: Normal models

Box-Cox Transformation









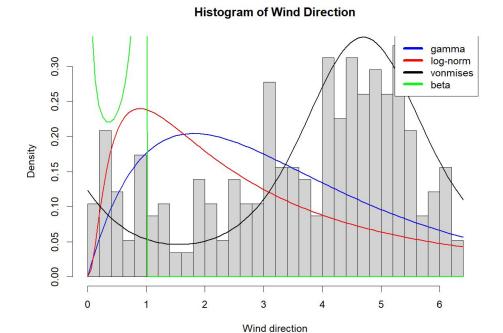
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Wind direction: non-normal models

log-normal Beta Circular Gamma (vonMises) (dgamma) (dbeta) (dlnorm) Negative log-likelihood function Optimization (nlminb) Extract parameters AIC and uncertainty

| Von Mises model | | |
|-----------------|------------------------|--|
| μ | 4.696011 | |
| CI μ | [4.474360, 4.917661] | |
| К | 1.000000 | |
| Cl ĸ | [0.8059838, 1.1940162] | |
| AIC | 1042.121 | |



Von Mises is the most appropriate model for **wind direction**.



Wind speed: non-normal models

"A new circular probability distribution which is based on **GvM** (generalization of the von Mises) is proposed. The new distribution is used to construct a joint probability distribution which is applied to fit joint distribution of linear and circular variables such as wind speed and wind direction. The results of several numerical experiments show that compared with the existing distribution models, the new circular distribution and the new constructed joint distribution in the paper can provide higher degree of the fit for the wind data under study"

(Qin et al., 2010)

www.ccsenet.org/jmr

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A New Circular Distribution and Its Application to Wind Data

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GOALS: ASSIGNMENT 2

Formulate model for predicting wind power



- 1. Consider non-normal models.
- 2. Present parameters of the final model and their uncertainty.
- 3. Interpretation of parameters and series expansions.
- 4. Graphical representation of predictions.



Normal model: $\hat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2; \quad \lambda = 0.2620668$

Im(formula = y.trans ~ dat\$ws30 + I(dat\$ws30^2), data = dat,
family = "Gaussian")

Residuals:

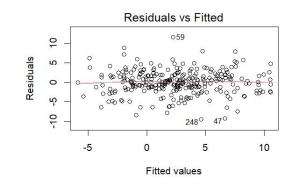
Min 1Q Median 3Q Max -9.5458 -1.4894 0.0397 1.7834 11.5948

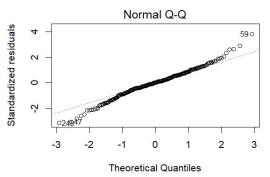
Coefficients:

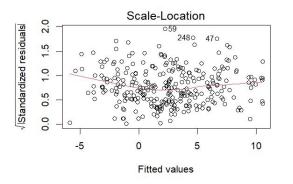
Estimate Std. Error t value Pr(>|t|)
(Intercept) -7.454787 0.735443 -10.136 < 2e-16 ***
dat\$ws30 1.410711 0.138145 10.212 < 2e-16 ***
I(dat\$ws30^2) -0.027604 0.005685 -4.855 1.98e-06 ***

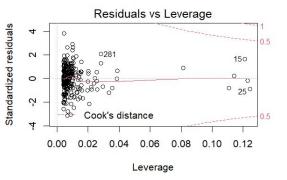
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.035 on 285 degrees of freedom Multiple R-squared: 0.587, Adjusted R-squared: 0.5841 F-statistic: 202.6 on 2 and 285 DF, p-value: < 2.2e-16











Normal model:

$$\widehat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2; \quad \lambda = 0.2620668$$

Series of expansions $\begin{cases} \widehat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \beta_3 ws^3 \\ \widehat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \beta_3 ws^3 + \beta_4 ws^4 \end{cases}$

| | df | AIC |
|---------|----|----------|
| Model 1 | 4 | 1461.802 |
| Model 2 | 5 | 1463.800 |
| Model 3 | 6 | 1465.755 |

Analysis of Variance Table

Model 1: y.trans ~ dat\$ws30 + I(dat\$ws30^2)

Model 2: y.trans ~ dat\$ws30 + I(dat\$ws30^2) + I(dat\$ws30^3)

Model 3: y.trans ~ dat\$ws30 + I(dat\$ws30^2) + I(dat\$ws30^3) + I(dat\$ws30^4)

Res.Df RSS Df Sum of Sq Pr(>Chi)

1 285 2625.4

2 284 2625.4 1 0.01904 0.9639

3 283 2625.0 1 0.41022 0.8334

| Model 1 | Estimate | CI |
|---------|-------------|----------------------------|
| βο | -7.45478717 | [-8.90237687, -6.00719746] |
| β1 | 1.41071121 | [1.13879808, 1.68262433] |
| β2 | -0.02760426 | [-0.03879467, -0.01641385] |



Normal model:
$$\hat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \beta_3 wd$$

Including variable Wind direction

Call: $Im(formula = v.trans \sim dat$ws30 + I(dat$ws30^2) + dat$wd30.$ data = dat.

family = "Gaussian")

Residuals:

Min 1Q Median 3Q Max -9.5542 -1.4635 0.0329 1.7768 11.5957

Significative p-value when chisq test

Coefficients:

Std. Error t value Pr(>|t|) Estimate -7.501578 0.796976 -9.413 < 2e-16 *** (Intercept) dat\$ws30 dat\$wd30

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.04 on 284 degrees of freedom Multiple R-squared: 0.5871, Adjusted R-squared: 0.5827 F-statistic: 134.6 on 3 and 284 DF, p-value: < 2.2e-16

AIC = 1463.778

$\hat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \beta_3 wd + \beta_4 wd^2$

Call:

Im(formula = y.trans ~ dat\$ws30 + I(dat\$ws30^2) + dat\$wd30 + I(dat\$wd30^2), data = dat, family = "Gaussian")

Residuals:

Min 1Q Median 3Q Max -9.8995 -1.6671 0.0294 1.7429 10.9922

Coefficients:

Std. Error t value / Pr(>|t|) Estimate -8.648971 0.867302 -9.972 < 2e-16 *** (Intercept) dat\$ws30 1.382843 0.136811 10.108 < 2e-16 *** I(dat\$ws30^2) -0.026749 0.005617 -4.762 3.07e-06 *** datŚwd30 1.250221 0.410146 3.048 0.00252 ** Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.995 on 283 degrees of freedom Multiple R-squared: 0.6007, Adjusted R-squared: 0.5951 F-statistic: 106.5 on 4 and 283 DF, p-value: < 2.2e-16

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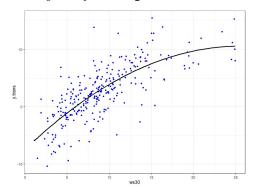
AIC = 1456.084



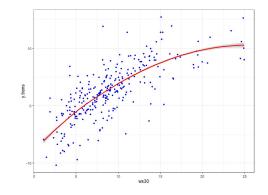
Normal model: Parameter uncertainty of the two best normal models

| | Model 1 (without wd) | Model 2 (with wd) |
|-------|----------------------------|-----------------------------|
| βΟ | -7. 45478717 | -8.6489713 |
| СІ ВО | [-8.90237687, -6.00719746] | [-10.35615292, -6.94178972] |
| β1 | 1.41071121 | 1.3828433 |
| СІ β1 | [1.13879808, 1.68262433] | [1.11354616, 1.65214036] |
| β2 | -0.02760426 | -0.0267491 |
| СІ β2 | [-0.03879467, -0.01641385] | [-0.03780599, -0.01569222] |
| β3 | - | 1.2502209 |
| СІ βЗ | - | [0.44289656, 2.05754528] |
| B4 | - | -0.1970357 |
| СІ β4 | - | [-0.32164310, -0.07242825] |

$$\hat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2; \quad \lambda = 0.2620668$$



$$\widehat{y}^{(\lambda)} = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \beta_3 wd + \beta_4 wd^2$$



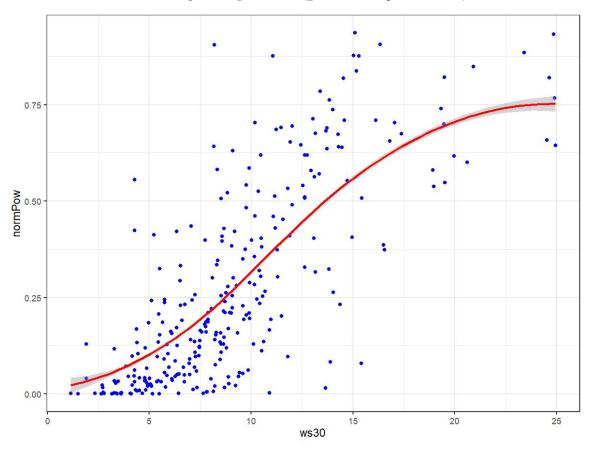


Non-normal model: Beta regression

| | Estimate value | CI |
|----|----------------|----------------------------|
| βΟ | -4.418622074 | [-4.98896968, -3.84827447] |
| β1 | 0.402275796 | [0.31733284, 0.48721875] |
| β2 | -0.007969534 | [-0.01135872, -0.00458035] |
| β3 | 0.349155870 | [0.10717155, 0.59114020] |
| B4 | -0.052658876 | [-0.08983082, -0.01548693] |
| Φ | 5.298284 | [4.42934017, 6.16722725] |

Best model! AIC = -484.3728

$$y = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \beta_3 wd + \beta_4 wd^2$$



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Project 1: Wind power data



GOALS: ASSIGNMENT 3



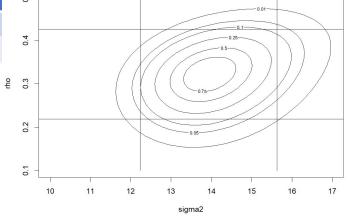
Analysis of autocorrelation AR(1)



AR(1) model:
$$\varepsilon_i = \phi \varepsilon_{i-1} + u_i$$
; $u_i \sim N(0, \sigma_u^2)$

$$Y^{(\,0.2)} = \beta_0 + \beta_1 ws + \beta_2 ws^2 + \varepsilon \quad ; \ \varepsilon \sim N \big(\,0, \ \sigma^2 \big)$$
 Extract residuals
$$\varepsilon = \begin{bmatrix} \varepsilon_1 & \varepsilon_2 \\ . & . \\ \varepsilon_{n-1} & \varepsilon_n \end{bmatrix}$$
 Fit to the model
$$\begin{bmatrix} \varepsilon_i, \ \varepsilon_{i+1} \end{bmatrix}^T \sim N (\,0, \ \Sigma) \; ; \quad \Sigma = \sigma^2 \begin{bmatrix} \ 1 & p \\ p & 1 \end{bmatrix}$$

| Parameter | Estimate value | CI |
|------------|----------------|------------------------|
| σ^2 | 13.93593 | [12.24160 , 15.63025] |
| р | 0.3222538 | [0.2185932, 0.4259144] |



Hypothesis testing: $H_0: \rho = 0$

Wilk's likelihood ratio statistic

$$W = 2\log\left(\frac{L(\rho_0)}{L(\widehat{\rho})}\right) \longrightarrow X^2 \qquad \text{p-value} = 1.468168e-07$$

Wald test

$$z = \frac{\widehat{\rho} - \rho_0}{se(\widehat{\rho})} \longrightarrow N(0,1) \longrightarrow p\text{-value} = 9.658089e-09$$

Reject null hypothesis

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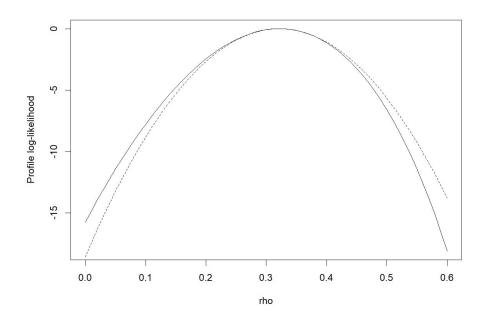
$$\left[\varepsilon_{i}, \ \varepsilon_{i+1}\right]^{T} \sim N(0, \ \Sigma); \quad \Sigma = \sigma^{2} \left[\begin{array}{c} 1 \ p \\ p \ 1 \end{array}\right]$$

Analytical method

$$I(\sigma^{2}, \rho) = \begin{pmatrix} \frac{n}{\sigma^{4}} & -\frac{n\rho}{\sigma^{2}(1-\rho^{2})} \\ -\frac{n\rho}{\sigma^{2}(1-\rho^{2})} & \frac{n(1+\rho^{2})}{(1-\rho^{2})^{2}} \end{pmatrix} \longrightarrow \begin{bmatrix} 1.477781 & -7.405631 \\ -7.405631 & 394.481927 \end{bmatrix}$$

Numerical method

Profile likelihood and quadratic approximation



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AR(1) model: $\varepsilon_i = \phi \varepsilon_{i-1} + u_i$; $u_i \sim N(0, \sigma_u^2)$

Call:

ar1

arima(x = y.trans, order = c(1, 0, 0), xreg = xreg)

Coefficients:

ar1 intercept ws30 ws30sq 0.3252 -6.7491 1.6170 -0.0282 s.e. 0.0559 0.9351 0.1731 0.0074

sigma^2 estimated as 12.44: log likelihood = -771.69, aic = 1553.39

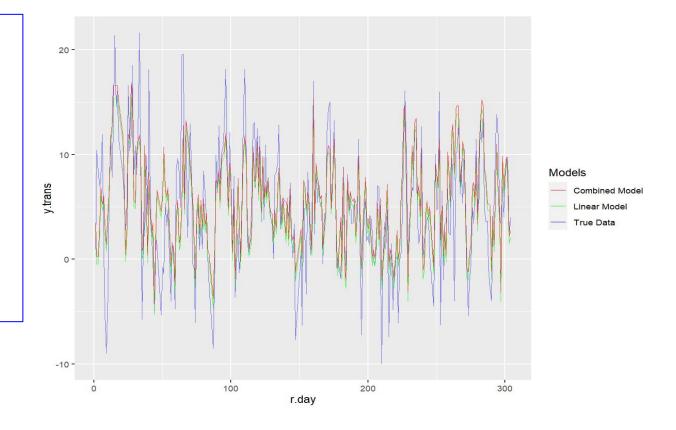
2.5 % 97.5 % 0.21558497 0.4347553 Intercept -8.58184971 -4.9163993 1.27766793 1.9563241

ws30 -0.04262742 -0.0138018 ws30sq

AIC linear model: 1583.305

AIC combined model: 1553.390

Better fit than linear model!



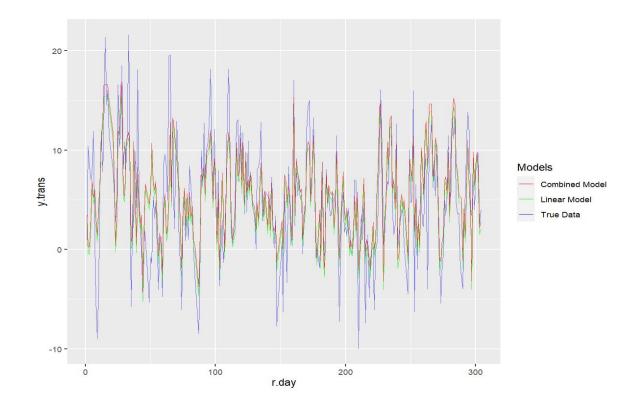
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AR(1) model:
$$\varepsilon_i = \phi \varepsilon_{i-1} + u_i$$
; $u_i \sim N(0, \sigma_u^2)$

| MAE (Mean Absolute Error) | Linear model | Combined model |
|---------------------------|--------------|----------------|
| Long term | 2.796916 | 2.848373 |
| Short term (last 3 days) | 10.85968 | 9.804569 |

AR(1) model more suitable for short term and Linear model more suitable for long term.





References

Pawitan Y. In All Likelihood: Statistical Modelling and Inference Using Likelihood. OUP Oxford; 2001. (Oxford science publications)

Code for the project can be found at <u>Statistical Modelling</u>

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