

Interaction analysis

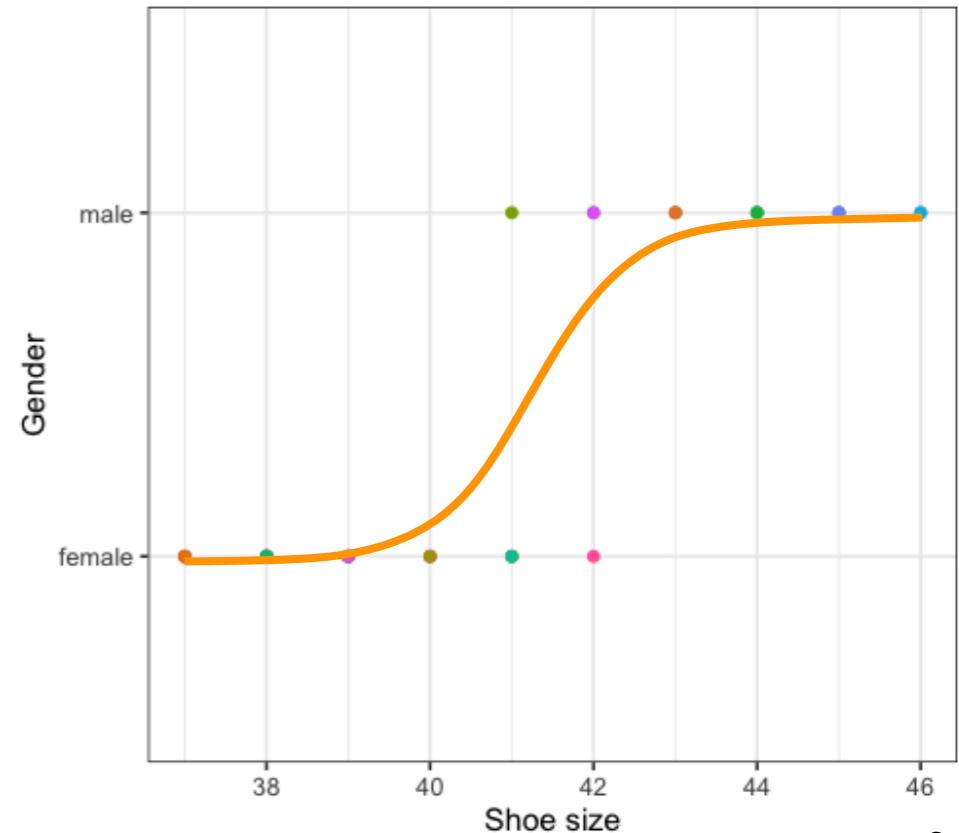
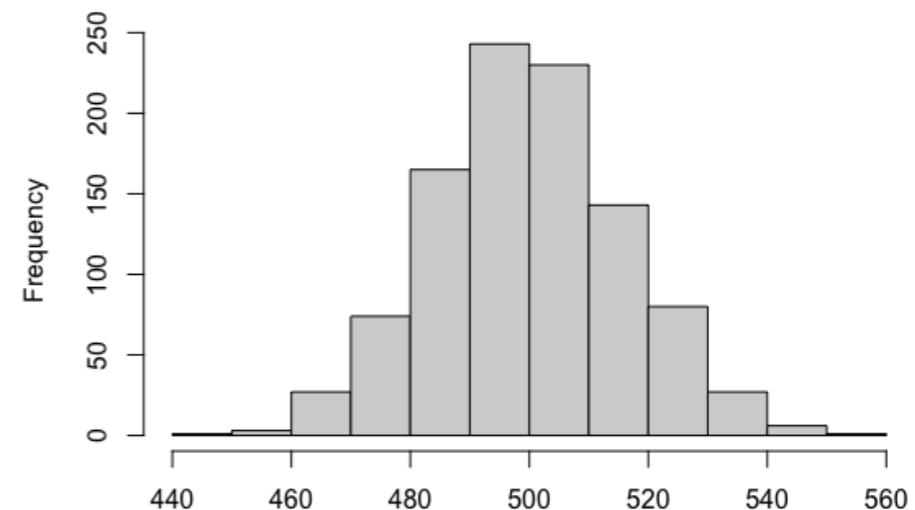
Methods 1, E2021 - Lecture 11
Tuesday 23/11/2021
Fabio Trecca

QUIZ
TIME



Recap: Bernoulli trials and the binomial distribution

- $p + q = 1$
- $p = 1 - q$
- $q = 1 - p$
- $odds = \frac{p}{1 - p}$
- $P_x = \binom{n}{k} p^x q^{n-x}$
- conditional probability = $P(Y_i | X_i)$



Recap: Linear vs logistic regression

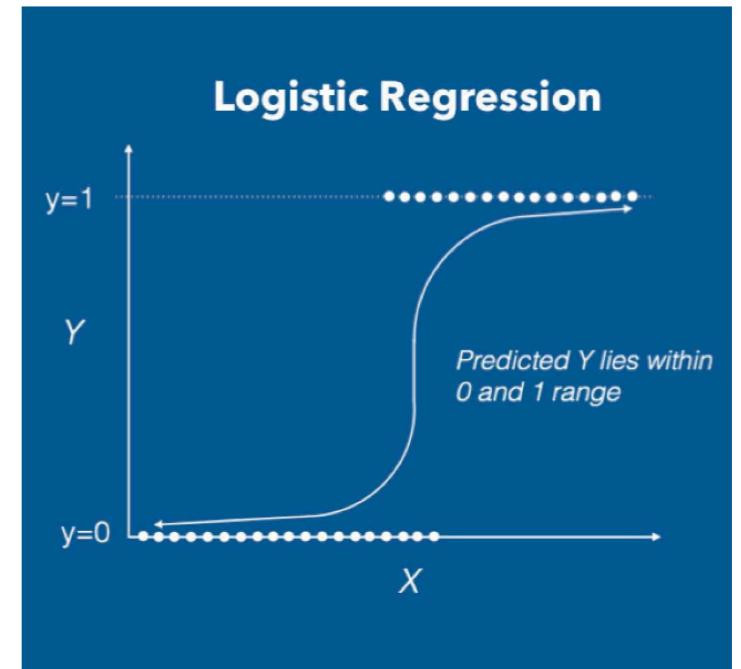
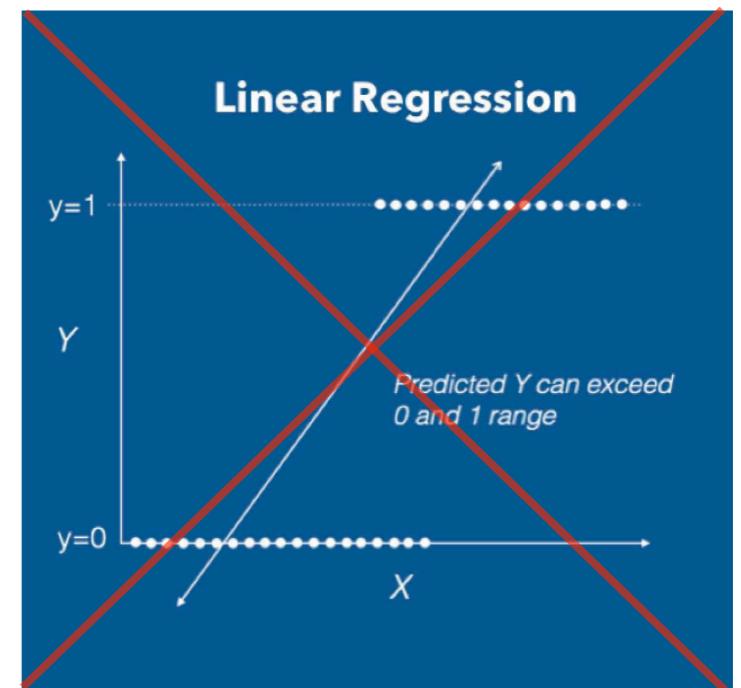
- **Linear regression:**

- gives the predicted mean value of an outcome variable at a particular value of a predictor variable

- **Logistic regression:**

- gives the conditional probability that an outcome variable equals one at a particular value of a predictor variable

- conditional probability = $P(Y_i | X_i)$



Recap: Logistic regression equation

- $P(Y_i) = \frac{e^{(\beta_0 + \beta_1 X_i + \varepsilon_i)}}{1 + e^{(\beta_0 + \beta_1 X_i + \varepsilon_i)}}$
- where $\beta_0 + \beta_1 X_i + \varepsilon_i$ = simple linear regression

Recap: Interpreting the output of logistic regression (2)

```
> summary(glm(gender_N ~ shoesize, family = binomial(link = logit),  
data))
```

Call:
glm(formula = gender_N ~ shoesize, family = binomial(link = logit),
data = data)

Deviance Residuals:

Min	10	Median	30	Max
-1.66609	-0.05431	-0.01526	0.01807	1.81759

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-105.570	49.470	-2.134	0.0328 *
shoesize	2.540	1.192	2.131	0.0331 *

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 55.0432 on 43 degrees of freedom
Residual deviance: 8.8056 on 42 degrees of freedom
AIC: 12.806

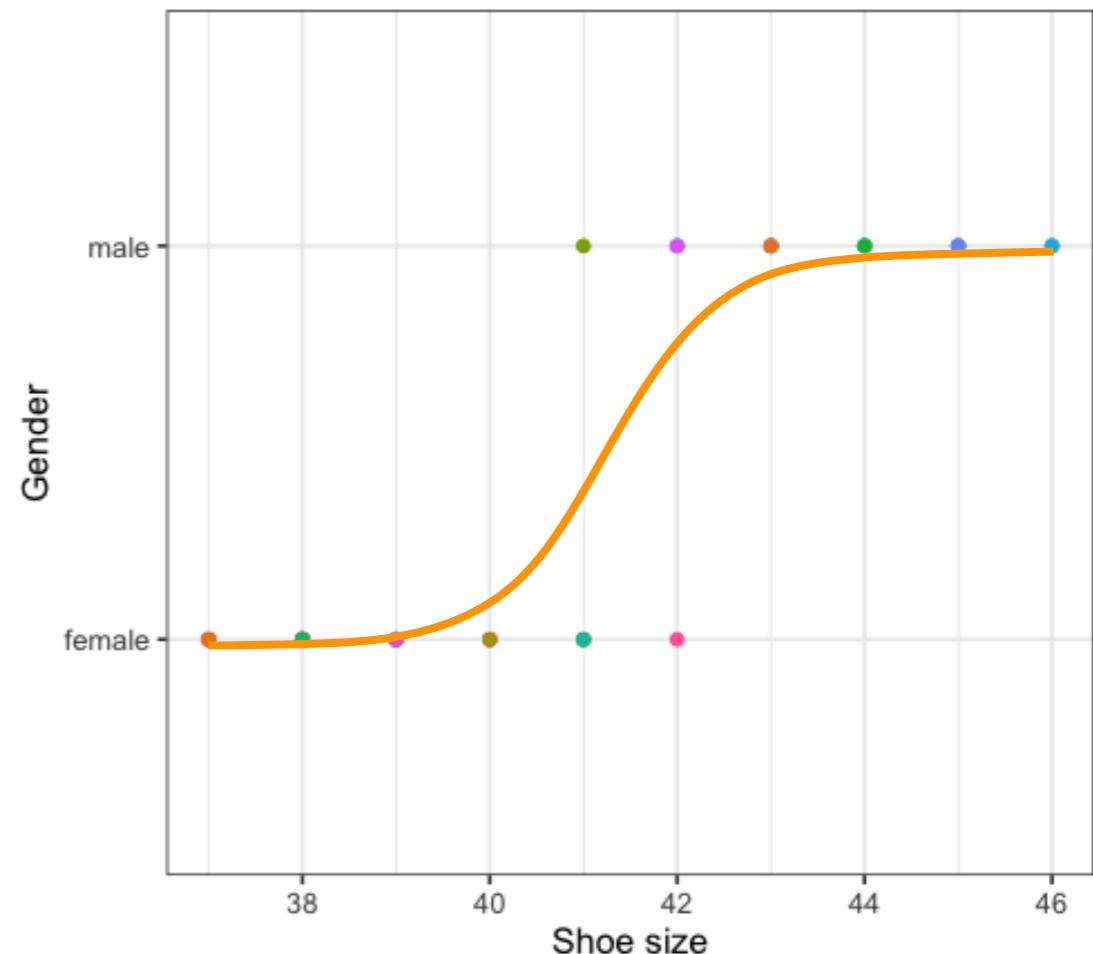
Number of Fisher Scoring iterations: 8

```
> boot::inv.logit(-105.57+(36*2.54))  
[1] 0
```

```
> boot::inv.logit(-105.57+(41*2.54))  
[1] 0.1930987
```

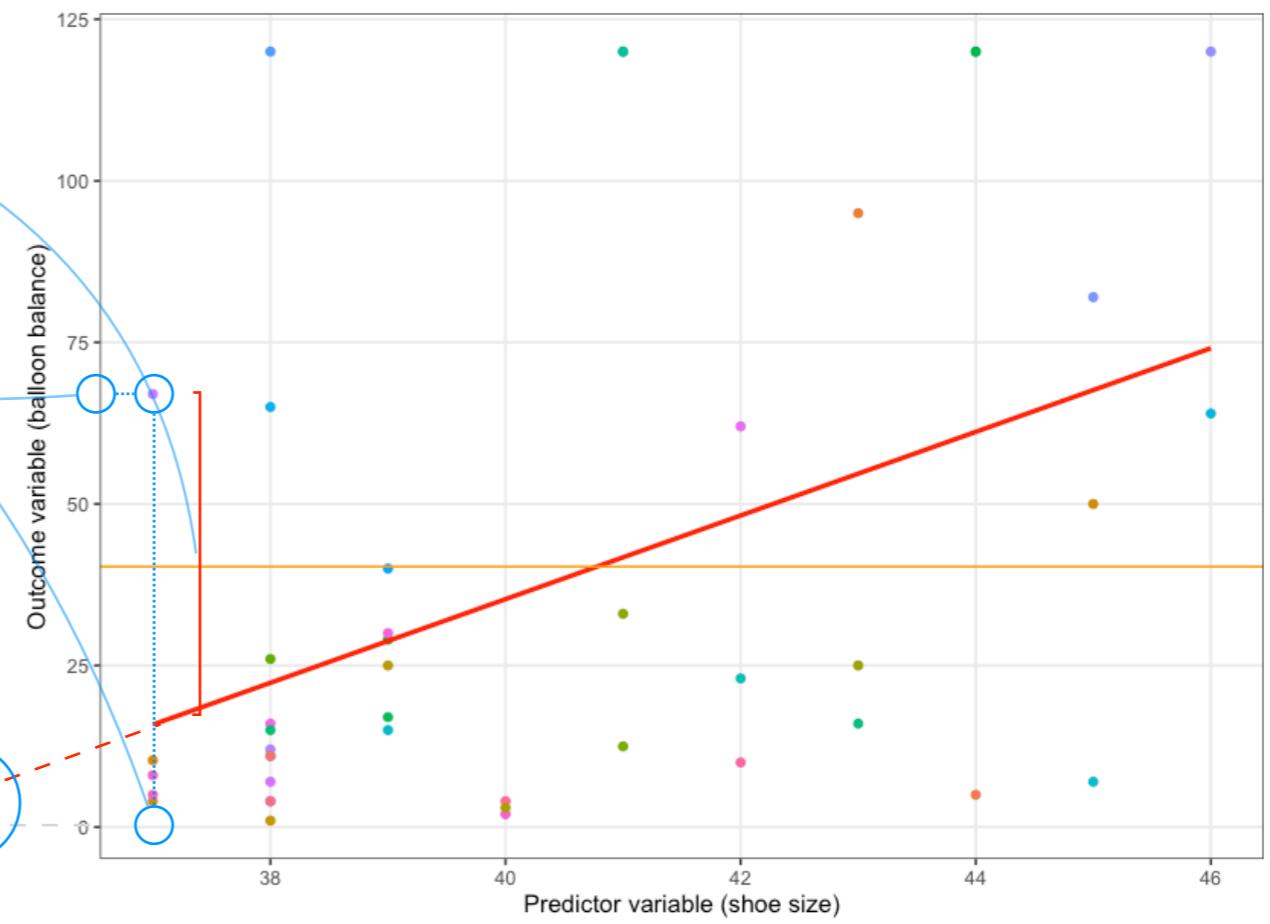
```
> boot::inv.logit(-105.57+(43*2.54))  
[1] 0.9746673
```

```
> boot::inv.logit(-105.57+(46*2.54))  
[1] 0.9999873
```



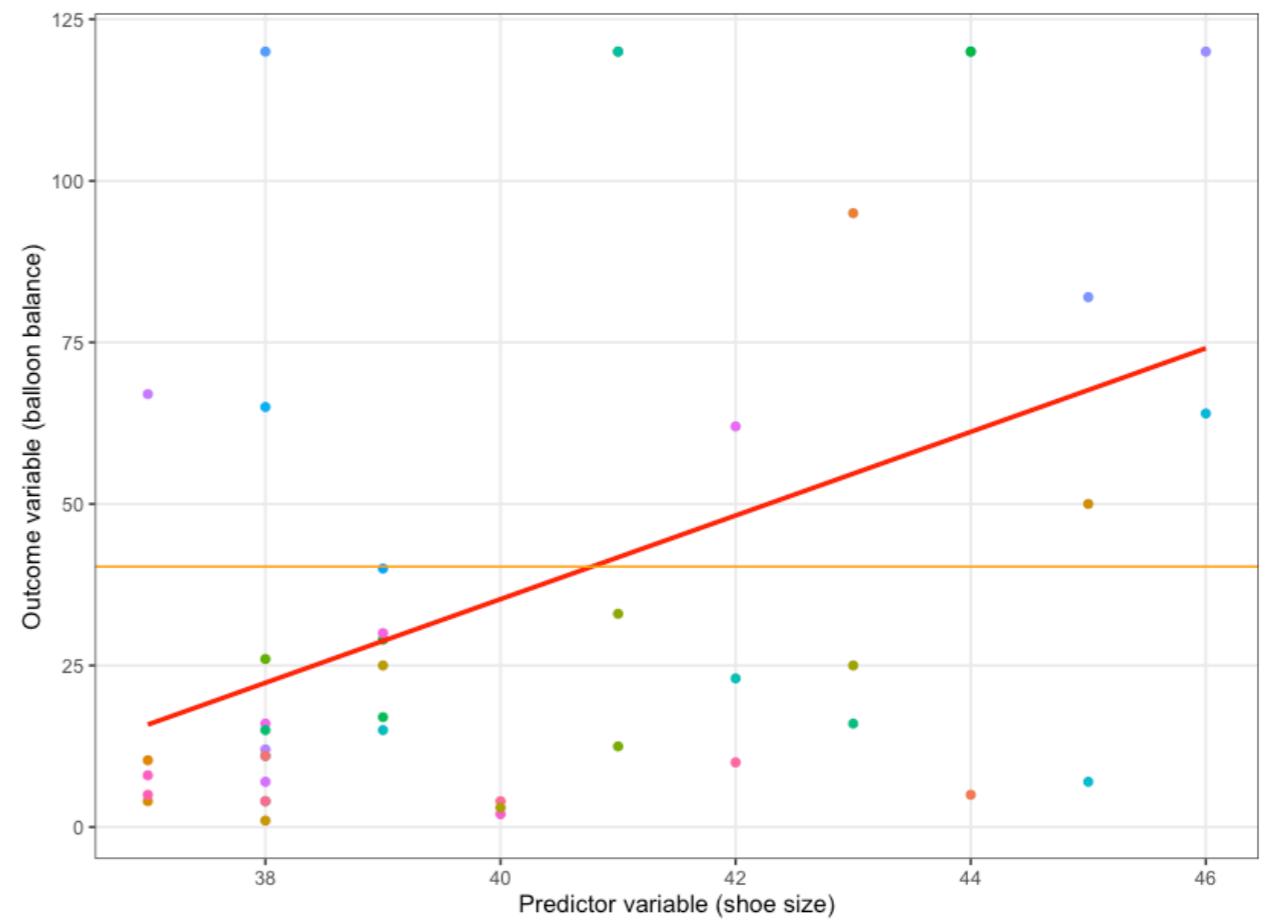
Recap: Interpreting the linear regression formula (1)

- $$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$



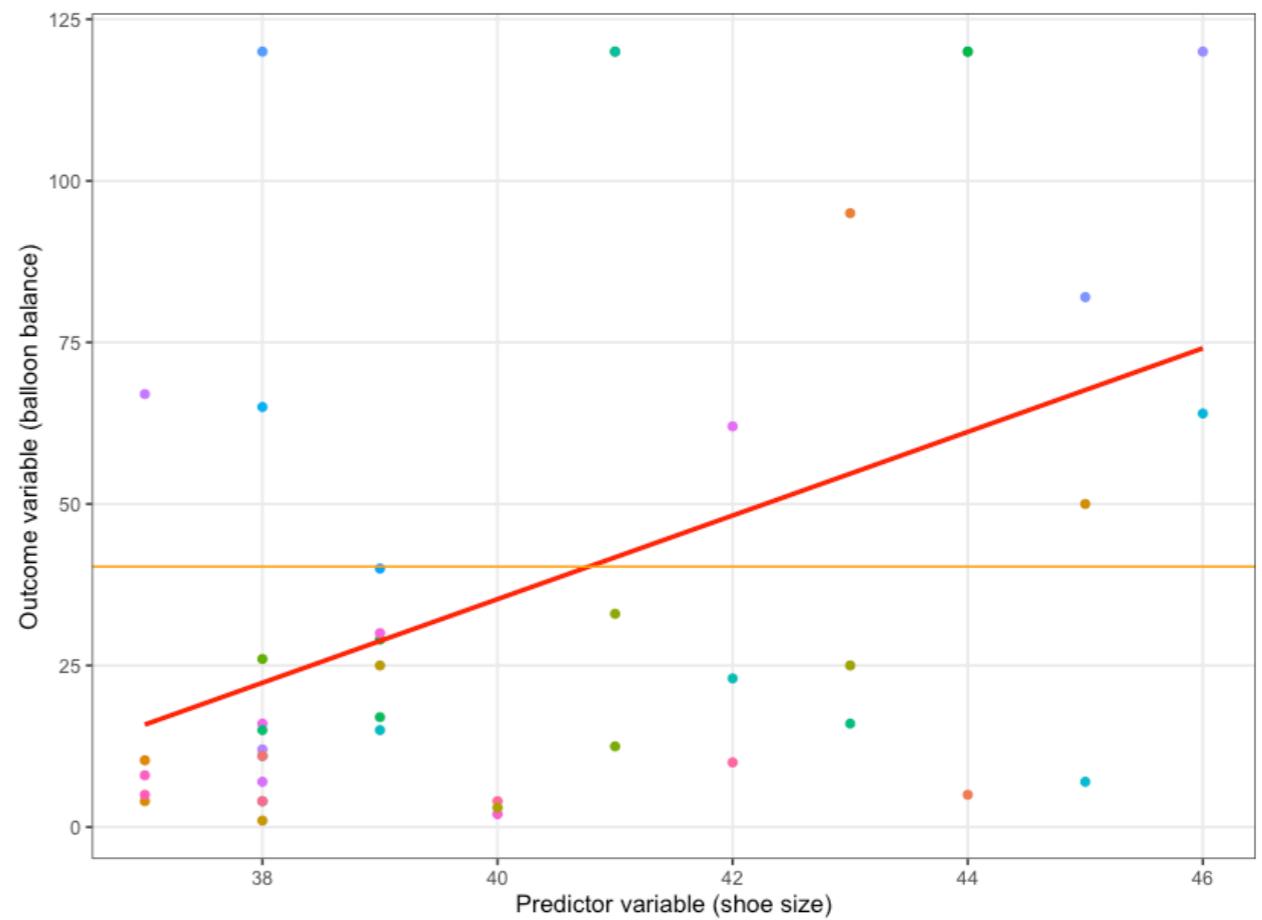
Recap: Interpreting the linear regression formula (2)

- $Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i =$
- $= -223.612 + 6.472X_i + \varepsilon_i$



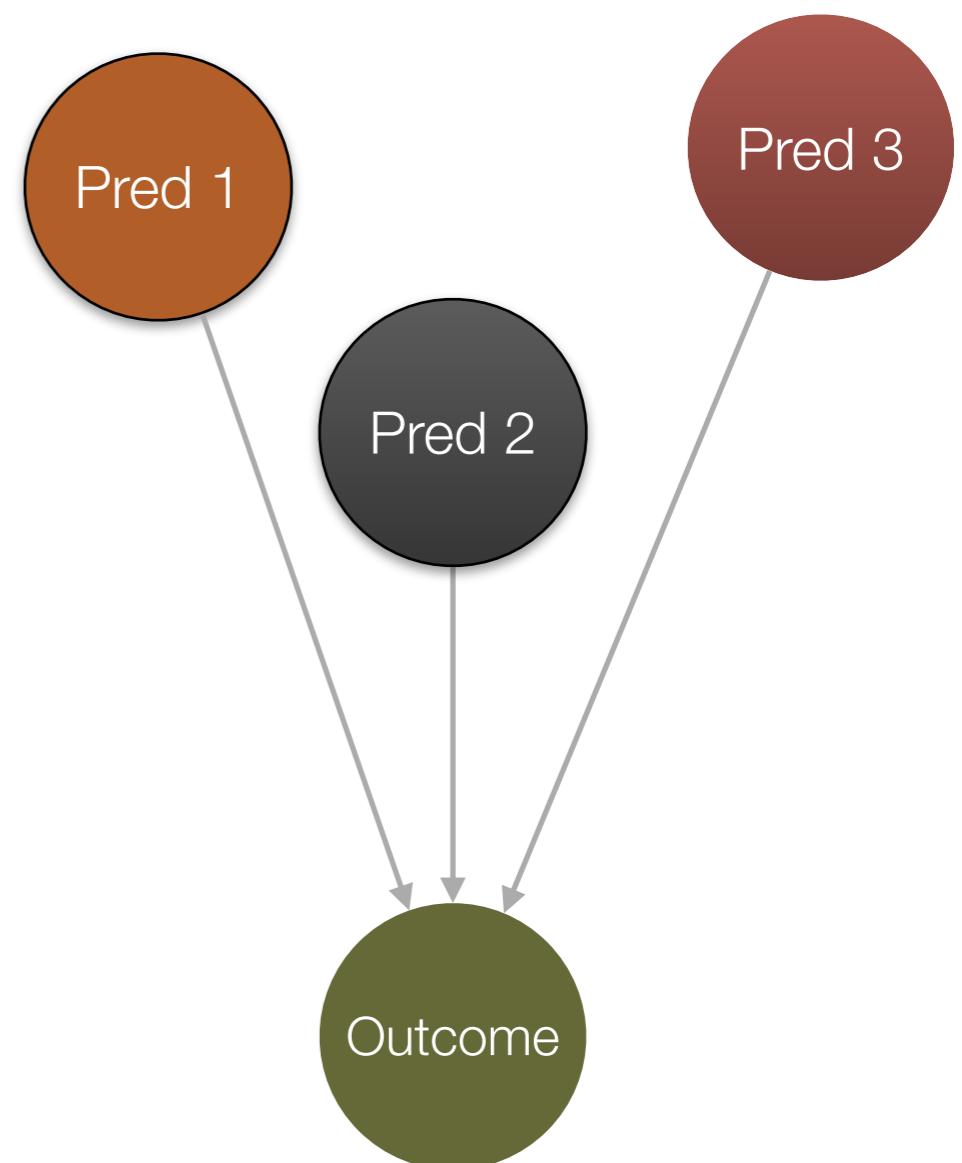
Recap: Interpreting the linear regression formula (3)

- $Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i =$
- $= -223.612 + 6.472 X_i + \varepsilon_i$
- $X_{Fabio} = 44$
- $Y_{Fabio} = -223.612 + 6.472 \times 44 + \varepsilon_{Fabio}$
- $Y_{Fabio} = 61.146$



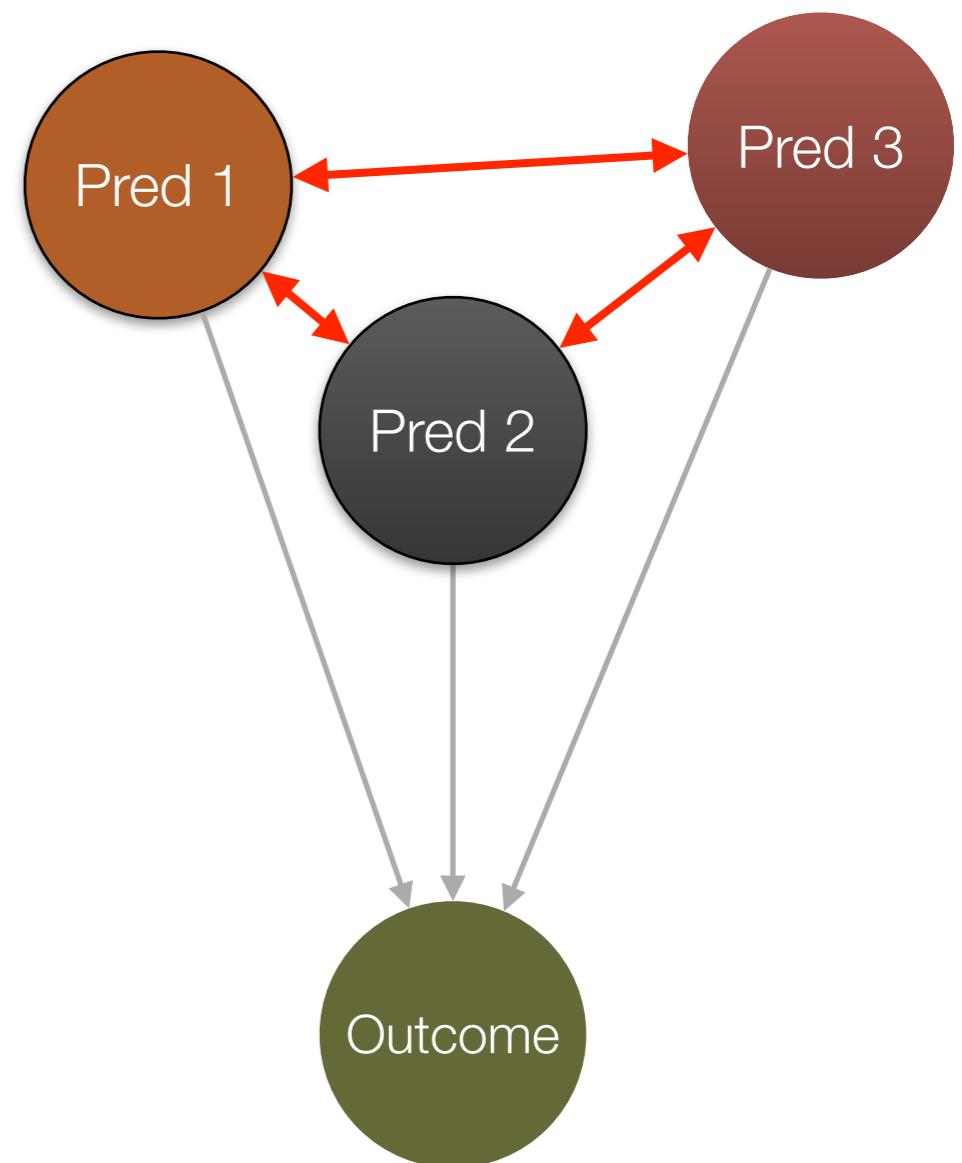
Interaction analysis

- Until now we have only looked at the effect of one predictor at the time
- Interaction analysis allows us to see how 2+ predictor variables influence the outcome variable **together**



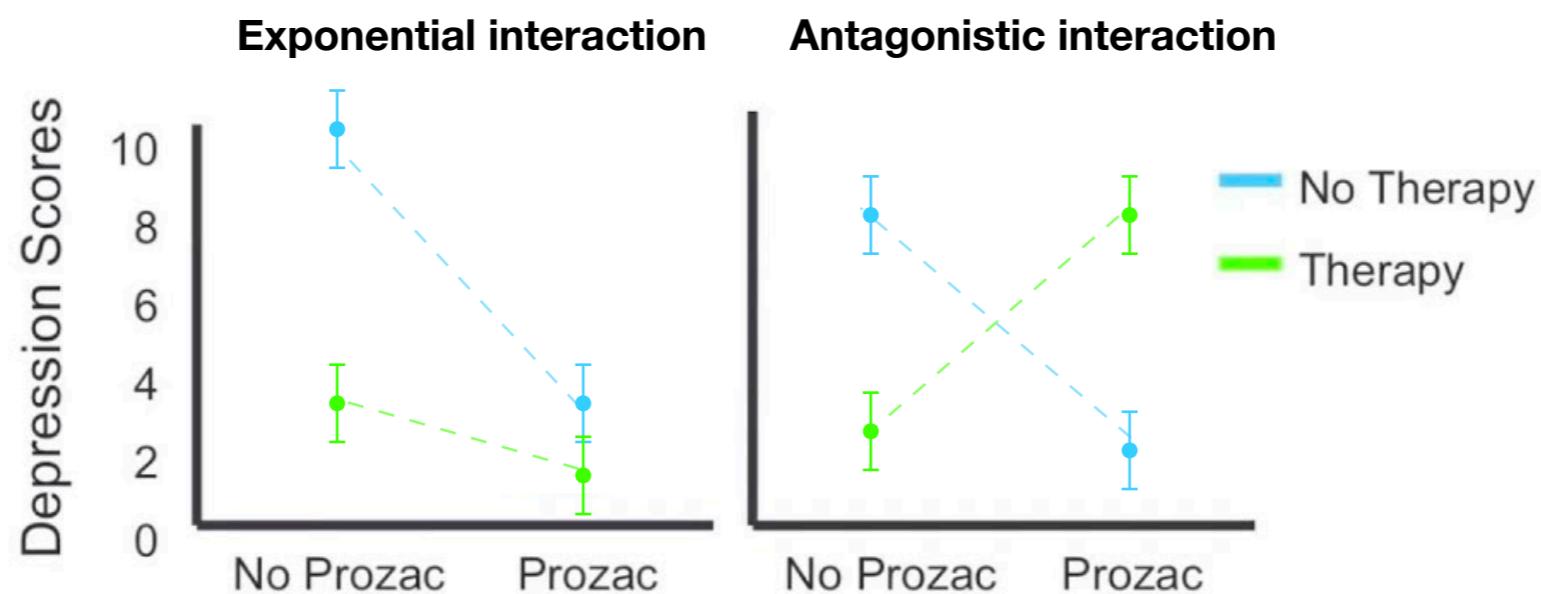
Interaction analysis

- Until now we have only looked at the effect of one predictor at the time
- Interaction analysis allows us to see how 2+ predictor variables influence the outcome variable **together**



Main effects vs Interaction effects

- Main effect: the effect of a predictor on an outcome variable
- Interaction effect: the effect of one predictor on an outcome variable as a function of another predictor variable



Interaction effects in linear regression

- $Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3(x_{1i} \times x_{2i}) + \varepsilon_i$

Interaction effects in linear regression

- $$Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 (x_{1i} \times x_{2i}) + \varepsilon_i$$
 intercept

Interaction effects in linear regression

- $$Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 (x_{1i} \times x_{2i}) + \varepsilon_i$$

intercept slope for
the first
predicto
r

Interaction effects in linear regression

- $$Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 (x_{1i} \times x_{2i}) + \varepsilon_i$$

intercept slope for slope for
the first the second
predictor predictor
r

Interaction effects in linear regression

- $$Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 (x_{1i} \times x_{2i}) + \varepsilon_i$$

intercept slope for slope for slope for the
 the first the second multiplicative effect
 predictor predictor between the two
 predictors

Three types of two-way interactions

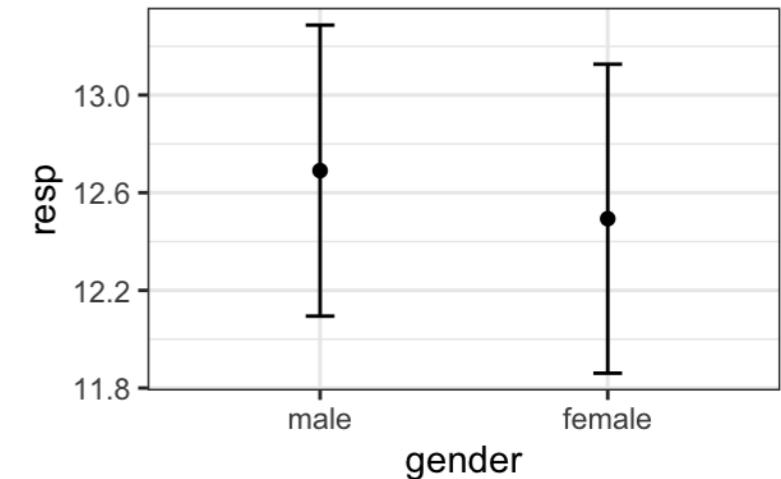
- Two continuous predictors
- Two categorical predictors
- One continuous and one categorical predictor
- High-order interactions: three-way interactions and above can be very hard to interpret and are generally avoided

Two categorical predictors (1)

- Does the effect of one categorical predictor depend on the levels of another categorical predictor?
- Eg: sex (male vs female), condition (1 vs 2), continuous response
- 2*2 categorical design
- Since the lines cross, we know that this is an antagonistic interaction

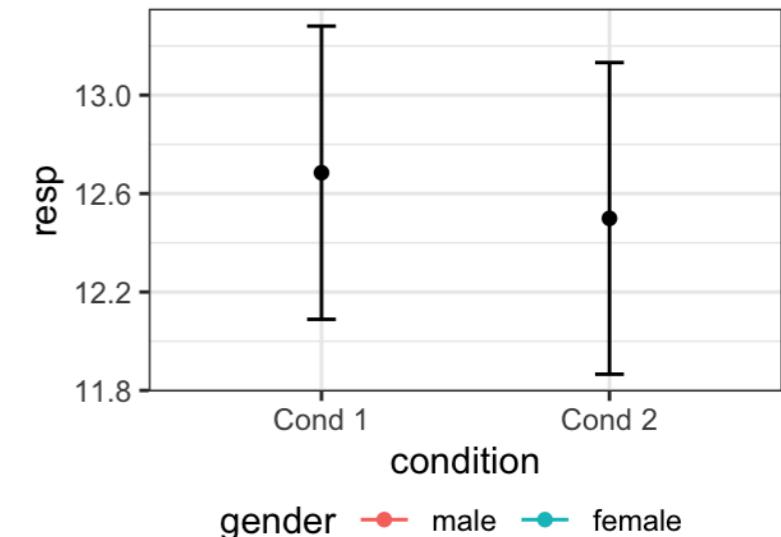
Main effect of sex

	sex	resp
1	male	12.69
2	female	12.49



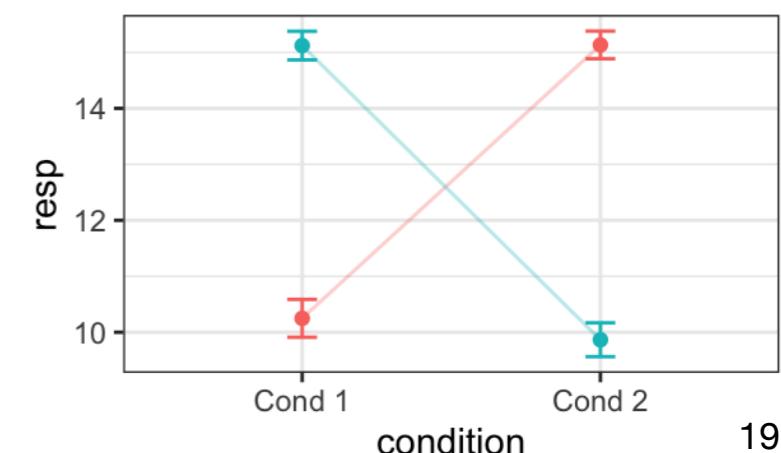
Main effect of condition

	condition	resp
1	Cond 1	12.68
2	Cond 2	12.49

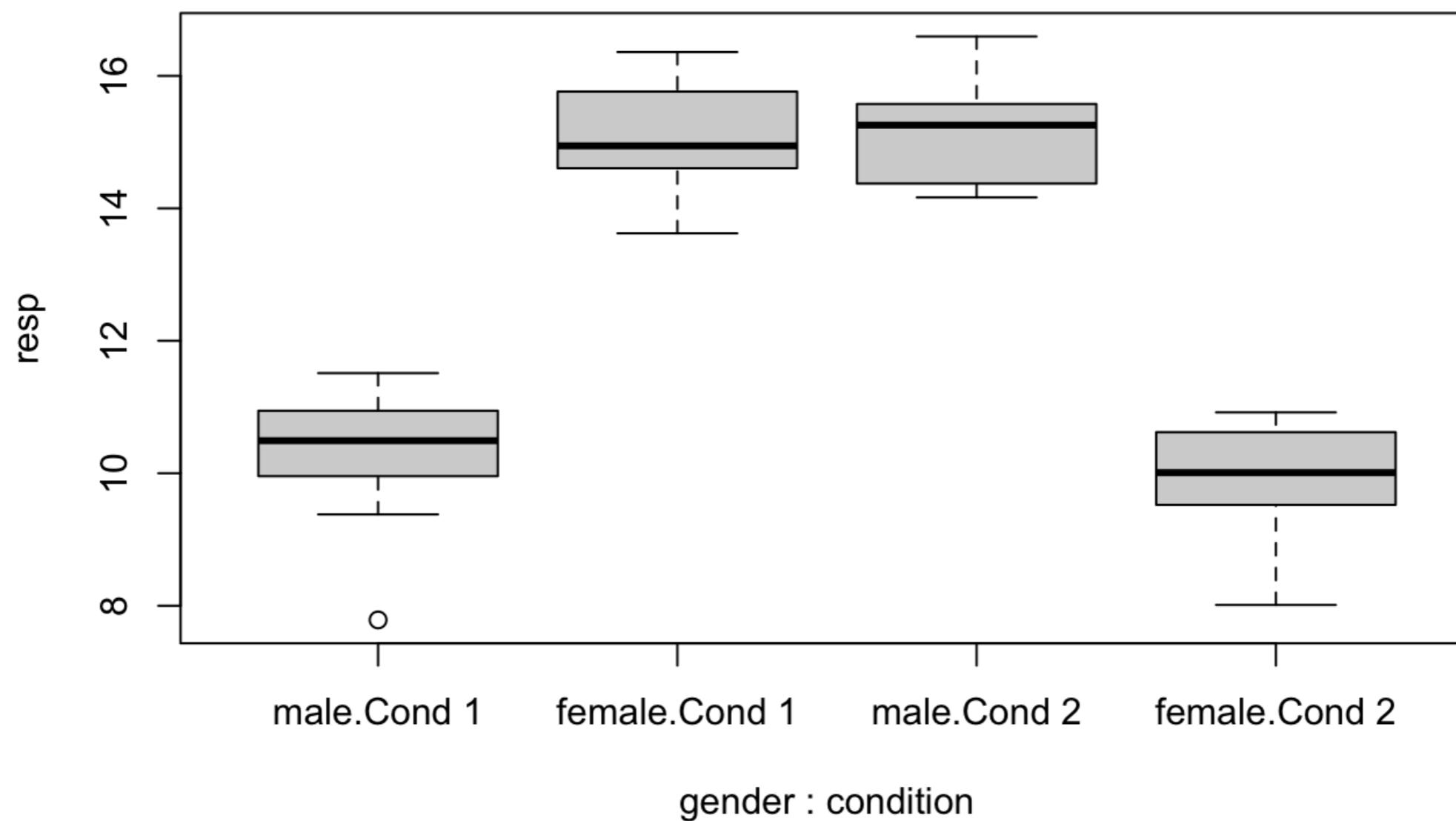


Sex * Condition

	Cond 1	Cond 2
male	10.24	15.13
female	15.12	9.86



Two categorical predictors (2)



Testing the interaction in R

- `summary(lm(resp ~ condition * gender, data = data))`

Call:

```
lm(formula = resp ~ condition * gender, data = dat)
```

Residuals:

Min	1Q	Median	3Q	Max
-2.4635	-0.4570	0.1093	0.6152	1.4631

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	10.2488	0.2881	35.57	< 2e-16 ***
conditionCond 2	4.8834	0.4074	11.98	3.98e-14 ***
genderfemale	4.8719	0.4074	11.96	4.26e-14 ***
conditionCond 2:genderfemale	-10.1378	0.5762	-17.59	< 2e-16 ***

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 0.9111 on 36 degrees of freedom

Multiple R-squared: 0.8961, Adjusted R-squared: 0.8874

F-statistic: 103.5 on 3 and 36 DF, p-value: < 2.2e-16

Interpreting the output

- β_0 (**Intercept**): the value of Y when condition and gender are in their reference (baseline) level
- β_1 **condition**: the change in Y when condition goes from “Cond 1” to “Cond 2” in the reference level of gender (“male”)
- β_2 **gender**: the change in Y when gender goes from “male” to “female” in the reference level of condition (“Cond 1”)
- β_4 **condition*gender**: the additional effect of “Cond 2” for females
- $\beta_1 + \beta_4$: the simple effect of “Cond 2” for females

Simple effects (“marginal effects”) in a categorical by categorical interaction (1)

- `library(emmeans)`
- `model <- lm(resp ~ condition * gender, data = data)`
- `emmeans(model, ~ condition * gender)`

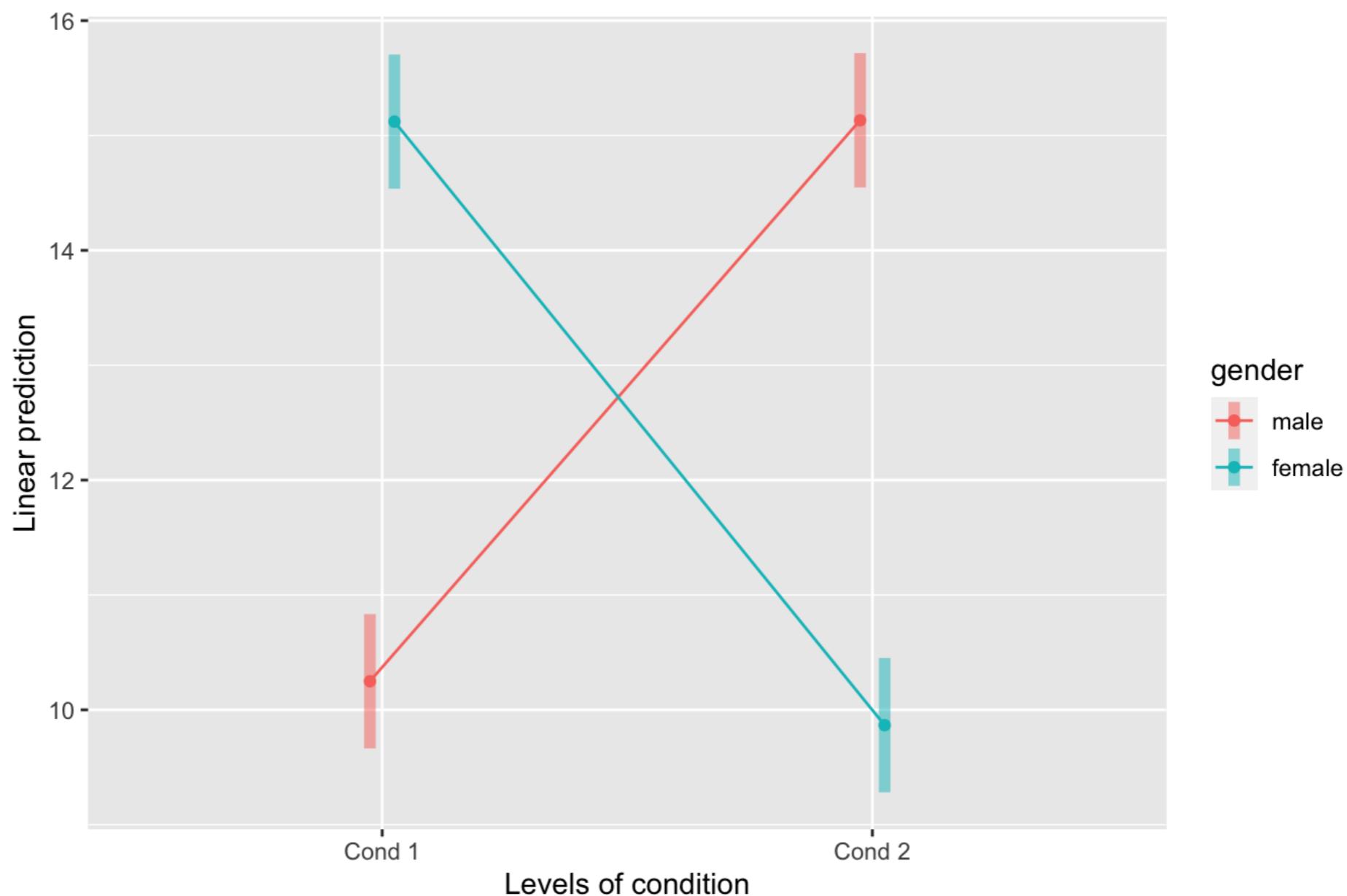
gender	condition	emmmean	SE	df	lower.CL	upper.CL
male	Cond 1	10.25	0.288	36	9.66	10.8
female	Cond 1	15.12	0.288	36	14.54	15.7
male	Cond 2	15.13	0.288	36	14.55	15.7
female	Cond 2	9.87	0.288	36	9.28	10.5

Confidence level used: 0.95

- **Note:** These are not descriptive statistics for the raw data, but rather the model predictions (i.e., the data points that fall on the regression lines estimated by the model)

Simple effects (“marginal effects”) in a categorical by categorical interaction (2)

- `library(emmeans)`
- `emmip(model, gender ~ condition, CIs = TRUE)`



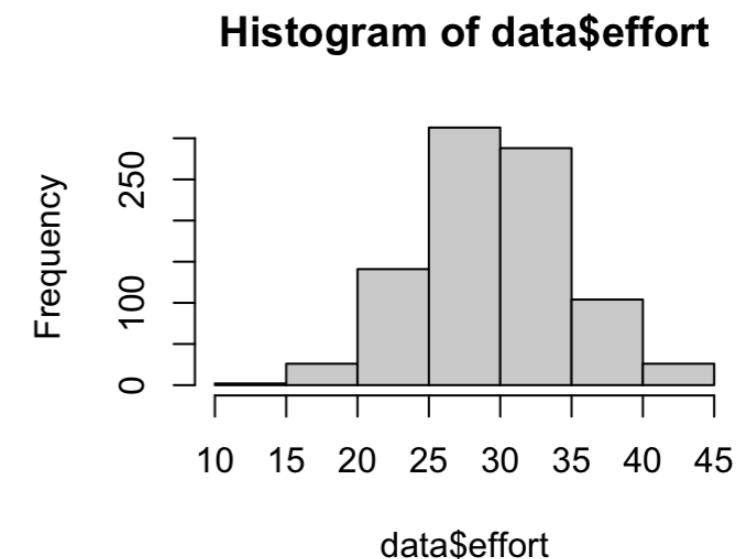
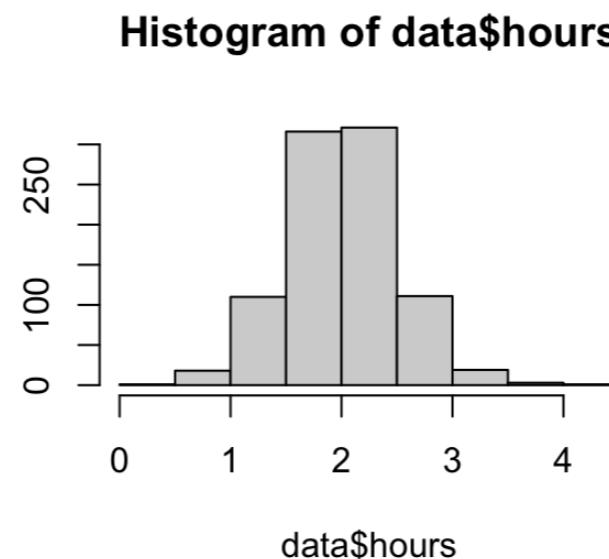
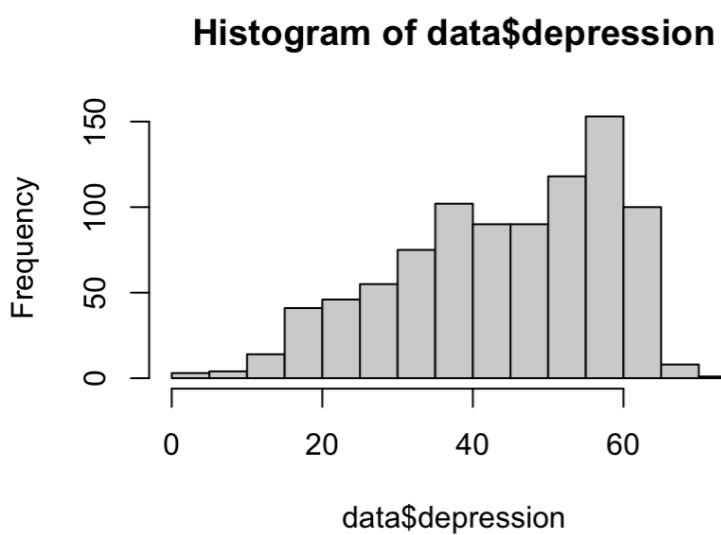
Two continuous predictors (2)

- **Does the effect of one continuous predictor depend on the levels of another continuous predictor?**
- Eg: amount of HIIT (hours), effort, depression scale

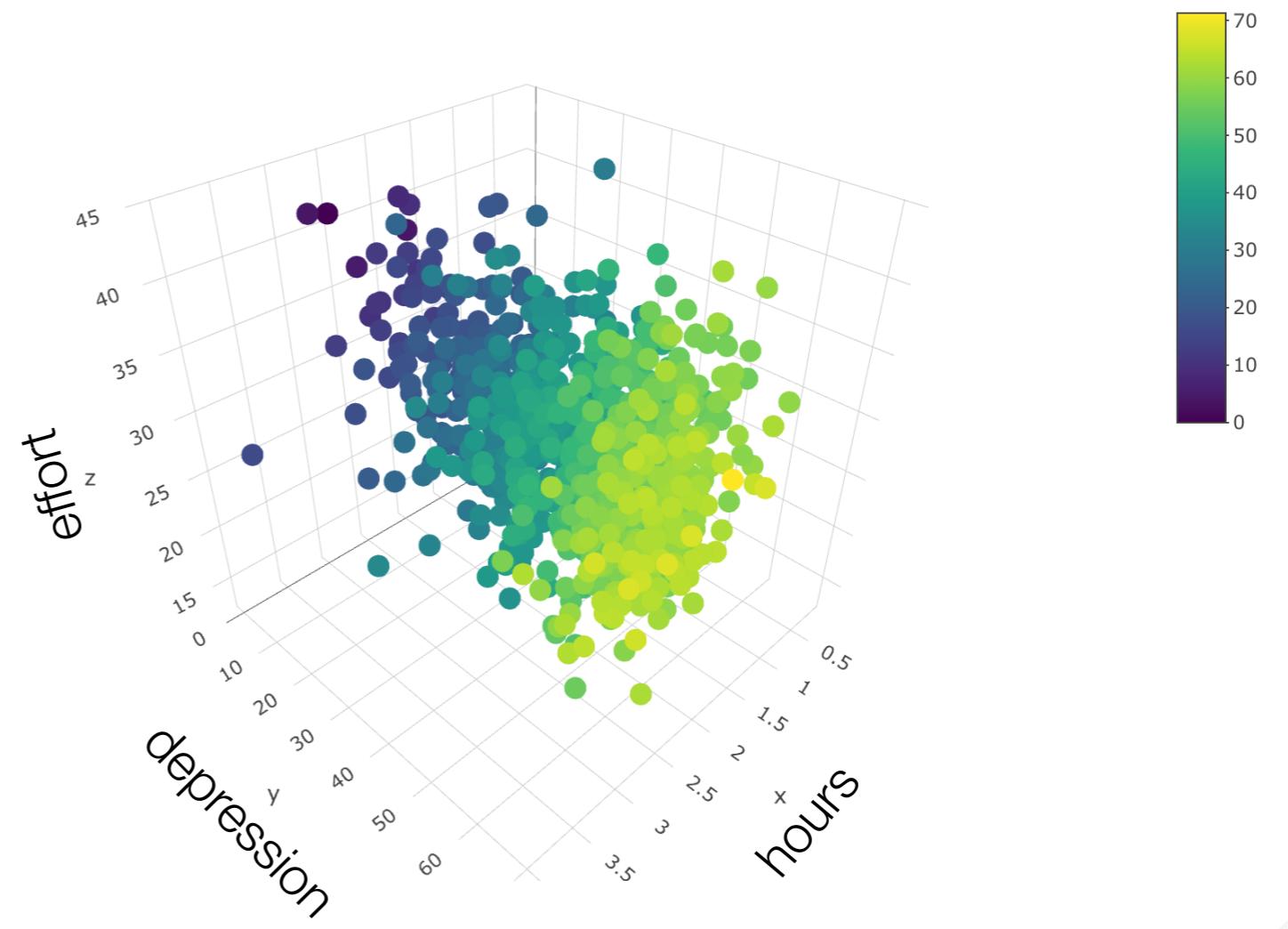
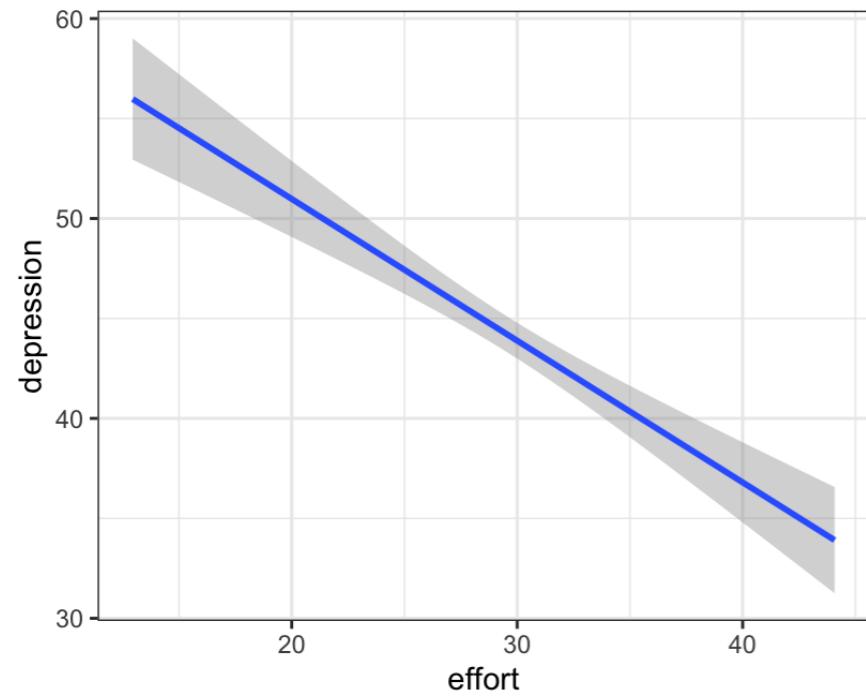
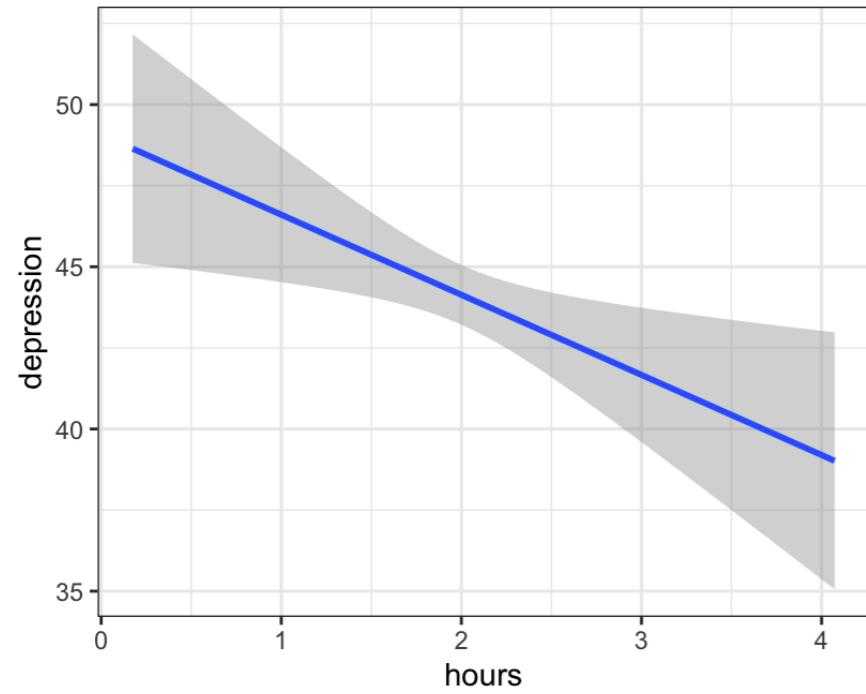
```
> summary(data$depression)
   Min. 1st Qu. Median      Mean 3rd Qu.      Max.
0.00    34.10   46.27    44.13   55.89    71.29

> summary(data$hours)
   Min. 1st Qu. Median      Mean 3rd Qu.      Max.
0.1751  1.6764  2.0051    2.0024  2.3375    4.0722

> summary(data$effort)
   Min. 1st Qu. Median      Mean 3rd Qu.      Max.
12.95   26.26   29.63    29.66   33.10    44.08
```



Two continuous predictors (3)



Testing the interaction in R

- `summary(lm(depression ~ hours * effort, data))`

Call:

```
lm(formula = depression ~ hours * effort, data = data)
```

Residuals:

Min	1Q	Median	3Q	Max
-34.51	-11.13	1.78	10.60	29.52

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	46.35182	11.60362	3.995	7.01e-05 ***
hours	9.37568	5.66392	1.655	0.0982 .
effort	0.08028	0.38465	0.209	0.8347
hours:effort	-0.39335	0.18750	-2.098	0.0362 *

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 13.56 on 896 degrees of freedom

Multiple R-squared: 0.07818, Adjusted R-squared: 0.07509

F-statistic: 25.33 on 3 and 896 DF, p-value: 9.826e-16

Interpreting the output

- β_0 (**Intercept**): The predicted value of Y when hours = 0 and effort = 0
- β_1 (**hours**): The change in Y when hours increases by 1 unit and effort = 0
- β_2 (**effort**): The change in Y when effort increases by 1 unit and hours = 0
- β_3 (**hours*effort**): The change in the slope of hours for every one unit increase in effort (or vice versa)

Simple effects (“marginal effects”) in a continuous by continuous interaction (1)

- To estimate marginal effects, we need to look at the slope of **hours** for specific levels of **effort**
- We code **high effort** as effort that is $2SD > \text{median}$, **medium effort** as the median of effort, and **low effort** as effort that is $2SD < \text{median}$

```
high_effort <- round(median(data$effort) + 2*sd(data$effort), digits = 2)
medium_effort <- round(median(data$effort), digits = 2)
low_effort <- round(median(data$effort) - 2*sd(data$effort), digits = 2)

mylist <- list(effort=c(low_effort,medium_effort,high_effort))

emtrends(model, ~effort, var="hours",at=mylist)
```

effort	hours.trend	SE	df	lower.CL	upper.CL
19.4	1.76	2.164	896	-2.48	6.012
29.6	-2.28	0.915	896	-4.08	-0.483
39.9	-6.33	2.105	896	-10.46	-2.196

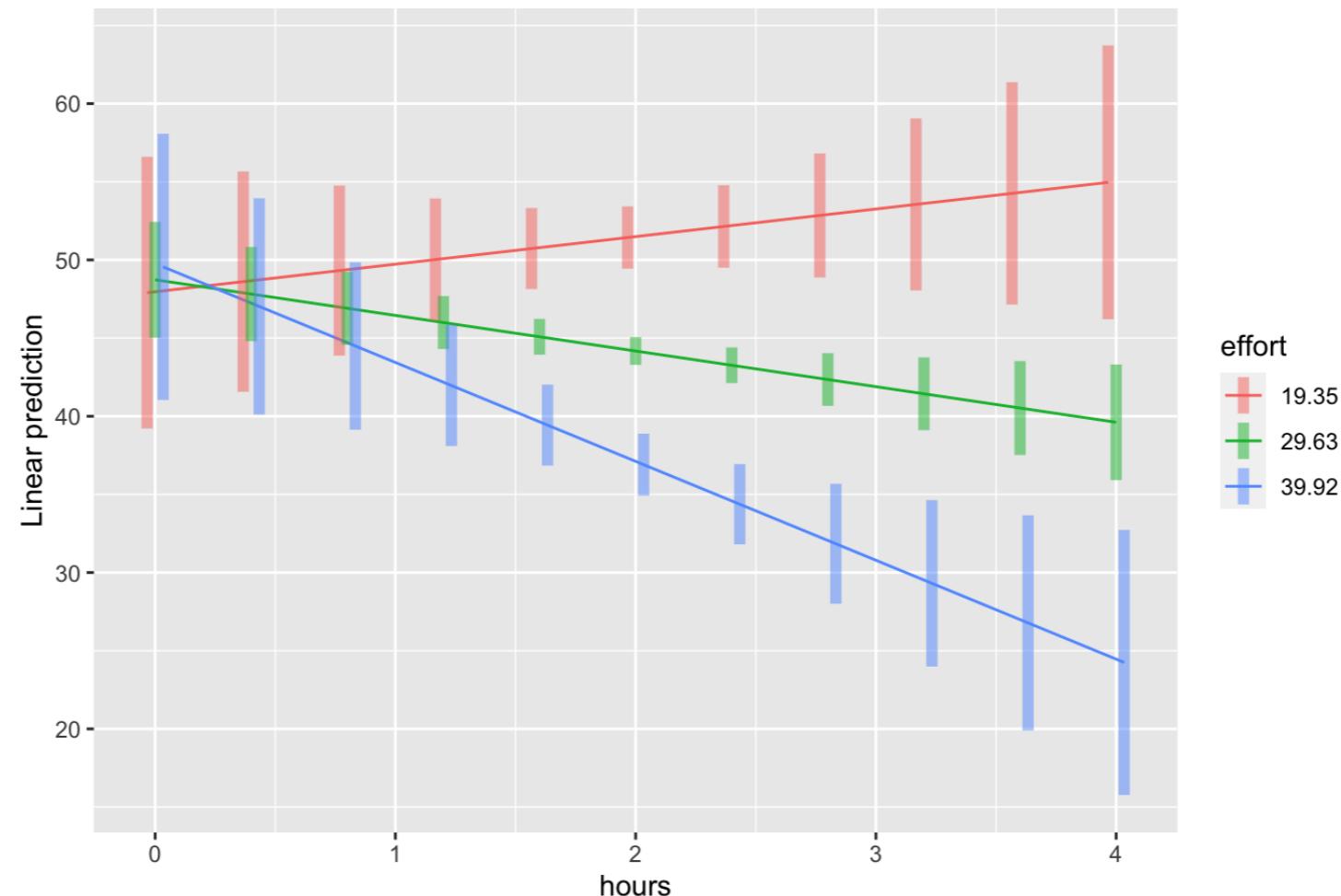
Results are averaged over the levels of: hours
Confidence level used: 0.95

Simple effects (“marginal effects”) in a continuous by continuous interaction (2)

- For plotting, we want **hours** to be on the x-axis and **effort** to be separate the lines
- For **hours** we create a sequence of values to span a reasonable range of hours (we use a sequence of 0.4 increments)
- **effort** is coded as earlier

Simple effects (“marginal effects”) in a continuous by continuous interaction (3)

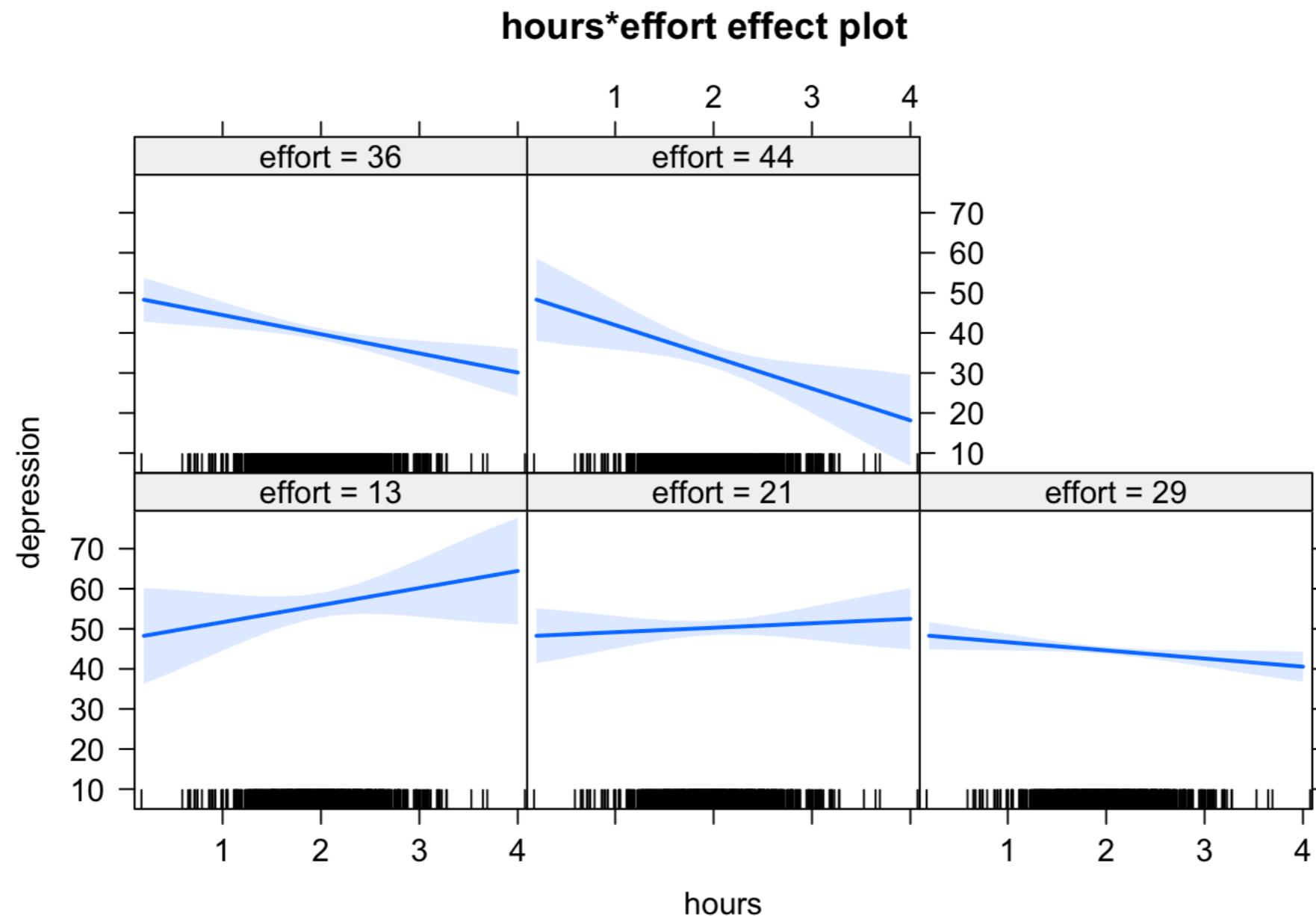
```
mylist <- list(hours=seq(0,4,by=0.4),effort=c(low_effort,medium_effort,high_effort))  
emmip(model,effort~hours,at=mylist, CIs=TRUE)
```



Note: the lines cross each other, so this is also an antagonistic interaction

An alternative method (easier but not as nice)

```
plot(allEffects(model))
```



Is there a significant difference between the slopes?

```
> emtrends(model, pairwise ~effort, var="hours",at=mylist, adjust="none")
```

\$emtrends

effort	hours.trend	SE	df	lower.CL	upper.CL
19.4	1.76	2.164	896	-2.48	6.012
29.6	-2.28	0.915	896	-4.08	-0.483
39.9	-6.33	2.105	896	-10.46	-2.196

Results are averaged over the levels of: hours

Confidence level used: 0.95

\$contrasts

contrast	estimate	SE	df	t.ratio	p.value
19.35 - 29.63	4.04	1.93	896	2.098	0.0362
19.35 - 39.92	8.09	3.86	896	2.098	0.0362
29.63 - 39.92	4.05	1.93	896	2.098	0.0362

Results are averaged over the levels of: hours

One continuous and one categorical predictor

- Does the effect of a continuous predictor depend on the levels of a categorical predictor?

- Eg: amount of HIIT (hours), vegan diet, depression scale

- 1 = vegan
- 0 = non-vegan

```
> table(data$vegan)
```

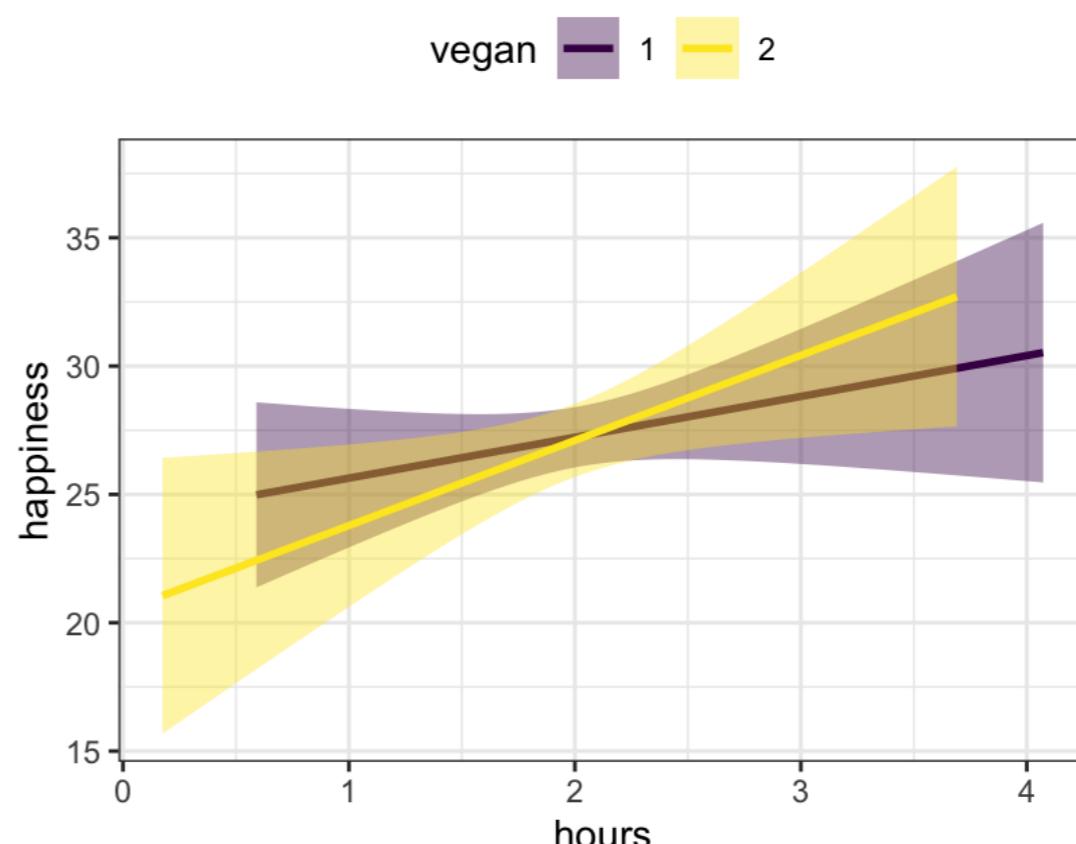
1	2
450	450

```
> summary(data$happiness)
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.00	15.39	25.02	27.16	37.19	71.29

```
> summary(data$hours)
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.1751	1.6764	2.0051	2.0024	2.3375	4.0722



Testing the interaction in R

```
> summary(lm(happiness ~ hours * vegan, data = data))
```

Call:

```
lm(formula = happiness ~ hours * vegan, data = data)
```

Residuals:

Min	1Q	Median	3Q	Max
-27.118	-11.350	-1.963	10.001	42.376

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	24.044	2.805	8.571	<2e-16 ***
hours	1.591	1.352	1.177	0.240
vegan2	-3.571	3.915	-0.912	0.362
hours:vegan2	1.724	1.898	0.908	0.364

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 14.06 on 896 degrees of freedom

Multiple R-squared: 0.008433, Adjusted R-squared: 0.005113

F-statistic: 2.54 on 3 and 896 DF, p-value: 0.05523

Interpreting the output

- β_0 (**Intercept**): The predicted value of Y when hours = 0 and vegan = 0 (reference level)
- β_1 (**hours**): The change in Y when hours increases by 1 unit and vegan = 0
- β_2 (**vegan**): The change in Y when vegan goes from 0 to 1 and hours = 0
- β_3 (**hours*vegan**): The change in the slope of hours when vegan goes from 0 to 1

Simple effects (“marginal effects”) in a continuous by categorical interaction (1)

```
> emtrends(model, ~ vegan, var="hours")
```

	vegan	hours.trend	SE	df	lower.CL	upper.CL
1		1.59	1.35	896	-1.063	4.25
2		3.32	1.33	896	0.702	5.93

Confidence level used: 0.95

```
> emtrends(model, pairwise ~ vegan, var="hours")
```

\$emtrends

	vegan	hours.trend	SE	df	lower.CL	upper.CL
1		1.59	1.35	896	-1.063	4.25
2		3.32	1.33	896	0.702	5.93

Confidence level used: 0.95

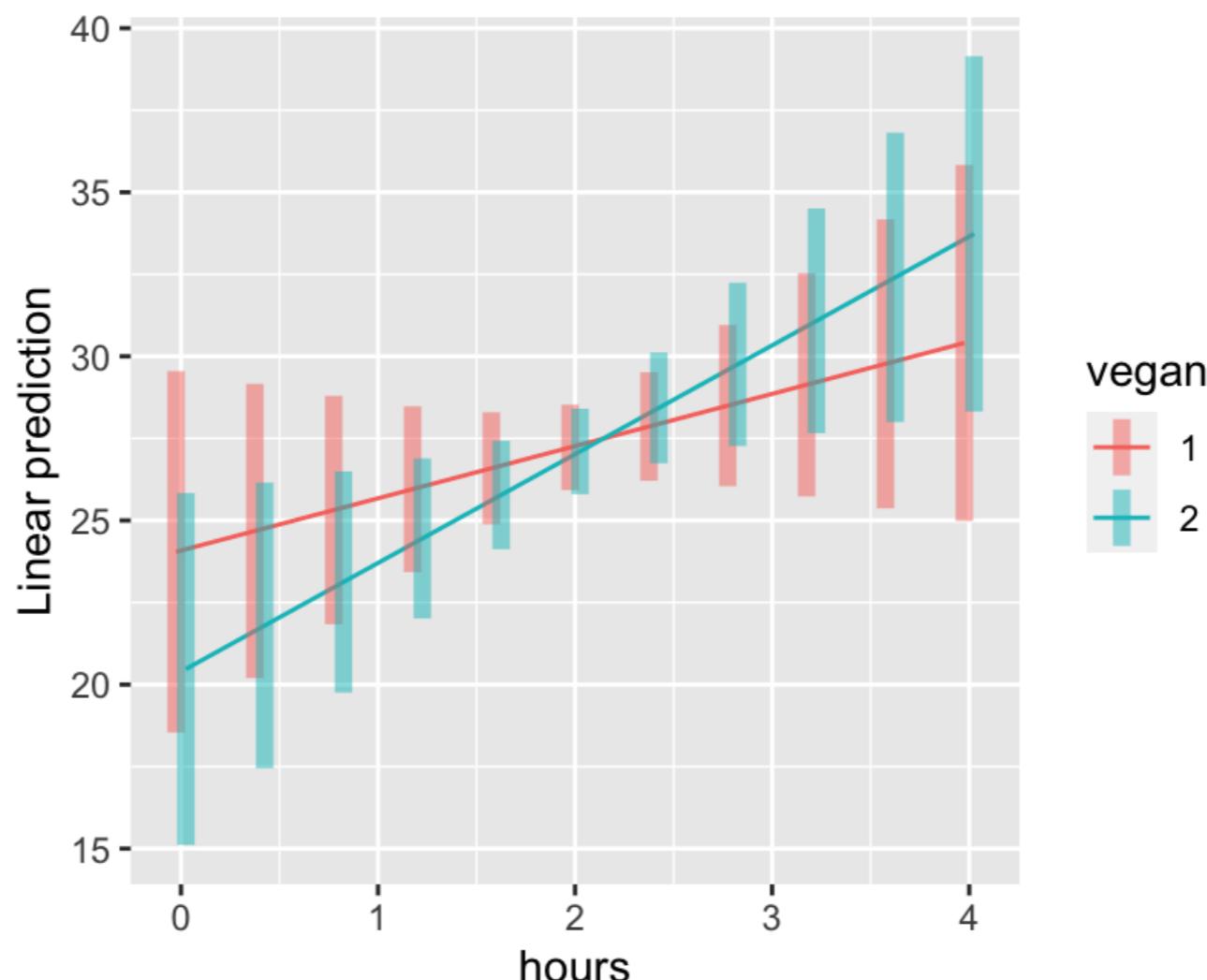
\$contrasts

	contrast	estimate	SE	df	t.ratio	p.value
1 - 2		-1.72	1.9	896	-0.908	0.3639

Simple effects (“marginal effects”) in a continuous by continuous interaction (2)

- To plot the data, we use the same “trick” as before by picking some specific values of **hours**

```
mylist <- list(hours = seq(0,4, by = 0.4))
emmip(model, vegan ~ hours, at = mylist, CIs = TRUE)
```



Take-home message

- Regression is a very powerful tool
- It allows us to measure the effect of predictor variables on an outcome variable
- Sometimes, multiple predictors may interact with each other
- This means that the effect of one predictor on the outcome variable is different for different levels of another predictor
- Interaction analysis helps us unearth these interactive effects
- We rarely go beyond a three-way interaction because of interpretability issues

A word on the textbook

- The book is your friend... use it!
- Methods 2 will be math-heavy
- This book does a decent job at giving you the basic mathematical notions that will prepare you for what comes next
- Remember that my lectures complement the book, they do not replace it
- You need both lectures and the book to understand the topics fully

