

Methods 4 - 10

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BSc Programme in Cognitive Science

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Generalized Linear Models

Linear Models: Expected value is additive (“linear”) combination of parameters

$$Y_i \sim \text{Normal}(\mu_i, \sigma)$$
$$\mu_i = \alpha + \beta_X X_i + \beta_Z Z_i$$

Generalized Linear Models:
Expected value is **some function** of an additive combination of parameters

$$Y_i \sim \text{Bernoulli}(p_i)$$
$$f(p_i) = \alpha + \beta_X X_i + \beta_Z Z_i$$

Generalized Linear Models:

Expected value is **some function** of an additive combination of parameters

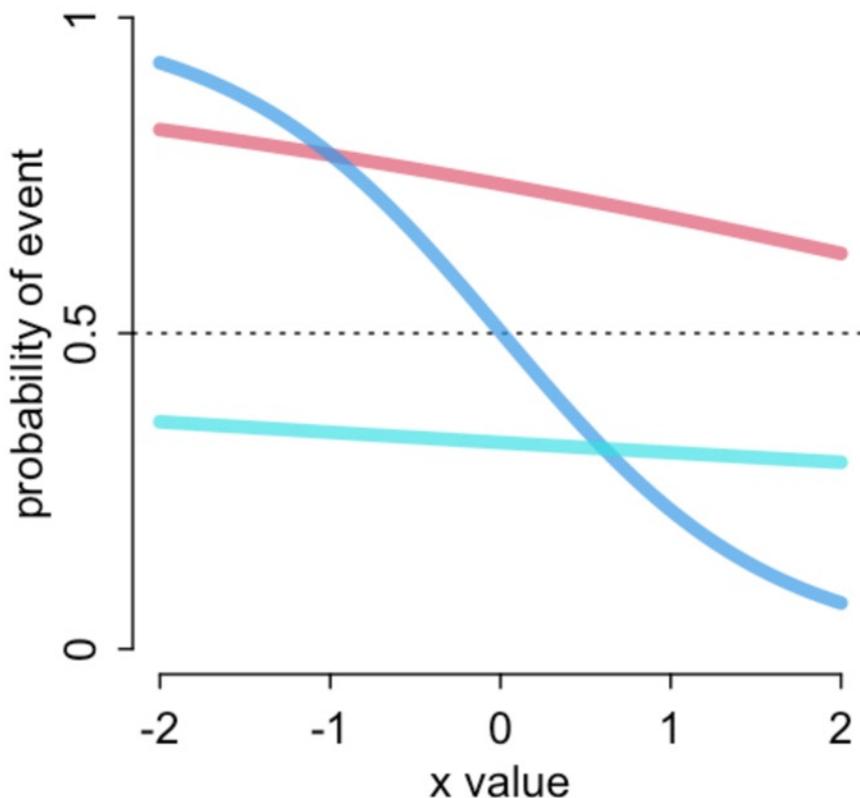
Uniform changes in predictor not uniform changes in prediction

All predictor variables interact, moderate one another

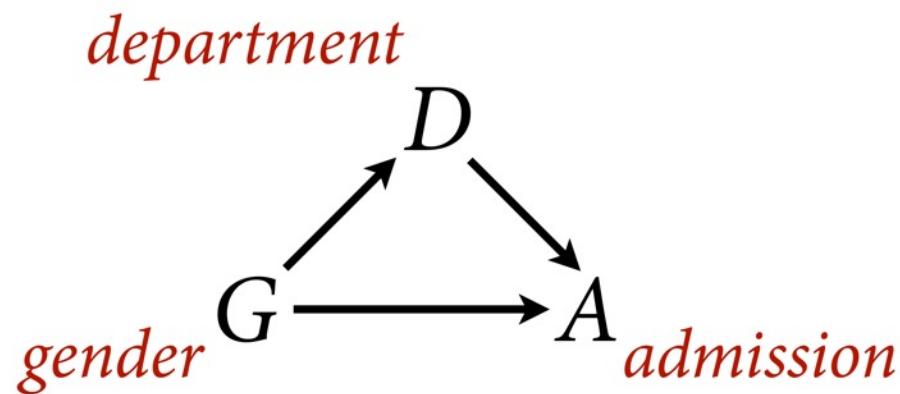
Influences predictions & uncertainty of predictions

$$\text{logit}(p_i) = \alpha + \beta x_i$$

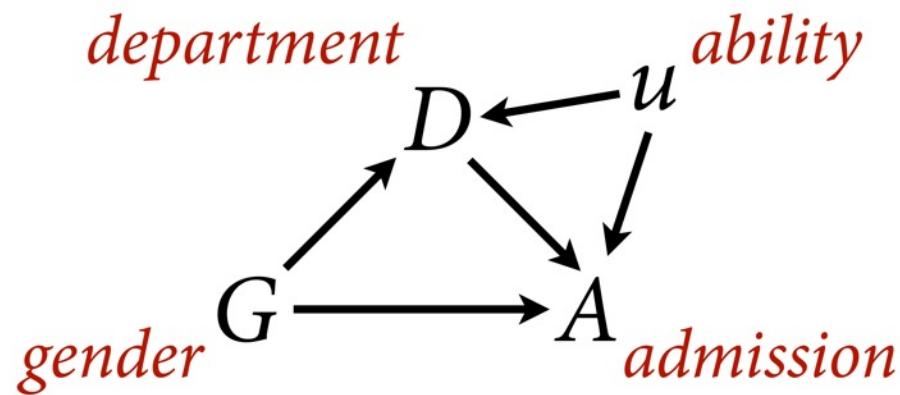
$$n = 0$$



Confounded Admissions



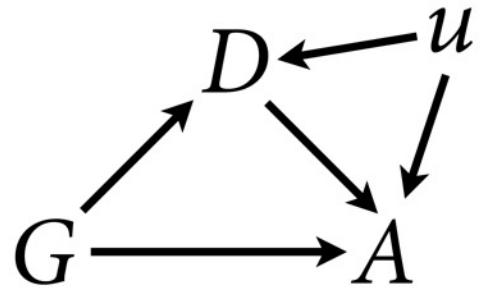
Confounded Admissions



```

set.seed(17)
N <- 2000 # number of applicants
# even gender distribution
G <- sample( 1:2 , size=N , replace=TRUE )
# sample ability, high (1) to average (0)
u <- rbern(N,0.1)
# gender 1 tends to apply to department 1, 2 to 2
# and G=1 with greater ability tend to apply to 2 as well
D <- rbern( N , ifelse( G==1 , u*0.5 , 0.8 ) ) + 1
# matrix of acceptance rates [dept,gender]
accept_rate_u0 <- matrix( c(0.1,0.1,0.1,0.3) , nrow=2 )
accept_rate_u1 <- matrix( c(0.2,0.3,0.2,0.5) , nrow=2 )
# simulate acceptance
p <- sapply( 1:N , function(i)
  ifelse( u[i]==0 , accept_rate_u0[D[i],G[i]] ,
  accept_rate_u1[D[i],G[i]] ) )
A <- rbern( N , p )

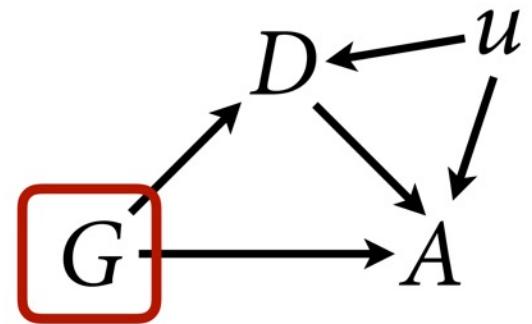
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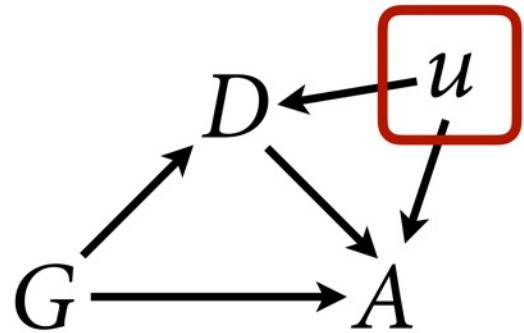
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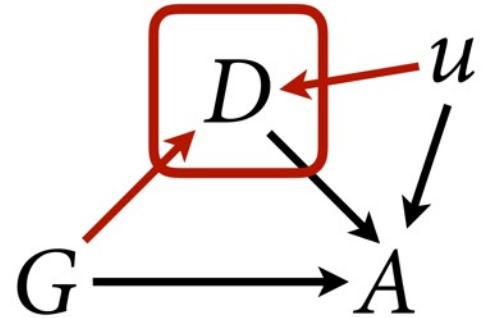
```



```

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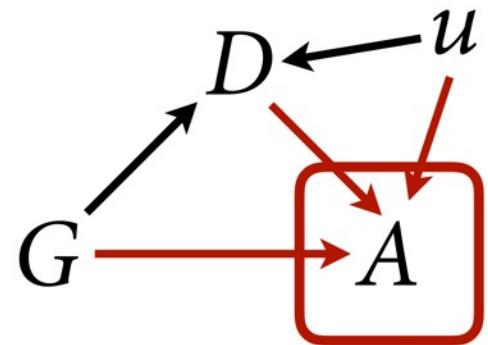
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  ifelse( u[i]==0 , accept_rate_u0[D[i],G[i]] ,
  accept_rate_u1[D[i],G[i]] ) )
A <- rbern( N , p )

```



```

> accept_rate_u0
      [,1] [,2]
[1,]  0.1  0.1
[2,]  0.1  0.3

```

```

> accept_rate_u1
      [,1] [,2]
[1,]  0.2  0.2
[2,]  0.3  0.5

```

```

dat_sim <- list( A=A , D=D , G=G )

# total effect gender
m1 <- ulam(
  alist(
    A ~ bernoulli(p),
    logit(p) <- a[G],
    a[G] ~ normal(0,1)
  ), data=dat_sim , chains=4 , cores=4 )

# direct effects - now confounded!
m2 <- ulam(
  alist(
    A ~ bernoulli(p),
    logit(p) <- a[G,D],
    matrix[G,D]:a ~ normal(0,1)
  ), data=dat_sim , chains=4 , cores=4 )

```

Overvej lige om ngoet af det her
skal i din model i port 3

total effect shows disadvantage

```

> precis(m1,depth=2)
      mean   sd  5.5% 94.5% n_eff Rhat4
a[1] -2.10 0.10 -2.26 -1.93  1297     1
a[2] -0.86 0.07 -0.97 -0.76  1008     1

```

direct effect confounded

```

> precis(m2,depth=3)
      mean   sd  5.5% 94.5% n_eff Rhat4
a[1,1] -2.18 0.11 -2.35 -2.01  2083     1
a[1,2] -0.99 0.30 -1.49 -0.51  2408     1
a[2,1] -1.97 0.21 -2.31 -1.65  2335     1
a[2,2] -0.65 0.07 -0.77 -0.53  2260     1

```

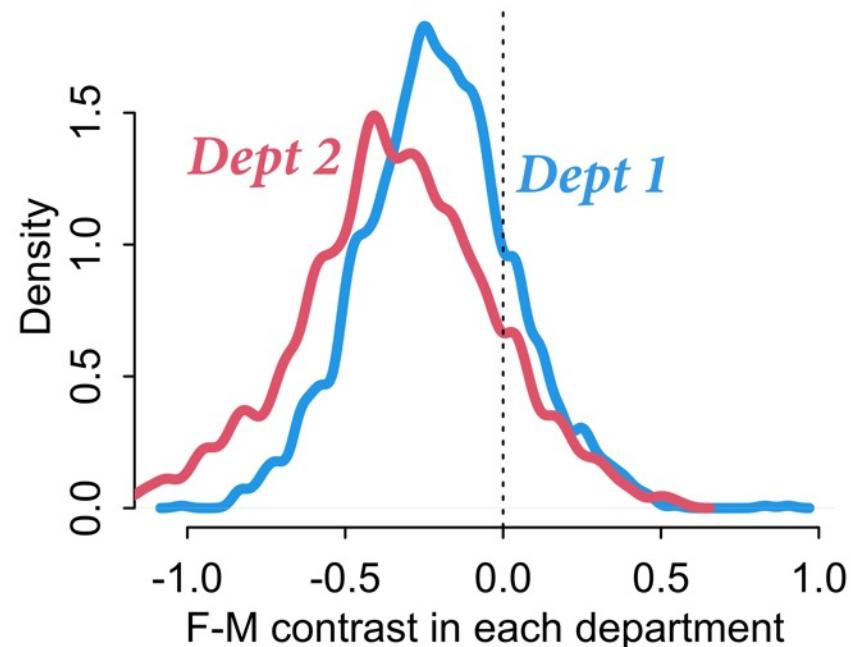
Let's look at the contrasts...

```
post2 <- extract.samples(m2)

post2$fm_contrast_D1 <-
  post2$a[,1,1] - post2$a[,2,1]

post2$fm_contrast_D2 <-
  post2$a[,1,2] - post2$a[,2,2]
```

Tekst



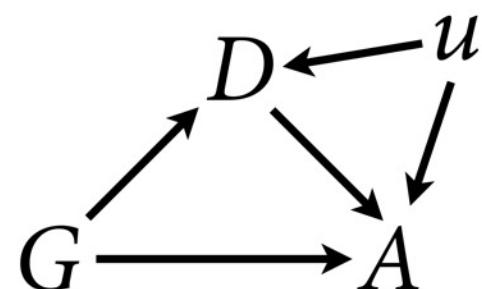
How can the confound hide discrimination?

You guessed it: Collider bias

Stratifying by D opens non-causal path through u

Can estimate **total** causal effect of G ,
but this isn't what we want

Cannot estimate **direct** effect of D or G



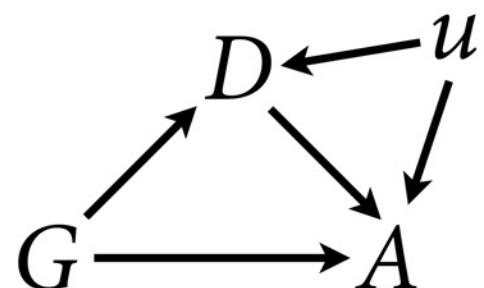
You guessed it: Collider bias

More intuitive explanation:

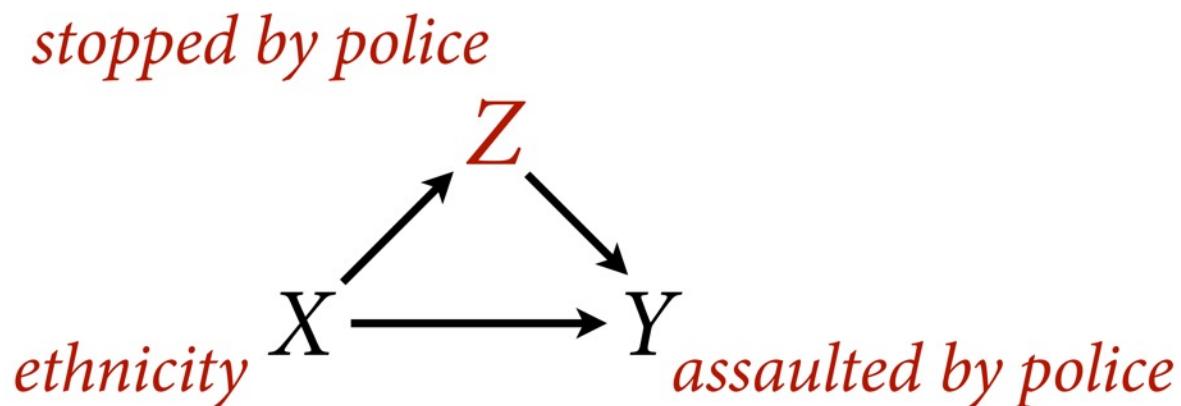
High ability G1s apply to discriminatory department anyway

G1s in that department are higher ability on average than G2s

High ability compensates for discrimination => masks evidence

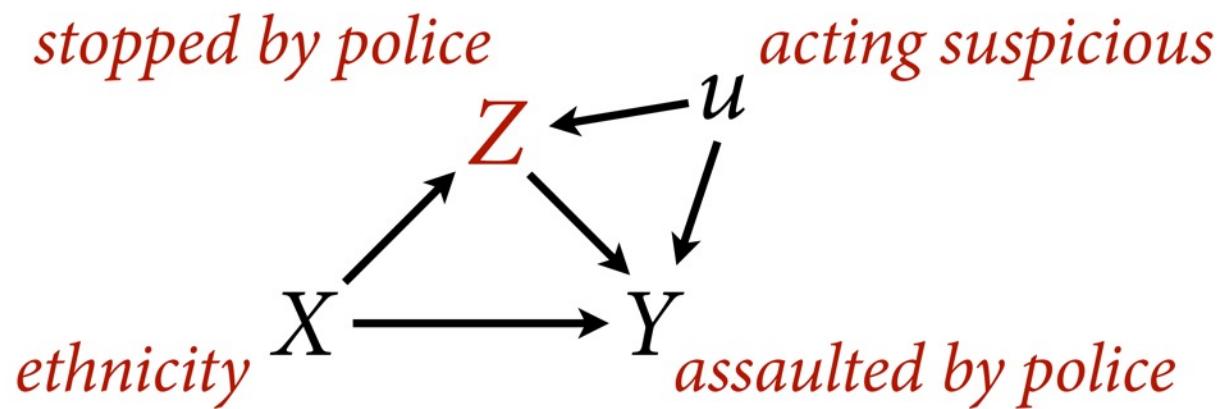


Policing confounds



Knox Lowe & Mummolo 2020 Administrative Records Mask Racially Biased Policing

Policing confounds



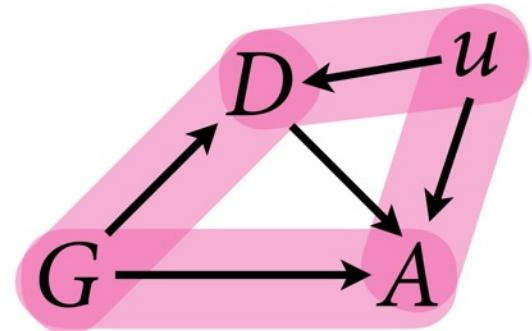
Data on police stops are confounded by lack of data on who wasn't stopped: forced conditioning on Z

Knox Lowe & Mummolo 2020 Administrative Records Mask Racially Biased Policing

Imagine we could measure the confound

```
# if we could measure u
dat_sim$u <- u
m3 <- ulam(
  alist(
    A ~ bernoulli(p),
    logit(p) <- a[G,D] + buA*u,
    matrix[G,D]:a ~ normal(0,1),
    buA ~ normal(0,1)
  ), data=dat_sim ,
constraints=list(buA="lower=0") ,
  chains=4 , cores=4 )

post3 <- extract.samples(m3)
post3$fm_contrast_D1 <-
  post3$a[,1,1] - post3$a[,2,1]
post3$fm_contrast_D2 <-
  post3$a[,1,2] - post3$a[,2,2]
```



Need to block non-causal
path through u

Imagine we could measure the confound

```
# if we could measure u
dat_sim$u <- u
m3 <- ulam(
  alist(
    ^ ~ bernoulli(p),
    logit(p) ~ a[G,D] + buA*u,
    matrix[C,D]:u ~ normal(0,1),
    buA ~ normal(0,1)
  ), data=dat_sim ,
  constraints=list(buA="lower=0") ,
  chains=4 , cores=4 )

post3 <- extract.samples(m3)
post3$fm_contrast_D1 <-
  post3$a[,1,1] - post3$a[,2,1]
post3$fm_contrast_D2 <-
  post3$a[,1,2] - post3$a[,2,2]
```

Add u to linear model

Imagine we could measure the confound

```
# if we could measure u
dat_sim$u <- u
m3 <- ulam(
  alist(
    A ~ bernoulli(p),
    logit(p) <- a[G,D] + buA*u,
    matrix[G,D]:a ~ normal(0,1),
    buA ~ normal(0,1)
  ), data=dat_sim ,
  constraints=list(buA="lower=0") ,
  chains=1 , cores=1 )

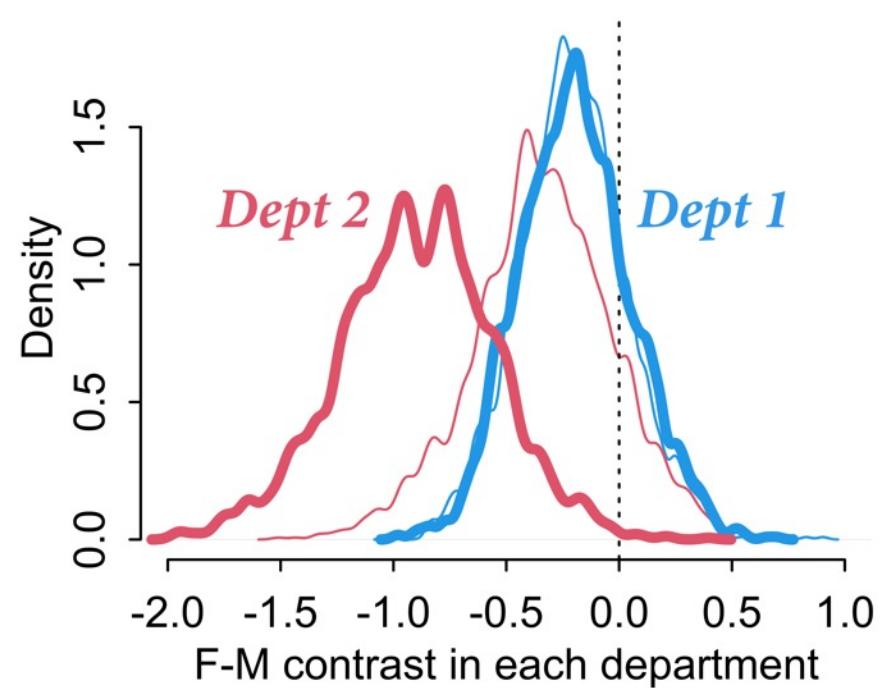
post3 <- extract.samples(m3)
post3$fm_contrast_D1 <-
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  post3$a[,1,2] - post3$a[,2,2]
```

Constrain effect of u to +

Imagine we could measure the confound

```
# if we could measure u
dat_sim$u <- u
m3 <- ulam(
  alist(
    A ~ bernoulli(p),
    logit(p) <- a[G,D] + buA*u,
    matrix[G,D]:a ~ normal(0,1),
    buA ~ normal(0,1)
  ), data=dat_sim ,
constraints=list(buA="lower=0") ,
  chains=4 , cores=4 )

post3 <- extract.samples(m3)
post3$fm_contrast_D1 <-
  post3$a[,1,1] - post3$a[,2,1]
post3$fm_contrast_D2 <-
  post3$a[,1,2] - post3$a[,2,2]
```



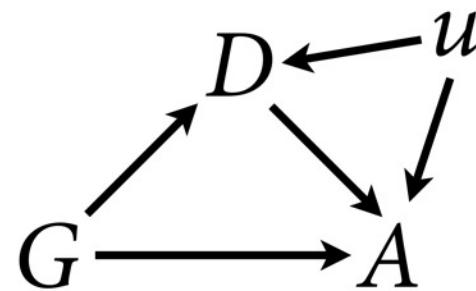
De-confounding

What can be done?

Experiments

Sensitivity analysis

Measure proxies of confound

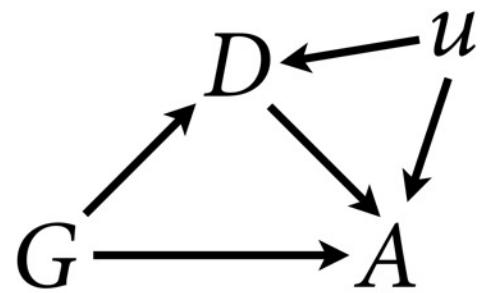


Sensitivity analysis

What are the implications of what we don't know?

Assume confound exists, model its consequences for different strengths/kinds of influence

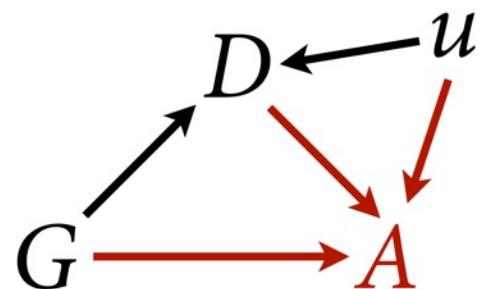
How strong must the confound be to change conclusions?



Sensitivity analysis

$$A_i \sim \text{Bernoulli}(p_i)$$

$$\text{logit}(p_i) = \alpha[G_i, D_i] + \beta_{G[i]} u_i$$



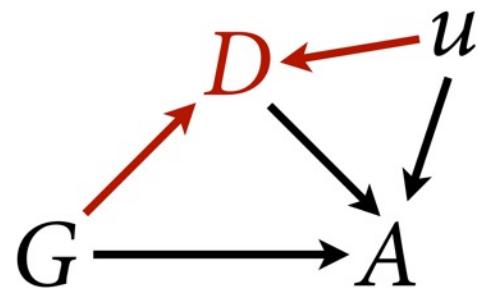
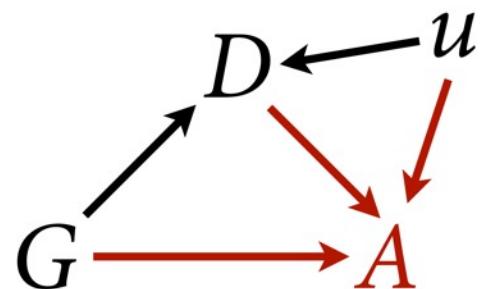
Sensitivity analysis

$$A_i \sim \text{Bernoulli}(p_i)$$

$$\text{logit}(p_i) = \alpha[G_i, D_i] + \beta_{G[i]} u_i$$

$$(D_i = 2) \sim \text{Bernoulli}(q_i)$$

$$\text{logit}(q_i) = \delta[G_i] + \gamma_{G[i]} u_i$$



```

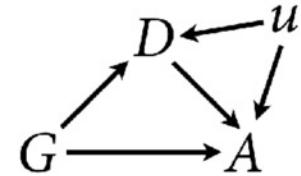
datl$D1 <- ifelse(datl$D==1,1,0)
datl$N <- length(datl$D)
datl$b <- c(1,1)
datl$g <- c(1,0)

mGDU <- ulam(
  alist(
    # A model
    A ~ bernoulli(p),
    logit(p) <- a[G,D] + b[G]*u[i],
    matrix[G,D]:a ~ normal(0,1),

    # D model
    D1 ~ bernoulli(q),
    logit(q) <- delta[G] + g[G]*u[i],
    delta[G] ~ normal(0,1),

    # declare unobserved u
    vector[N]:u ~ normal(0,1)
  ), data=datl , chains=4 , cores=4 )

```



$$A_i \sim \text{Bernoulli}(p_i)$$

$$\text{logit}(p_i) = \alpha[G_i, D_i] + \beta_{G[i]} u_i$$

$$(D_i = 2) \sim \text{Bernoulli}(q_i)$$

$$\text{logit}(q_i) = \delta[G_i] + \gamma_{G[i]} u_i$$

$$u_j \sim \text{Normal}(0,1)$$

```

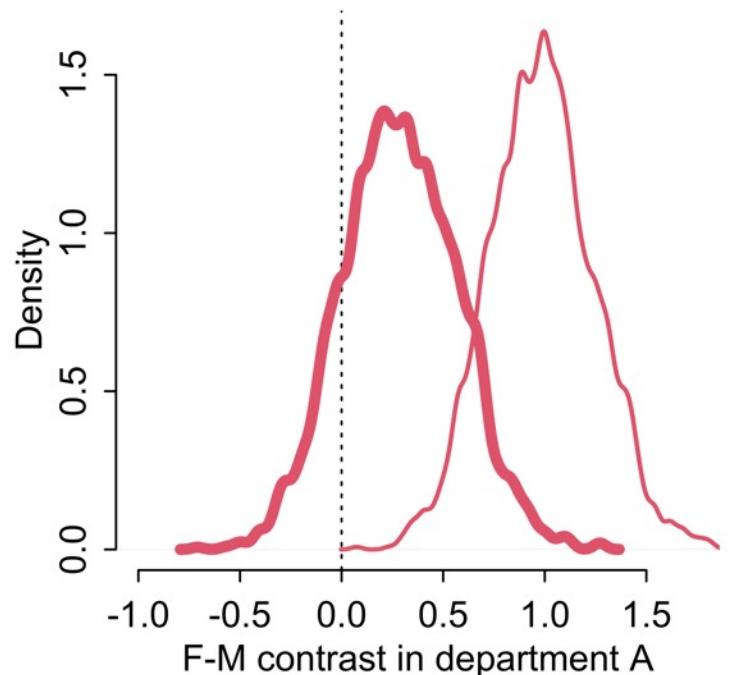
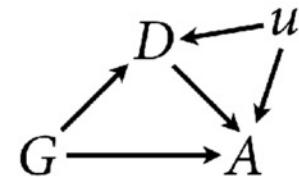
datl$D1 <- ifelse(datl$D==1,1,0)
datl$N <- length(datl$D)
datl$b <- c(1,1)
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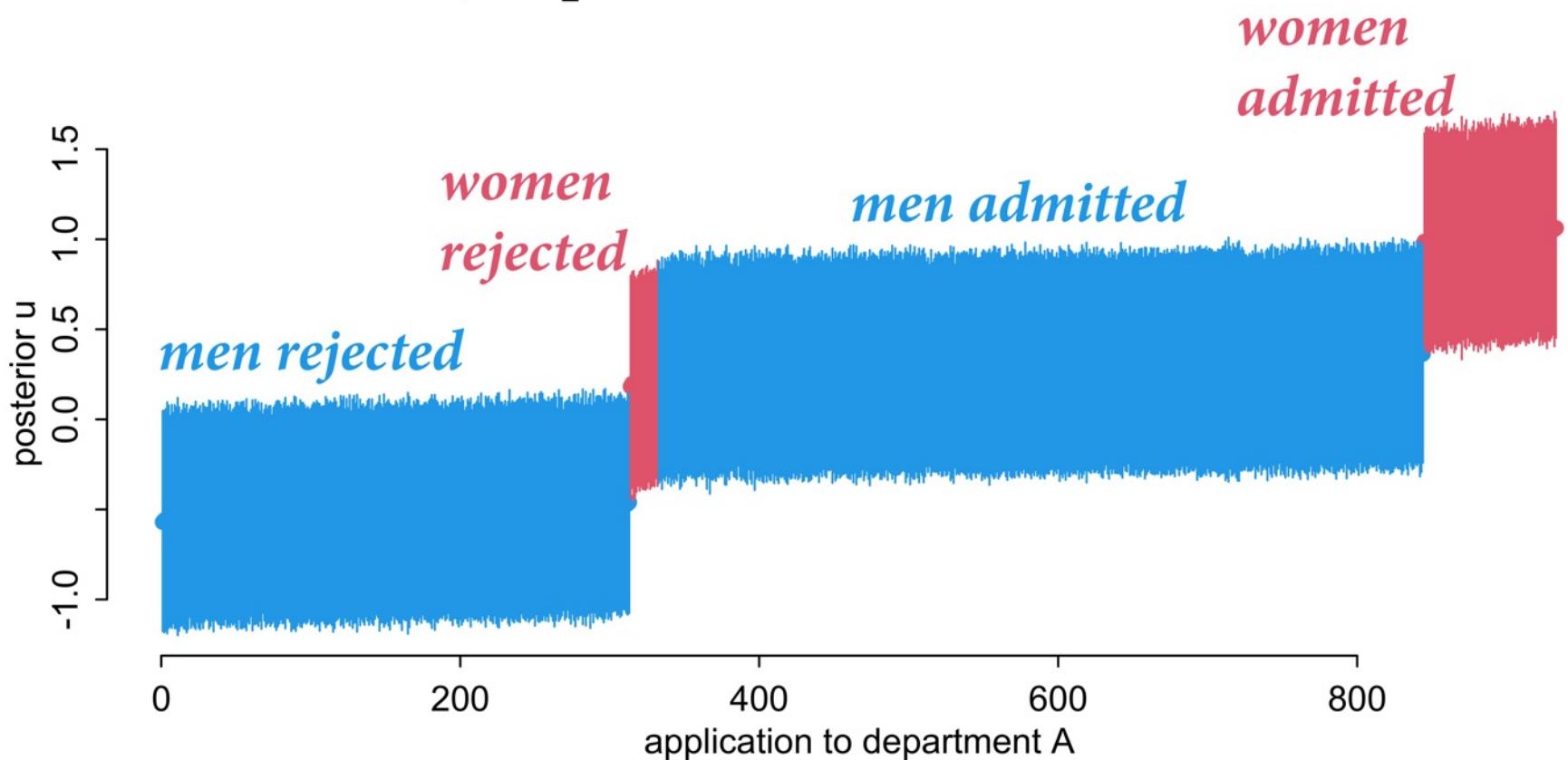
    # D model
    D1 ~ bernoulli(q),
    logit(q) <- delta[G] + g[G]*u[i],
    delta[G] ~ normal(0,1),

    # declare unobserved u
    vector[N]:u ~ normal(0,1)
  ), data=datl , chains=4 , cores=4 )

```



Posterior u values, Department A



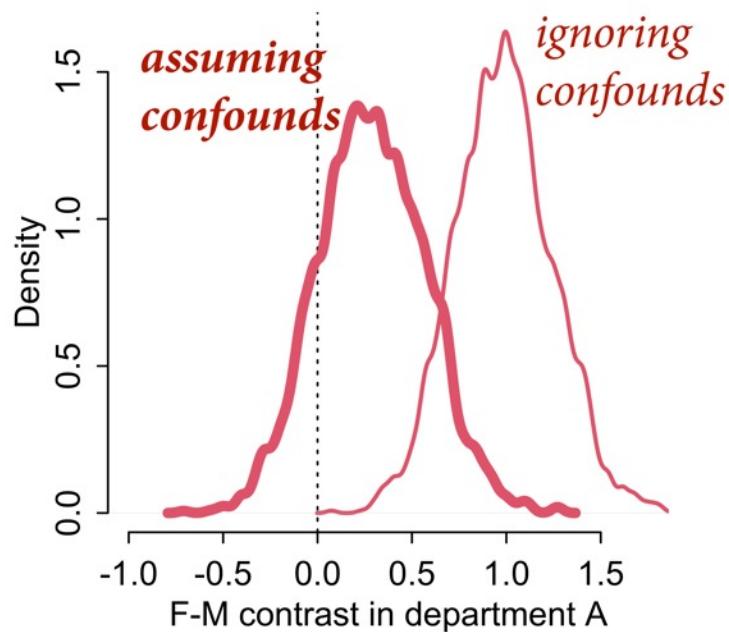
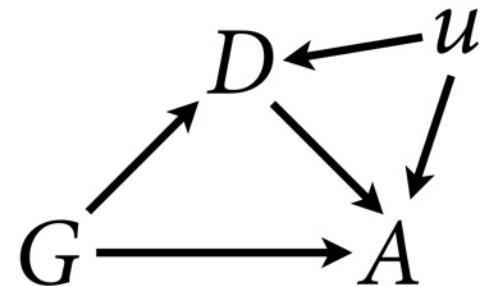
Sensitivity analysis

What are the implications of what we don't know?

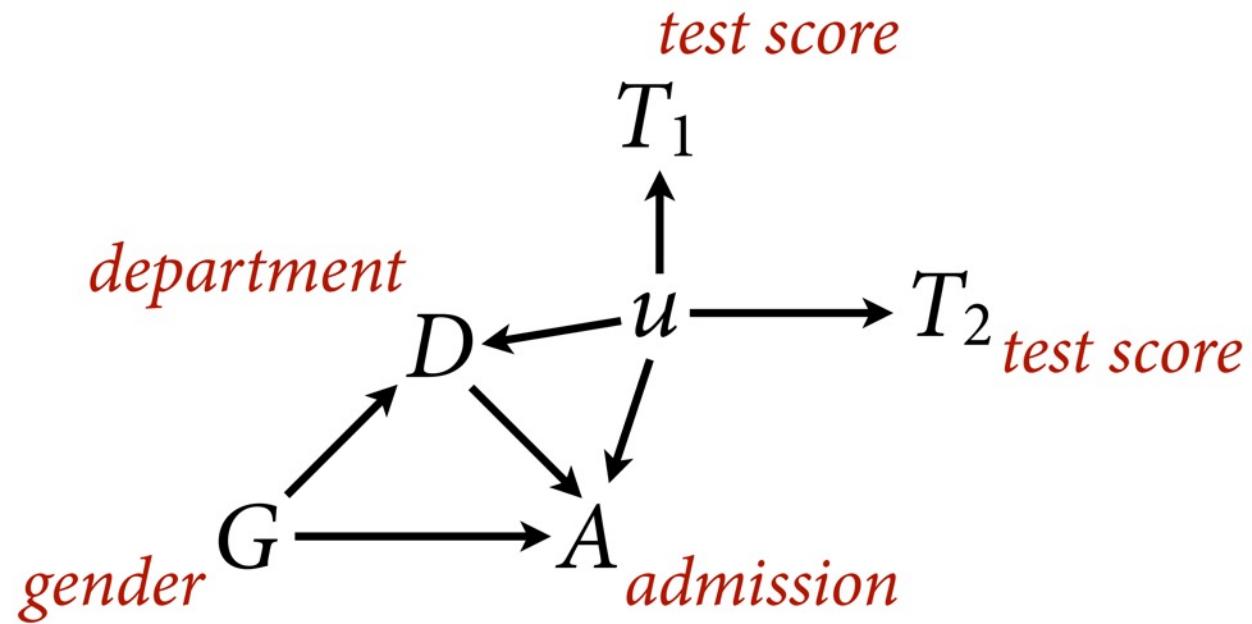
Somewhere between pure simulation and pure analysis

Vary confound strength over range and show how results change –or– vary other effects and estimate confound strength

Confound often persist — don't pretend



De-confounding



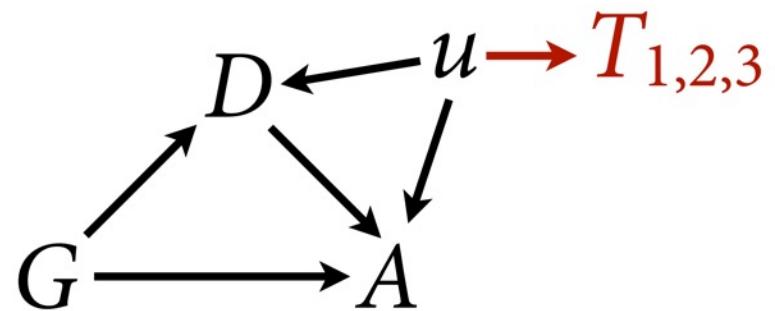
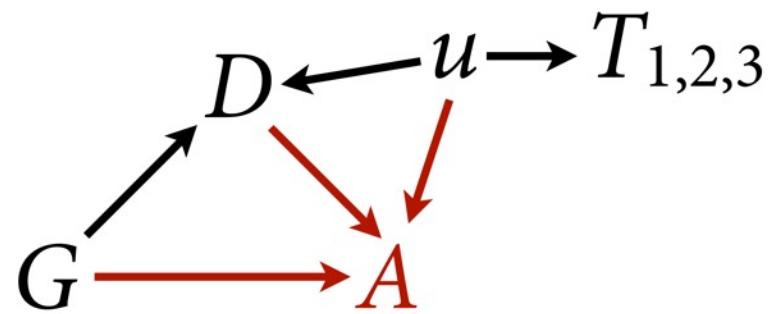
De-confounding

$$A_i \sim \text{Bernoulli}(p_i)$$

$$\text{logit}(p_i) = \alpha[G_i, D_i] + \beta_{G[i]} u_i$$

$$T_{i,j} \sim \text{Normal}(u_i, \tau_j)$$

$$u_k \sim \text{Normal}(0, 1)$$

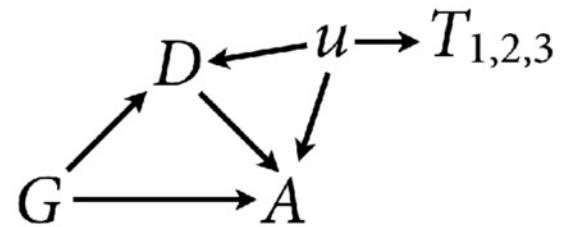


```

T1 <- rnorm(N,u,0.1)
T2 <- rnorm(N,u,0.5)
T3 <- rnorm(N,u,0.25)

m4 <- ulam(
  alist(
    # A model
    A ~ bernoulli(p),
    logit(p) <- a[G,D] + b*u[i],
    matrix[G,D]:a ~ normal(0,1),
    b ~ normal(0,1),
    # u and T model
    vector[N]:u ~ normal(0,1),
    T1 ~ normal(u,tau[1]),
    T2 ~ normal(u,tau[2]),
    T3 ~ normal(u,tau[3]),
    vector[3]:tau ~ exponential(1)
  ),
  data=dat_sim , chains=4 , cores=4 ,
  constraints=list(b="lower=0") )

```



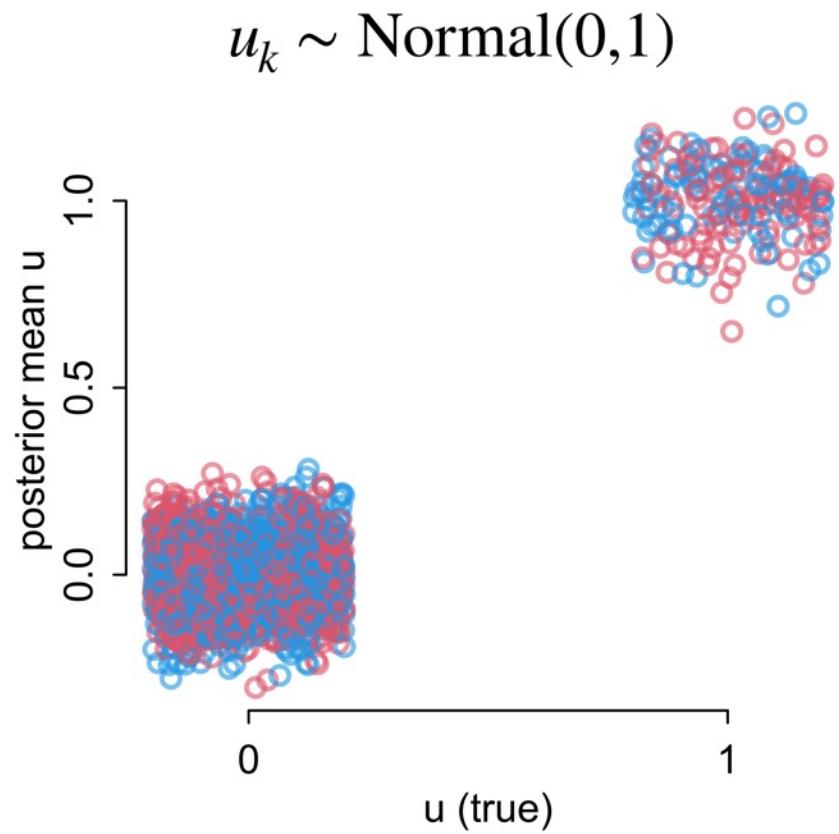
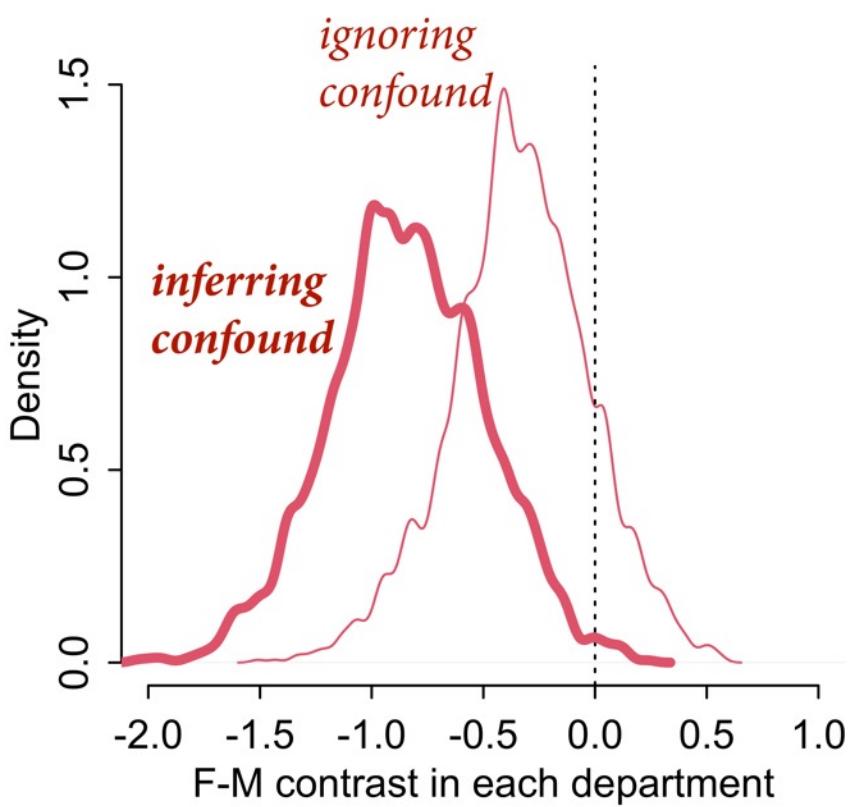
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$$\text{logit}(p_i) = \alpha[G_i, D_i] + \beta_{G[i]} u_i$$

$$u_k \sim \text{Normal}(0,1)$$

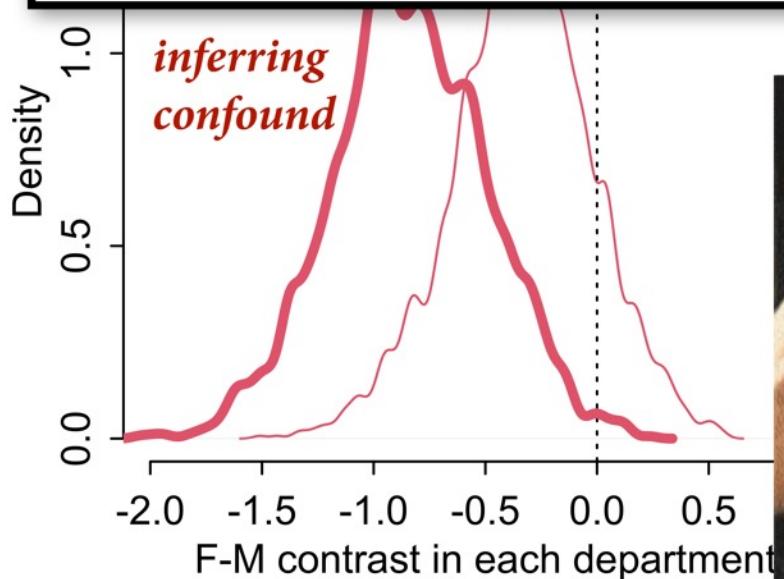
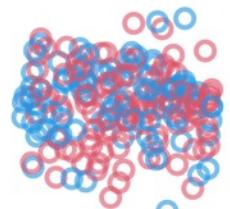
$$T_{i,j} \sim \text{Normal}(u_i, \tau_j)$$

*This model samples inefficiently; learn to fix later in course



*More parameters (2008)
than observations (2000)!*

Normal(0,1)



Oceanic tool complexity

How is technological complexity related to population size?

To social structure?

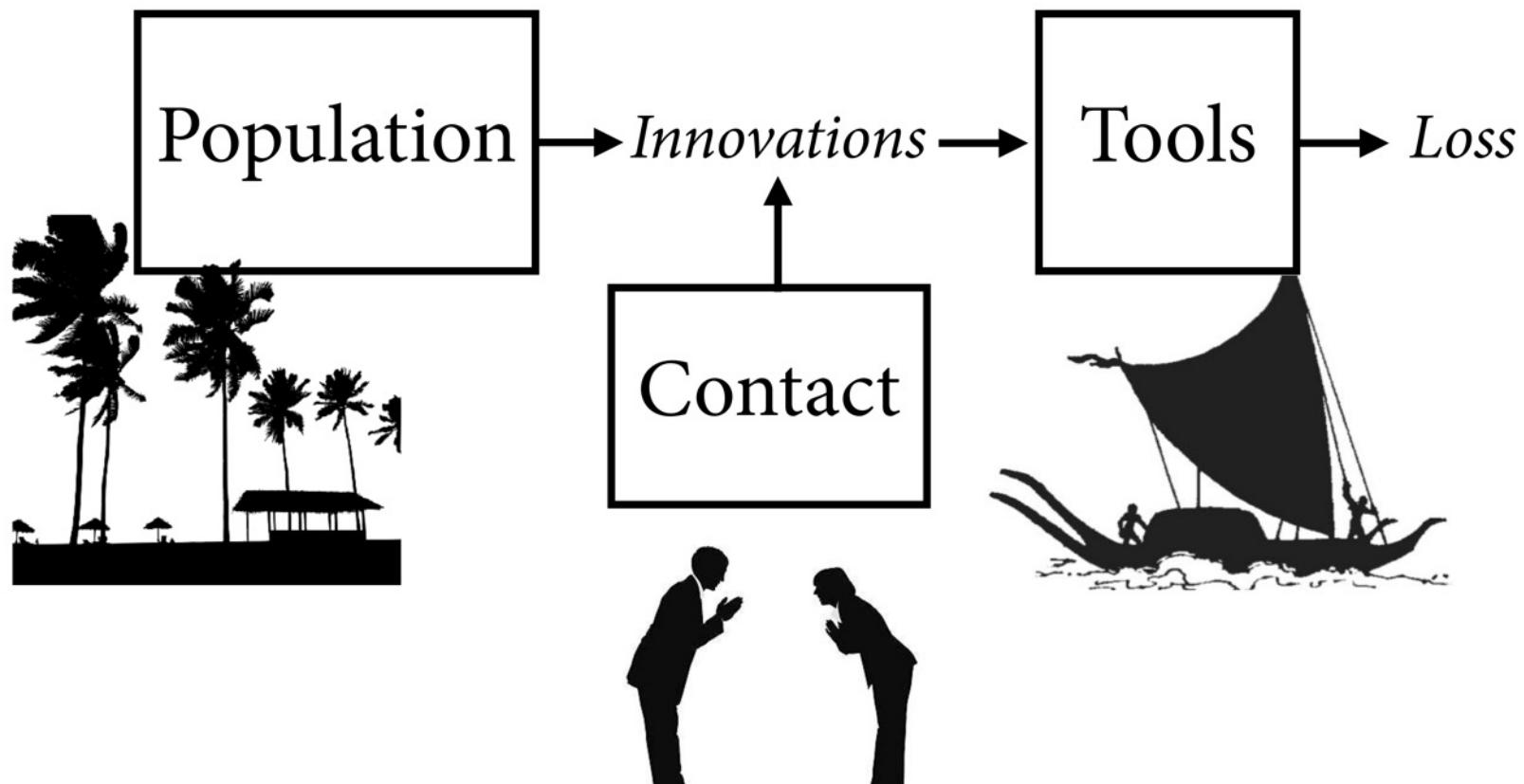
data(Kline)

	culture	population	contact	total_tools	mean_TU
1	Malekula	1100	low	13	3.2
2	Tikopia	1500	low	22	4.7
3	Santa Cruz	3600	low	24	4.0
4	Yap	4791	high	43	5.0
5	Lau Fiji	7400	high	33	5.0
6	Trobriand	8000	high	19	4.0
7	Chuuk	9200	high	40	3.8
8	Manus	13000	low	28	6.6
9	Tonga	17500	high	55	5.4
10	Hawaii	275000	low	71	6.6

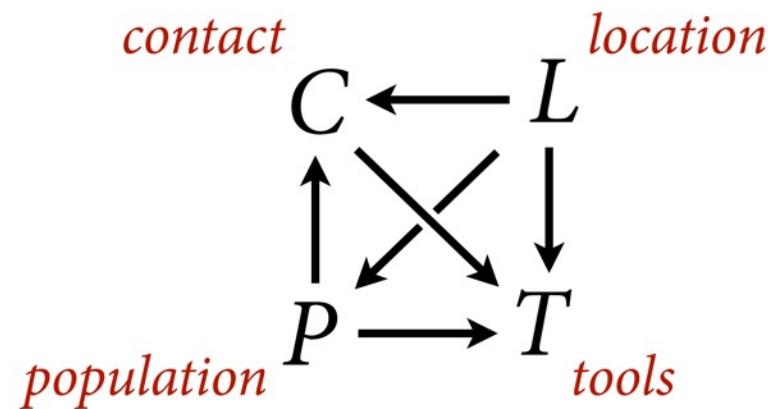
Estimand: Causal influence of **population size** and **contact** on **total tools**

Kline & Boyd 2010 Population size predicts technological complexity in Oceania

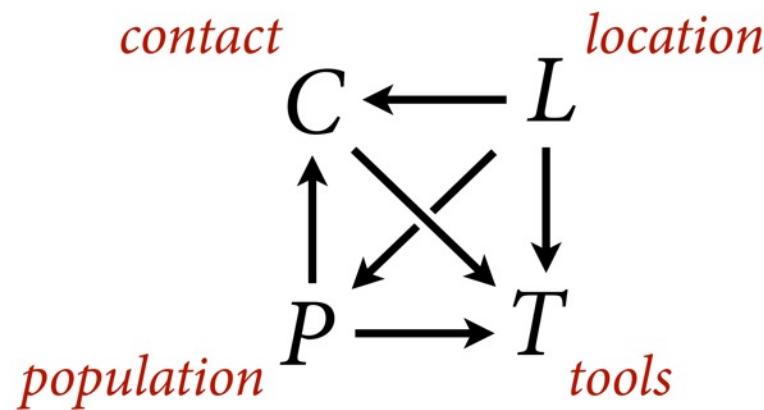
Technological complexity



Technological complexity



Technological complexity



Adjustment set for P : L

Also want to stratify by C ,
to study interaction

Modeling tools

Tool count is not binomial: No maximum

Poisson distribution: Very high maximum and very low probability of each success

Here: Many many possible technologies, very few realized in any one place

Poisson link is log

Poisson distribution takes shape
from expected value

Must be positive

Exponential scaling can be
surprising!

$$Y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha + \beta x_i$$

$$\lambda_i = \exp(\alpha + \beta x_i)$$

Poisson (poison) priors

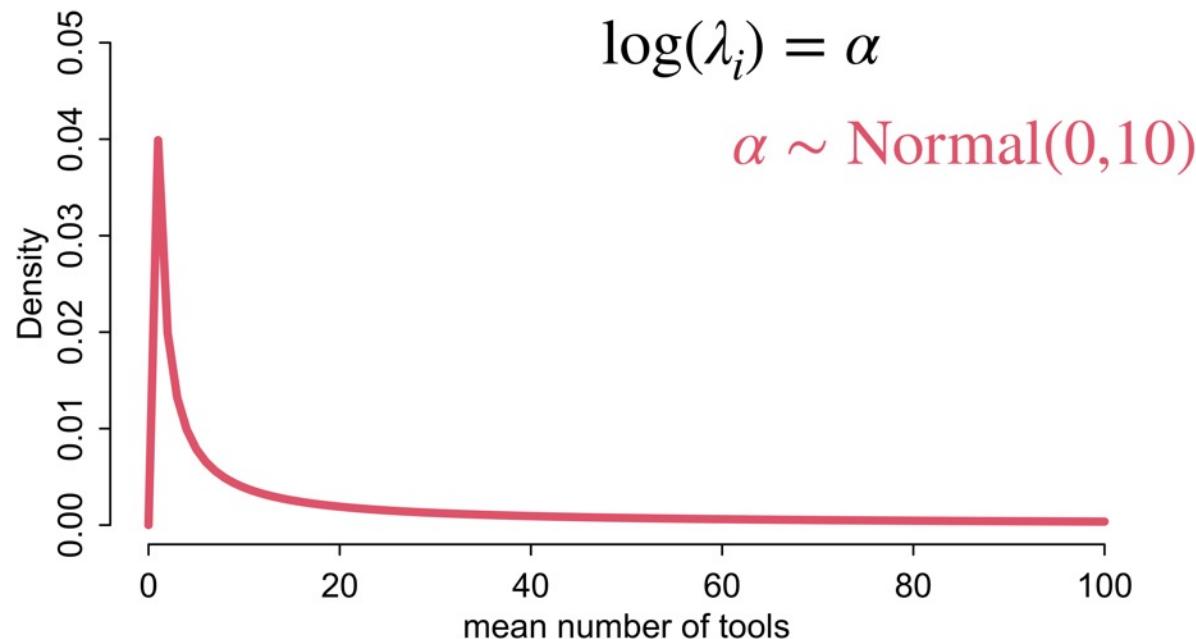
Exponential scaling can be surprising

$$\log(\lambda_i) = \alpha$$

$$\alpha \sim \text{Normal}(0, 10)$$

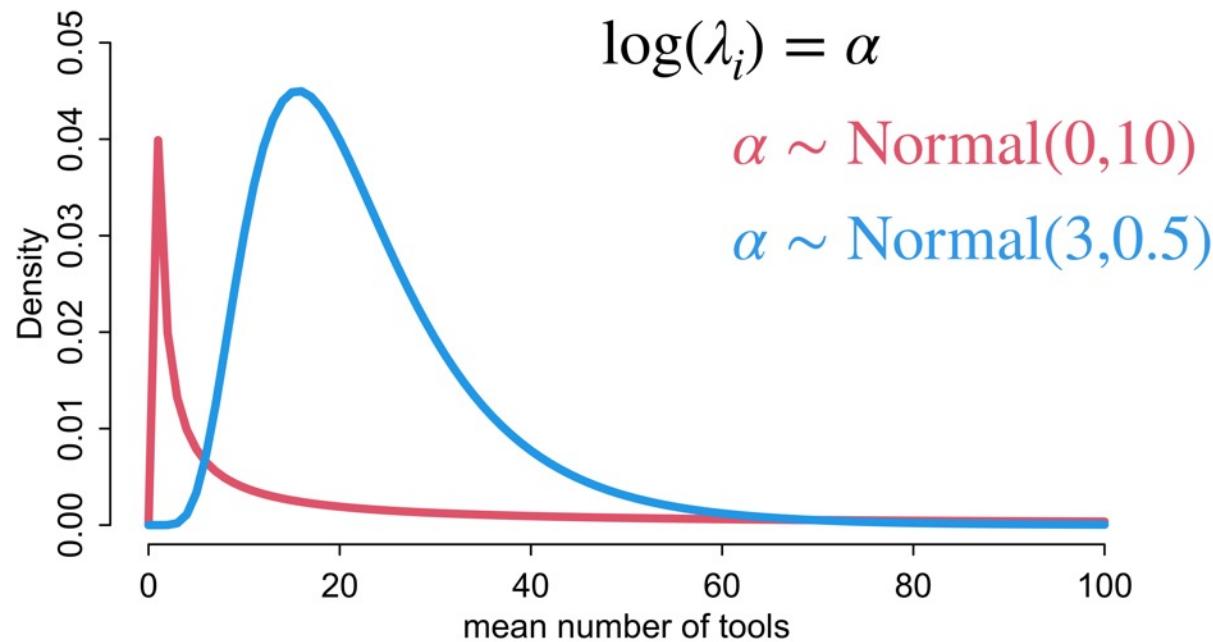
Poisson (poison) priors

Exponential scaling can be surprising



Poisson (poison) priors

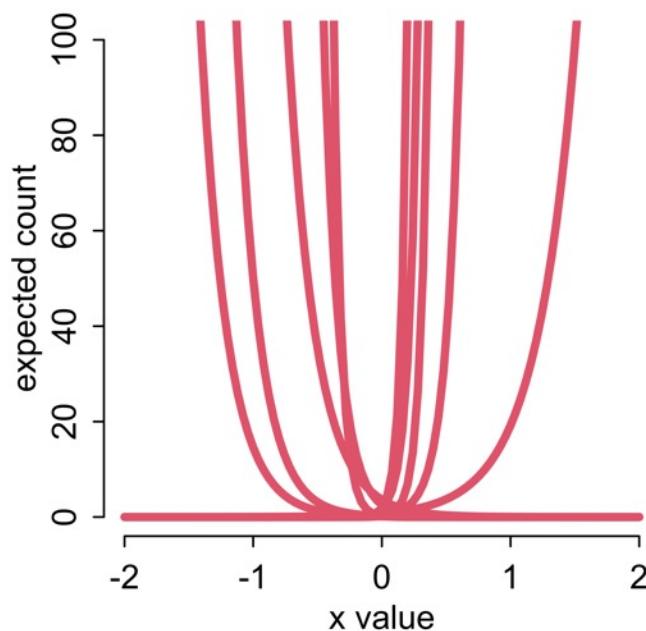
Exponential scaling can be surprising



Poisson priors

$\alpha \sim \text{Normal}(0,1)$

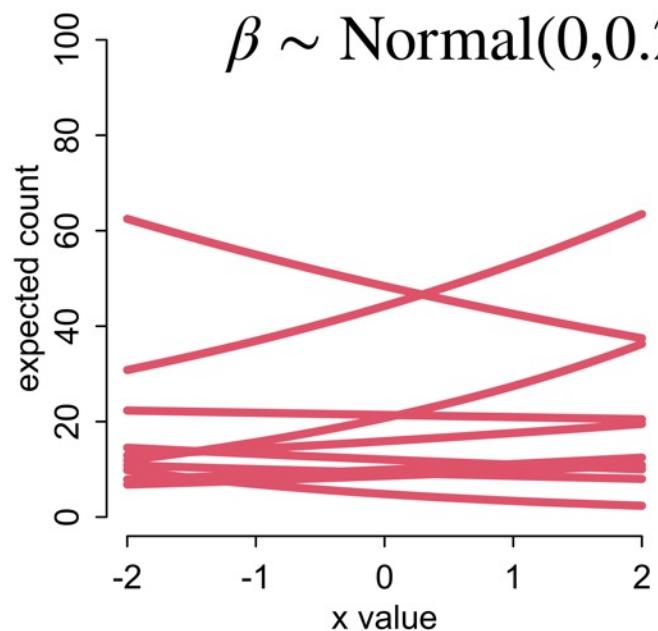
$\beta \sim \text{Normal}(0,10)$



$$\log(\lambda_i) = \alpha + \beta x_i$$

$\alpha \sim \text{Normal}(3,0.5)$

$\beta \sim \text{Normal}(0,0.2)$



```

data(Kline)
d <- Kline
d$P <- scale( log(d$population) )
d$contact_id <- ifelse( d$contact=="high" , 2 , 1 )
dat <- list(
  T = d$total_tools ,
  P = d$P ,
  C = d$contact_id )

# intercept only
m11.9 <- ulam(
  alist(
    T ~ dpois( lambda ),
    log(lambda) <- a,
    a ~ dnorm( 3 , 0.5 )
  ), data=dat , chains=4 , log_lik=TRUE )
# interaction model
m11.10 <- ulam(
  alist(
    T ~ dpois( lambda ),
    log(lambda) <- a[C] + b[C]*P,
    a[C] ~ dnorm( 3 , 0.5 ),
    b[C] ~ dnorm( 0 , 0.2 )
  ), data=dat , chains=4 , log_lik=TRUE )

compare( m11.9 , m11.10 , func=PSIS )

```

$$Y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha_{C[i]} + \beta_{C[i]} \log(P_i)$$

$$\alpha_j \sim \text{Normal}(3,0.5)$$

$$\beta_j \sim \text{Normal}(0,0.2)$$

```

data(Kline)
d <- Kline
d$P <- scale( log(d$population) )
d$contact_id <- ifelse( d$contact=="high" , 2 , 1 )
dat <- list(
  T = d$total_tools ,
  P = d$P ,
  C = d$contact_id )

# intercept only
m11.9 <- ulam(
  alist(
    T ~ dpois( lambda ),
    log(lambda) <- a,
    a ~ dnorm( 3 , 0.5 )
  ), data=dat , chains=4 , log_lik=TRUE )
# interaction model
m11.10 > compare( m11.9 , m11.10 , func=PSIS )
Some Pareto k values are high (>0.5). Set pointwise=TRUE to inspect individual points.
Some Pareto k values are high (>0.5). Set pointwise=TRUE to inspect individual points.
      PSIS      SE dPSIS   dSE pPSIS weight
m11.10  85.9 13.50    0.0    NA    7.3      1
m11.9   141.3 33.69   55.4 33.13    8.0      0
), update=update , chains=4 , log_lik=TRUE )

compare( m11.9 , m11.10 , func=PSIS )

```

$$Y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha_{C[i]} + \beta_{C[i]} \log(P_i)$$

$$\alpha_j \sim \text{Normal}(3,0.5)$$

$$\beta_j \sim \text{Normal}(0,0.2)$$

```

k <- PSIS( m11.10 , pointwise=TRUE )$k
plot( dat$P , dat$T , xlab="log population (std)" ,
ylab="total tools" ,
col=ifelse( dat$C==1 , 4 , 2 ) , lwd=4+4*normalize(k) ,
ylim=c(0,75) , cex=1+normalize(k) )

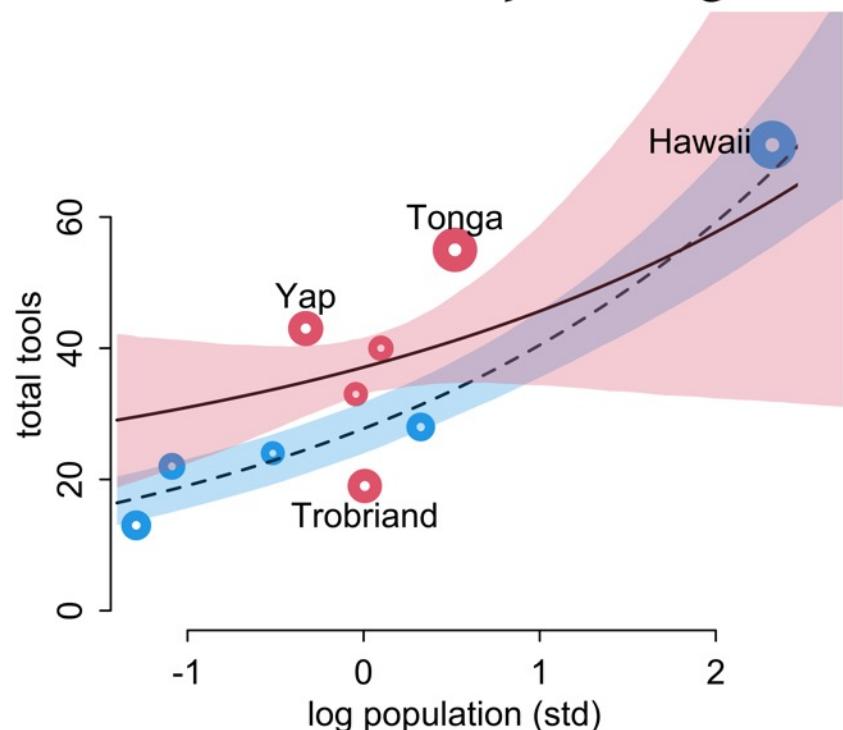
# set up the horizontal axis values to compute predictions
at
P_seq <- seq( from=-1.4 , to=3 , len=100 )

# predictions for C=1 (low contact)
lambda <- link( m11.10 , data=data.frame( P=P_seq , C=1 ) )
lmu <- apply( lambda , 2 , mean )
lci <- apply( lambda , 2 , PI )
lines( P_seq , lmu , lty=2 , lwd=1.5 )
shade( lci , P_seq , xpd=TRUE , col=col.alpha(4,0.3) )

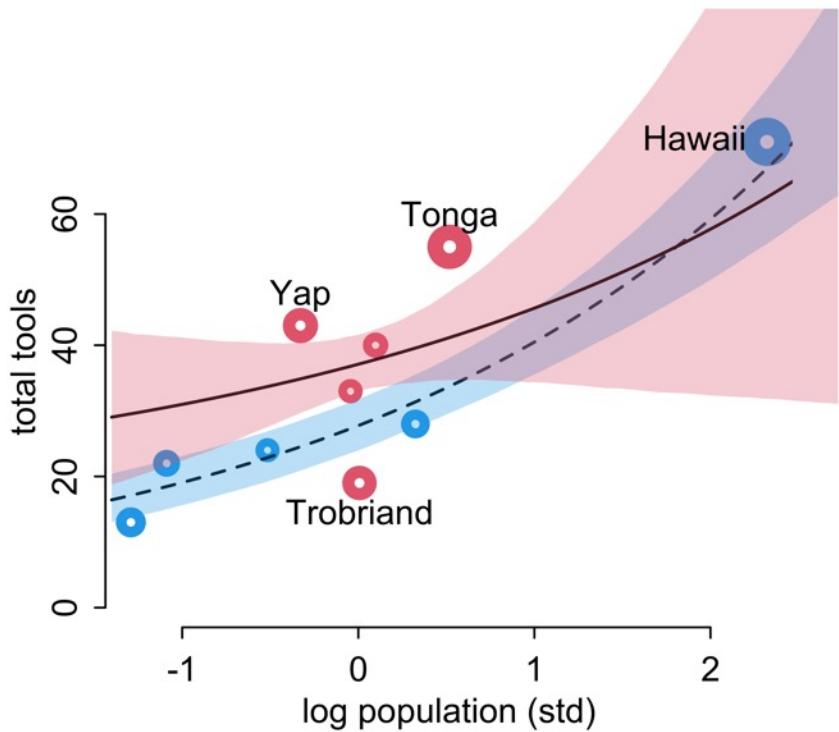
# predictions for C=2 (high contact)
lambda <- link( m11.10 , data=data.frame( P=P_seq , C=2 ) )
lmu <- apply( lambda , 2 , mean )
lci <- apply( lambda , 2 , PI )
lines( P_seq , lmu , lty=1 , lwd=1.5 )
shade( lci , P_seq , xpd=TRUE , col=col.alpha(2,0.3) )

```

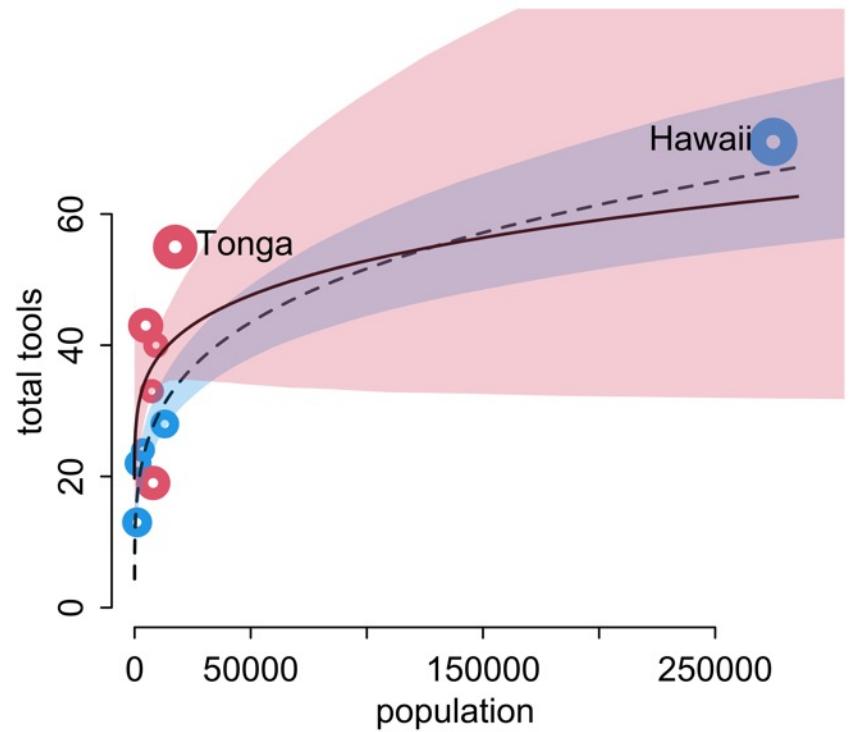
Points scaled by leverage

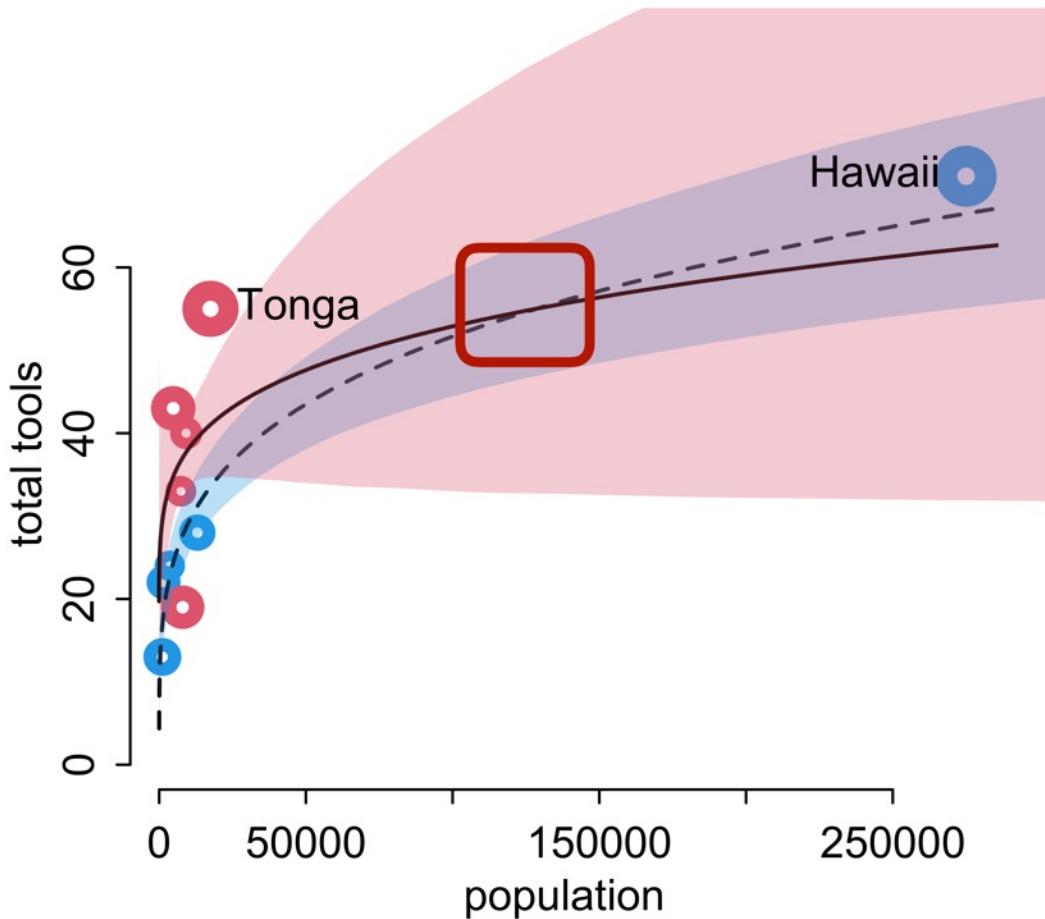


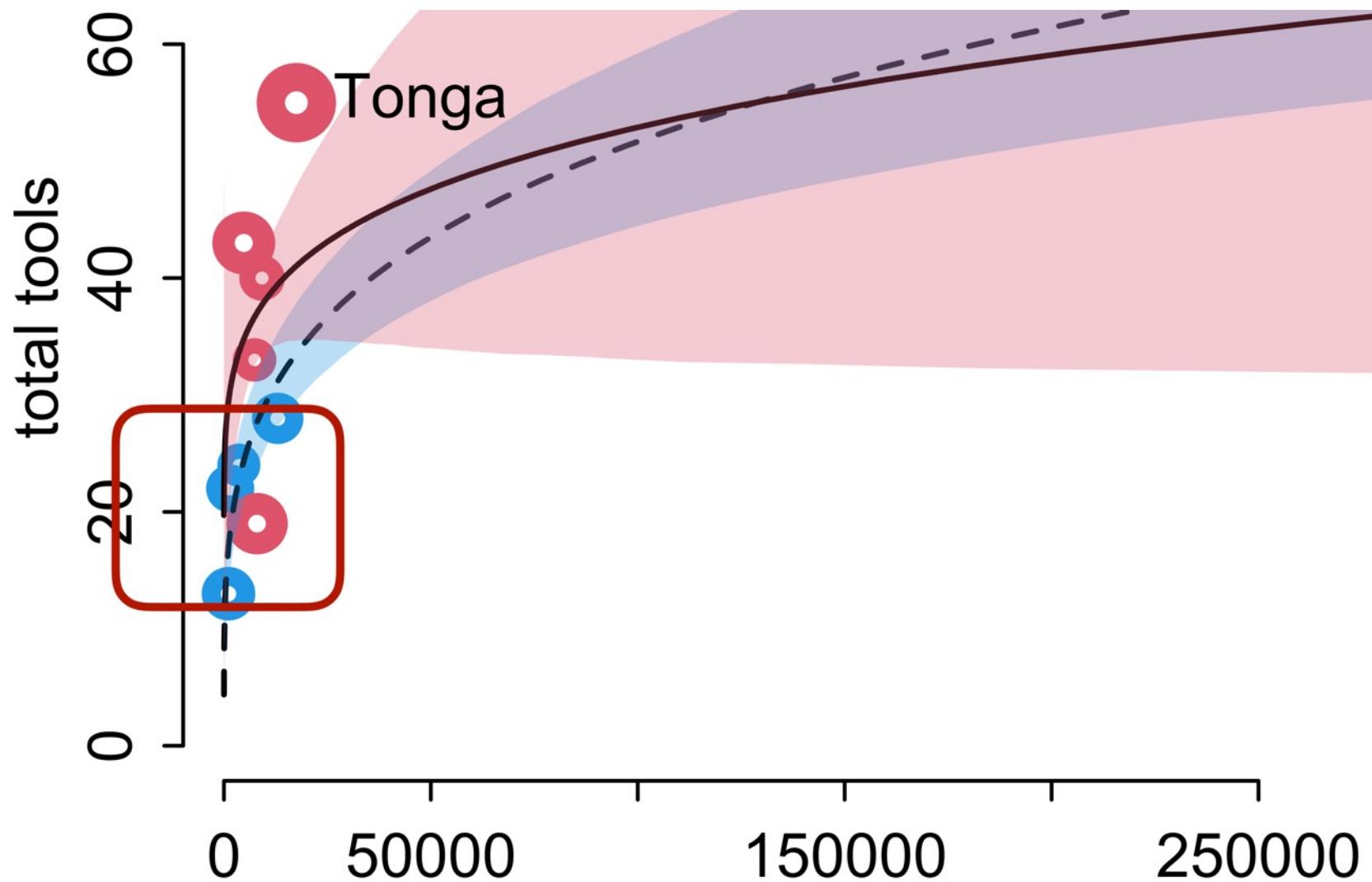
log scale



Natural scale



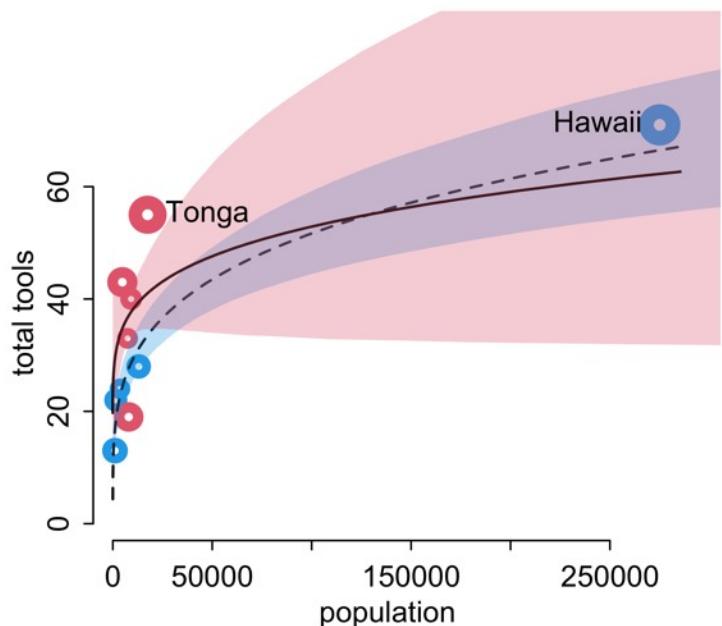




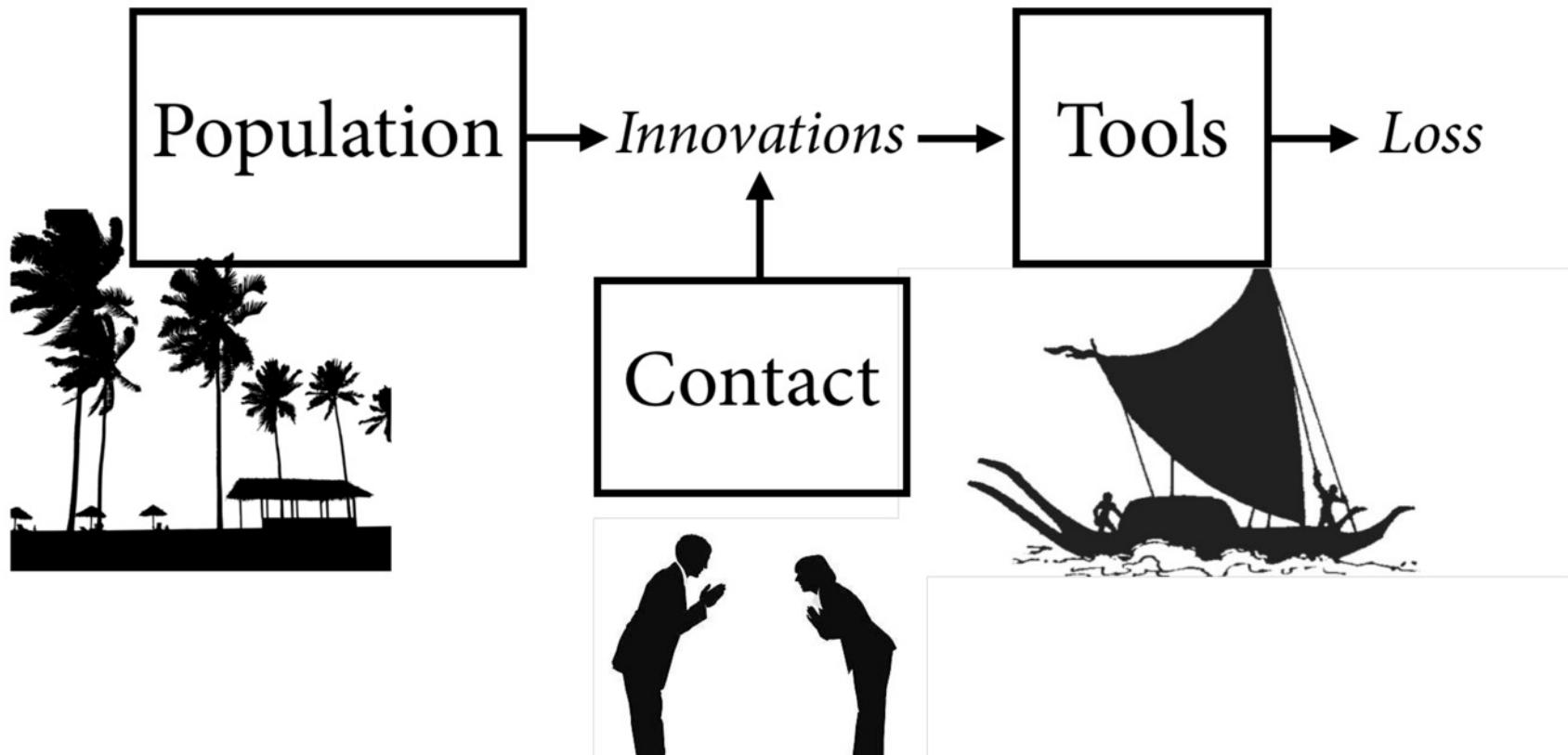
Oceanic tools

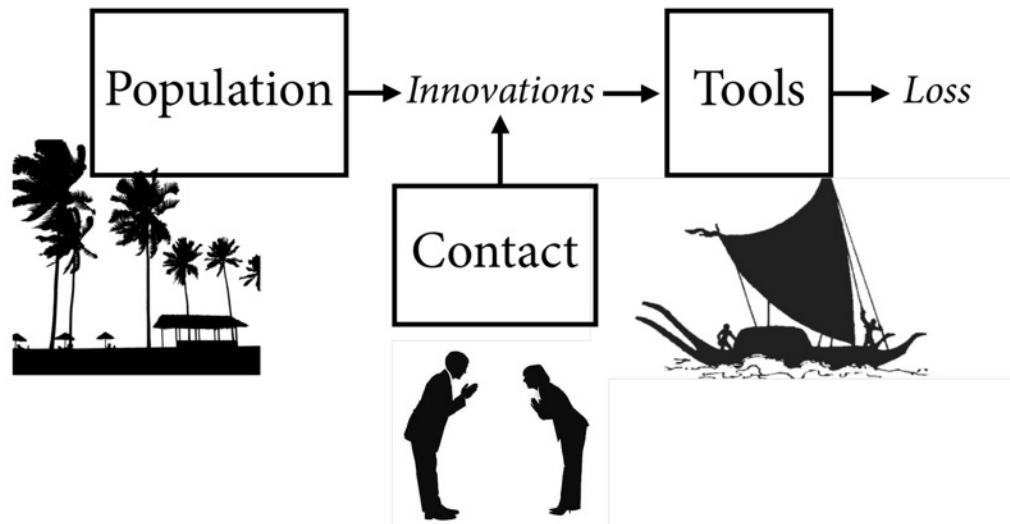
Two immediate ways to improve the model

- (1) Use a robust model:
gamma-Poisson (neg-binomial)
- (2) Use a more principled scientific
model



Technological complexity





$$\Delta T = \alpha P^\beta - \gamma T$$

$$\Delta T = \alpha P^\beta - \gamma T$$

Diagram illustrating the components of the equation:

- change in tools* points to the term αP^β .
- innovation rate* points to the term αP^β .
- diminishing returns (elasticity)* points to the term γT .
- rate of loss* points to the term γT .

*diminishing returns
depend upon contact*

$$\Delta T = \alpha_C P^{\beta_C} - \gamma T$$

*innovation depends
upon contact*

$$\Delta T = \alpha_C P^{\beta_C} - \gamma T = 0$$

Solve for T

$$\Delta T = \alpha_C P^{\beta_C} - \gamma T = 0$$

Solve for T

$$\hat{T} = \frac{\alpha_C P^{\beta_C}}{\gamma}$$

$$\hat{T} = \frac{\alpha_C P^{\beta_C}}{\gamma}$$

$$T_i \sim \text{Poisson}(\lambda_i)$$
$$\lambda_i = \hat{T}$$

```
# innovation/loss model

dat2 <- list( T=d$total_tools, P=d$population,
C=d$contact_id )

m11.11 <- ulam(
  alist(
    T ~ dpois( lambda ),
    lambda <- exp(a[C])*P^b[C]/g,
    a[C] ~ dnorm(1,1),
    b[C] ~ dexp(1),
    g ~ dexp(1)
  ), data=dat2 , chains=4 , cores=4 )
```

All parameters must be positive

Two ways to do this

(1) use `exp()`

(2) use appropriate prior

```

# innovation/loss model

dat2 <- list( T=d$total_tools, P=d$population,
C=d$contact_id )

m11.11 <- ulam(
  alist(
    T ~ dpois( lambda ),
    lambda <- exp(a[C])*P^b[C]/g,
    a[C] ~ dnorm(1,1),
    b[C] ~ dexp(1),
    g ~ dexp(1)
  ), data=dat2 , chains=4
)

```

All parameters must be positive

Two ways to do this

(1) use `exp()`

```

> precis(m11.11,2)
      mean   sd  5.5% 94.5% n_eff Rhat4
a[1] 0.85 0.68 -0.26  1.90   698     1
a[2] 0.93 0.83 -0.39  2.31   902     1
b[1] 0.26 0.03  0.21  0.32  1149     1
b[2] 0.29 0.10  0.12  0.45   711     1
g    1.11 0.70  0.32  2.43   862     1

```

appropriate prior

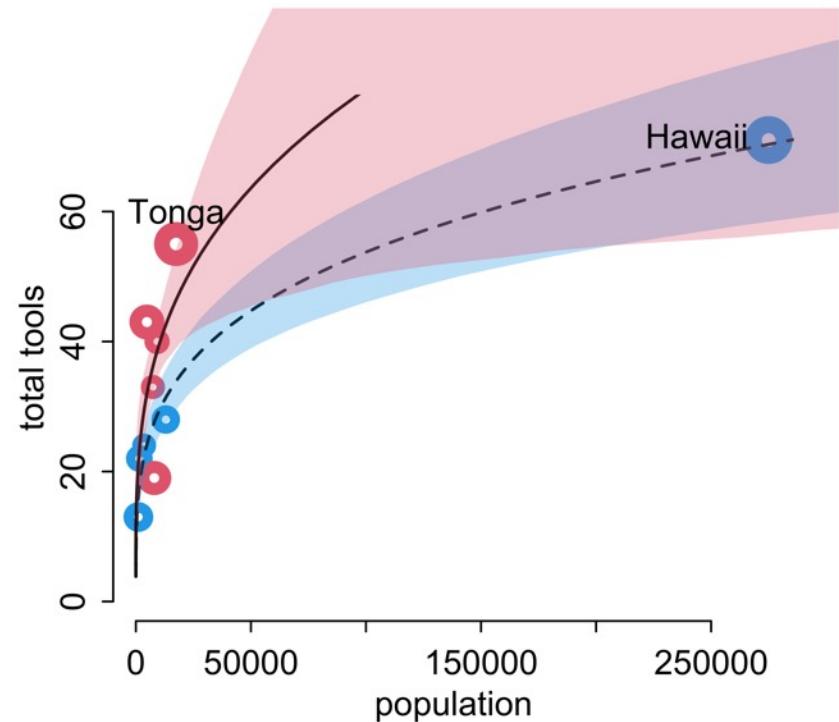
```

# innovation/loss model

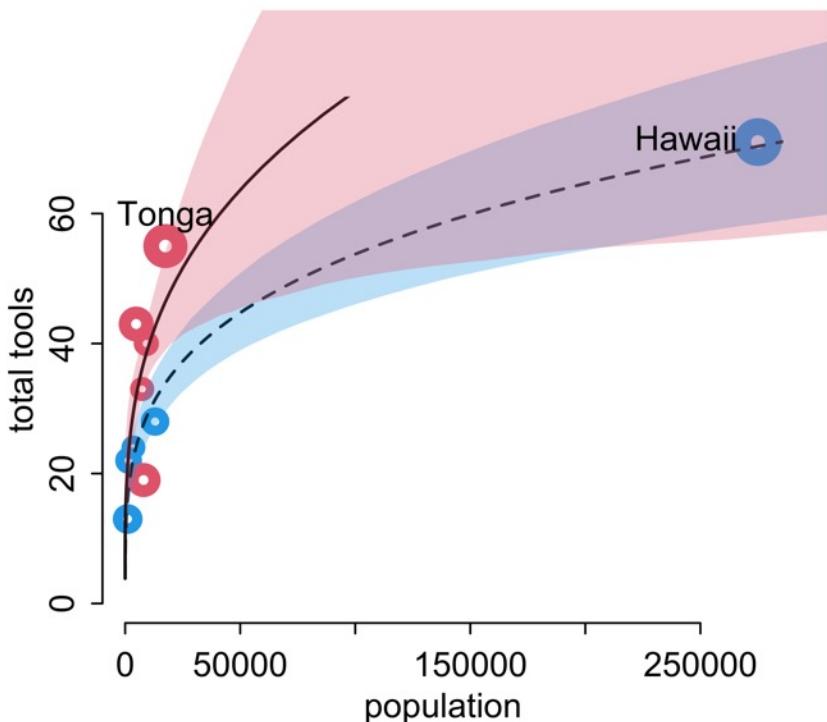
dat2 <- list( T=d$total_tools, P=d$population,
C=d$contact_id )

m11.11 <- ulam(
  alist(
    T ~ dpois( lambda ),
    lambda <- exp(a[C])*P^b[C]/g,
    a[C] ~ dnorm(1,1),
    b[C] ~ dexp(1),
    g ~ dexp(1)
  ), data=dat2 , chains=4 , cores=4 )

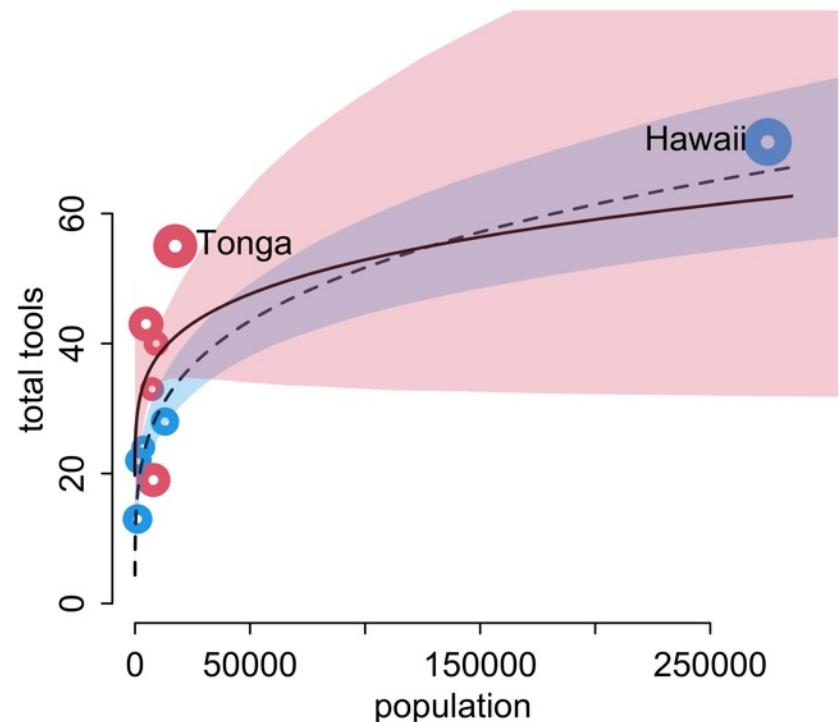
```



Innovation/loss model



Generalized linear model



Still have to deal with location as confound

Count GLMs

Before you see the values, you know a count is zero or positive

Maximum entropy priors: Binomial, Poisson, and extensions

More event types: Multinomial and categorical

Robust regressions:
Beta-binomial, gamma-Poisson (neg-binomial)

Examples to come



Trolley Problems

data(Trolley)

331 individuals (age, gender, edu)

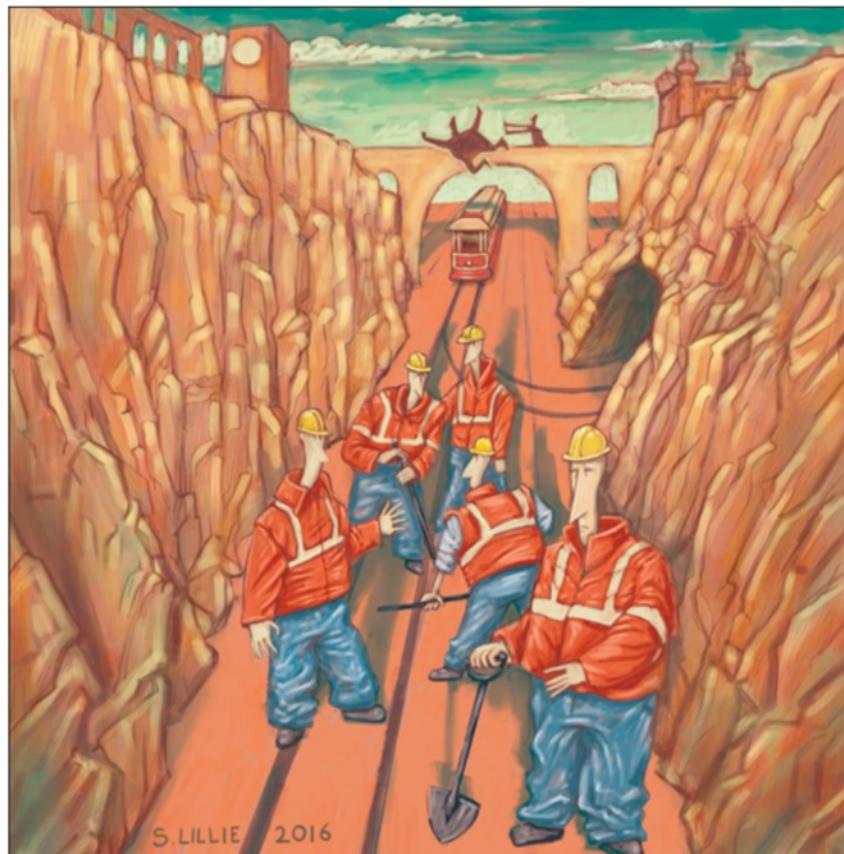
Voluntary participation (online)

30 different trolley problems

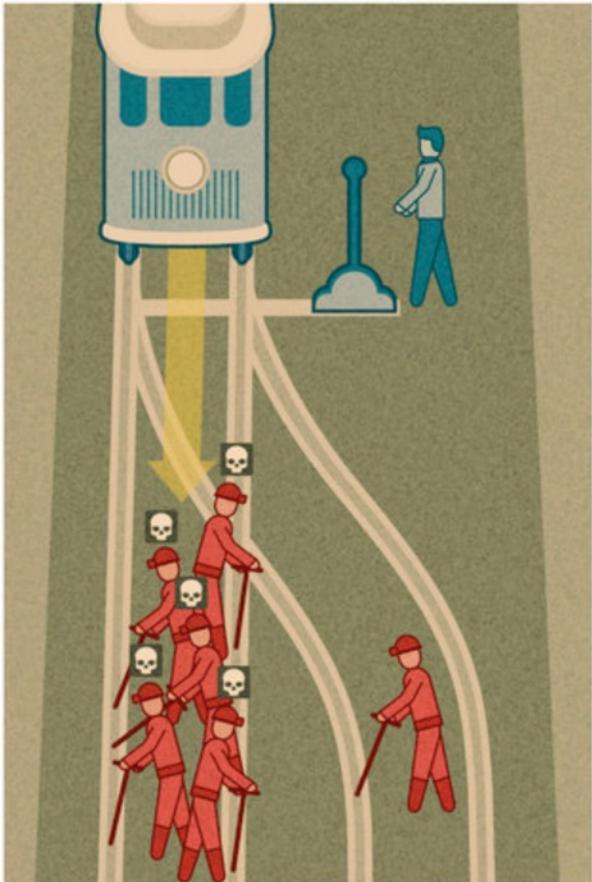
action / intention / contact

9930 responses:

How appropriate (from 1 to 7)?



Action



Trolley Problems

data(Trolley)

331 individuals (age, gender, edu)

Voluntary participation (online)

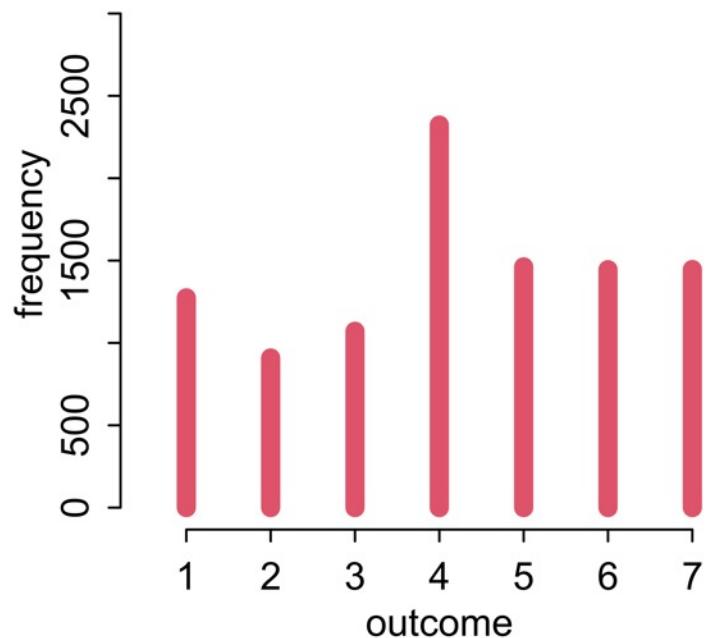
30 different trolley problems

action / intention / contact

9930 responses:

How appropriate (from 1 to 7)?

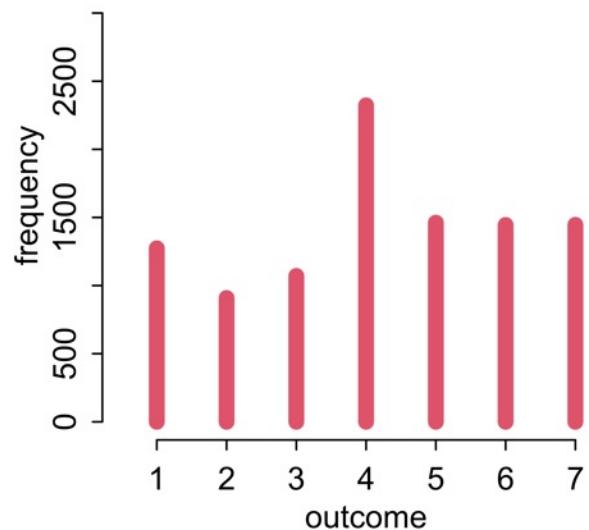
Ordered categorical



Estimand: How do **action**, **intention**, **contact** influence **response** to a trolley story?

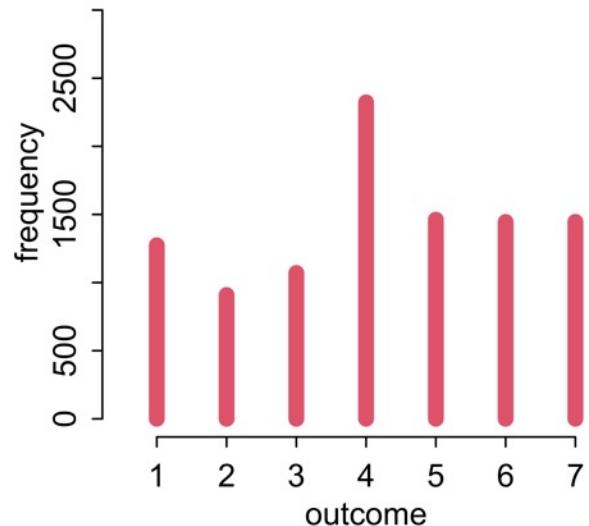
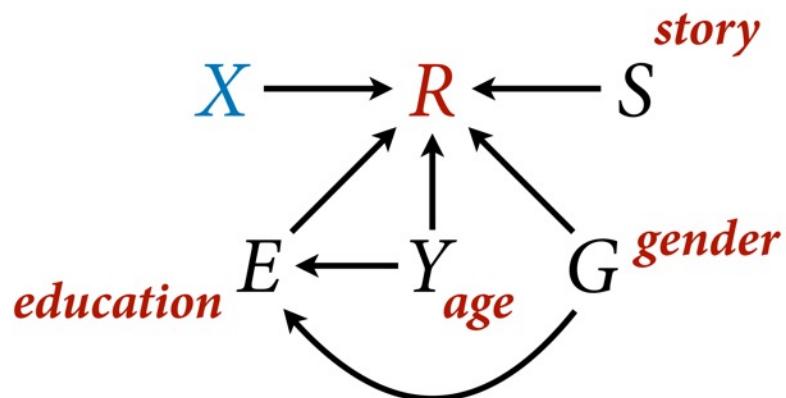
treatment

$X \longrightarrow R$
response



Estimand: How do **action**, **intention**, **contact** influence **response** to a trolley story?

How are influences of A/I/C associated with other variables?



Ordered categories

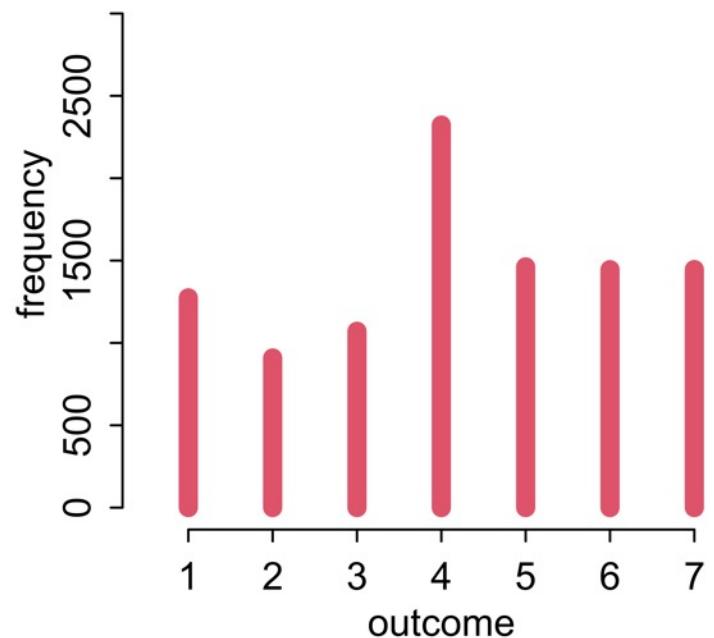
Categories: Discrete types

cat, dog, chicken

Ordered categories: Discrete types
with ordered relationships

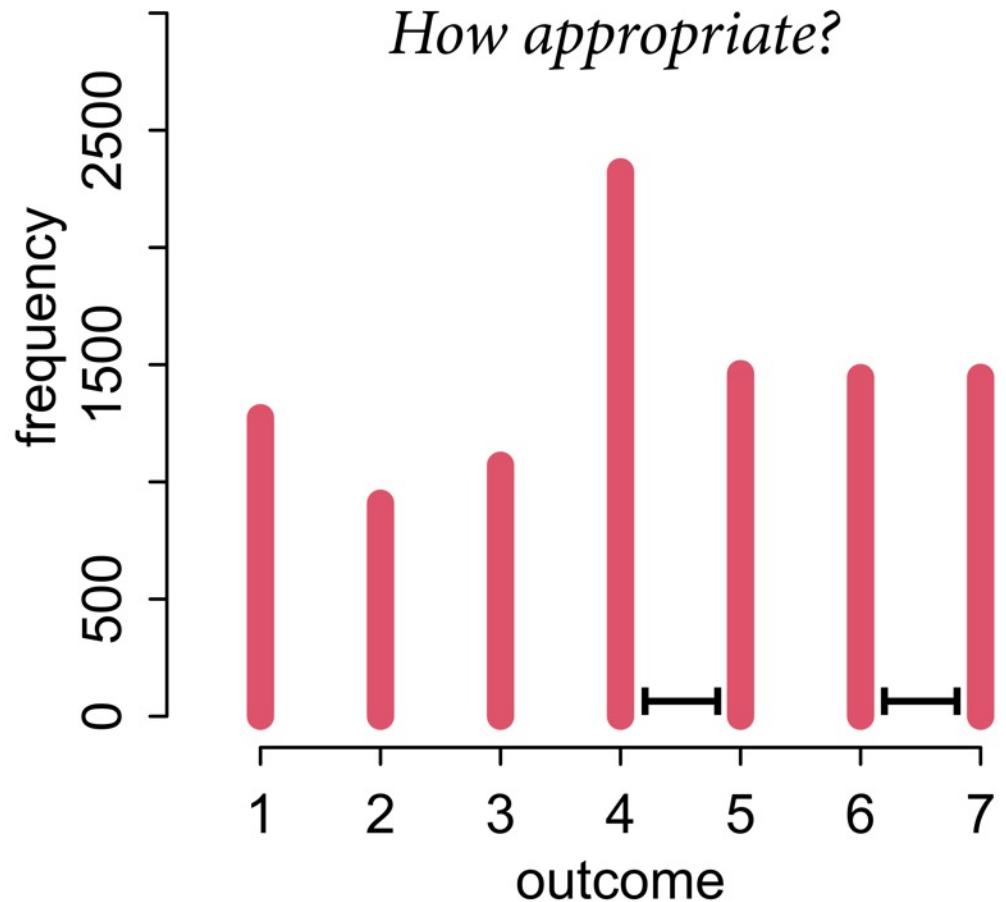
bad, good, excellent

Ordered categorical



Distance between values
not constant

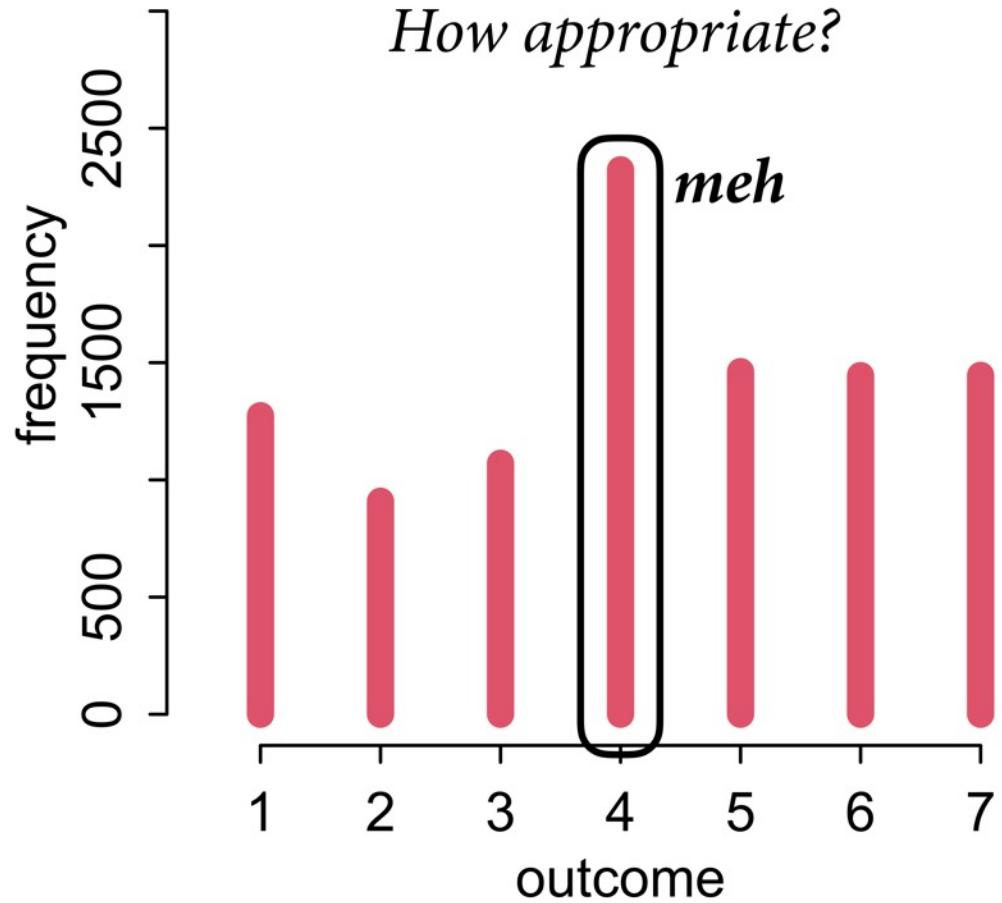
Probably much easier to
go from 4 to 5 than from
6 to 7



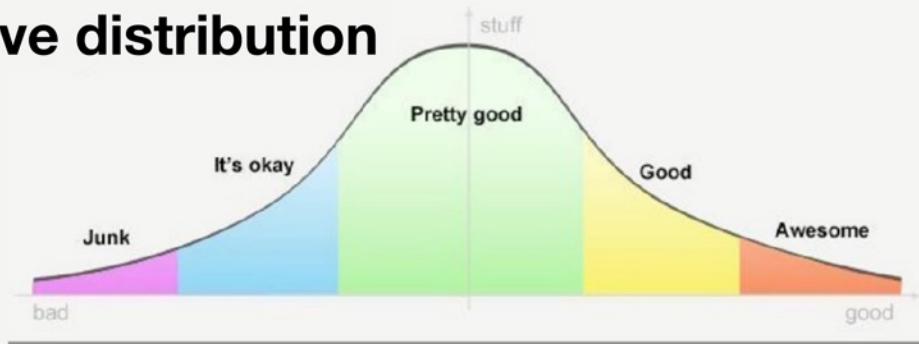
Anchor points common

Not everyone shares the same anchor points

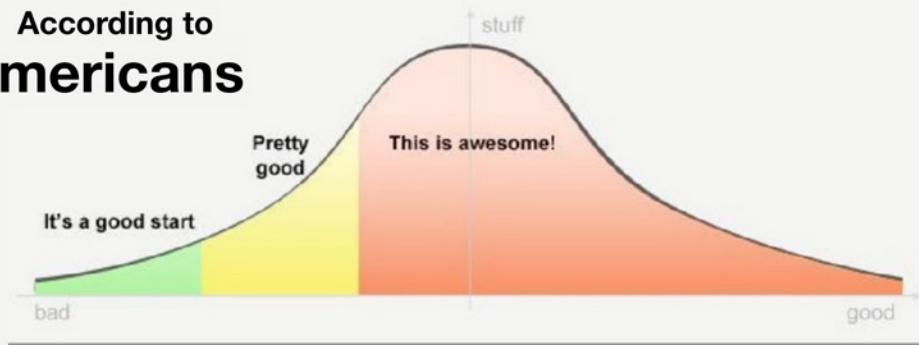
How appropriate?



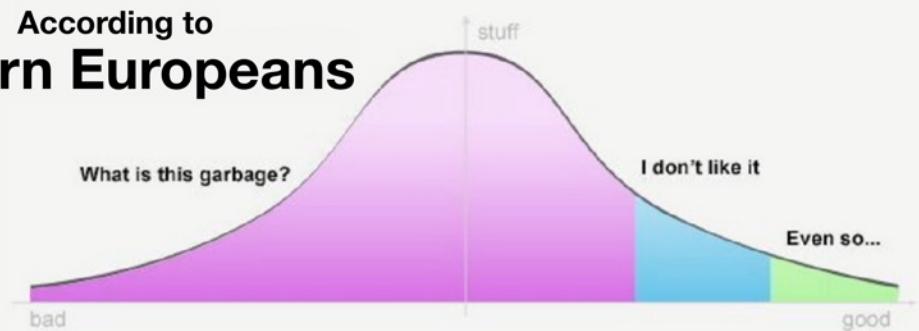
Objective distribution

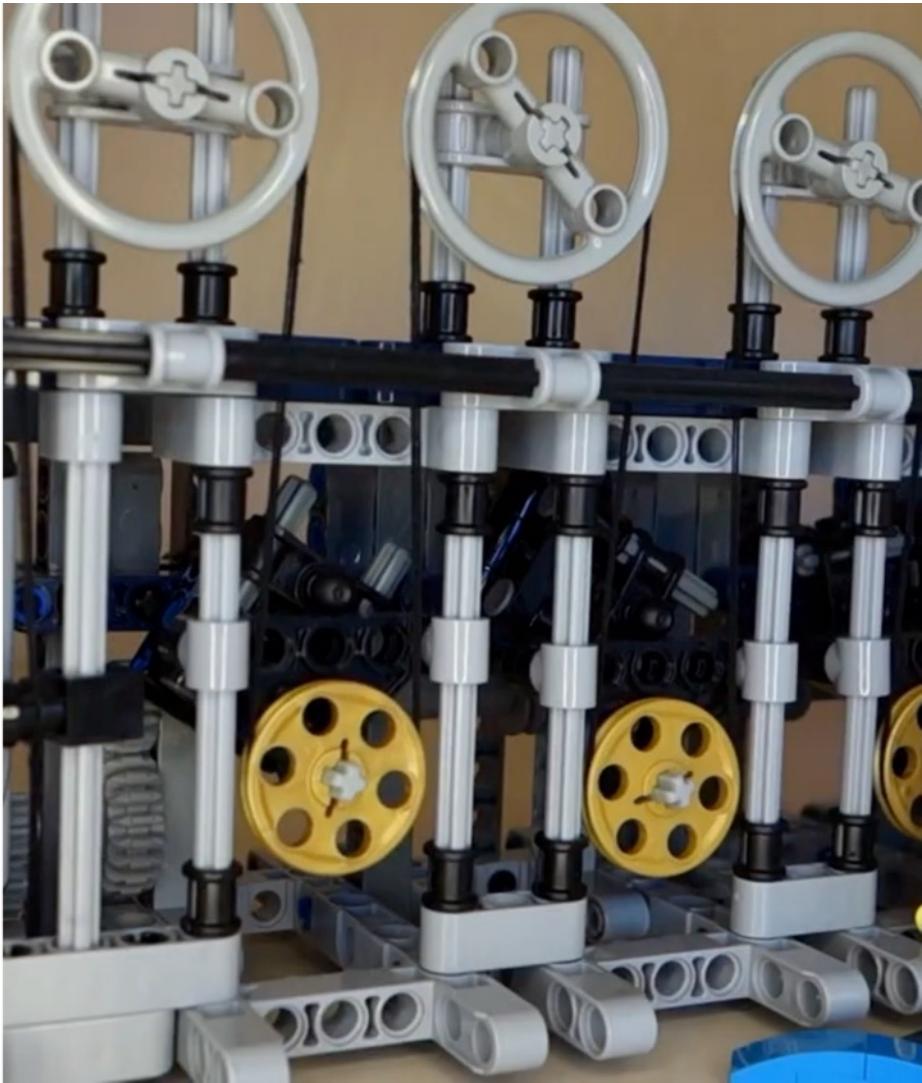
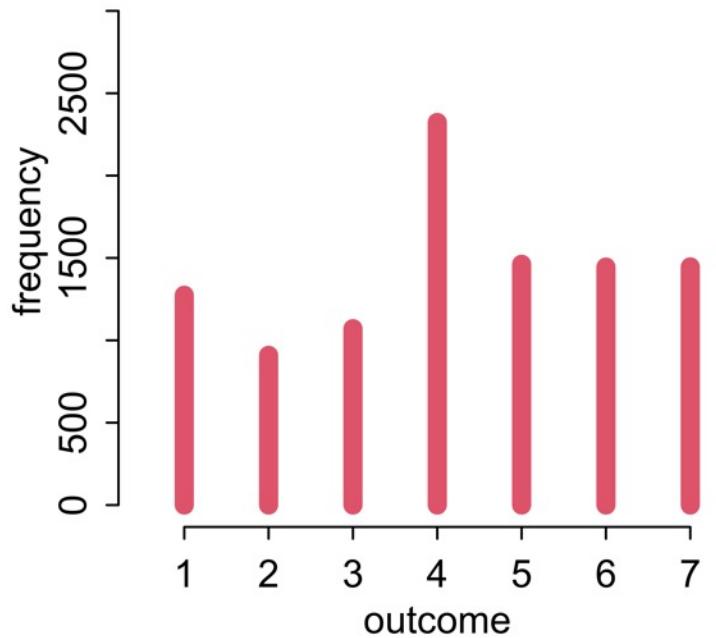


According to Americans

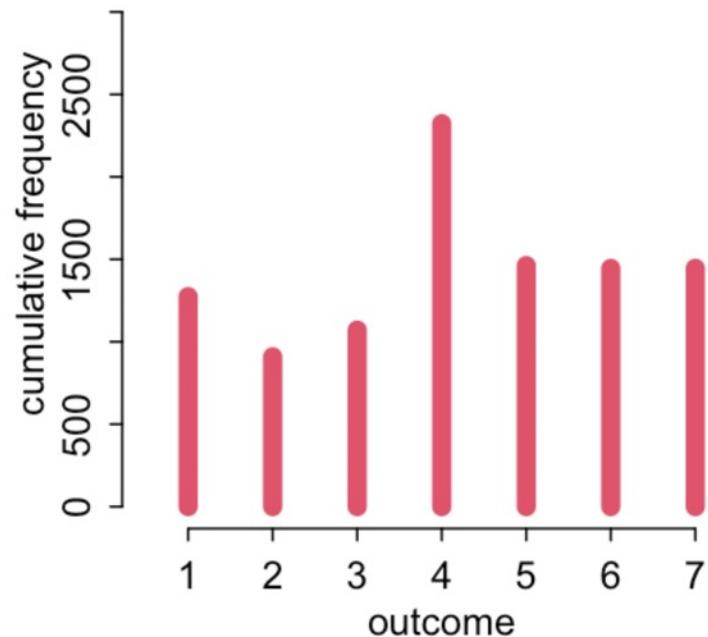
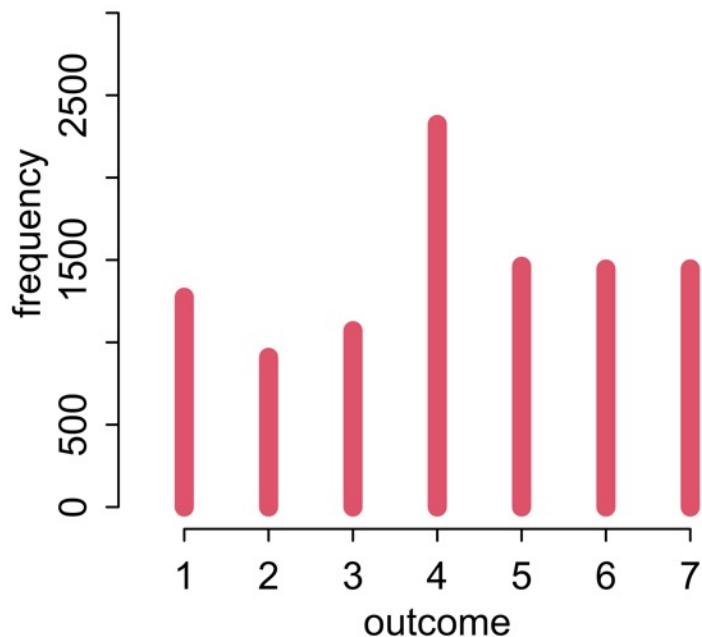


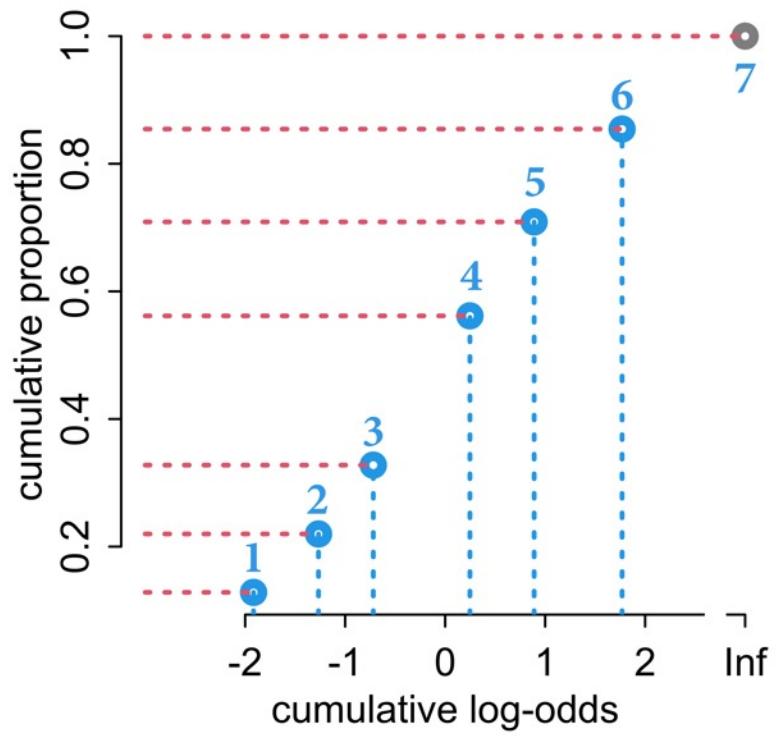
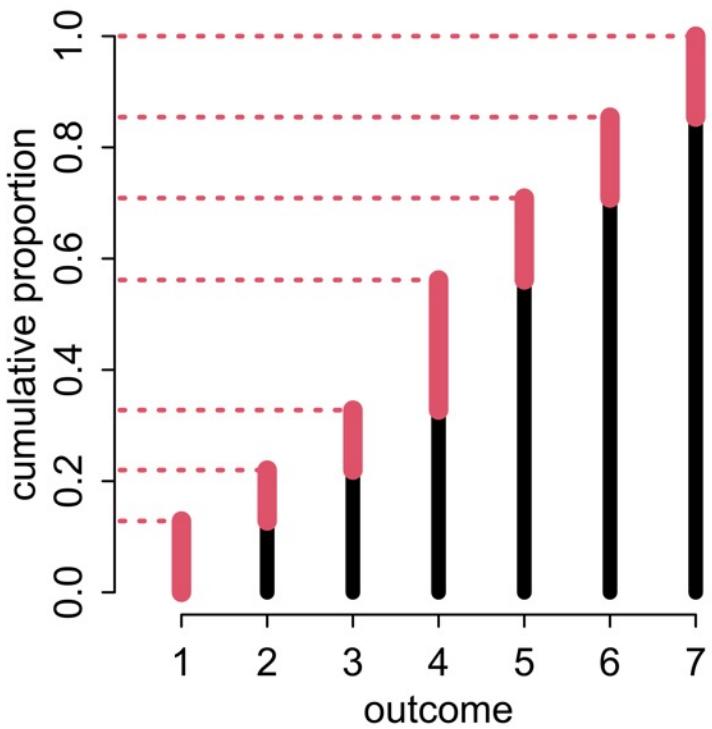
According to Eastern Europeans

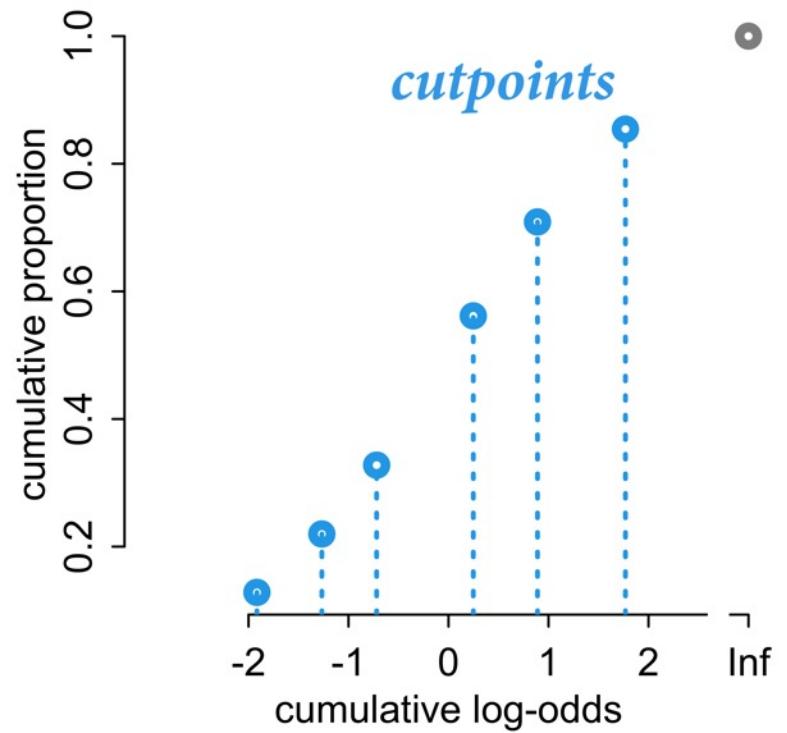
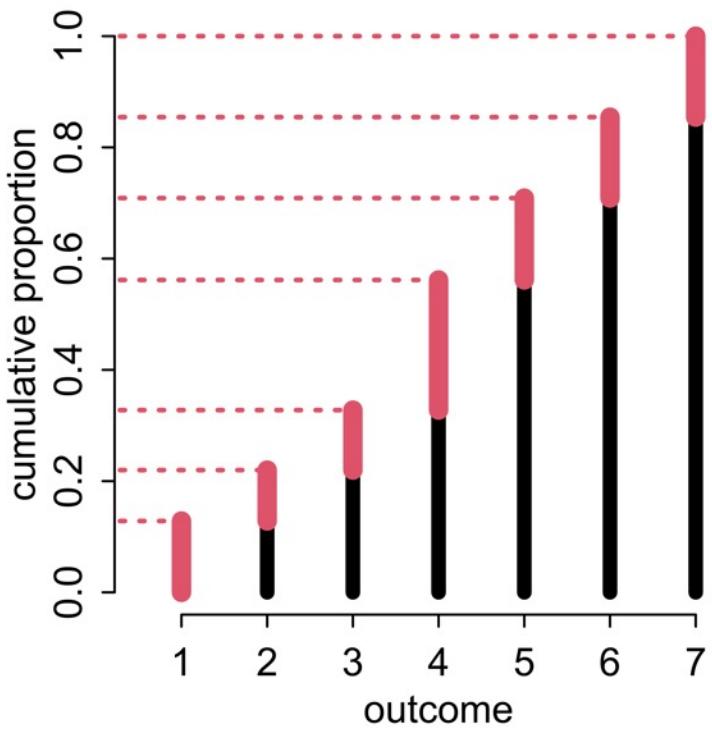


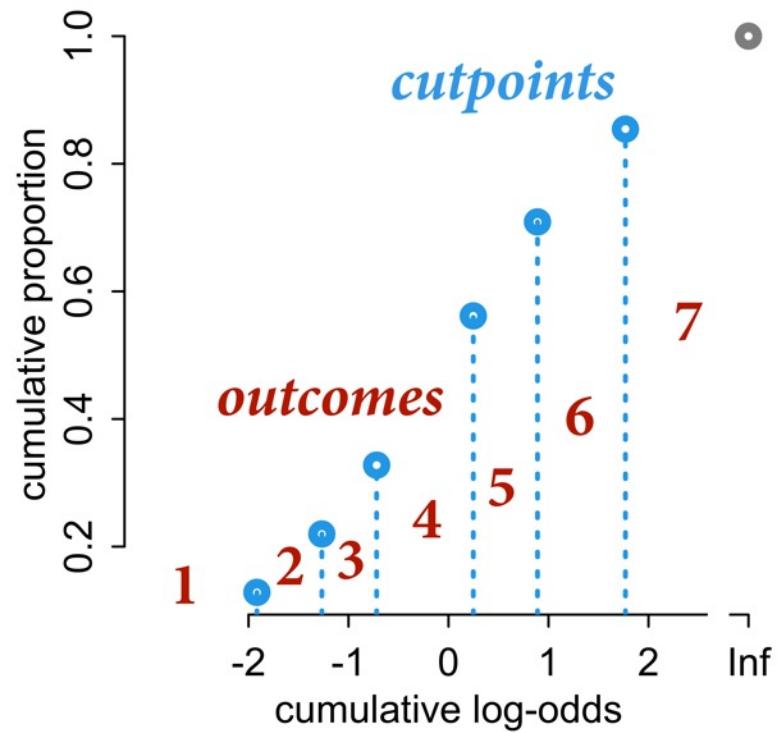
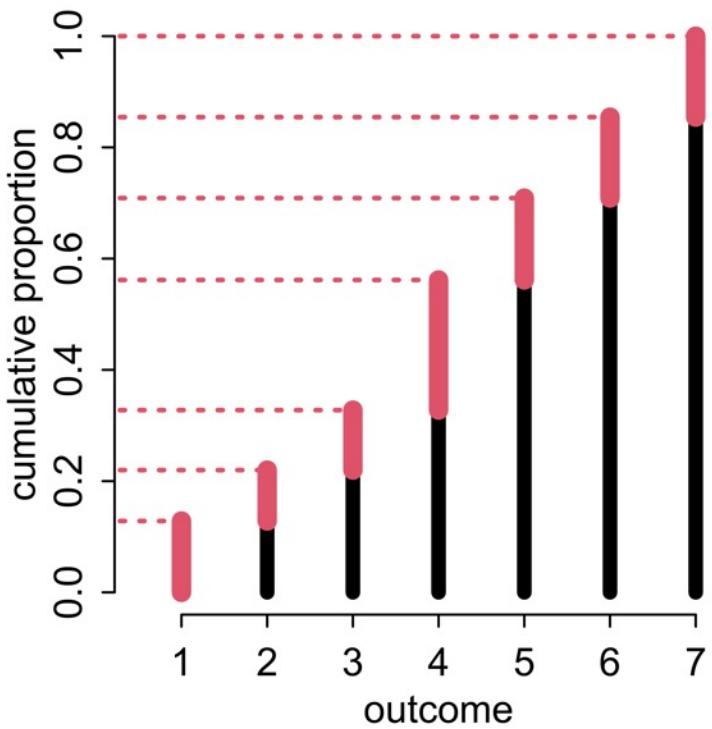


Ordered = Cumulative

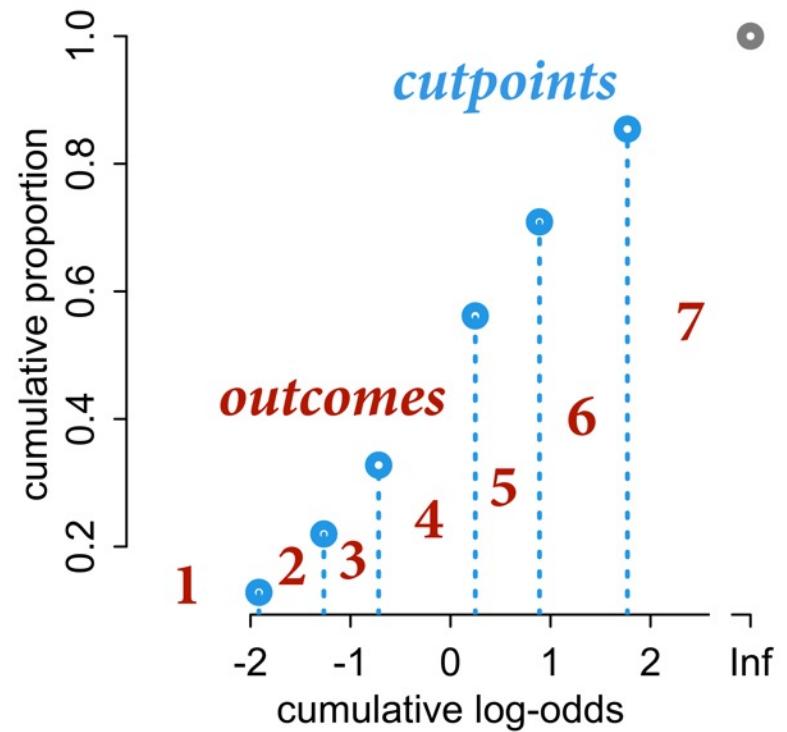




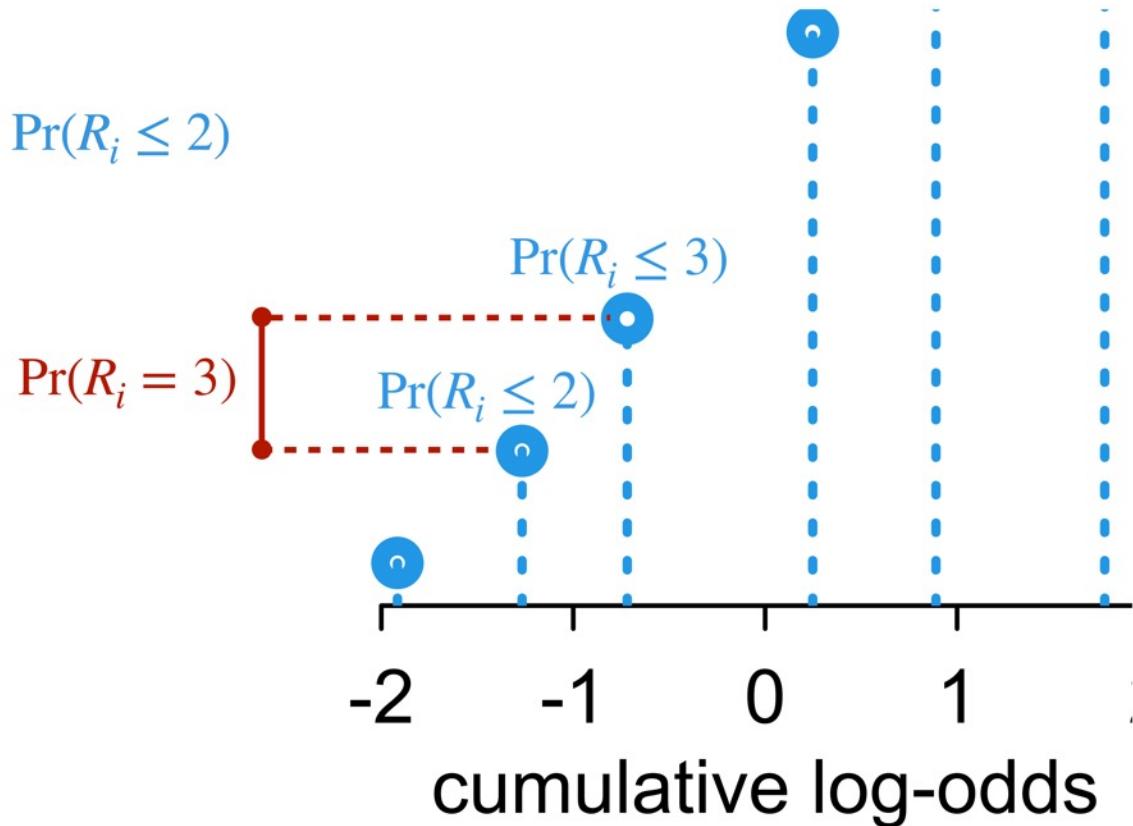




$$\Pr(R_i = k) = \Pr(R_i \leq k) - \Pr(R_i \leq k - 1)$$



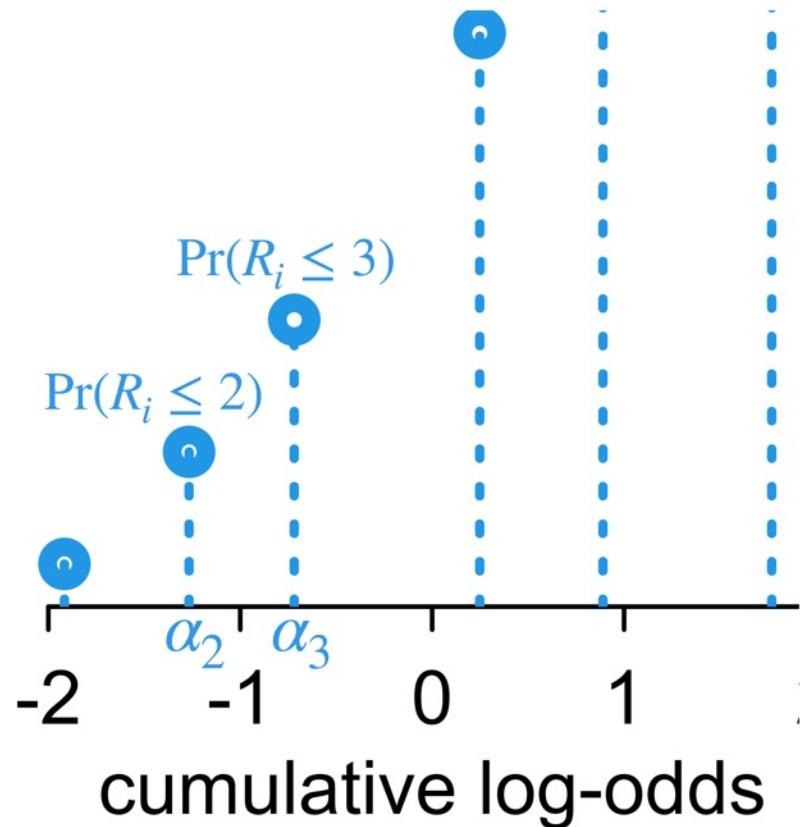
$$\Pr(R_i = 3) = \Pr(R_i \leq 3) - \Pr(R_i \leq 2)$$



$$\Pr(R_i = 3) = \Pr(R_i \leq 3) - \Pr(R_i \leq 2)$$

$$\frac{\log \frac{\Pr(R_i \leq k)}{1 - \Pr(R_i \leq k)}}{cumulative\ log-odds} = \alpha_k$$

*cutpoint
(to estimate)*



Where's the GLM?

So far just estimating the histogram

How to make it a function of variables?

(1) Stratify cutpoints

(2) Offset each cutpoint by value of linear model ϕ_i

Where's the GLM?

So far just estimating the histogram

How to make it a function of variables?

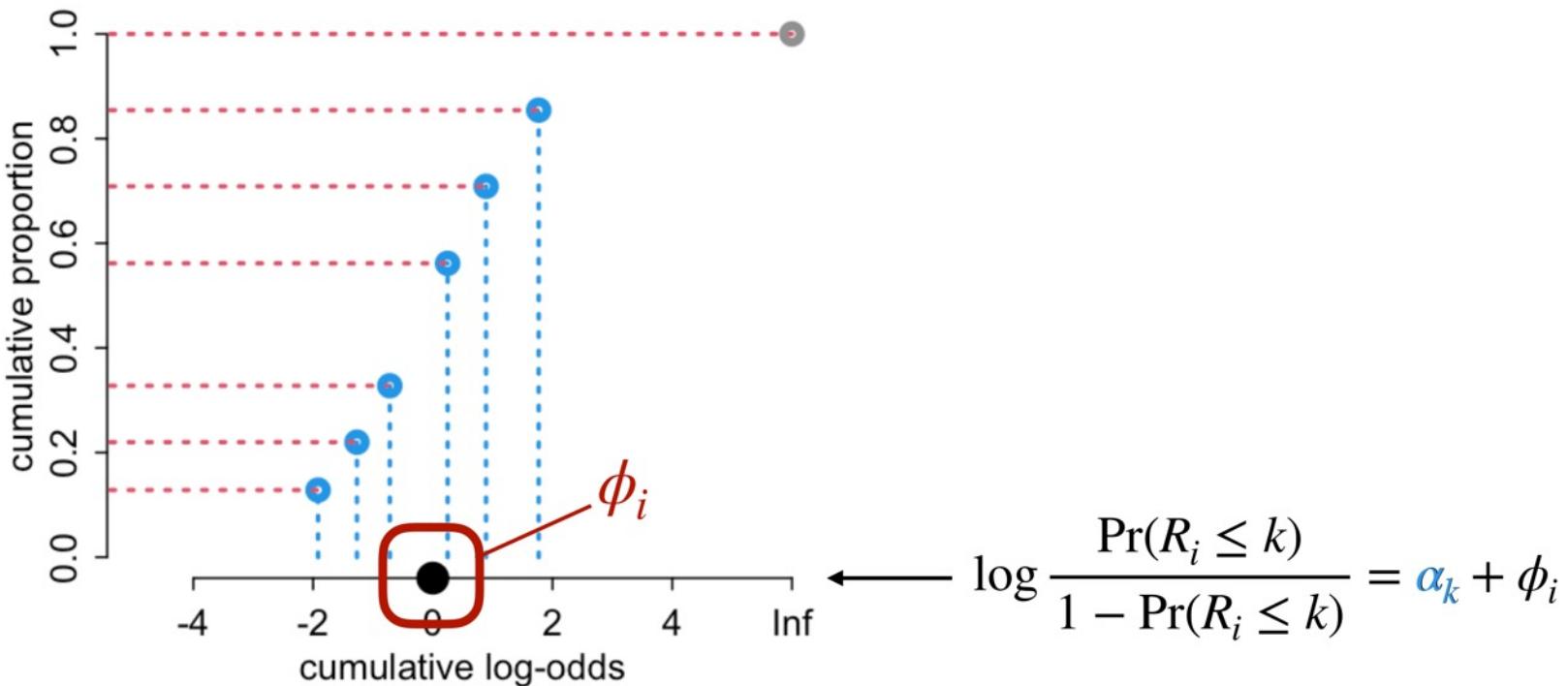
(1) Stratify cutpoints

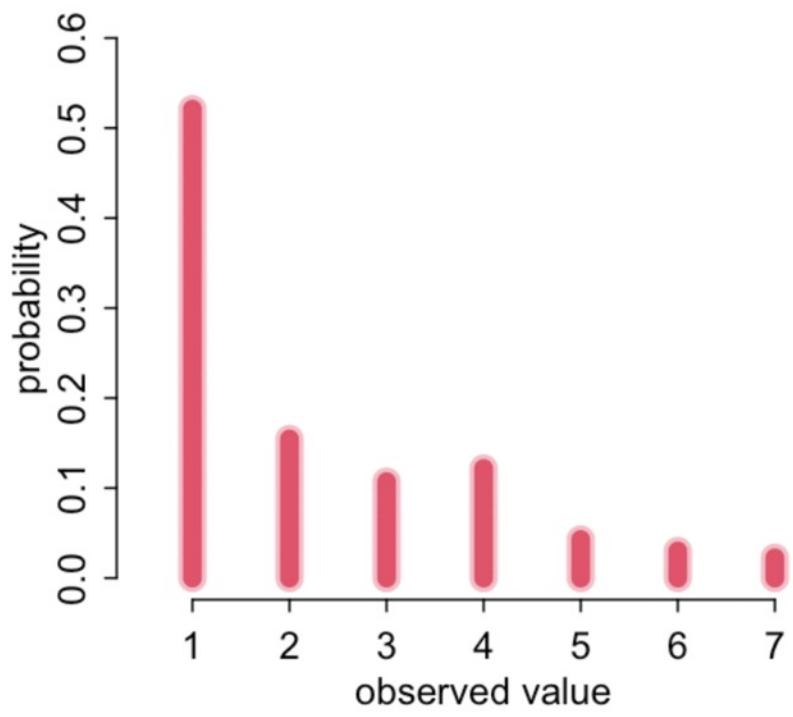
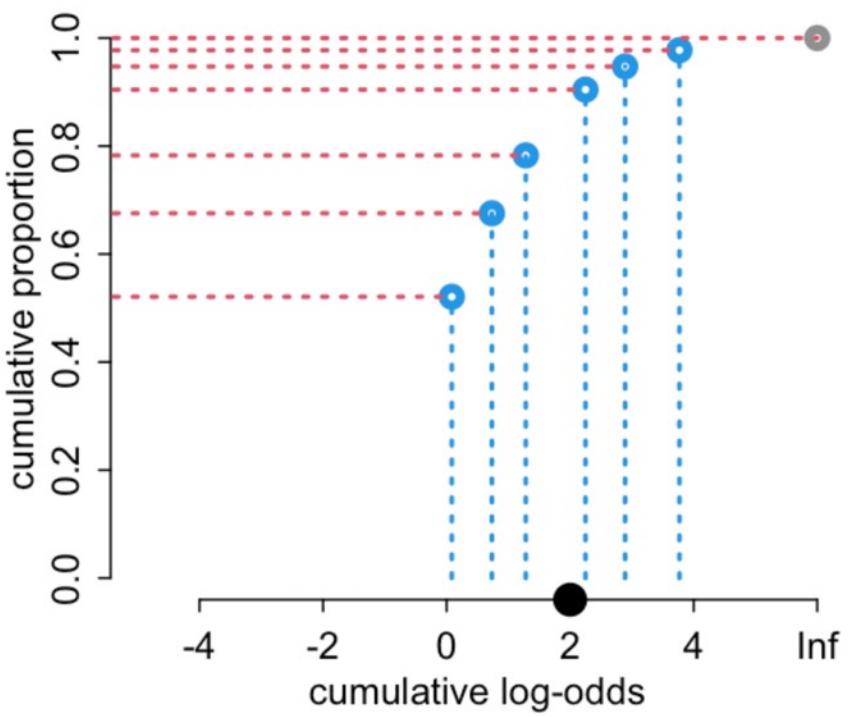
(2) Offset each cutpoint by value of linear model ϕ_i

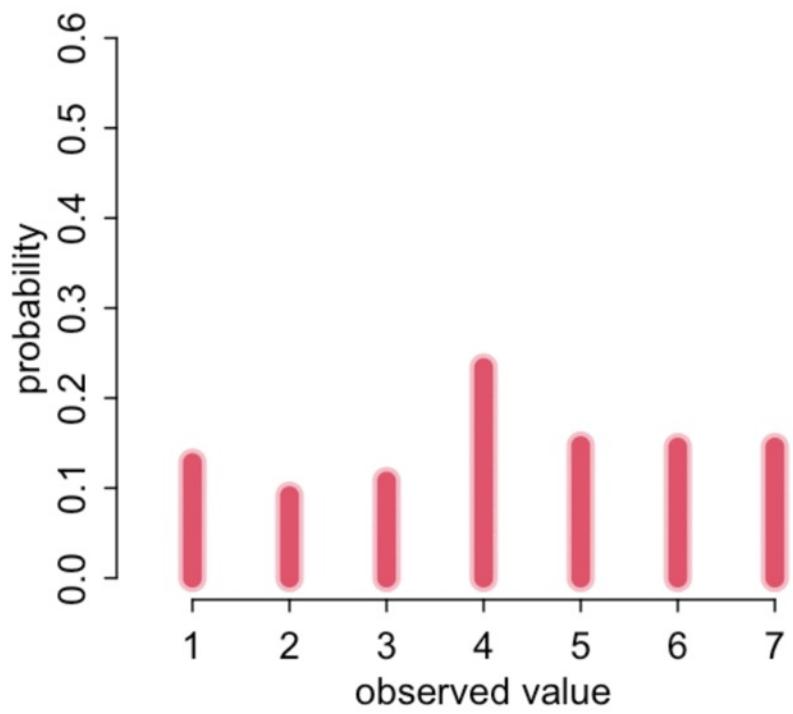
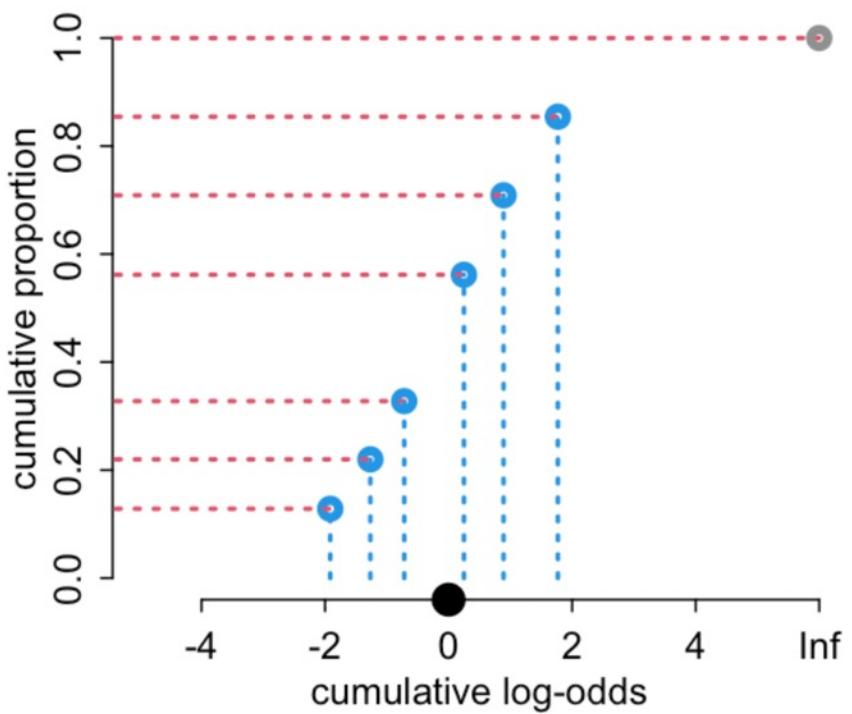
$$\phi_i = \beta x_i$$

$$\log \frac{\Pr(R_i \leq k)}{1 - \Pr(R_i \leq k)} = \alpha_k + \phi_i$$

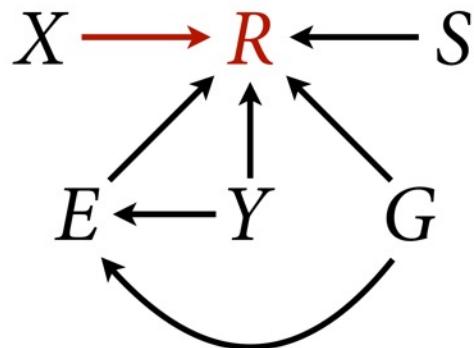
$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$







Start off easy:



$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_A A_i + \beta_C C_i + \beta_I I_i$$

$$\beta_- \sim \text{Normal}(0, 0.5)$$

$$\alpha_j \sim \text{Normal}(0, 1)$$

```

data(Trolley)
d <- Trolley
dat <- list(
  R = d$response,
  A = d$action,
  I = d$intention,
  C = d$contact
)
mRX <- ulam(
  alist(
    R ~ dordlogit(phi,alpha),
    phi <- bA*A + bI*I + bC*C,
    c(bA,bI,bC) ~ normal(0,0.5),
    alpha ~ normal(0,1)
  ) , data=dat , chains=4 , cores=4 )

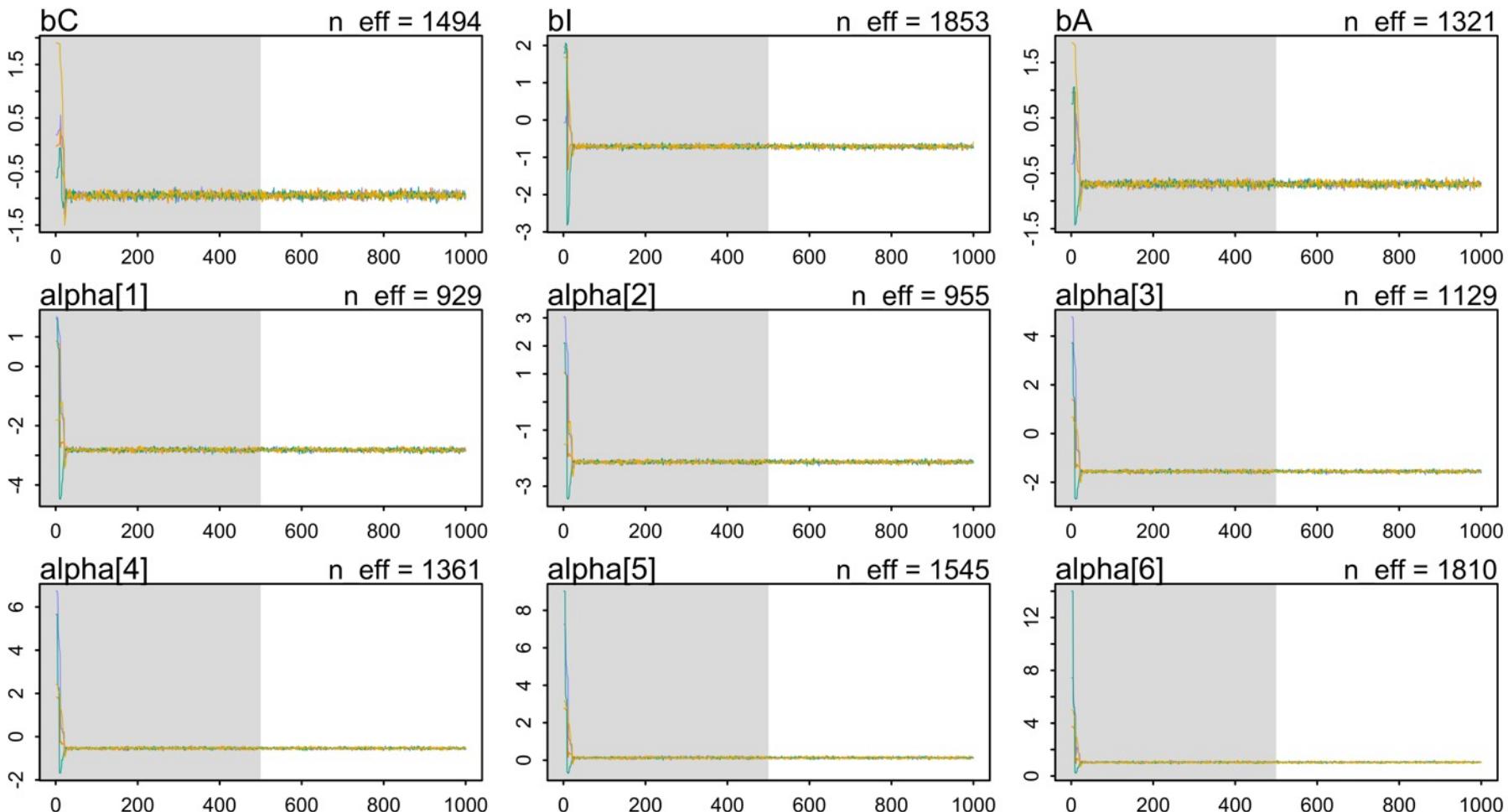
```

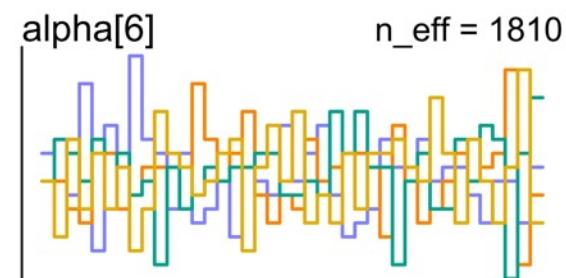
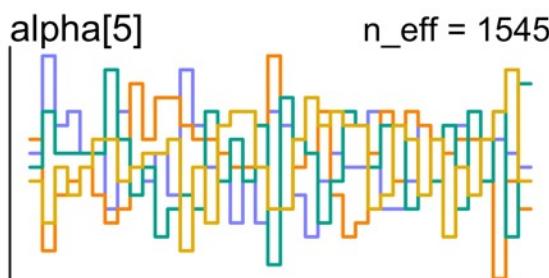
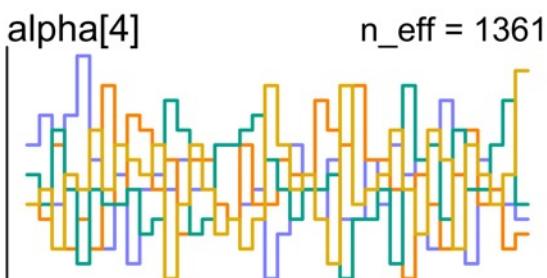
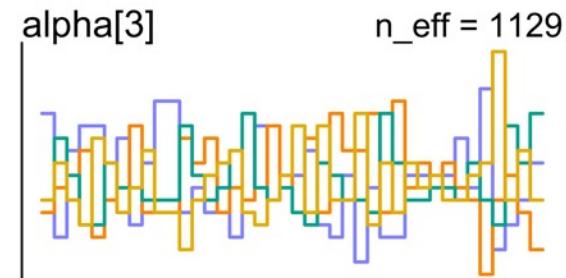
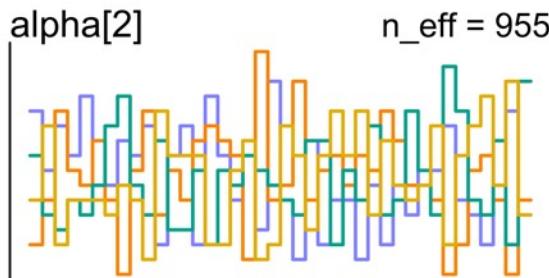
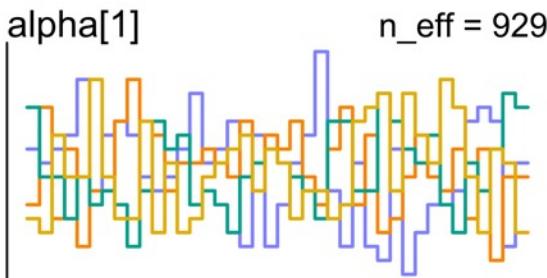
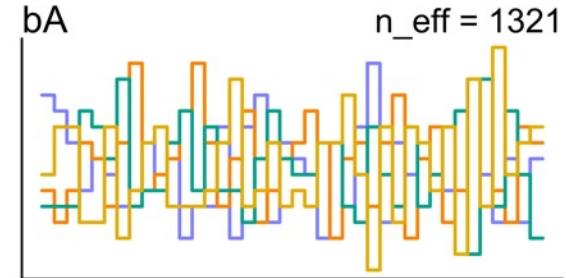
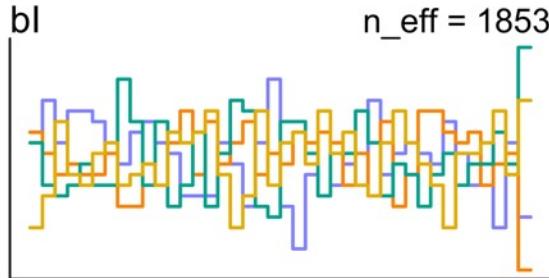
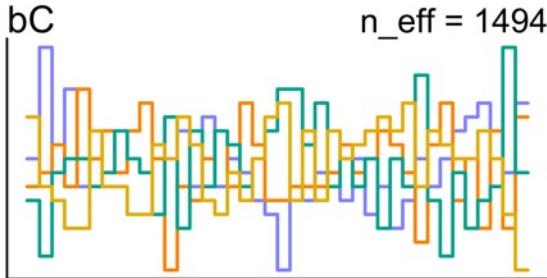
$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_A A_i + \beta_C C_i + \beta_I I_i$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

$$\alpha_j \sim \text{Normal}(0,1)$$





```

data(Trolley)
d <- Trolley
dat <- list(
  R = d$response,
  A = d$action,
  I = d$intention,
  C = d$contact
)

mRX <- ulam(
  alist(
    R ~ dordlogit(phi,alpha),
    phi <- bA*A + bI*I + bC*C,
    c(bA,bI,bC) ~ normal(0,0.5),
    alpha ~ normal(0,1)
  ) , data=dat , chains=4 , cores=4 )

```

```

> precis(mRX,2)
      mean   sd  5.5% 94.5% n_eff Rhat4
bC     -0.94 0.05 -1.02 -0.87  1494    1
bI     -0.71 0.04 -0.77 -0.65  1853    1
bA     -0.69 0.04 -0.76 -0.63  1321    1
alpha[1] -2.82 0.05 -2.89 -2.74   929    1
alpha[2] -2.14 0.04 -2.20 -2.07   955    1
alpha[3] -1.56 0.04 -1.62 -1.49  1129    1
alpha[4] -0.54 0.04 -0.59 -0.48  1361    1
alpha[5]  0.13 0.04  0.07  0.19  1545    1
alpha[6]  1.04 0.04  0.97  1.10  1810    1

```

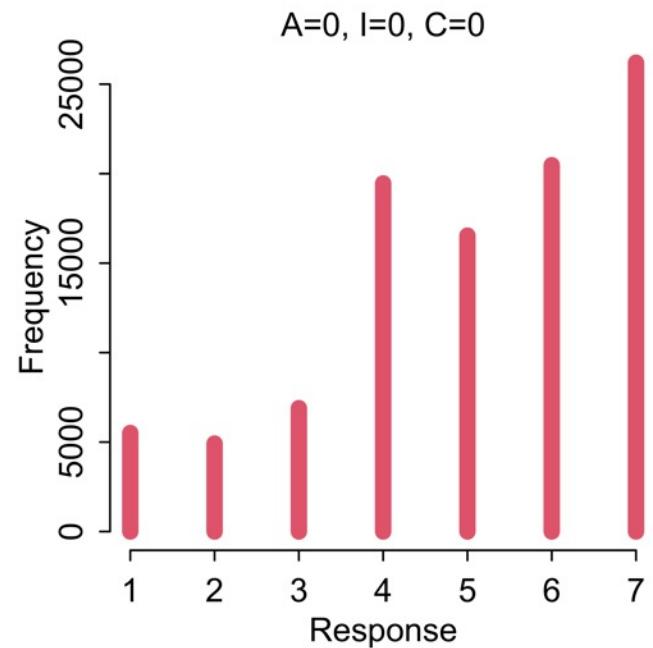
```

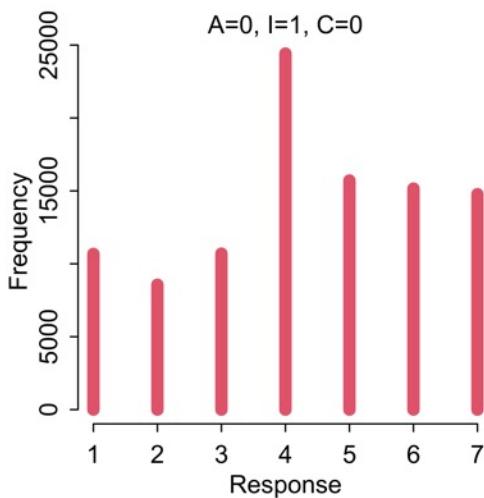
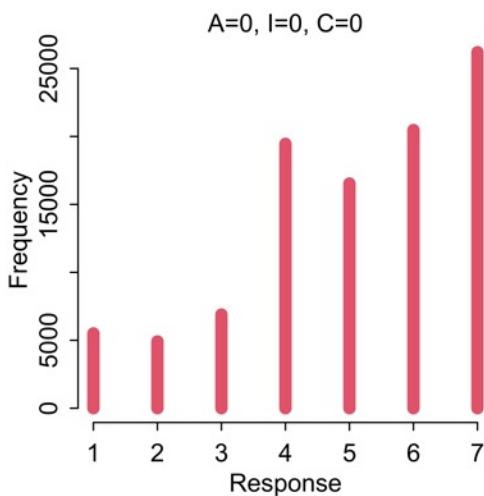
# plot predictive distributions for each treatment

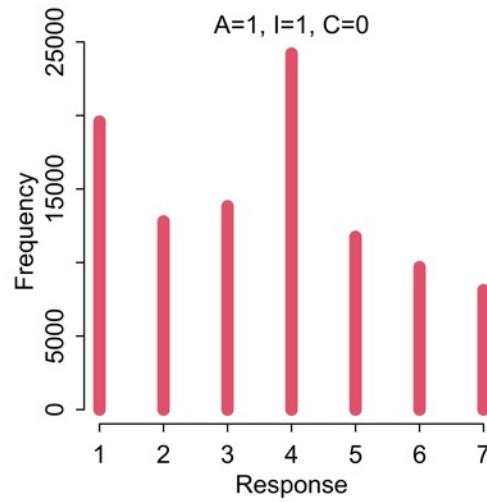
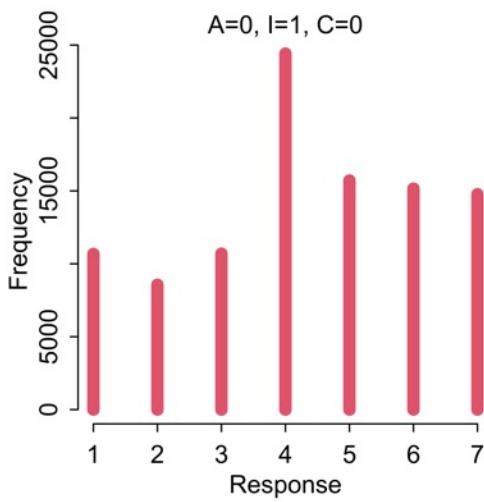
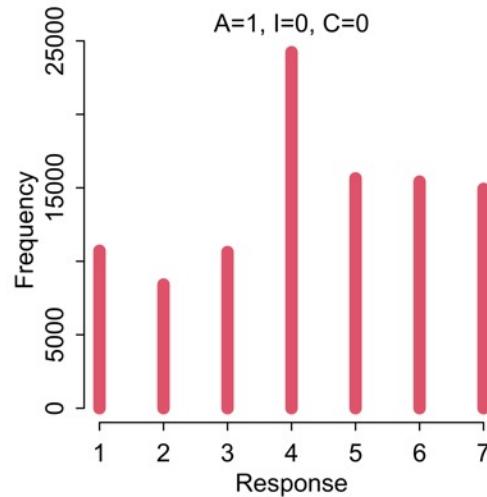
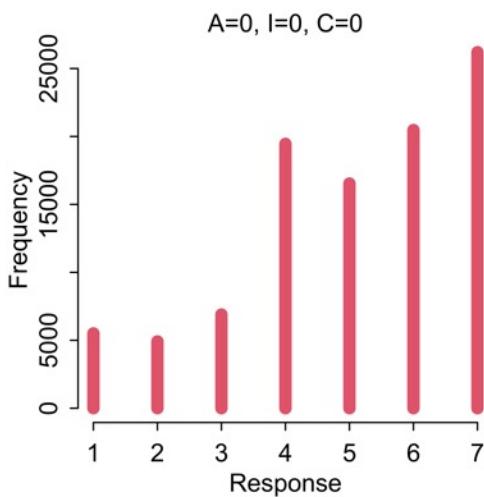
vals <- c(0,0,0)
Rsim <- mcreplicate( 100 ,
sim(mRX,data=list(A=vals[1],I=vals[2],C=vals[3])) ,
mc.cores=6 )

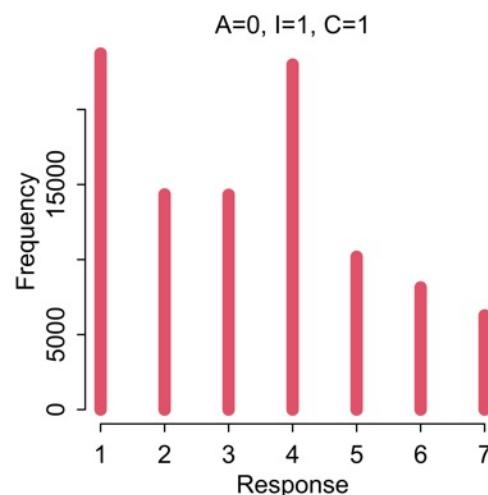
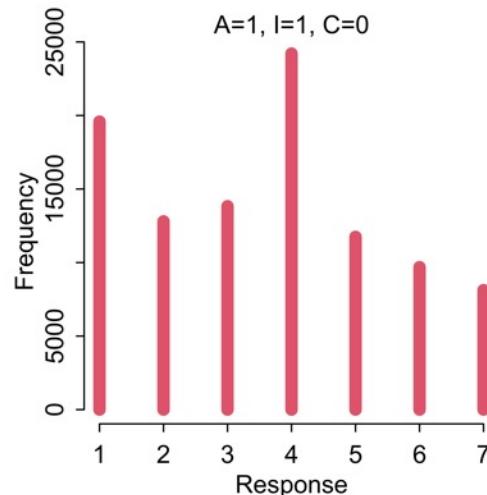
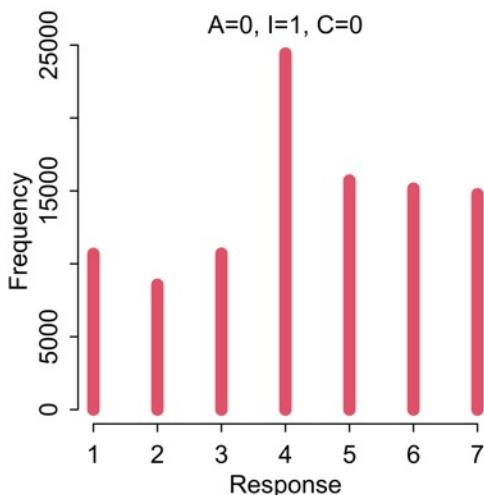
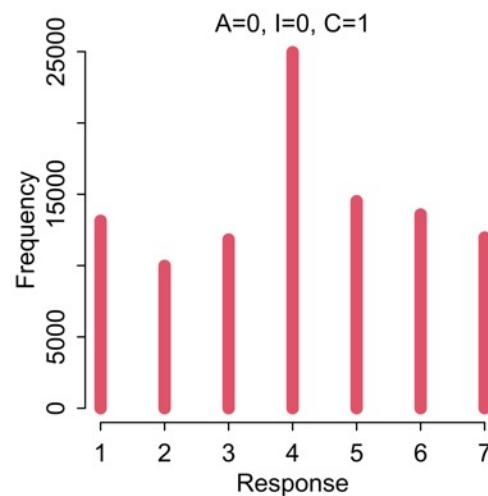
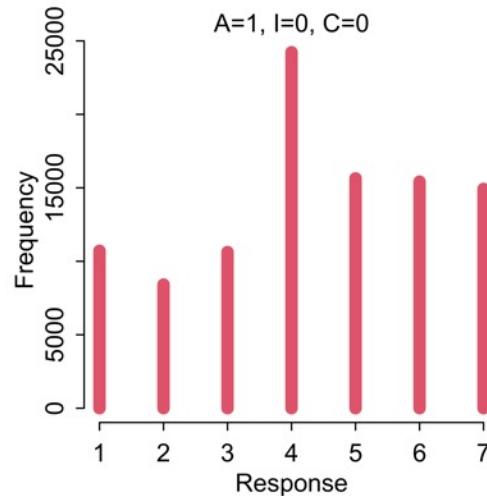
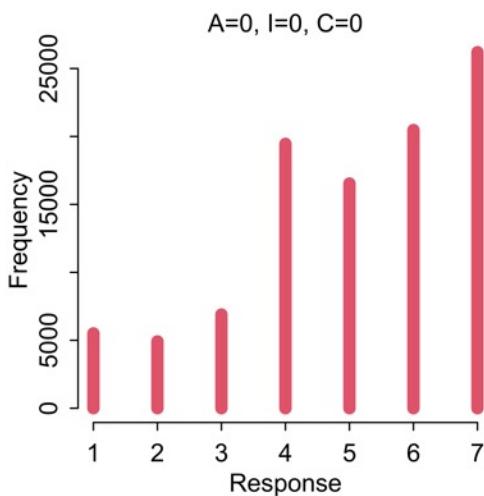
simplehist(as.vector(Rsim),lwd=8,col=2,xlab="Response")
mtext(concat("A=",vals[1]," I=",vals[2]," C=",
C=vals[3]))

```

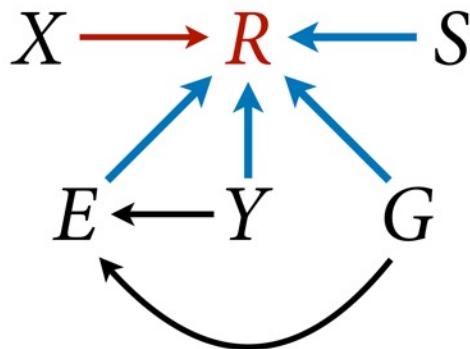








What about the competing causes?



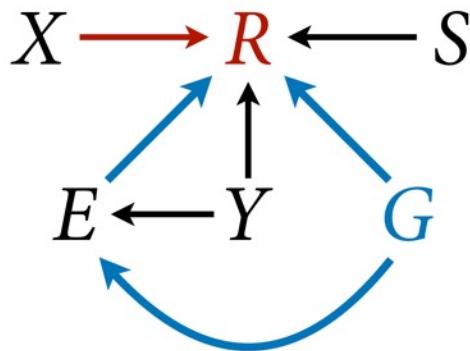
$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_A A_i + \beta_C C_i + \beta_I I_i$$

$$\beta \sim \text{Normal}(0, 0.5)$$

$$\alpha_j \sim \text{Normal}(0, 1)$$

Total effect of gender:



$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_{A,G[i]} A_i + \beta_{C,G[i]} C_i + \beta_{I,G[i]} I_i$$

$$\beta_- \sim \text{Normal}(0, 0.5)$$

$$\alpha_j \sim \text{Normal}(0, 1)$$

```

# total effect of gender
dat$G <- ifelse(d$male==1,2,1)
mRXG <- ulam(
  alist(
    R ~ dordlogit(phi,alpha),
    phi <- bA[G]*A + bI[G]*I + bC[G]*C,
    bA[G] ~ normal(0,0.5),
    bI[G] ~ normal(0,0.5),
    bC[G] ~ normal(0,0.5),
    alpha ~ normal(0,1)
  ) , data=dat , chains=4 , cores=4 )

```

$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_{A,G[i]} A_i + \beta_{C,G[i]} C_i + \beta_{I,G[i]} I_i$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

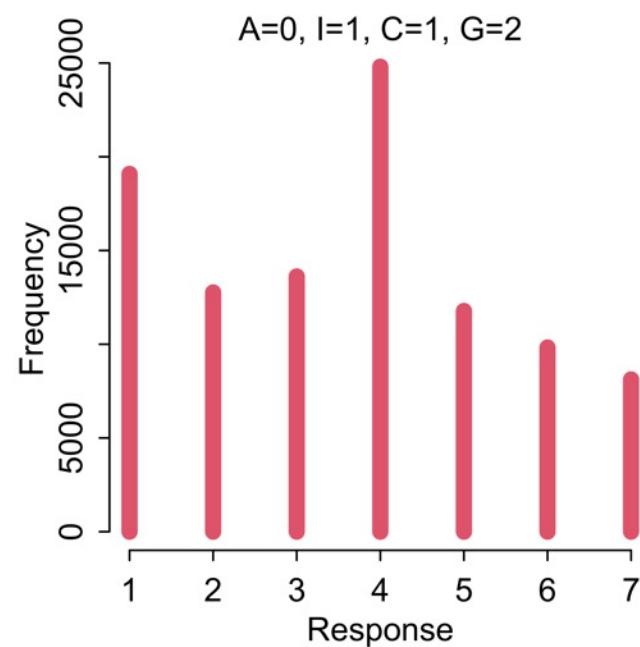
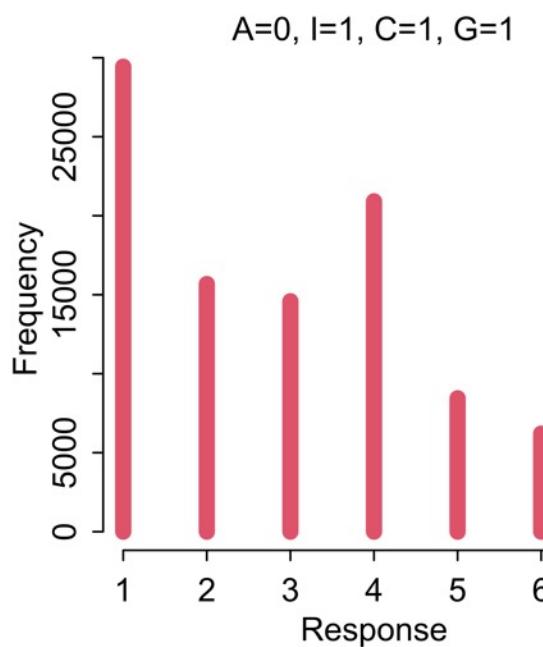
$$\alpha_j \sim \text{Normal}(0,1)$$

```

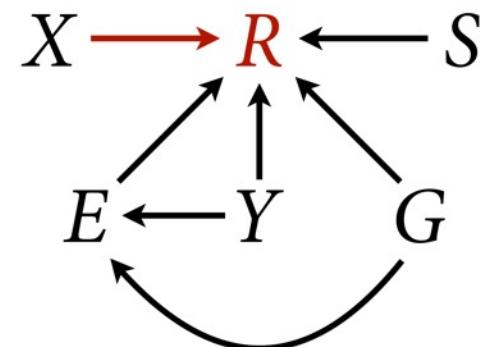
# total effect of gender
dat$G <- ifelse(d$male==1,2,1)
mRXG <- ulam(
  alist(
    R ~ dordlogit(phi,alpha),
    phi <- bA[G]*A + bI[G]*I + bC[G]*C,
    bA[G] ~ normal(0,0.5),
    bI[G] ~ normal(0,0.5),
    bC[G] ~ normal(0,0.5),
    alpha ~ normal(0,1)
  ) , data=dat , chains=4 , cores=4 )

```

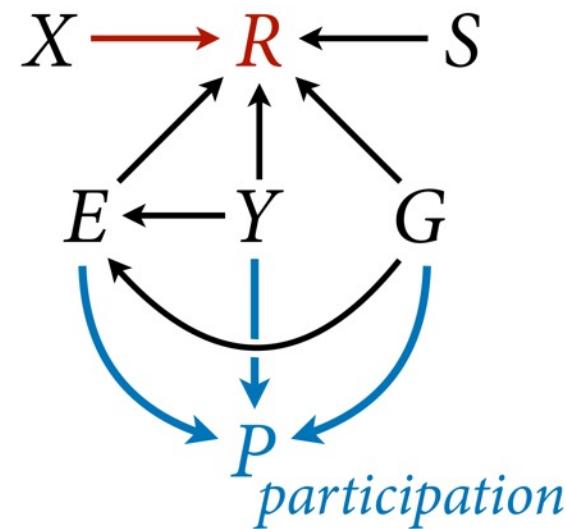
	> precis(mRXG,2)	mean	sd	5.5%	94.5%	n_eff	Rhat4
bA[1]	-0.88	0.05	-0.96	-0.80	1858	1.00	
bA[2]	-0.53	0.05	-0.61	-0.45	1724	1.00	
bI[1]	-0.90	0.05	-0.97	-0.82	2189	1.00	
bI[2]	-0.55	0.05	-0.63	-0.48	2382	1.00	
bC[1]	-1.06	0.07	-1.17	-0.95	2298	1.00	
bC[2]	-0.84	0.06	-0.94	-0.74	2000	1.00	
alpha[1]	-2.83	0.05	-2.90	-2.75	1054	1.01	
alpha[2]	-2.15	0.04	-2.21	-2.08	1104	1.00	
alpha[3]	-1.56	0.04	-1.62	-1.50	1076	1.00	
alpha[4]	-0.53	0.04	-0.59	-0.47	1080	1.00	
alpha[5]	0.14	0.04	0.09	0.20	1216	1.00	
alpha[6]	1.06	0.04	1.00	1.12	1532	1.00	



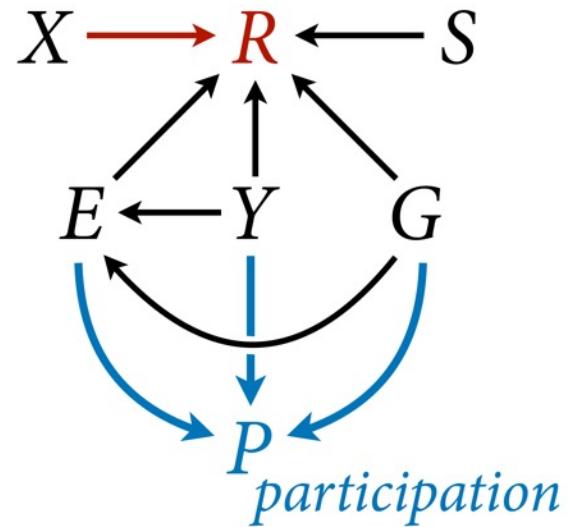
Hang on! This is a voluntary sample



Hang on! This is a **voluntary** sample



Hang on! This is a **voluntary** sample



Conditioning on P makes E, Y, G covary in sample

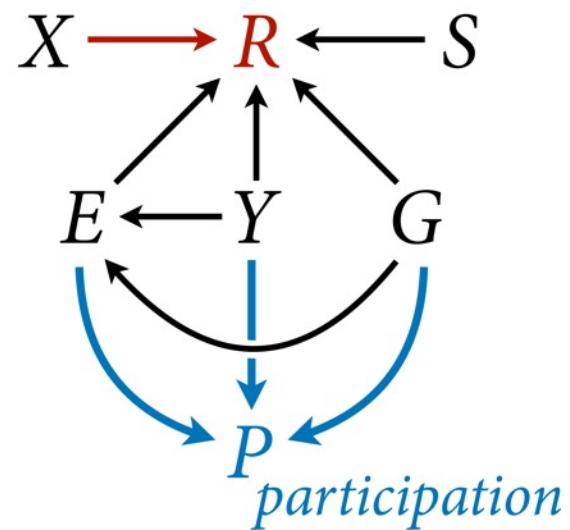
Endogenous selection

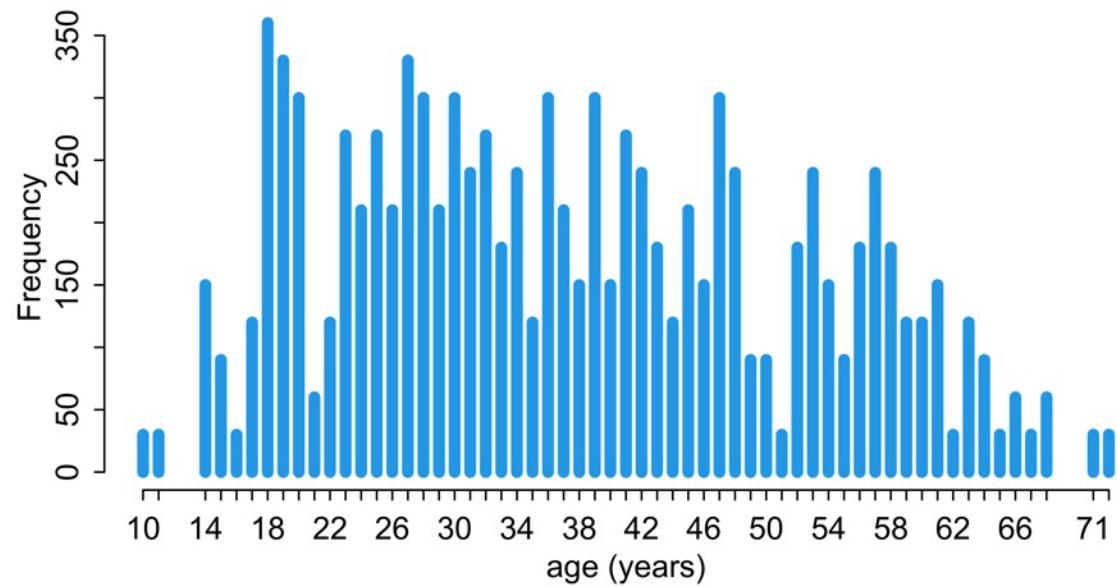
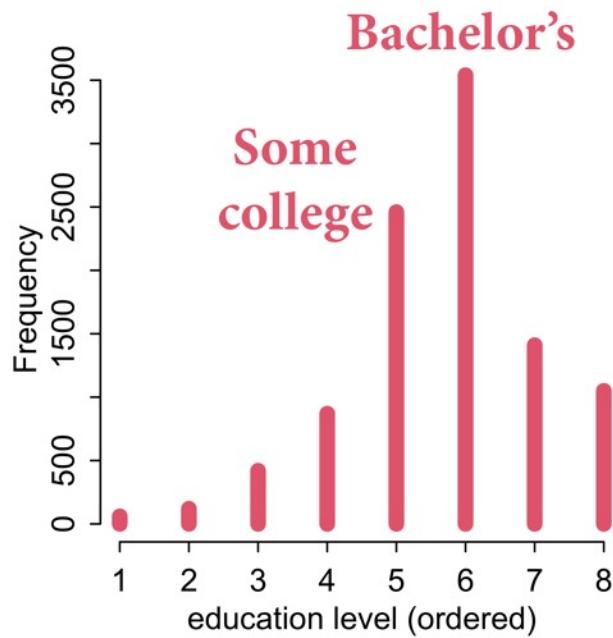
Sample is selected on a collider

Induces misleading associations among variables

Not possible here to estimate total effect of G , BUT can get direct effect

Need to stratify by E and Y and G





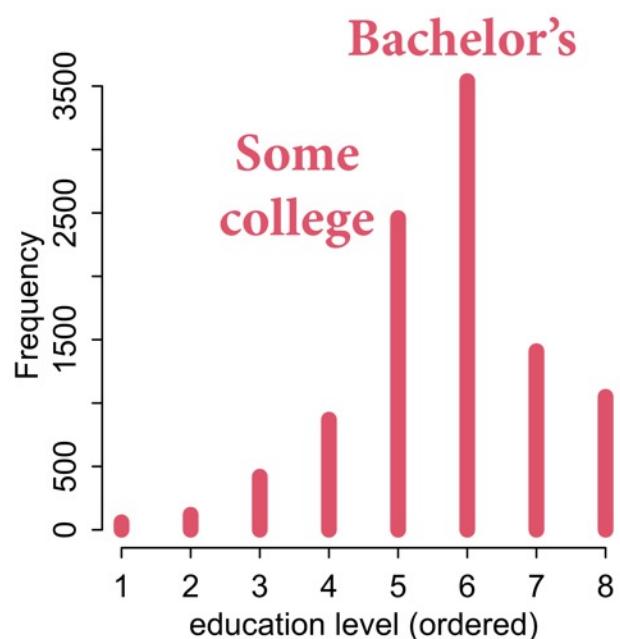
Ordered monotonic predictors

Education is an ordered category

Unlikely that each level has same effect

Want a parameter for each level

But how to enforce ordering, so that each level has larger (or smaller) effect than previous?



Ordered monotonic predictors

1 (elementary) $\phi_i = 0$

2 (middle school) $\phi_i = \delta_1$

3 (some high school) $\phi_i = \delta_1 + \delta_2$

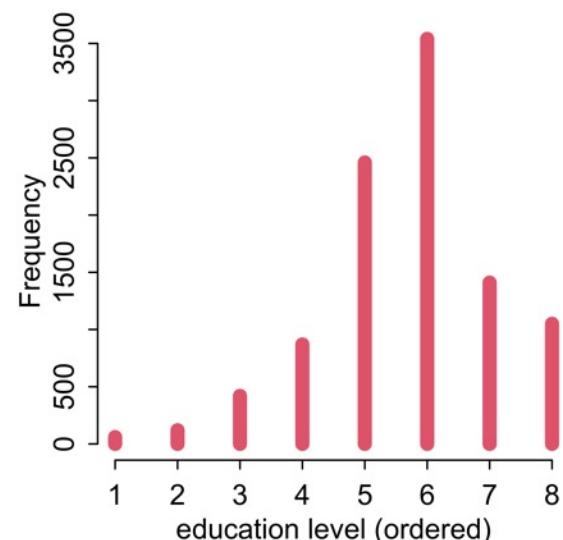
4 (high school) $\phi_i = \delta_1 + \delta_2 + \delta_3$

5 (some college) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4$

6 (college) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4 + \delta_5$

7 (master's) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4 + \delta_5 + \delta_6$

8 (doctorate) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4 + \delta_5 + \delta_6 + \delta_7$



Ordered monotonic predictors

1 (elementary) $\phi_i = 0$

2 (middle school) $\phi_i = \delta_1$

3 (some high school) $\phi_i = \delta_1 + \delta_2$

4 (high school) $\phi_i = \delta_1 + \delta_2 + \delta_3$

5 (some college) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4$

6 (college) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4 + \delta_5$

*maximum effect
of education*

7 (master's) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4 + \delta_5 + \delta_6$

8 (doctorate) $\phi_i = \delta_1 + \delta_2 + \delta_3 + \delta_4 + \delta_5 + \delta_6 + \delta_7 = \beta_E$

Ordered monotonic predictors

1 (elementary)

2 (middle school)

3 (some high school)

4 (high school)

5 (some college)

6 (college)

7 (master's)

8 (doctorate)

$$\delta_0 = 0$$

$$\sum_{j=0}^7 \delta_j = 1$$

Ordered monotonic predictors

1 (elementary)

2 (middle school)

3 (some high school)

4 (high school)

5 (some college)

6 (college)

7 (master's)

8 (doctorate)

$$\phi_i = \beta_E \sum_{j=0}^{E_i-1} \delta_j$$

education level
maximum effect *proportion of maximum effect*

Ordered monotonic *priors*

How do we set priors for the delta parameters?

delta parameters form a **simplex**

Simplex: vector that sums to 1

$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_E \sum_{j=0}^{E_i-1} \delta_j + \dots$$

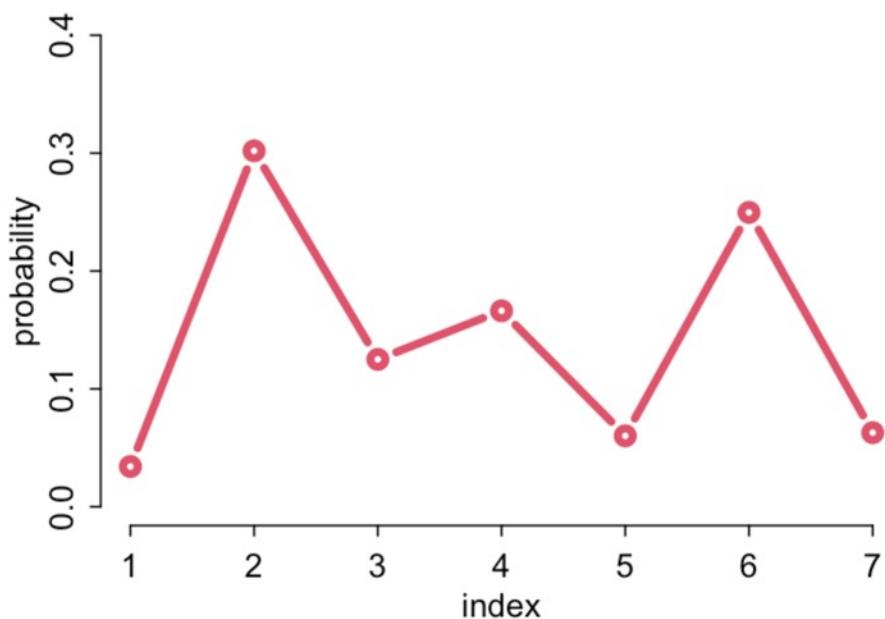
$$\alpha_j \sim \text{Normal}(0,1)$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

$$\delta_j \sim ?$$

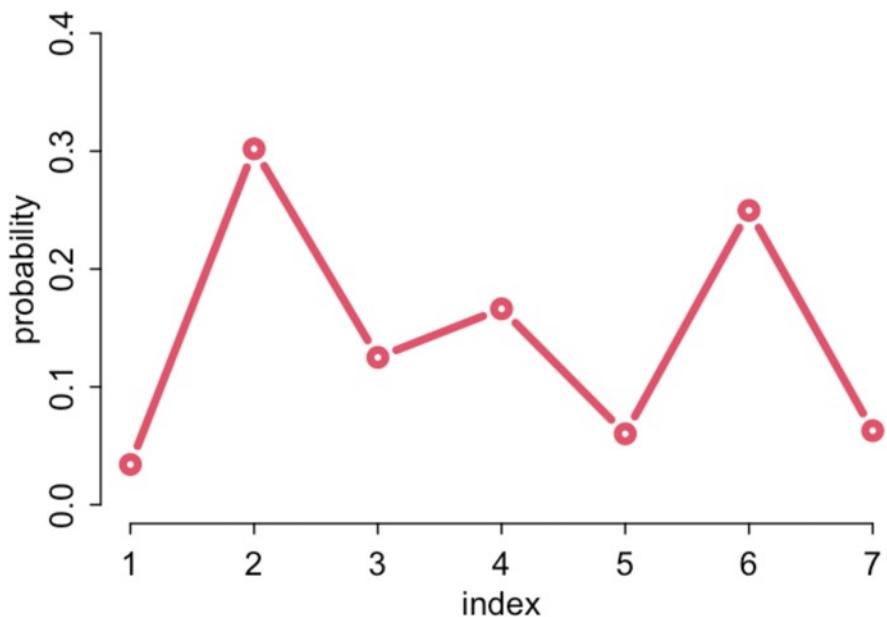
$\delta \sim \text{Dirichlet}(a)$

$a = [2,2,2,2,2,2,2]$



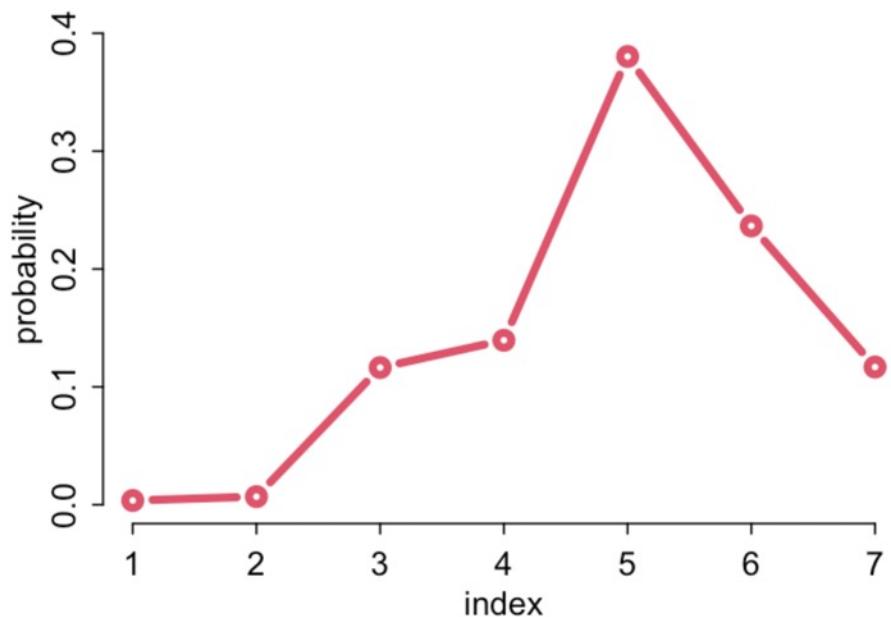
$\delta \sim \text{Dirichlet}(a)$

$a = [2,2,2,2,2,2,2]$



$\delta \sim \text{Dirichlet}(a)$

$a = [1,2,3,4,5,6,7]$



```

edu_levels <- c( 6 , 1 , 8 , 4 , 7 , 2 , 5 , 3 )
edu_new <- edu_levels[ d$edu ]

dat$E <- edu_new
dat$a <- rep(2,7) # dirichlet prior

mRXE <- ulam(
  alist(
    R ~ ordered_logistic( phi , alpha ),
    phi <- bE*sum( delta_j[1:E] ) +
      bA*A + bI*I + bC*C,
    alpha ~ normal( 0 , 1 ),
    c(bA,bI,bC,bE) ~ normal( 0 , 0.5 ),
    vector[8]: delta_j <- append_row( 0 , delta ),
    simplex[7]: delta ~ dirichlet( a )
  ), data=dat , chains=4 , cores=4 )

```

$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_E \sum_{j=0}^{E_i-1} \delta_j + \dots$$

$$\alpha_j \sim \text{Normal}(0,1)$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

$$\delta \sim \text{Dirichlet}(a)$$

```

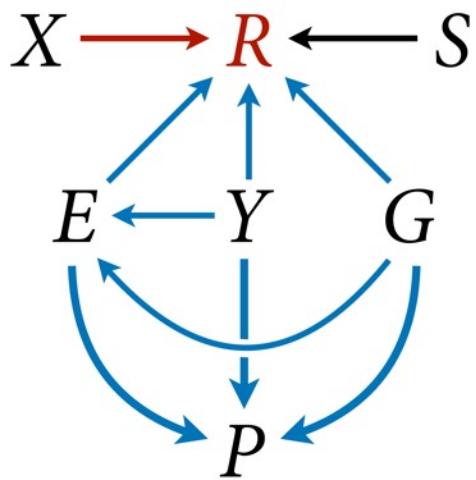
edu_levels <- c( 6 , 1 , 8 , 4 , 7 , 2 , 5 , 3 )
edu_new <- edu_levels[ d$edu ]

dat$E <- edu_new
dat$a <- rep(2,7) # dirichlet prior

mRXE <- ulam(
  alist(
    R ~ ordered_logistic( phi , alpha ),
    phi <- bE*sum( delta_j[1:E] ) +
      bA*A + bI*I + bC*C,
    alpha ~ normal( 0 , 1 ),
    c(bA,bI,bC,bE) ~ normal( 0 , 0.5 ),
    vector[8]: delta_j <- append_row( 0 ,
      simplex[7]: delta ~ dirichlet( a )
  ), data=dat , chains=4 , cores=4 )

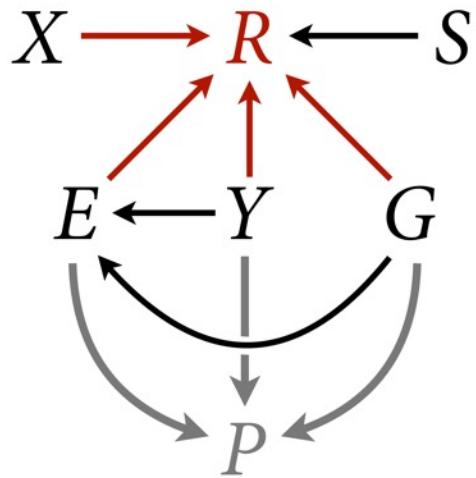
```

	> precis(mRXE,2)	mean	sd	5.5%	94.5%	n_eff	Rhat4
alpha[1]	-3.07	0.14	-3.32	-2.86	793	1	
alpha[2]	-2.39	0.14	-2.63	-2.17	804	1	
alpha[3]	-1.81	0.14	-2.05	-1.60	811	1	
alpha[4]	-0.79	0.14	-1.03	-0.57	799	1	
alpha[5]	-0.12	0.14	-0.36	0.10	804	1	
alpha[6]	0.79	0.14	0.54	1.00	831	1	
bE	-0.31	0.16	-0.57	-0.06	838	1	
bC	-0.96	0.05	-1.04	-0.88	1757	1	
bI	-0.72	0.04	-0.77	-0.66	1982	1	
bA	-0.70	0.04	-0.77	-0.64	1779	1	
delta[1]	0.22	0.13	0.05	0.47	1227	1	
delta[2]	0.14	0.09	0.03	0.31	2258	1	
delta[3]	0.20	0.11	0.05	0.38	2256	1	
delta[4]	0.17	0.09	0.04	0.34	1926	1	
delta[5]	0.04	0.05	0.01	0.12	945	1	
delta[6]	0.10	0.07	0.02	0.23	1870	1	
delta[7]	0.13	0.08	0.03	0.27	2335	1	



bE not interpretable

	mean	sd	5.5%	94.5%	n_eff	Rhat4
alpha[1]	-3.07	0.14	-3.32	-2.86	793	1
alpha[2]	-2.39	0.14	-2.63	-2.17	804	1
alpha[3]	-1.81	0.14	-2.05	-1.60	811	1
alpha[4]	-0.79	0.14	-1.03	-0.57	799	1
alpha[5]	-0.12	0.14	-0.36	0.10	804	1
alpha[6]	0.79	0.14	0.54	1.00	831	1
bE	-0.31	0.16	-0.57	-0.06	838	1
bC	-0.96	0.05	-1.04	-0.88	1757	1
bI	-0.72	0.04	-0.77	-0.66	1982	1
bA	-0.70	0.04	-0.77	-0.64	1779	1
delta[1]	0.22	0.13	0.05	0.47	1227	1
delta[2]	0.14	0.09	0.03	0.31	2258	1
delta[3]	0.20	0.11	0.05	0.38	2256	1
delta[4]	0.17	0.09	0.04	0.34	1926	1
delta[5]	0.04	0.05	0.01	0.12	945	1
delta[6]	0.10	0.07	0.02	0.23	1870	1
delta[7]	0.13	0.08	0.03	0.27	2335	1



$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\begin{aligned}\phi_i &= \beta_{E,G[i]} \sum_{j=0}^{E_i-1} \delta_j + \\ &\quad \beta_{A,G[i]} A_i + \beta_{I,G[i]} I_i + \beta_{C,G[i]} C_i + \\ &\quad \beta_{Y,G[i]} Y_i\end{aligned}$$

$$\alpha_j \sim \text{Normal}(0,1)$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

$$\delta \sim \text{Dirichlet}(a)$$

```

dat$Y <- standardize(d$age)

mRXEYGt <- ulam(
  alist(
    R ~ ordered_logistic( phi , alpha ),
    phi <- bE[G]*sum( delta_j[1:E] ) +
      bA[G]*A + bI[G]*I + bC[G]*C +
      bY[G]*Y,
    alpha ~ normal( 0 , 1 ),
    bA[G] ~ normal( 0 , 0.5 ),
    bI[G] ~ normal( 0 , 0.5 ),
    bC[G] ~ normal( 0 , 0.5 ),
    bE[G] ~ normal( 0 , 0.5 ),
    bY[G] ~ normal( 0 , 0.5 ),
    vector[8]: delta_j <- append_row( 0 , delta ),
    simplex[7]: delta ~ dirichlet( a )
  ), data=dat , chains=4 , cores=4 , threads=2 )

```

$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_{E,G[i]} \sum_{j=0}^{E_i-1} \delta_j +$$

$$\beta_{A,G[i],i} A_i + \beta_{I,G[i]} I_i + \beta_{C,G[i]} C_i +$$

$$\beta_{Y,G[i]} Y_i$$

$$\alpha_j \sim \text{Normal}(0,1)$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

$$\delta \sim \text{Dirichlet}(a)$$

```

dat$Y <- standardize(d$age)

mRXEYGt <- ulam(
  alist(
    R ~ ordered_logistic( phi , alpha ),
    phi <- bE[G]*sum( delta_j[1:E] ) +
      bA[G]*A + bI[G]*I + bC[G]*C +
      bY[G]*Y,
    alpha ~ normal( 0 , 1 ),
    bA[G] ~ normal( 0 , 0.5 ),
    bI[G] ~ normal( 0 , 0.5 ),
    bC[G] ~ normal( 0 , 0.5 ),
    bE[G] ~ normal( 0 , 0.5 ),
    bY[G] ~ normal( 0 , 0.5 ),
    vector[8]: delta_j <- append_row( 0 , delta ),
    simplex[7]: delta ~ dirichlet( a )
  ), data=dat , chains=4 , cores=4 , threads=2 )

```

4 chains times 2 threads each = 8 cores

$$R_i \sim \text{OrderedLogit}(\phi_i, \alpha)$$

$$\phi_i = \beta_{E,G[i]} \sum_{j=0}^{E_i-1} \delta_j +$$

$$\beta_{A,G[i],i} A_i + \beta_{I,G[i]} I_i + \beta_{C,G[i]} C_i +$$

$$\beta_{Y,G[i]} Y_i$$

$$\alpha_j \sim \text{Normal}(0,1)$$

$$\beta_- \sim \text{Normal}(0,0.5)$$

$$\delta \sim \text{Dirichlet}(a)$$

```

dat$Y <- standardize(d$age)

mRXEYGt <- ulam(
  alist(
    R ~ ordered_logistic( phi , alpha ),
    phi <- bE[G]*sum( delta_j[1:E] ) +
      bA[G]*A + bI[G]*I + bC[G]*C +
      bY[G]*Y,
    alpha ~ normal( 0 , 1 ),
    bA[G] ~ normal( 0 , 0.5 ),
    bI[G] ~ normal( 0 , 0.5 ),
    bC[G] ~ normal( 0 , 0.5 ),
    bE[G] ~ normal( 0 , 0.5 ),
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    vector[8]: delta_j <- append_row( 0 , delta ),
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  ), data=dat , chains=4 , cores=4 , threads=2 )

```

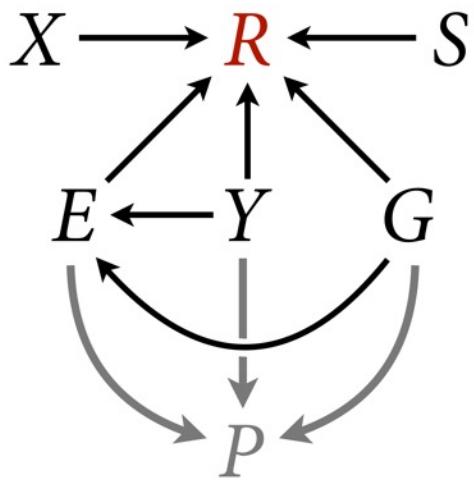
4 chains times 2 threads each = 8 cores

1 thread each

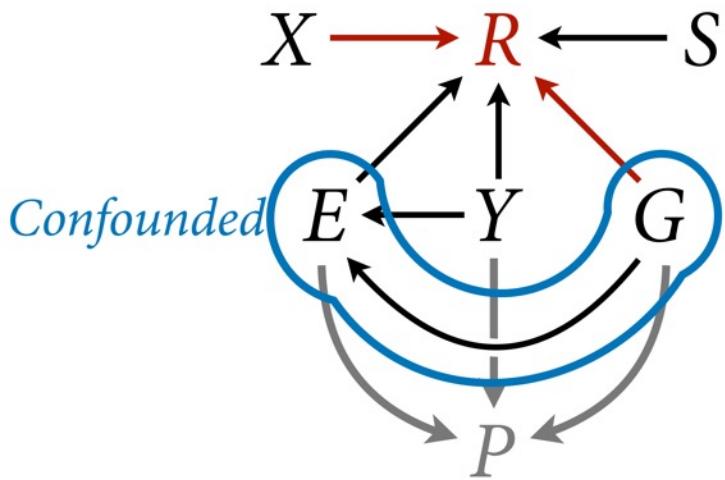
Sampling durations (minutes):			
	warmup	sample	total
chain:1	6.53	3.99	10.52
chain:2	7.33	2.66	9.99
chain:3	6.88	3.70	10.58
chain:4	6.40	2.63	9.03

2 threads each

Sampling durations (minutes):			
	warmup	sample	total
chain:1	4.41	1.80	6.21
chain:2	4.69	1.87	6.56
chain:3	5.14	1.56	6.70
chain:4	4.21	1.84	6.05

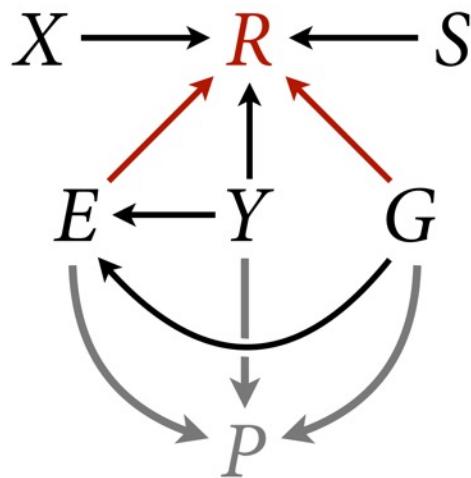


	mean	sd	5.5%	94.5%	n_eff	Rhat4
alpha[1]	-2.89	0.10	-3.06	-2.73	729	1
alpha[2]	-2.21	0.10	-2.37	-2.06	728	1
alpha[3]	-1.62	0.10	-1.78	-1.47	724	1
alpha[4]	-0.58	0.10	-0.74	-0.43	729	1
alpha[5]	0.11	0.10	-0.05	0.26	726	1
alpha[6]	1.03	0.10	0.87	1.18	746	1
bA[1]	-0.56	0.06	-0.65	-0.47	1932	1
bA[2]	-0.81	0.05	-0.90	-0.73	2013	1
bI[1]	-0.66	0.05	-0.74	-0.58	2539	1
bI[2]	-0.76	0.05	-0.84	-0.68	2283	1
bC[1]	-0.77	0.07	-0.88	-0.65	2029	1
bC[2]	-1.09	0.07	-1.20	-0.99	2012	1
bE[1]	-0.63	0.14	-0.85	-0.42	810	1
bE[2]	0.41	0.14	0.19	0.62	795	1
bY[1]	0.00	0.03	-0.05	0.05	2740	1
bY[2]	-0.13	0.03	-0.18	-0.09	1426	1
delta[1]	0.15	0.08	0.04	0.31	1759	1
delta[2]	0.15	0.09	0.04	0.30	2440	1
delta[3]	0.29	0.11	0.11	0.46	2001	1
delta[4]	0.08	0.05	0.02	0.17	2414	1
delta[5]	0.06	0.04	0.01	0.14	1087	1
delta[6]	0.24	0.07	0.13	0.34	2301	1
delta[7]	0.04	0.02	0.01	0.08	2755	1

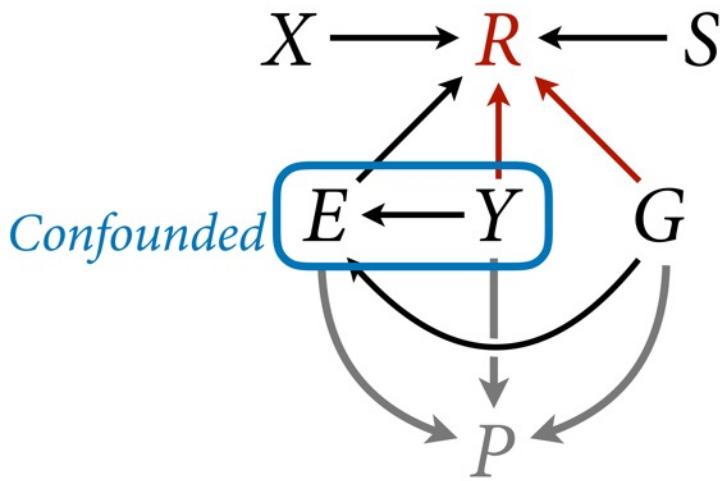


*Only
direct
effect G*

```
> precis(mRXYGt,2)
      mean   sd  5.5% 94.5% n_eff Rhat4
alpha[1] -2.89 0.10 -3.06 -2.73    729     1
alpha[2] -2.21 0.10 -2.37 -2.06    728     1
alpha[3] -1.62 0.10 -1.78 -1.47    724     1
alpha[4] -0.58 0.10 -0.74 -0.43    729     1
alpha[5]  0.11 0.10 -0.05  0.26    726     1
alpha[6]  1.03 0.10  0.87  1.18    746     1
bA[1]   -0.56 0.06 -0.65 -0.47   1932     1
bA[2]   -0.81 0.05 -0.90 -0.73  2013     1
bI[1]   -0.66 0.05 -0.74 -0.58  2539     1
bI[2]   -0.76 0.05 -0.84 -0.68  2283     1
bC[1]   -0.77 0.07 -0.88 -0.65  2029     1
bC[2]   -1.09 0.07 -1.20 -0.99  2012     1
bE[1]   -0.63 0.14 -0.85 -0.42   810     1
bE[2]    0.41 0.14  0.19  0.62   795     1
bY[1]    0.00 0.03 -0.05  0.05  2740     1
bY[2]   -0.13 0.03 -0.18 -0.09  1426     1
delta[1]  0.15 0.08  0.04  0.31  1759     1
delta[2]  0.15 0.09  0.04  0.30  2440     1
delta[3]  0.29 0.11  0.11  0.46  2001     1
delta[4]  0.08 0.05  0.02  0.17  2414     1
delta[5]  0.06 0.04  0.01  0.14  1087     1
delta[6]  0.24 0.07  0.13  0.34  2301     1
delta[7]  0.04 0.02  0.01  0.08  2755     1
```

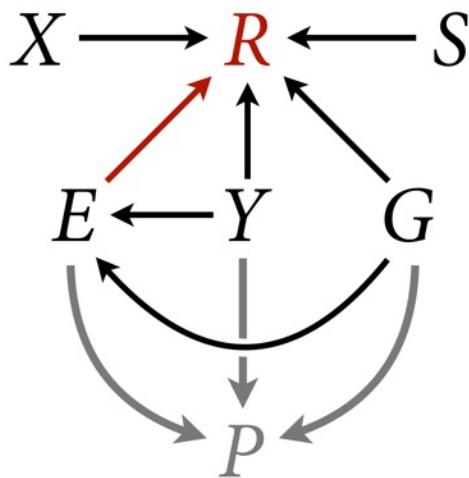


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*Only
direct
effect*

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bC[2]	-1.09	0.07	-1.20	-0.99	2012	1
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Model mRXEYG2t stratifies by G, in Lecture 11 script

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bC[1]	-0.77	0.07	-0.88	-0.65	2029	1
bC[2]	-1.09	0.07	-1.20	-0.99	2012	1
bE[1]	-0.63	0.14	-0.85	-0.42	810	1
bE[2]	0.41	0.14	0.19	0.62	795	1
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delta[5]	0.06	0.04	0.01	0.14	1087	1
delta[6]	0.24	0.07	0.13	0.34	2301	1
delta[7]	0.04	0.02	0.01	0.08	2755	1

Complex causal effects

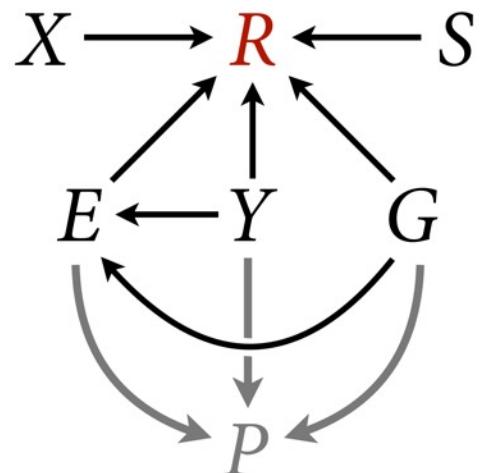
Causal effects (predicted consequences of intervention) require **marginalization**

Example: Causal effect of E requires distribution of Y and G to average over

Problem 1: Should not marginalize over **this** sample—**cursed P !** Post-stratify to new target.

Problem 2: Should not set all Y to same E

Example: Causal effect of Y requires effect of Y on E , which we cannot estimate (P again!)



Complex causal effects

Causal effects (predicted consequences of intervention)

Example of Y and

Problem sample

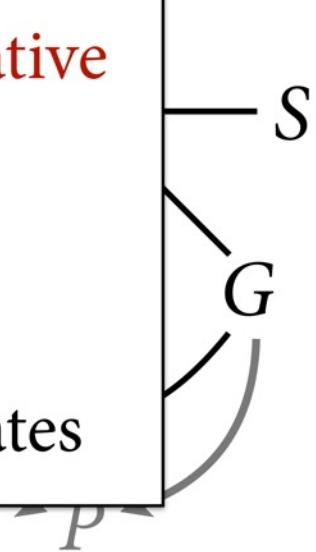
Problem

No matter how complex, still just a **generative simulation using posterior samples**

Need generative model to plan estimation

Need generative model to compute estimates

Example: Causal effect of Y requires effect of Y on E , which we cannot estimate (P again!)

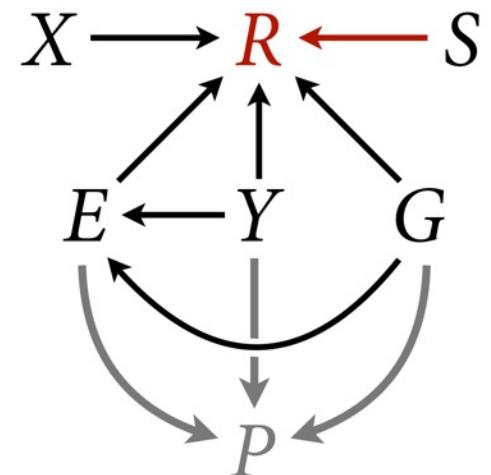


Repeat observations

30 stories (S)

```
> table(d$story)
```

aqu	boa	box	bur	car	che	pon	rub	sha	shi	spe	swi
662	662	1324	1324	662	662	662	662	662	662	993	993



Repeat observations

30 stories (S)

```
> table(d$story)
```

aqu	boa	box	bur	car	che	pon	rub	sha	shi	spe	swi
662	662	1324	1324	662	662	662	662	662	662	993	993

331 individuals (U)

```
> table(d$id)
```

96;434 96;445 96;451 96;456 96;458 96;466 96;467 96;474 96;480 96;481 96;497
30 30 30 30 30 30 30 30 30 30 30 30
96;498 96;502 96;505 96;511 96;512 96;518 96;519 96;531 96;533 96;538 96;547
30 30 30 30 30 30 30 30 30 30 30 30
96;550 96;553 96;555 96;558 96;560 96;562 96;566 96;570 96;581 96;586 96;591
30 30 30 30 30 30 30 30 30 30 30 30

