

STATISTICAL MODELLING: Theory and practice

Project 3: Financial data

GOALS: ASSIGNMENT 1

- Present the data
- Fit and asses **normal model**
- Present a **new hypothetically better model**
- Discuss which model is better

The financial data set

Weekly returns from Exchange Traded Fund (ETF)

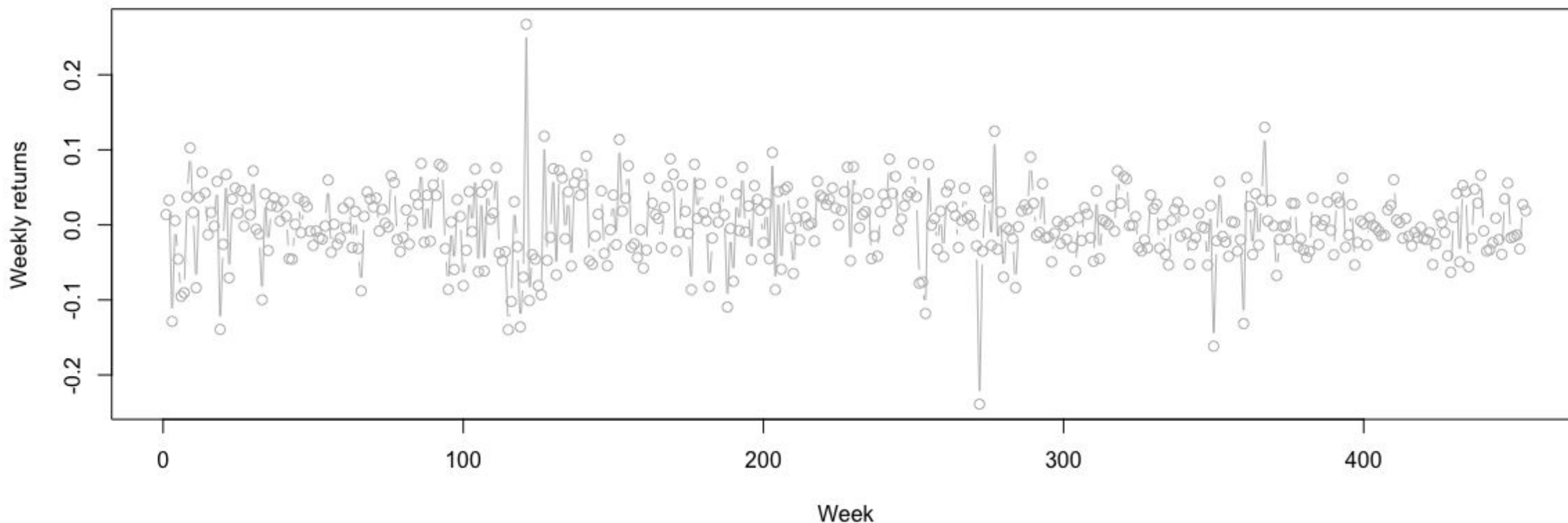
$$\text{weekly returns} = \frac{\text{final price}}{\text{initial price}} - 1$$

Data set

	time	SLV
1	2006-5-5	0.01376
2	2006-5-12	0.03286
3	2006-5-19	-0.12863
...
452	2015-4-24	-0.03213
453	2015-5-1	0.02722
454	2015-5-8	0.01875

Summary statistics of weekly returns

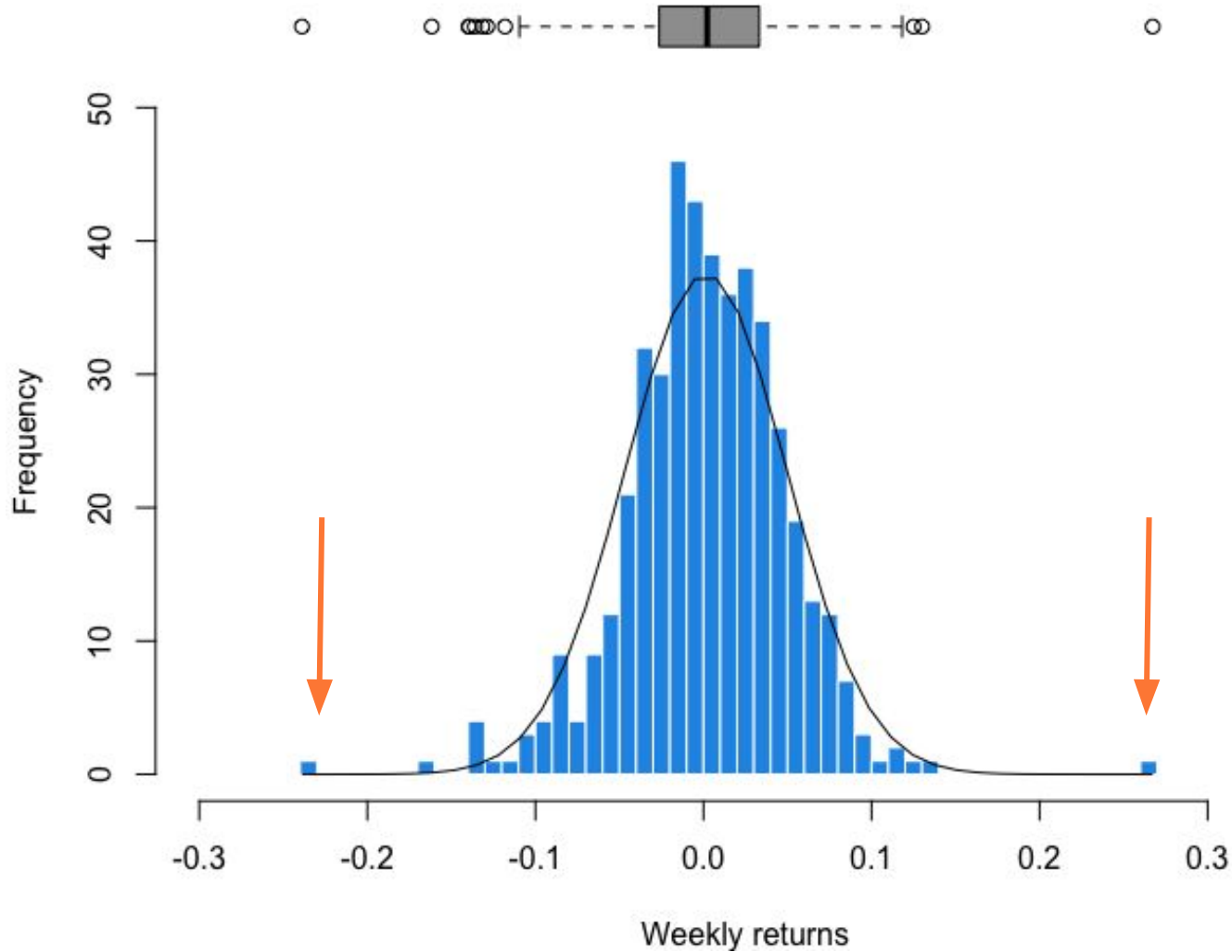
	SLV
Min. :	-0.238893
1st Qu.:	-0.026350
Median :	0.002226
Mean :	0.001468
3rd Qu.:	0.033122
Max. :	0.267308



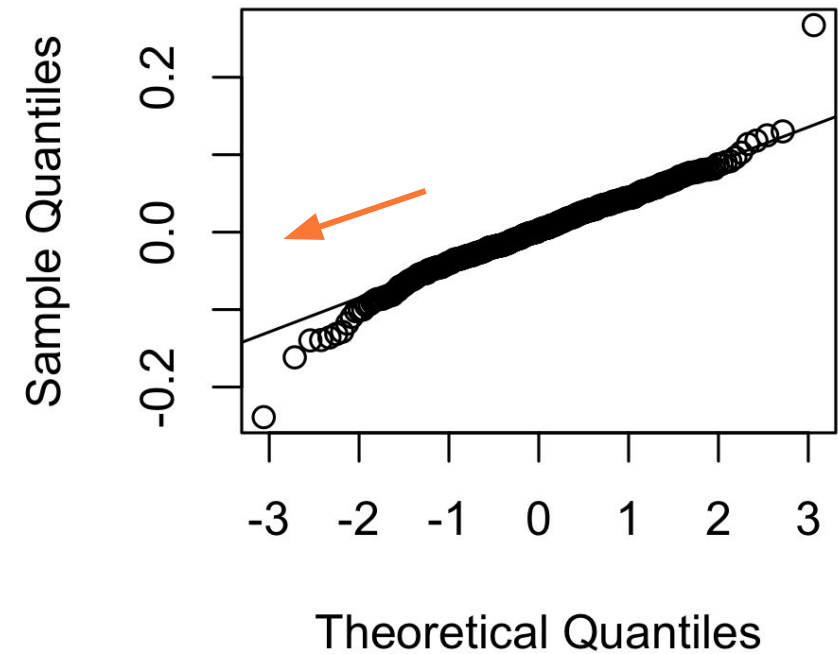
Fit to normal distribution

$$f_0(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

$$\log L(\mu, \sigma^2) = -\frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_i (x_i - \mu)^2,$$



Normal Q-Q Plot



Normal distribution

$$f_0(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

$$\log L(\mu, \sigma^2) = -\frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_i (x_i - \mu)^2,$$

Normal Model

	Est.	2.5%	97.5%
mu	0.001467	-0.00297	0.00591
sigma	0.04830	0.04385	0.05274

New model hypothesis: Cauchy

Cauchy distribution for heavy tails

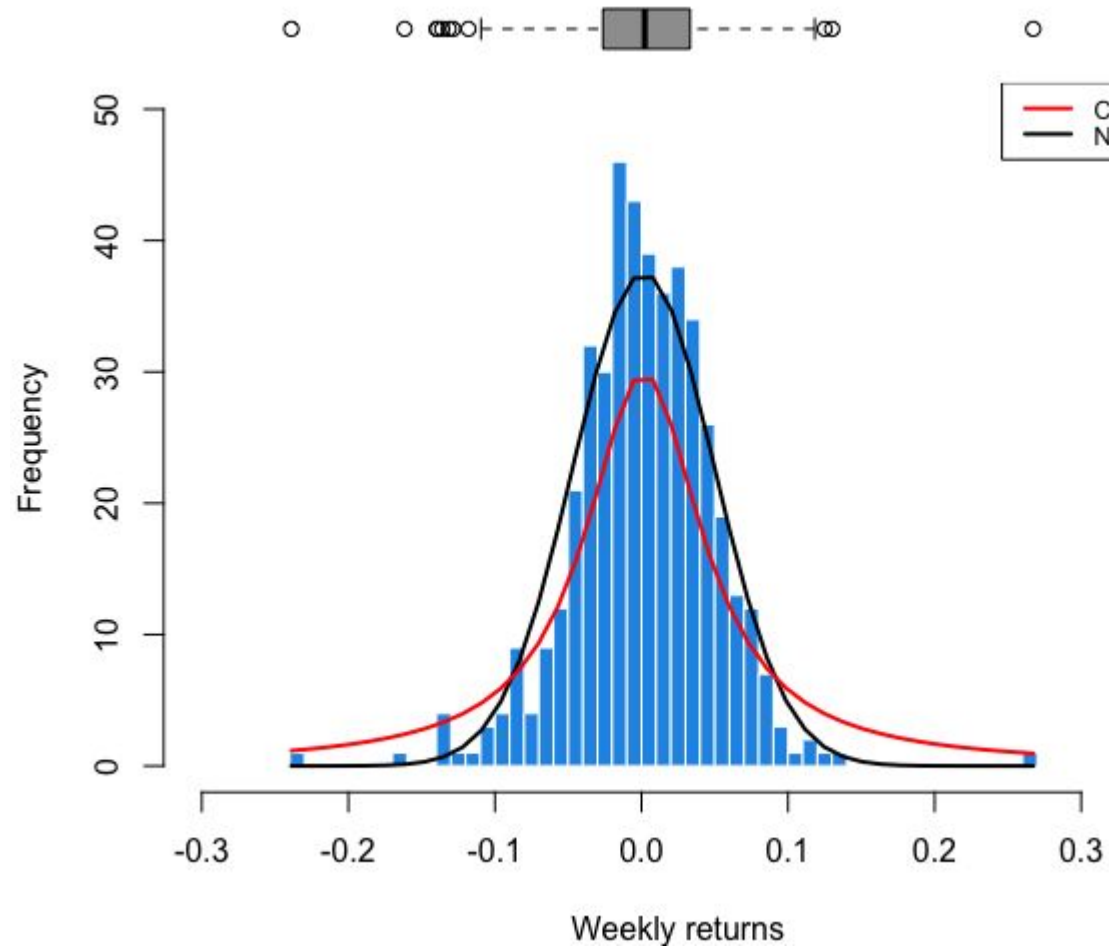
$$f_0(x) = \frac{1}{\pi(1+x^2)}$$

$$L(\mu, \sigma) = \prod_i \frac{1}{\sigma} \left\{ 1 + \frac{(x_i - \mu)^2}{\sigma^2} \right\}^{-1}$$

Cauchy Model

	Est.	2.5%	97.5%
mu	0.002653	-0.00144	0.00579
sigma	0.027	0.0229	0.0301

NORMAL vs CAUCHY



Model	AIC
Normal	-1460
Cauchy	-1363.414

Cauchy distribution could be more suitable for finance data analysis because of the **heavy tails probabilities**, which decay much more slowly.

This would need further analysis in order to make a final decision on the model.

GOALS: ASSIGNMENT 3

Mixture models

- 1) Fit a **normal mixture model** :
 - 2 components
 - 3 components
- 2) **Compare models**
- 3) Report **confidence interval** for the parameters
- 4) **Profile likelihood** of one of the variance parameters.
- 5) **Reparametrize** the model to obtain one maximum

HMM models

- 1) Fit **normal Hidden Markov Model** with 2 and 3 states
- 2) Find CI 95% for **working parameters** and report **natural parameters** and their CI 95%
- 3) Plot long term distribution and 1-step ahead distribution - Forecasting
- 4) Discuss how to do short term prediction

Unconstrained optimizer : `nlm`

Natural parameters

$$\sigma_i = \exp(\rho_i), i = 1, \dots, m$$

$$\delta_i = \frac{\exp(\tau_i)}{1 + \sum_{j=2}^m \exp(\tau_j)}, i = 2, \dots, m$$

$$\delta_1 = 1 - \sum_{j=2}^m \delta_j$$

Working parameters

$$\rho_i = \log(\sigma_i), i = 1, \dots, m$$

$$\tau_i = \log\left(\frac{\delta_i}{1 - \sum_{j=2}^m \delta_j}\right), i = 2, \dots, m$$

2 Components :

$$\delta_1 N(\mu_1, \sigma_1^2) + \delta_2 N(\mu_2, \sigma_2^2)$$

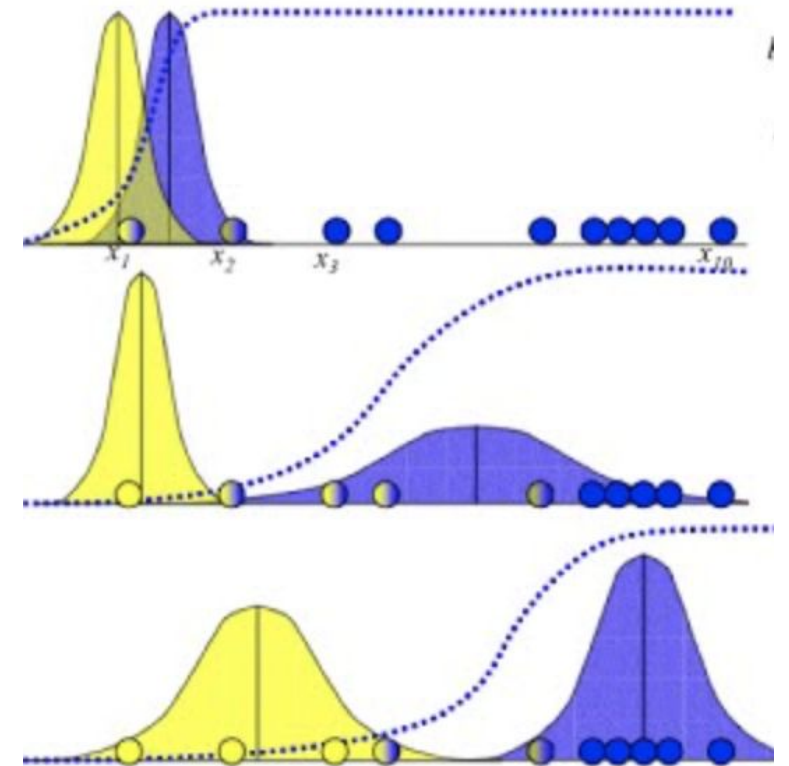
3 Components :

$$\delta_1 N(\mu_1, \sigma_1^2) + \delta_2 N(\mu_2, \sigma_2^2) + \delta_3 N(\mu_3, \sigma_3^2)$$

Likelihood (m components)

$$\log L(\theta; y) = \sum_i \log \sum_{m=1}^M \delta_m N_m(y_i | \mu_m, \sigma_m^2)$$

EM algorithm for mixture models



V. Lavrenko, 2014

Fit a normal mixture model

Unconstrained optimizer : `nlm`

Natural parameters

$$\sigma_i = \exp(\rho_i), i = 1, \dots, m$$

$$\delta_i = \frac{\exp(\tau_i)}{1 + \sum_{j=2}^m \exp(\tau_j)}, i = 2, \dots, m$$

$$\delta_1 = 1 - \sum_{j=2}^m \delta_j$$

Working parameters

$$\rho_i = \log(\sigma_i), i = 1, \dots, m$$

$$\tau_i = \log\left(\frac{\delta_i}{1 - \sum_{j=2}^m \delta_j}\right), i = 2, \dots, m$$

Model	m	AIC
Normal	1	-1460
Normal	2	-1489.644
Normal	3	-1484.256

2 Components :

$$\delta_1 N(\mu_1, \sigma_1^2) + \delta_2 N(\mu_2, \sigma_2^2)$$

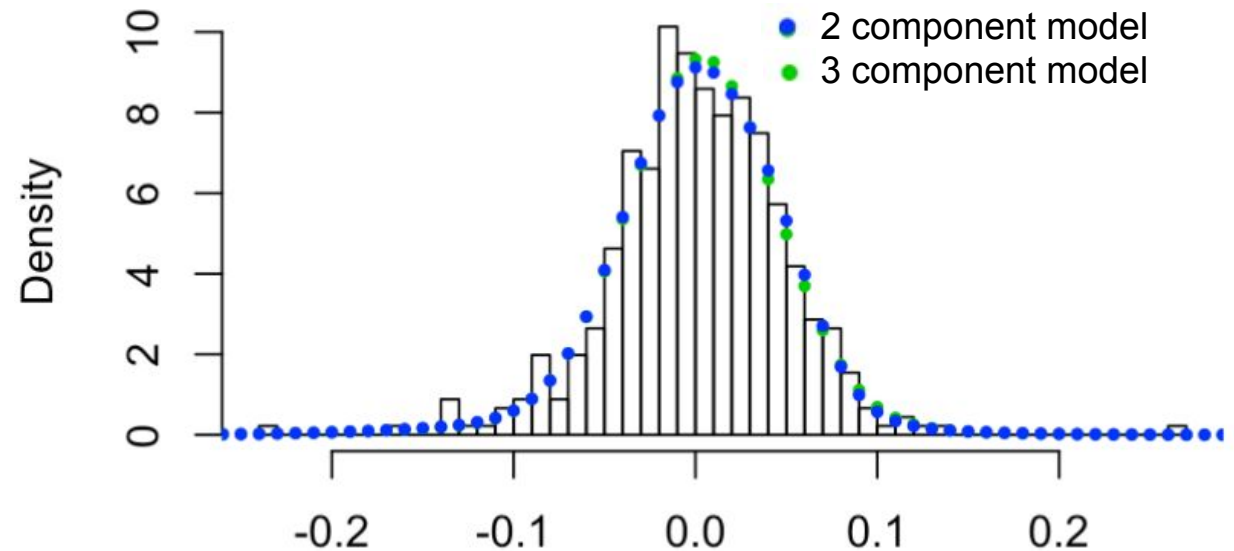
3 Components :

$$\delta_1 N(\mu_1, \sigma_1^2) + \delta_2 N(\mu_2, \sigma_2^2) + \delta_3 N(\mu_3, \sigma_3^2)$$

Likelihood (m components)

$$\log L(\theta; y) = \sum_i \log \sum_{m=1}^M \delta_m N_m(y_i | \mu_m, \sigma_m^2)$$

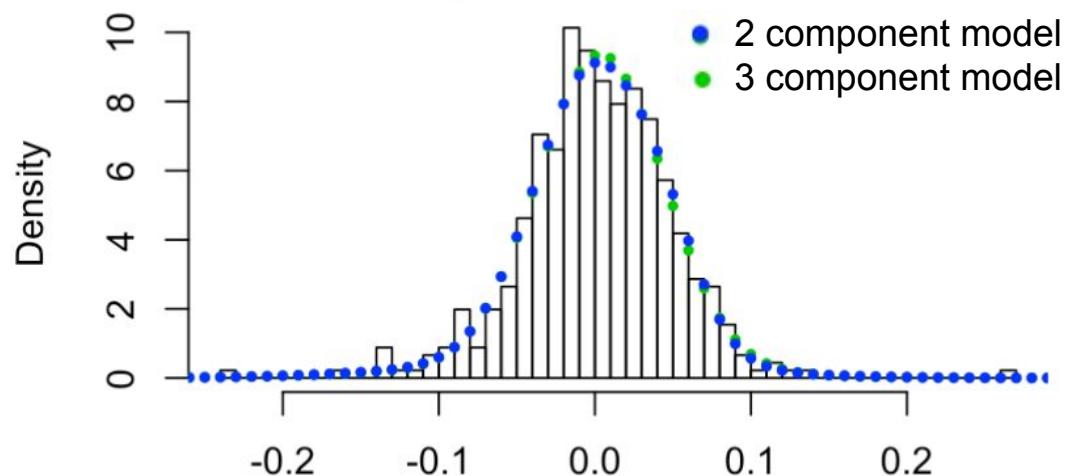
Estimated normal distributions



MIXTURE MODELS

Compare models and report CI

Model	m	AIC
Normal	1	-1460
Normal	2	-1489.644
Normal	3	-1484.256



Parameter estimation and CI for m=2

Wald confidence intervals of working parameters:

$$CI(\sigma_i) = \exp\left(\hat{\rho}_i \pm z_{1-\frac{\alpha}{2}} \cdot se(\hat{\rho}_i)\right), i = 1, \dots, k$$

Wald interval simulation from distribution

$$\hat{\theta} \sim N(\theta, \mathcal{I}^{-1}(\theta))$$

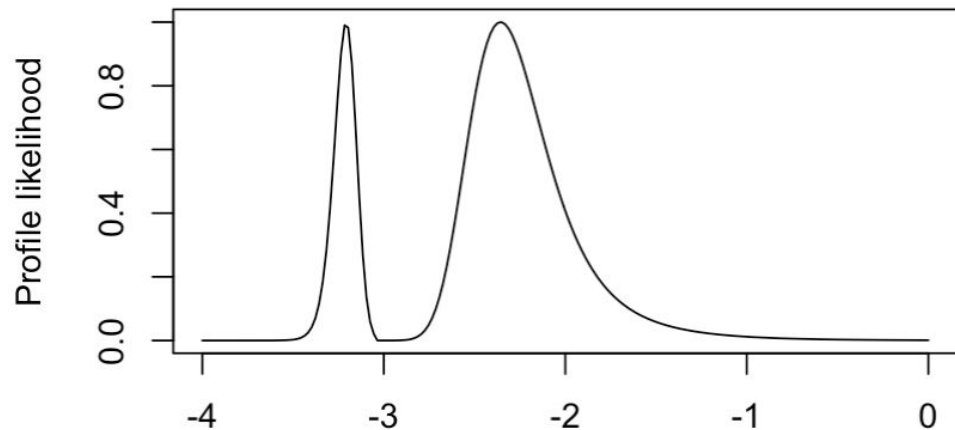
CI from quantiles of 100.000 samples from $N(\hat{\theta}, I^{-1}(\hat{\theta}))$
Transformed back to natural deltas.

	Parameter(N)	Confidence interval (0.025 - 0.975)
μ_1	0.0039	[-0.0007570256 , - 0.0086514630]
μ_2	-0.0251	[-0.06722030 , 0.01692244]
σ_1	0.04046	[0.03592759 , 0.04557654]
σ_2	0.09472	[0.06251529 , 0.14351169]
δ_1	0.9147814	[0.7210504 , 0.9778096]
δ_2	0.08521855	[0.02219040 , 0.27894959]

Profile Likelihood - Nuisance parameter

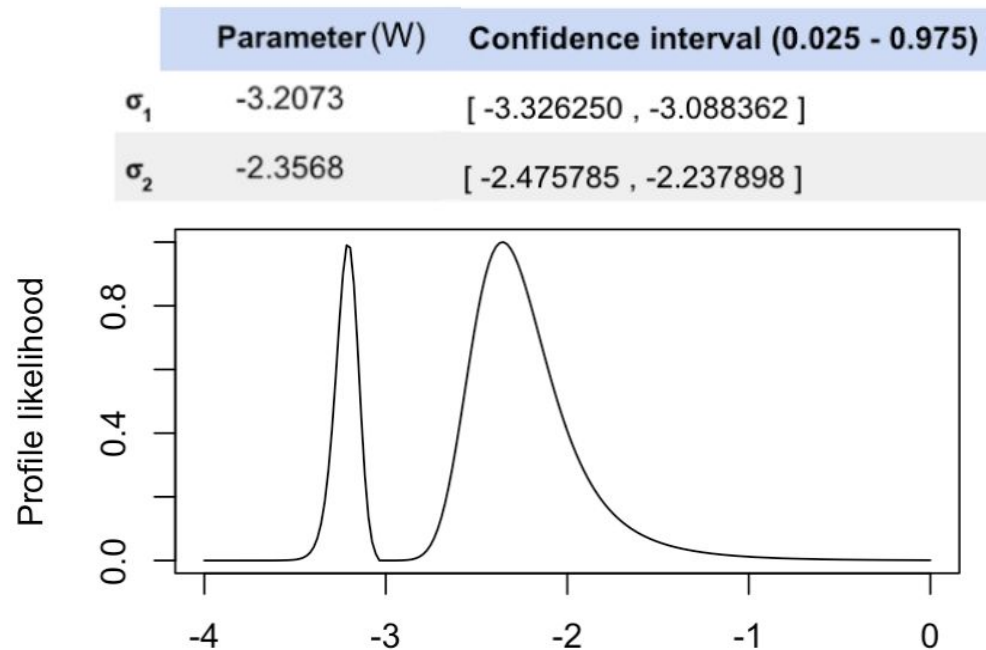
$$\log L(\hat{\mu}_1, \hat{\mu}_2, \sigma_1^2, \hat{\sigma}_2^2; y) = \sum_i \log \delta_1 N(y_i | \hat{\mu}_1, \sigma_1^2) + \log \delta_2 N(y_i | \hat{\mu}_2, \hat{\sigma}_2^2)$$

	Parameter (W)	Confidence interval (0.025 - 0.975)
σ_1	-3.2073	[-3.326250 , -3.088362]
σ_2	-2.3568	[-2.475785 , -2.237898]



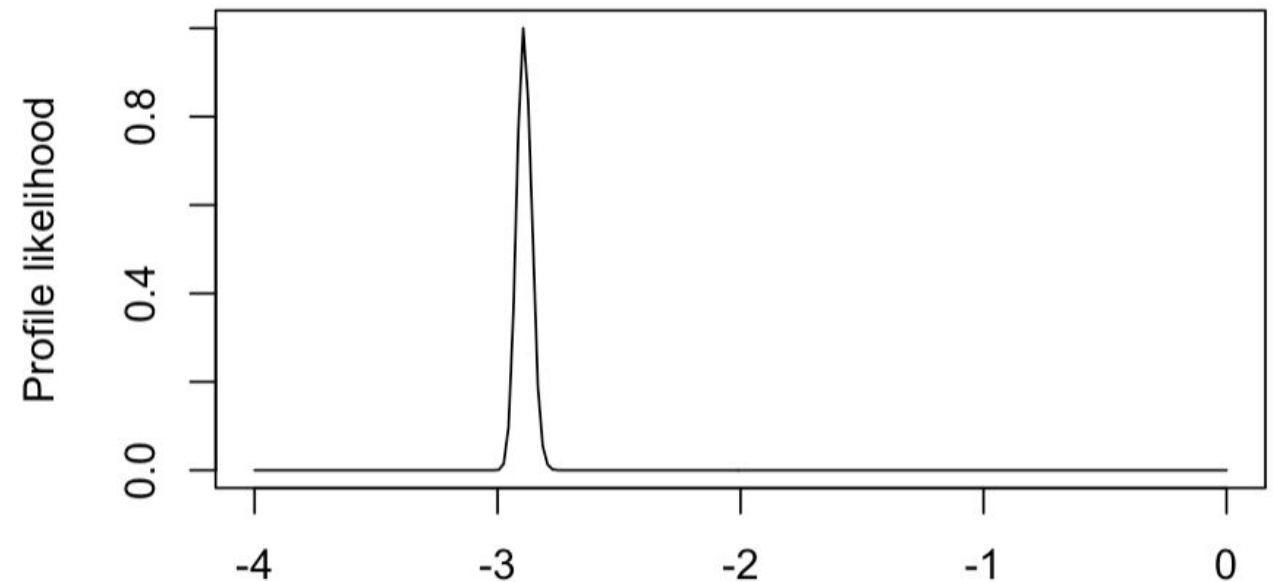
Profile Likelihood - Nuisance parameter

$$\log L(\hat{\mu}_1, \hat{\mu}_2, \sigma_1^2, \hat{\sigma}_2^2; y) = \sum_i \log \delta_1 N(y_i | \hat{\mu}_1, \sigma_1^2) + \log \delta_2 N(y_i | \hat{\mu}_2, \hat{\sigma}_2^2)$$



Profile Likelihood - Reparametrization

$$\log L(\hat{\mu}_1, \hat{\mu}_2, \sigma_1^2, \hat{\sigma}_2^2 + \sigma_1^2; y) = \sum_i \log \delta_1 N(y_i | \hat{\mu}_1, \sigma_1^2) + \log \delta_2 N(y_i | \hat{\mu}_2, \hat{\sigma}_2^2 + \sigma_1^2)$$



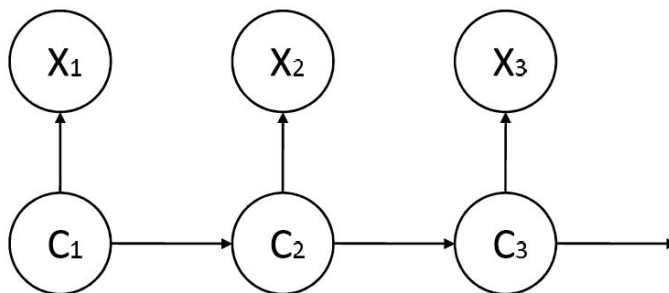
HMM Normal models

Natural to working parameters

$$\mu_t = \mu$$

$$\sigma_t = \log(\sigma)$$

$$\tau_{ij} = \log\left(\frac{\gamma_{ij}}{1 - \sum_{k \neq i} \gamma_{ik}}\right), i = 1, \dots, m, j = 2, \dots, m$$



Working to natural parameters

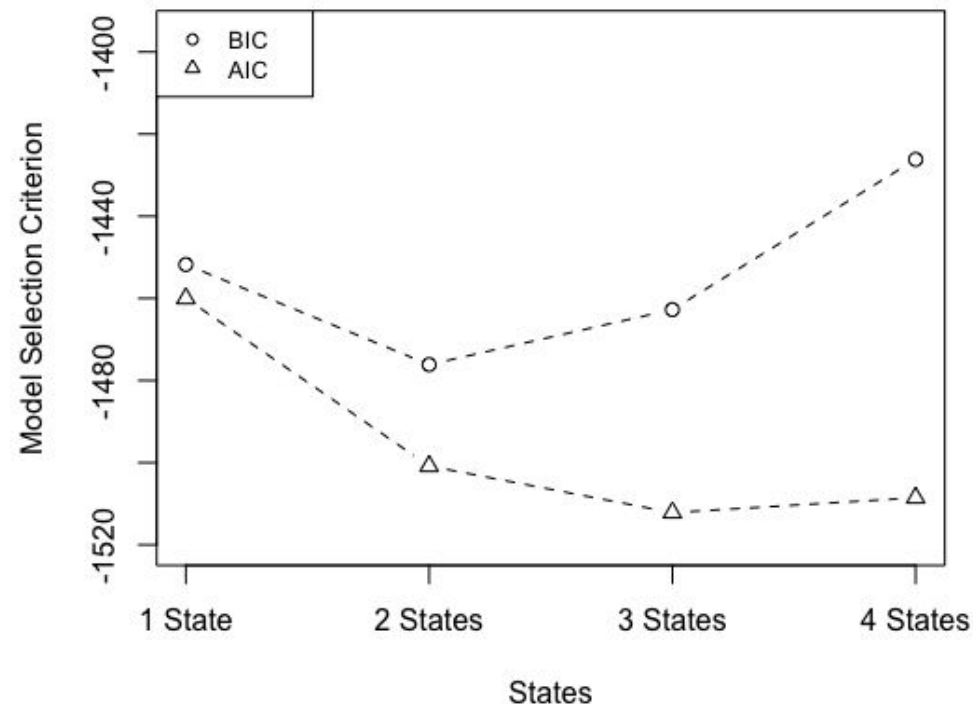
$$\mu = \mu_t$$

$$\sigma = \exp(\sigma_t)$$

$$\gamma_{ij} = \frac{\rho_{ij}}{1 + \sum_{k \neq i} \exp(\tau_{ik})}, i, j = 1, \dots, m$$

where

$$\rho_{ij} = \begin{cases} \exp(\tau_{ik}) & i \neq j \\ 1 & i = j \end{cases}$$



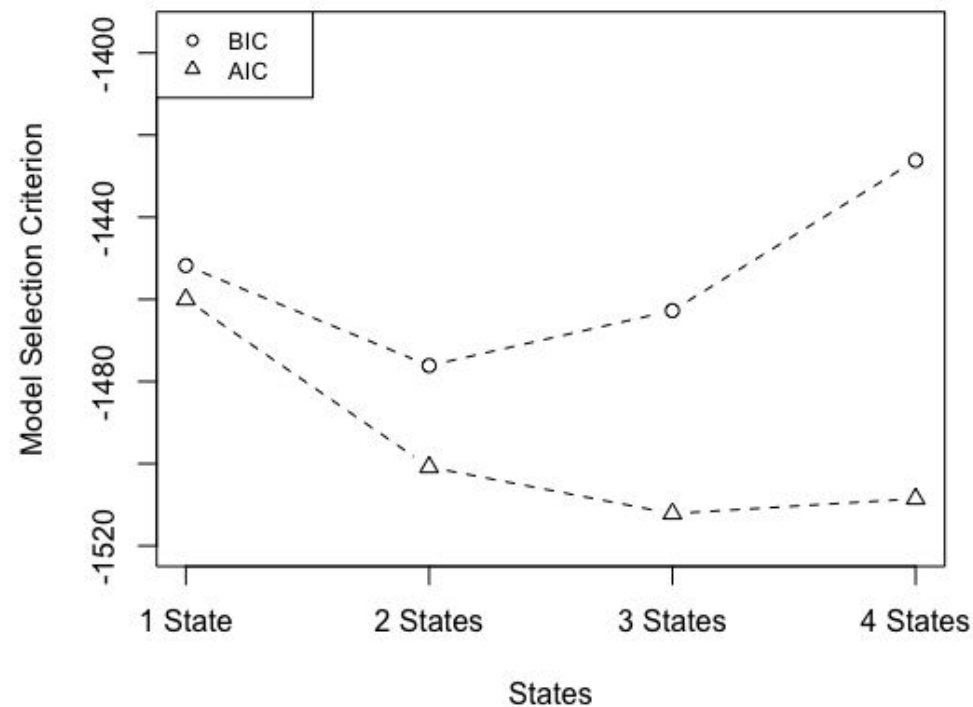
# states	Degrees of freedom	Log-Likelihood	AIC
1	2	731.9998	-1460.000
2	6	756.4172	-1500.834
3	12	768.0791	-1512.158
4	20	774.2719	-1508.544

```

norm.HMM.mllk <- function(parvect,x,m,...)
{
  # print(parvect)
  if(m==1) return(-sum(dnorm(x, parvect[1], exp(parvect[2]), log=TRUE)))
  n      <- length(x)
  pn     <- norm.HMM.pw2pn(m,parvect)

  allprobs <- matrix(nrow = n, ncol = m)
  for (j in 1:m){
    allprobs[,j] = dnorm(x, pn$mu[j], pn$sigma2[j])
  }
  allprobs <- ifelse(!is.na(allprobs),allprobs,1)
  lscale  <- 0
  foo     <- pn$delta
  for (i in 1:n)
  {
    foo <- foo%*%pn$gamma*allprobs[i,]
    sumfoo <- sum(foo)
    lscale <- lscale+log(sumfoo)
    foo <- foo/sumfoo
  }
  mllk <- -lscale
  mllk
}

```



# states	Degrees of freedom	Log-Likelihood	AIC
1	2	731.9998	-1460.000
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3	12	768.0791	-1512.158
4	20	774.2719	-1508.544

HMM Model with 3 states

Working parameters

	Estimate	2.5%	97.5%
mu1	0.0119	0.0039	0.01993
mu2	-0.0026	-0.0078	0.0026
mu3	-0.0332	-6.634e-02	-5.89e-05
sigma21	-3.1104	-3.268	-2.9528
sigma22	-3.5034	-3.6368	-3.37
sigma23	-2.4825	-2.7552	-2.21
tau21	-28.3362	NaN	NaN
tau31	-1.018	-2.1914	0.1555
tau12	-4.0342	-5.3756	-2.6927
tau32	-19.4691	-21.3123	-17.626
tau13	-3.0867	-4.908	-1.2654
tau23	-3.8643	-5.271	-2.4576

Natural parameters

$$\mu_1 = 0.0119 \quad \mu_2 = -0.0026 \quad \mu_3 = -0.0332$$

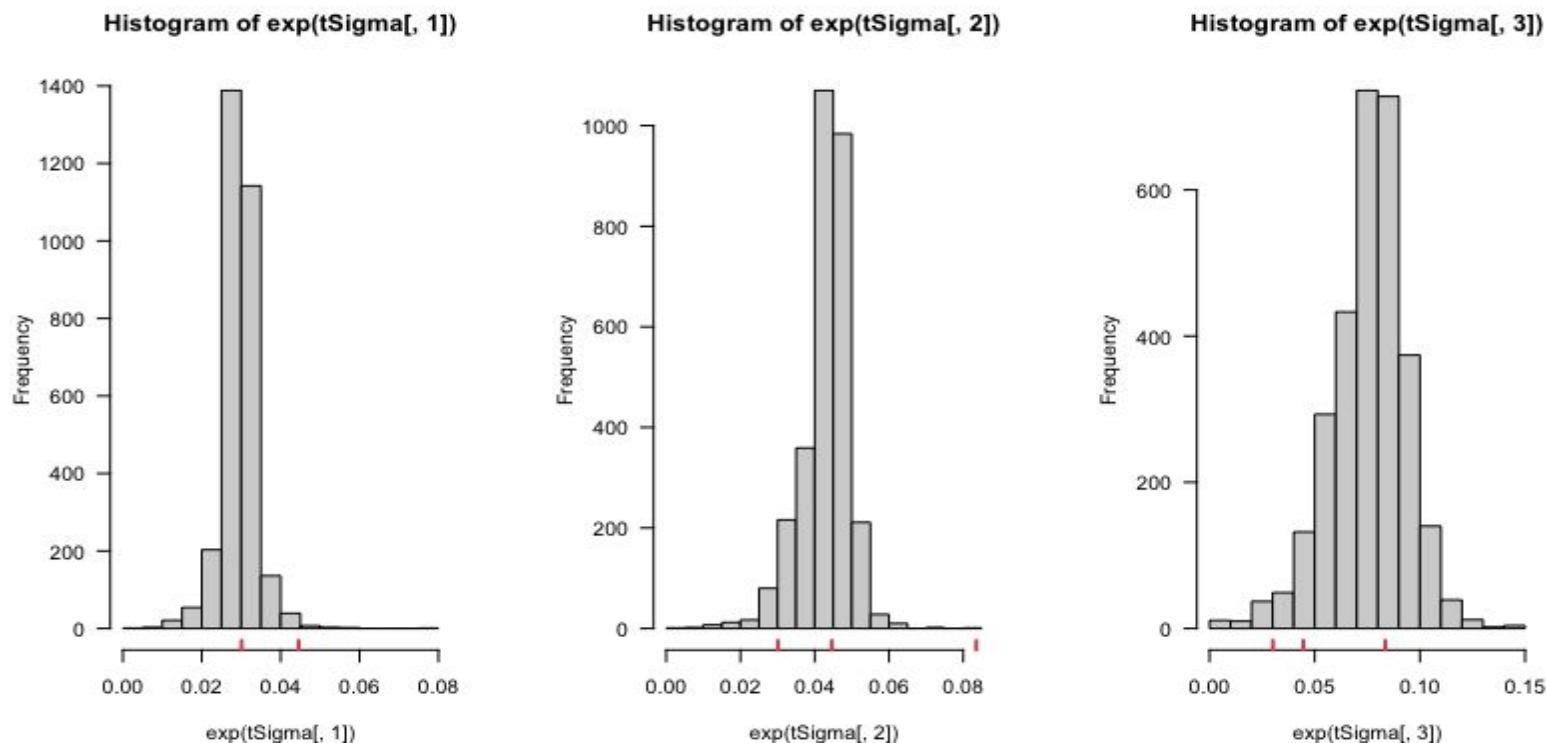
$$\sigma_1 = 0.0446 \quad \sigma_2 = 0.03 \quad \sigma_3 = 0.0835$$

$$\delta = [0.4915 \quad 0.3982 \quad 0.1103]$$

$$\tau_1 = \begin{bmatrix} 0.9404 & 0.0166 & 0.0429 \\ 0.0000 & 0.9795 & 0.0205 \\ 0.2654 & 0.0000 & 0.7346 \end{bmatrix}$$

HMM Model with 3 states

NATURAL PARAMETERS CI 95% BOOTSTRAP WITH K= 3000



STEPS:

1. Generate a sample from the MLE
2. Fit new model to the sample
3. Store the MLE of the parameters estimated in the new distribution

Repeat 1-3 k times

	MLE	2.5%	97.5%
σ_1	0.04458	0.0198	0.0389
σ_2	0.03	0.0291	0.0534
σ_3	0.08353	0.034	0.1082

	MLE	2.5%	97.5%
μ_1	0.01193	-0.0110	0.0119
μ_2	-0.00258	-0.0048	0.0311
μ_3	-0.03319	-0.1495	0.0088

Make short term predictions

- Using **Viterbi algorithm** to decode state sequence until now, and predict next week's state
- Look at previous observations which were observed after the **same state transition** as the upcoming one

References

Pawitan Y. In All Likelihood: Statistical Modelling and Inference Using Likelihood. OUP Oxford; 2001. (Oxford science publications)

Code for the project can be found at [Statistical Modelling](#)

DTU



Long term and 1-step ahead model forecast

```
normal.HMM.forecast <- function(xf, h=1, m, x, mod){
  n <- length (x)
  nxf <- length (xf)
  dxf <- matrix (0, nrow =h, ncol = nxf)
  foo <- mod$delta * dnorm (x[1] , mod$mu, mod$sigma2 )
  sumfoo <- sum (foo)
  lscale <- log ( sumfoo )
  foo <- foo / sumfoo
  for (i in 2:n){
    foo <- foo %*% mod$gamma * dnorm(x[i], mod$mu, mod$sigma2)
    sumfoo <- sum( foo)
    lscale <- lscale + log ( sumfoo )
    foo<- foo / sumfoo
  }
  for (i in 1:h)
  {
    foo <- foo %*% mod$gamma
    for (j in 1: m ) dxf[i ,] <- dxf[i ,] +
      foo [j]* dnorm (xf , mod$mu[j], mod$sigma2)
  }
  return ( dxf)
}
```

We wanted to do 1-step ahead with the function on the left, to compute the marginal distribution of the data.

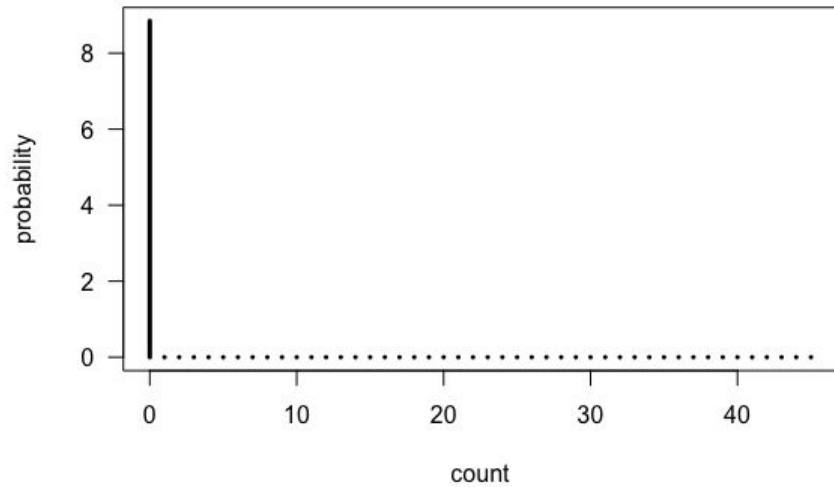
To do 1-step ahead we would set the $h=1$, 1 year ahead and we would use our 3-states model (mod3s) with the argument stationary = TRUE.

On the other hand to do the the long term prediction, we use our initial 3-state model, this time not taking the stationary argument (fit3). We use our mu, and sigma2 and extract the marginal distribution with the code below:

```
m <- -3
xf <- 0:45
mu <- fit3$mu
sigma2 <- fit3$sigma2
delta <- solve (t( diag (m)- fit3$gamma +1) ,rep (1,m))
dstat <- numeric ( length (xf))
for (j in 1:m) dstat <- dstat + delta [j]* dnorm(xf , mu[j], sigma2[j])
```

Long term and 1-step ahead model forecast

Financial series: forecast distribution for 455



1 step ahead

Long term prediction

