07 | Linear Regression

Ivan Corneillet

Data Scientist



Learning Objectives

After this lesson, you should be able to:

- Define simple linear regression
- Build a linear regression model using statsmodels
- Evaluate model fit using statistical analysis (t-tests, p-values, t-values, confidence intervals)

Here's what's happening today:

- Simple Linear Regression
 - Interpreting the regression's coefficients
 - Are these coefficients significant?
 - Common Regression Assumptions
 - Model Fit and R^2
- The Normal Distribution
 - ► The 68 90 95 99.7 Rule

- Hypothesis Testing
 - Two-Tail Hypothesis Testing
 - t-values
 - p-values
 - Confidence Intervals
- The Student's t-distribution



Simple Linear Regression

Simple Linear Regression

• The simple linear regression model captures a linear relationship between a single feature variable *x* and a response variable *y*

$$y = \beta_0 + \beta_1 \cdot x + \varepsilon$$

- y is the **response** variable (what we want to predict);
 also called *dependent* variable, *endogenous* variable, or
 regressand
- *x* is the **feature** variable (what we use to train the model); also called *explanatory* variable, *independent* variable, *exogenous* variable, or *regressor*
- β_0 and β_1 are the **regression's coefficients**; also called the *model's parameters*
 - β_0 is the line's intercept; β_1 is the line's slope
- ε is the **error** term; also called the residual

Simple Linear Regression (cont.)

• Given $x = (x^{(1)}, x^{(2)}, ..., x^{(n)}), y =$ $(y^{(1)}, y^{(2)}, ..., y^{(n)}), \text{ and } \varepsilon = (\varepsilon^{(1)}, \varepsilon^{(2)}, ..., \varepsilon^{(n)}),$ we can formulate the linear model as

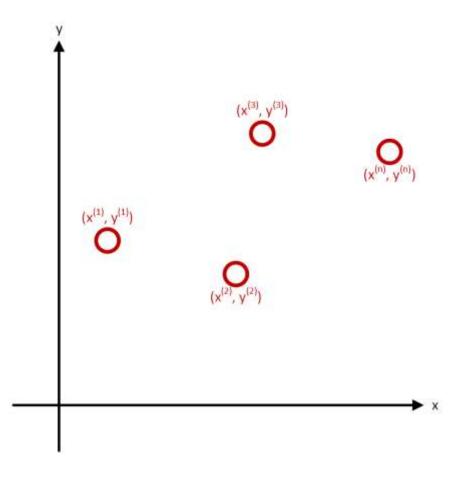
$$y^{(i)} = \beta_0 + \beta_1 \cdot x^{(i)} + \varepsilon^{(i)}$$

• In words, this equation says that for each sample i, $y^{(i)}$ can be explained by $\beta_0 + \beta_1 \cdot x^{(i)}$

- In our Python environment, x and y represent pandas Series and $x^{(i)}$ and $y^{(i)}$ their values at index i
- E.g. (SF housing dataset),
 - x is the property's size (df.Size)
 - y is the property's sale price (df.SalePrice)

Simple Linear Regression (cont.)

- ε is a "white noise" disturbance which we do
 not observe
 - ho ho models how the observations deviate from the exact slope-intercept relation
- We **do not observe** the constants β_0 or β_1 either, so we have to estimate them



Simple Linear Regression (cont.)

• E.g. (SF housing dataset),

$$SalePrice = \hat{\beta}_0 + \hat{\beta}_1 \cdot Size$$

How to interpret *statsmodels* report?

Dep. Variable:	SalePrice	K-squared:	0.236
Model:	OLS	Adj. R-squared:	0.235
Method:	Least Squares	r-statistic:	297.4
Date:		Prob (F-statistic):	2.67e-59
Time:		Log-Likelihood:	-1687.9
No. Observations:	967	AIC:	3380.
Df Residuals:	965	BIC:	3390.
Df Model:	1		
Covariance Type:	nonrobust		

The model's fit

Is the model's fit significant?

The estimated coefficients $\widehat{m{eta}}_0$ (the intercept) and $\widehat{m{eta}}_1$ (the slope; "size")

	coef	std err	•	P> t	[95.0% Conf. In.
Intercept	0.1551).084	1.842	0.066	-0.010 0.320
ize	0.749	0.043	1, 246	0.000	0.664 0.835

Omnibus:	1842.865	Durbin-Watson:	1.704
Prob(Omnibus):	0.000	Jarque-Bera (JB):	3398350.943
Skew:	13.502	Prob(JB):	0.00
Kurtosis:	292.162	Cond. No.	4.40

Are these estimated significant? (i.e., are they meaningful?; do they make sense?)



Simple Linear Regression

Interpreting the regression's coefficients \hat{eta}

Interpreting the regression's coefficients

• What a the unit of ? • [] = [Sale Price] = \$M • How to interpret ?? • [] = [Sale Price] = \$M	· How to interpret ??? · Po = . 155 (\$M) = \$155k	How to interpret \$ 19 (dm/1000 sqft)
---	---	---------------------------------------



Simple Linear Regression

Are the regression's coefficients $\hat{\beta}$ significant?

Are the regression's coefficients $\hat{\beta}$ significant?

The β coefficients follow a normal distribution:

$$\mu_{\beta} \sim N\left(\beta, (X^T X)^{-1} \sigma^2\right)$$

$$(or)$$

$$\mu_{\beta_j} \sim N(\beta_j, v_j \sigma^2)$$

$$(X^T X)^{-1} = \begin{pmatrix} v_0 = v_{0,0} & \cdots & v_{0,j} & \cdots \\ \vdots & \ddots & & & \\ v_{j,0} & v_j = v_{j,j} & & \ddots \end{pmatrix}$$

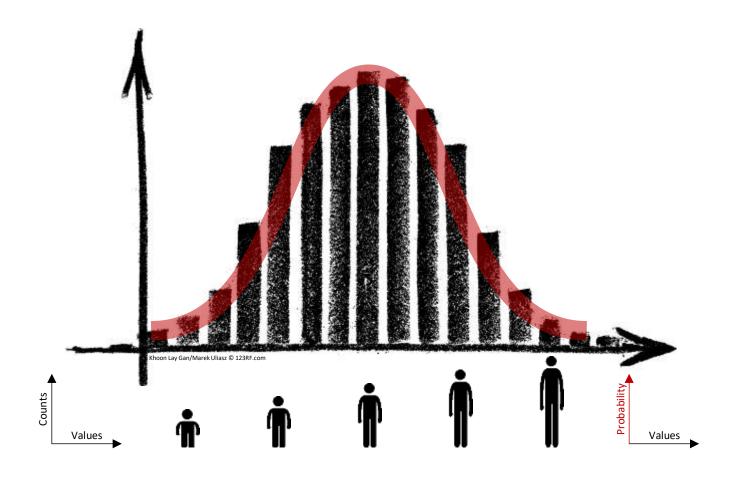
$$\vdots & \ddots & & & \\ \vdots & \ddots & & \\ \vdots & \ddots & & & \\ \vdots & \ddots & \\ \vdots & \vdots & \ddots & \\ \vdots & \ddots & \\ \vdots &$$

 $(v_j \text{ is the } j^{th} \text{ diagonal element of } (X^T X)^{-1})$

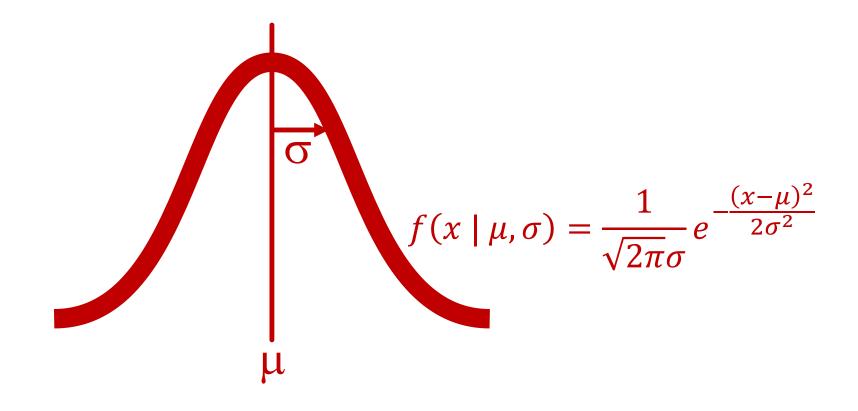


The Normal Distribution

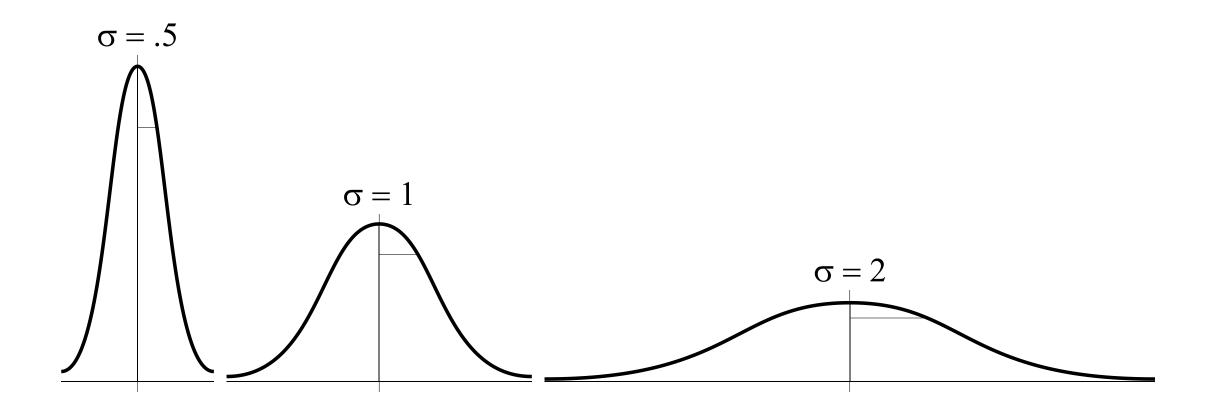
People's height follows a bell shape distribution. (For men in the US, the average height is around 70 inches (5'10) with a standard deviation of 4 inches; few people are shorter than 67 inches; few are as tall as 73 inches)



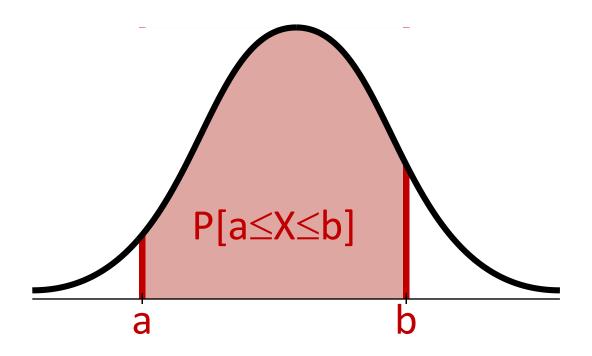
The Normal Distribution



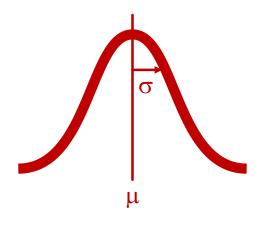
This bell-shaped curve is a probability density function (PDF): The area under the curve is always 1 (for any σ) (cont.)



The area under the curve is called a Cumulative Distribution Function (CDF)

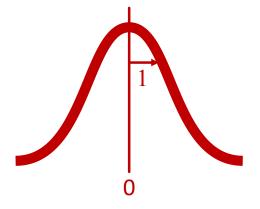


The Standard Normal Distribution ($\mu=0;\sigma=1$)



$$X \sim N(\mu, \sigma)$$

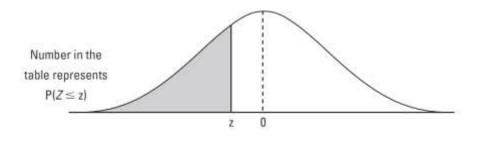
$$X = \mu + \sigma \cdot Z$$



$$Z = \frac{X - \mu}{\sigma}$$

$$Z \sim N(0,1)$$

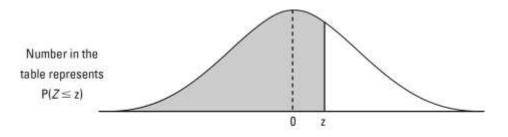
The Standard Normal Distribution Table



z	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	80.0	0.09
-3.6	.0002	.0002	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001
-3.5	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002	.0002
-3.3	.0005	.0005	.0005	.0004	.0004	.0004	.0004	.0004	.0003	.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0005	.0005	.0005
-3.1	.0010	.0009	.0009	.0009	.0008	.0008	.0008	.0008	.0007	.0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9	.0019	.0018	.0018	.0017	.0016	.0016	.0015	.0015	.0014	.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	.0033	.0032	.0031	.0030	.0029	.0028	.0027	.0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	.0060	.0059	.0057	.0055	.0054	.0052	.0051	.0049	.0048
-2.4	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107	.0104	.0102	.0099	.0096	.0094	.0091	.0089	.0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110

-2.1	.0179	.0174	.0170	.0166	.0162	.0158	.0154	.0150	.0146	.0143
-2.0	.0228	.0222	.0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
-1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
-1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
-1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379
-0.9	.1841	.1814	.1788	.1762	.1736	.1711	.1685	.1660	.1635	.1611
-0.8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
-0.7	.2420	.2389	.2358	.2327	.2296	.2266	.2236	.2206	.2177	.2148
-0.6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
-0.5	.3085	.3050	.3015	.2981	.2946	.2912	.2877	.2843	.2810	.2776
-0.4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
-0.3	.3821	.3783	.3745	.3707	.3669	.3632	.3594	.3557	.3520	.3483
-0.2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
-0.1	.4602	.4562	.4522	.4483	.4443	.4404	.4364	.4325	.4286	.4247
-0.0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641

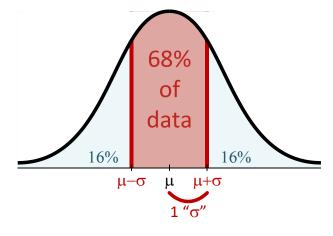
The Standard Normal Distribution Table (cont.)

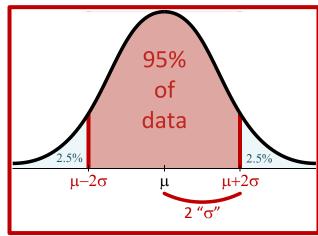


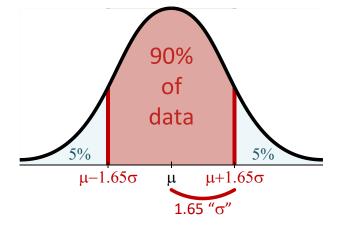
Z	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
0.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
0.1	.5398	.5438	.5478	.5517	.5557	.5596	.5636	.5675	.5714	.5753
0.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
0.3	.6179	.6217	.6255	.6293	.6331	.6368	.6406	.6443	.6480	.6517
0.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
0.5	.6915	.6950	.6985	.7019	.7054	.7088	.7123	.7157	.7190	.7224
0.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	.7549
0.7	.7580	.7611	.7642	.7673	.7704	.7734	.7764	.7794	.7823	.7852
8.0	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
0.9	.8159	.8186	.8212	.8238	.8264	.8289	.8315	.8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	.8621
1.1	.8643	.8665	.8686	.8708	.8729	.8749	.8770	.8790	.8810	.8830
1.2	.8849	.8869	.8888	.8907	.8925	.8944	.8962	.8980	.8997	.9015
1.3	.9032	.9049	.9066	.9082	.9099	.9115	.9131	.9147	.9162	.9177
1.4	.9192	.9207	.9222	.9236	.9251	.9265	.9279	.9292	.9306	.9319

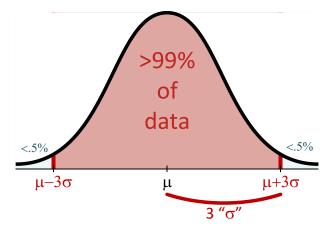
1.5	.9332	.9345	.9357	.9370	.9382	.9394	.9406	.9418	.9429	.9441
1.6	.9452	.9463	.9474	.9484	.9495	.9505	.9515	.9525	.9535	.9545
1.7	.9554	.9564	.9573	.9582	.9591	.9599	.9608	.9616	.9625	.9633
1.8	.9641	.9649	.9656	.9664	.9671	.9678	.9686	.9693	.9699	.9706
1.9	.9713	.9719	.9726	.9732	.9738	.9744	.9750	.9756	.9761	.9767
2.0	.9772	.9778	.9783	.9788	.9793	.9798	.9803	.9808	.9812	.9817
2.1	.9821	.9826	.9830	.9834	.9838	.9842	.9846	.9850	.9854	.9857
2.2	.9861	.9864	.9868	.9871	.9875	.9878	.9881	.9884	.9887	.9890
2.3	.9893	.9896	.9898	.9901	.9904	.9906	.9909	.9911	.9913	.9916
2.4	.9918	.9920	.9922	.9925	.9927	.9929	.9931	.9932	.9934	.9936
2.5	.9938	.9940	.9941	.9943	.9945	.9946	.9948	.9949	.9951	.9952
2.6	.9953	.9955	.9956	.9957	.9959	.9960	.9961	.9962	.9963	.9964
2.7	.9965	.9966	.9967	.9968	.9969	.9970	.9971	.9972	.9973	.9974
2.8	.9974	.9975	.9976	.9977	.9977	.9978	.9979	.9979	.9980	.9981
2.9	.9981	.9982	.9982	.9983	.9984	.9984	.9985	.9985	.9986	.9986
3.0	.9987	.9987	.9987	.9988	.9988	.9989	.9989	.9989	.9990	.9990
3.1	.9990	.9991	.9991	.9991	.9992	.9992	.9992	.9992	.9993	.9993
3.2	.9993	.9993	.9994	.9994	.9994	.9994	.9994	.9995	.9995	.9995
3.3	.9995	.9995	.9995	.9996	.9996	.9996	.9996	.9996	.9996	.9997
3.4	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9998
3.5	.9998	.9998	.9998	.9998	.9998	.9998	.9998	.9998	.9998	.9998
3.6	.9998	.9998	.9999	.9999	.9999	.9999	.9999	.9999	.9999	.9999

The 68 - 90 - 95 - 99.7 Rule (cont.)











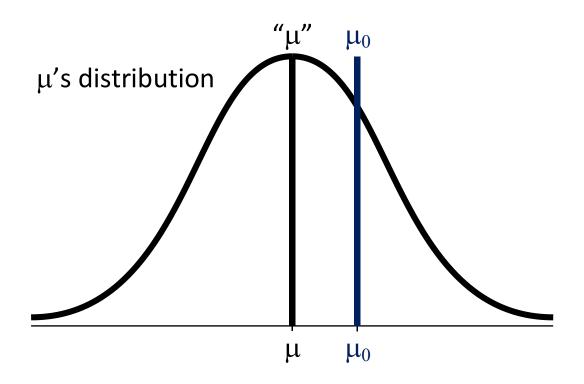
Hypothesis Testing

Hypothesis Testing

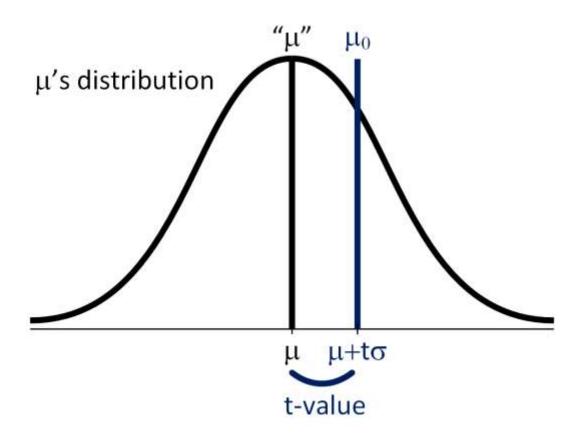
· A hypothesis is an assumption about the a population parameter. E.g.,

- μ_{β_0} = <a specific value, e.g. 0.155>
- $\mu_{\beta_1} = \langle \text{a specific value, e.g. } 0.750 \rangle$
- In both cases, we made a statement about a population parameter that may or may not be true
- The purpose of hypothesis testing is to make a statistical conclusion about
 rejecting or failing to reject such statement

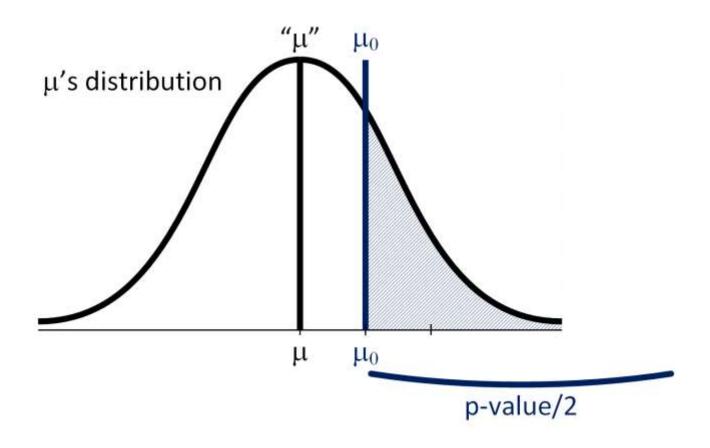
Two-Tail Hypothesis Testing (cont.)



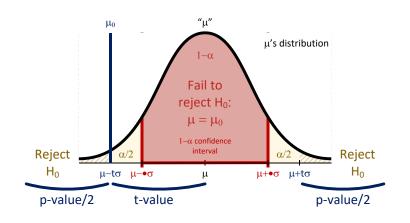
t-value measures the difference to μ_0 in σ . t-values of large magnitudes (either negative or positive) are less likely. The far left and right "tails" of the distribution curve represent instances of obtaining extreme values of t, far from μ

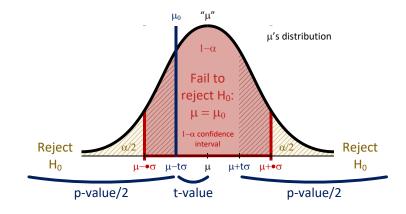


p-value determines the probability (assuming H_0 is true) of observing a more extreme test statistic in the direction of H_a than the one observed



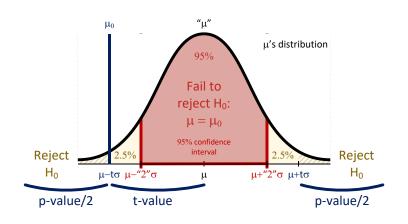
Two-Tail Hypothesis Testing (simplified) (cont.)

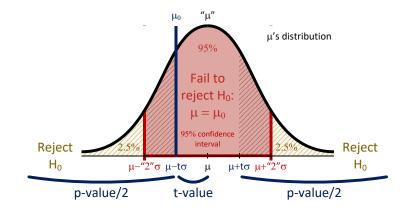




t-value	p-value	$1-\alpha$ Confidence Interval $([\mu-\cdot\sigma,\mu+\cdot\sigma])$	H_0/H_a	Outcome
<.	> α	μ_0 is inside	Did not find evidence that $\mu \neq \mu_0$: Fail to reject H_0	$\mu = \mu_0$
≥·	$\leq \alpha$	μ_0 is outside	Found evidence that $\mu \neq \mu_0$: Reject μ_0 :	$\mu \neq \mu_0$

Two-Tail Hypothesis Testing (simplified) ($\alpha = .05$) (cont.)





t-value	p-value	95% Confidence Interval $([\mu - 2\sigma, \mu + 2\sigma])$	H_0/H_a	Outcome
$<$ " \sim 2"(*) (*) (check t-table)	> .05	μ_0 is inside	Did not find that $\mu \neq \mu_0$: Fail to reject H_0	$\mu = \mu_0$
≥ "~2"	≤ .05	μ_0 is outside	Found evidence that $\mu \neq \mu_0$: Reject μ_0 :	$\mu \neq \mu_0$



Simple Linear Regression

Are the regression's coefficients $\hat{\beta}$ significant? (cont.)

What β_1 would make our multiple linear regression model useless?

• (the simple linear regression model again, without intercept to keep things simple)

$$y = \beta_1 \cdot x + \varepsilon$$

- Answer: If $\beta_1 = 0$, we don't have a linear model
 - (y = o isn't very exciting, is it?)

Is the regression's coefficient $\hat{\beta}$ significant?

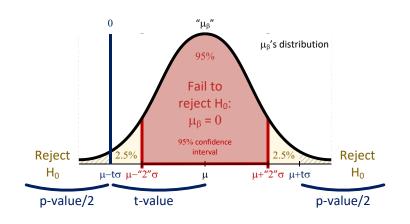
• The *null hypothesis* (H_0) represents the status quo; that the mean of the regression's coefficient β is equal to 0, i.e. that β is not significant:

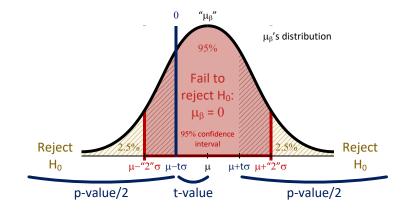
$$H_0: \mu_{\beta} = 0$$

• The *alternate hypothesis* (H_a) represents the opposite of the null hypothesis and holds true if the *null hypothesis* is found to be false; that the mean of the regression's coefficient β is not equal to 0, i.e. that β is significant:

$$H_a$$
: $\mu_\beta \neq 0$

Is the regression's coefficient $\hat{\beta}$ significant? (at the 5% significance level)





t-value	p-value	95% Confidence Interval $([\mu_{\beta}-2\sigma,\mu_{\beta}+2\sigma])$	H_0/H_a	Outcome
< "~2"(*) (*) (check t-table)	> .05	0 is inside	Did not find that $\mu_{\beta} \neq 0$: Fail to reject H_0	$\mu_{eta}=0$; the coefficient eta is not significant
≥ "~2"	≤ .05	0 is outside	Found evidence that $\mu_{\beta} \neq 0$: Reject H_0	$\mu_{\beta} \neq 0$; the coefficient β is significant

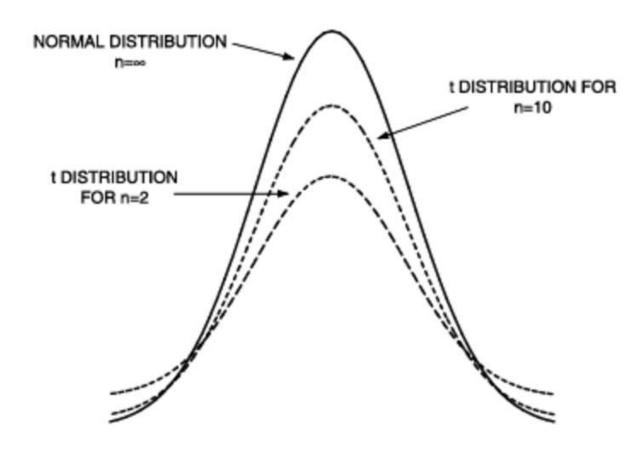
Details of statsmodels' coefficients table

	coef	std err	t	P> t	[95.0% Conf. Int.]
Feature variable j, e.g., "Intercept" or "Size"	$\widehat{oldsymbol{eta}}_{oldsymbol{j}}$	$\sqrt{v_j}\cdot\widehat{\pmb{\sigma}}$	$z_j = rac{\hat{eta}_j}{\sqrt{v_j} \cdot \hat{\sigma}}$ (or) $rac{coef}{std\ err}$	$2 \times$ area under the curve from the Student's tdistribution between $ t $ and $+ \infty$	\hat{eta}_j $\pm z_{lpha=.025} \cdot \hat{\sigma}$ (the value reported in the Student-t distribution table under the 5 th column for $lpha=.025$)



The Student's t-distribution

FYI | We simplified things a bit... t-values refer to the Student's t-distribution, not the normal distribution; the reason behind this is that we substituted $\hat{\sigma}$ for σ (and $\hat{\beta}$ for β)



FYI | We simplified things a bit... t-values refer to the Student's t-distribution, not the normal distribution; the reason behind this is that we substituted $\hat{\sigma}$ for σ (and $\hat{\beta}$ for β) (cont.)

 σ^2 is estimated by $\hat{\sigma}^2$

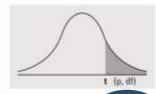
$$\hat{\sigma}^2 = \frac{1}{df} \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

with
$$df = \underbrace{n}_{number\ of\ samples} - \underbrace{k}_{number\ of\ parameters}$$

$$(intercept\ included)$$

Student's t-distribution table: as the sample size grows, the Student's t-distribution converges to a normal distribution

Numbers in each row of the table are values on a t-distribution with (df) degrees of freedom for selected right-tail (greater-than) probabilities (p).



df/p	0.40	0.25	0.10	0.05	0.025	0.01	0.005	0.0005
1	0.324920	1.000000	3.077684	6.313752	12.70620	31.82052	63.65674	636.6192
2	0.288675	0.816497	1.885618	2.919986	4.30265	6.96456	9.92484	31.5991
3	0.276671	0.764892	1.637744	2.353363	3.18245	4.54070	5.84091	12.9240
4	0270722	0.740697	1.533206	2.131847	2.77645	3.74695	4.60409	8.6103
5	0.267181	0.726687	1.475884	2.015048	2.57058	3.36493	4.03214	6.8688
6	0.264835	0.717558	1.439756	1.943180	2.44691	3.14267	3.70743	5.9588
7	0.263167	0.711142	1.414924	1.894579	2.36462	2.99795	3.49948	5.4079
8	0.261921	0.706387	1.396815	1.859548	2.30600	2.89646	3.35539	5.0413
9	0.260955	0.702722	1.383029	1.833113	2.26216	2.82144	3.24984	4.7809
10	0260185	0.699812	1.372184	1.812461	2.22814	2.76377	3.16927	4.5869
11	0259556	0.697445	1.363430	1.795885	2.20099	2.71808	3.10581	4.4370
12	0259033	0.695483	1.356217	1.782288	2.17881	2.68100	3.05454	43178
13	0.258591	0.693829	1.350171	1.770933	2.16037	2.65031	3.01228	4.2208

CI			80%	90%	95%	98%	99%	99.9%
z	0.253347	0.674490	1.281552	1.6446.1	1.95996	2.32635	2.57583	3.2905
30	0.255605	0.682756	1.310415	1.697721	2.04227	45726	2.75000	3.6460
29	0.255684	0.683044	1.311434	1.699127	2.04523	2.46202	2.75639	3.6594
28	0.255768	0.683353	1.312527	1.701131	2.04841	2.46714	2.76326	3.6739
27	0.255858	0.683685	1.313703	1.703288	2.05183	2.47266	2.77068	3.6896
26	0.255955	0.684043	1.314972	1.705618	2.05553	2.47863	2.77871	3.7066
25	0.256060	0.684430	1.316345	1.708141	2.05954	2.48511	2.78744	3.7251
24	0.256173	0.684850	1.317836	1.710882	2.06390	2.49216	2.79694	3.7454
23	0256297	0.685306	1.319460	1.713872	2.06866	2.49987	2.80734	3.7676
22	0256432	0.685805	1.321237	1.717144	2.07387	2.50832	2.81876	3.7921
21	0.256580	0.686352	1.323188	1.720743	2.07961	2.51765	2.83136	3.8193
20	0.256743	0.686954	1.325341	1.724718	2.08596	2.52798	2.84534	3.8495
19	0.256923	0.687621	1.327728	1.729133	2.09302	2.53948	2.86093	3.8834
18	0.257123	0.688364	1.330391	1.734064	2.10092	2,55238	2.87844	3.9216
17	0.257347	0.689195	1.333379	1.739607	2.10982	2.56693	2.89823	3.9651
16	0257599	0.690132	1.336757	1.745884	2.11991	2.58349	2.92078	4.0150
15	0.257885	0.691197	1.340606	1.753050	2.13145	2.60248	2.94671	4.0728
14	0.258213	0.692417	1.345030	1.761310	2.14479	2.62449	2.97684	4.1405



Simple Linear Regression

Are the regression's coefficients $\hat{\beta}$ significant? (cont.)

SalePrice as a function of Size (cont.)

Dep. Variable:	SalePrice	R-squared:	0.236
Model:	OLS	Adj. R-squared:	0.235
Method:	Least Squares	F-statistic:	297.4
Date:		Prob (F-statistic):	2.67e-58
Time:		Log-Likelihood:	-1687.9
No. Observations:	967	AIC:	3380.
Df Residuals:	965	BIC:	3390.
Df Model:	1		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	0.1551	0.084	1.842	0.066	-0.010 0.320
Size	0.7497	0.043	17.246	0.000	0.664 0.835

Omnibus:	1842.865	Durbin-Watson:	1.704
Prob(Omnibus):	0.000	Jarque-Bera (JB):	3398350.943
Skew:	13.502	Prob(JB):	0.00
Kurtosis:	292.162	Cond. No.	4.40

SalePrice
$$[\$M] = \underbrace{.155}_{\widehat{\beta}_0} + \underbrace{.750}_{\widehat{\beta}_1} \times Size [1,000 \ sqft]$$

(the slope is significant but not the intercept)

SalePrice ~ 0 + Size ('0' meaning the intercept is forced to 0) (cont.)

Dep. Variable:	SalePrice	R-squared:	0.565
Model:	OLS	Adj. R-squared:	0.565
Method:	Least Squares	F-statistic:	1255.
Date:		Prob (F-statistic):	7.83e-177
Time:		Log-Likelihood:	-1689.6
No. Observations:	967	AIC:	3381.
Df Residuals:	966	BIC:	3386.
Df Model:	1		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[95.0% Conf. Int.]
Size	0.8176	0.023	35.426	0.000	0.772 0.863

Omnibus:	1830.896	Durbin-Watson:	1.722
Prob(Omnibus):	0.000	Jarque-Bera (JB):	3370566.094
Skew:	13.300	Prob(JB):	0.00
Kurtosis:	291.005	Cond. No.	1.00

SalePrice
$$[\$M] = \underbrace{0}_{\widehat{\beta}_0} + \underbrace{\$10}_{\widehat{\beta}_1} \times Size [1,000 \ sqft]$$

(the slope is significant)

SalePrice ~ Size (with outliers removed) (cont.)

Dep. Variable:	SalePrice	R-squared:	0.200
Model:	OLS	Adj. R-squared:	0.199
Method:	Least Squares	F-statistic:	225.0
Date:		Prob (F-statistic):	1.41e-45
Time:		Log-Likelihood:	-560.34
No. Observations:	903	AIC:	1125.
Df Residuals:	901	BIC:	1134.
Df Model:	1		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	0.7082	0.032	22.152	0.000	0.645 0.771
Size	0.2784	0.019	15.002	0.000	0.242 0.315

Omnibus:	24.647	Durbin-Watson:	1.625
Prob(Omnibus):	0.000	Jarque-Bera (JB):	53.865
Skew:	0.054	Prob(JB):	2.01e-12
Kurtosis:	4.192	Cond. No.	4.70

$$SalePrice [\$M] =$$

$$.708 + .278 \times Size [1,000 \, sqft]$$
(was .155) (was .750)

(both intercept and slope are now significant)



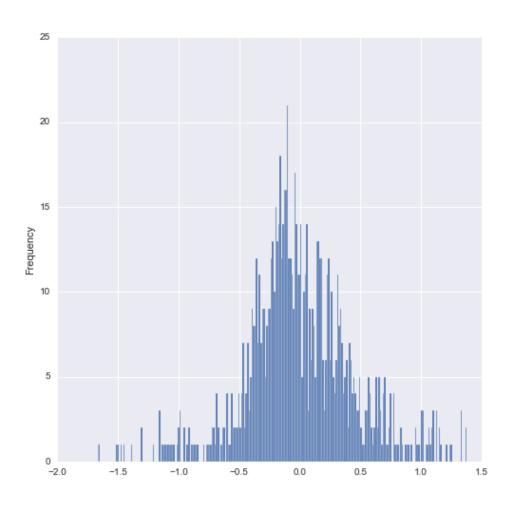
Simple Linear Regression

Common Regression Assumptions

Common Regression Assumptions (Part 1)

- The model is linear
 - x significantly explains y
- $\varepsilon \sim N(0, \cdot)$
 - Specifically, we expect ε to be 0 on average, i.e., $\mu_{\varepsilon} = 0$
- x and ε are independent
 - $\rho(x,\varepsilon)=0$

Is $\varepsilon \sim N(0, \cdot)$?

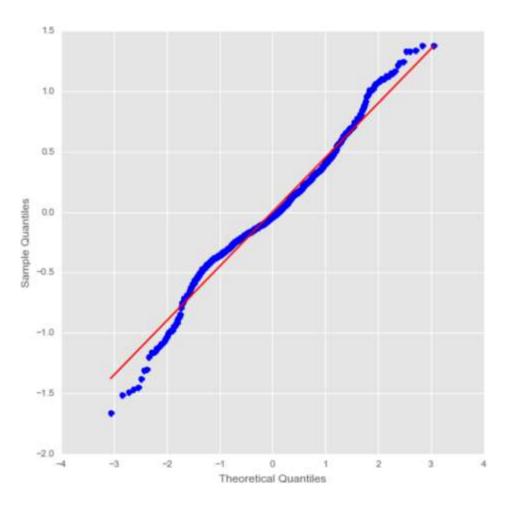


$\varepsilon \sim N(0, \cdot)$: .qqplot()

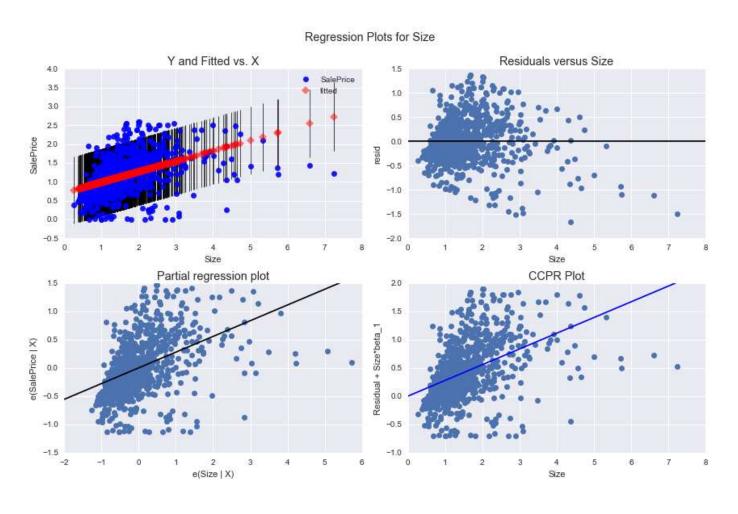
- "Quantile-Quantile (q-q) Plot"
- Graphical technique for determining if two datasets come from populations with a common distribution
- Plot of the quantiles of the first dataset (vertically) against the quantiles of the second's (horizontally)
- If unspecified, the second dataset will default to N(0, 1)

- If the two datasets come from a
 population with the same distribution, the
 points should fall approximately along a
 45-degree reference line
- The greater the departure from this reference line, the greater the evidence for the conclusion that the datasets have come from populations with different distributions

$\varepsilon \sim N(0, \cdot)$: .qqplot() (with line = 's') (cont.)



x and ε are independent: .plot_regress_exog()



x and ε are independent: .plot_regress_exog() (cont.)

- Scatterplot of observed values (y) compared to fitted values (ŷ) with confidence intervals against the regressor (x)
 - .plot_fit()

- "Residual Plot"
 - Scatterplot of the model's residuals $(\hat{\epsilon})$ against the regressor (x)

- "Partial Regression Plot" and "CCPR Plot (Component and Component-Plus-Residual)"
 - (useful for multiple regression)

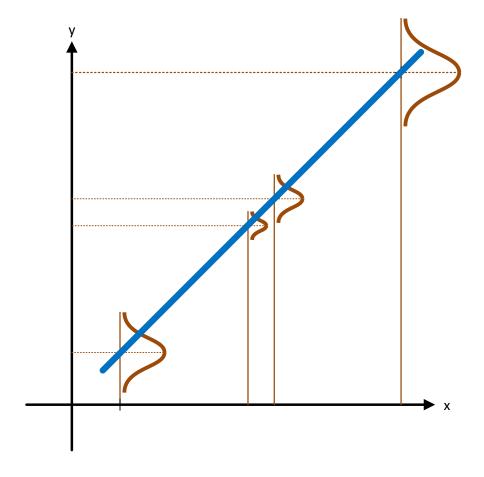


Simple Linear Regression

Assessing the model's fit with R^2

Fit and Inference

- The deviations of the data from the best fitting line are normally distributed about the line. Since $\mu_{\varepsilon}=0$, we "expect" that on average, the line will be correct
- How confident we are about how well the relationship holds depends on σ_{ε}^2



Assessing the model's fit with R^2

When a measure of how much of the total variation in y, $\sigma_y^2 = \beta^2 \sigma_x^2 + \sigma_\varepsilon^2$, is explained by the portion associated with the explanatory variable x, $\sigma_{\hat{y}}^2 = \beta^2 \sigma_x^2$; also called systematic variation (the variation explained by your model)

$$R^{2} = \frac{\sigma_{\hat{y}}^{2}}{\sigma_{y}^{2}} = \frac{\beta^{2} \sigma_{x}^{2}}{\beta^{2} \sigma_{x}^{2} + \sigma_{\varepsilon}^{2}}$$

•
$$0 \le R^2 \le 1$$
 (since $-1 \le \rho_{xy} \le 1$)

•
$$1 - R^2 = \frac{\sigma_{\varepsilon}^2}{\beta^2 \sigma_{\chi}^2 + \sigma_{\varepsilon}^2}$$
 is the idiosyncratic variation (the variation left unexplained by your model)

Assessing the model's fit with \mathbb{R}^2 (cont.)

When x significantly explains y	When x does not significantly explains y
☐ The fit is better	☐ The fit is worse
☐ The explained systematic variation dominates	☐ The unexplained idiosyncratic variation dominates
\square σ_{ε}^2 is low (and/or $\beta^2 \sigma_{\chi}^2$ is high)	\square σ_{ε}^2 is high (and/or $\beta^2 \sigma_{x}^2$ is low)
$\square R^2 = \frac{1}{1 + \underbrace{\frac{\sigma_{\mathcal{E}}^2}{\beta^2 \sigma_{\mathcal{X}}^2}}_{\cong 0}} $ is closer to 1	$\square R^2 = \frac{1}{1 + \underbrace{\frac{\sigma_{\mathcal{E}}^2}{\beta^2 \sigma_{\mathcal{X}}^2}}_{\gg 1}} \text{ is closer to 0}$

Slides © 2017 Ivan Corneillet Where Applicable Do Not Reproduce Without Permission