Visualizations of Outreach Effect in Kenyan ELections

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```
#Written 1
                           y_i = \beta_0 + \beta_1 local + SMS_i + \beta_2 mean\_agei + \epsilon_i
# Written 2
## stan_glm
  family:
##
                   gaussian [identity]
    formula:
                   perc_reg ~ treatment + mean_age
    observations: 100
    predictors:
##
##
                          Median MAD_SD
## (Intercept)
                          -0.1231 0.0617
## treatmentlocal + SMS
                           0.0264
                                   0.0076
## mean_age
                           0.0031 0.0015
##
## Auxiliary parameter(s):
##
         Median MAD_SD
## sigma 0.0375 0.0027
##
## * For help interpreting the printed output see ?print.stanreg
## * For info on the priors used see ?prior_summary.stanreg
```

The median value for the Intercept is the best estimate for beta0, which is the perc_reg, i.e. the percent of registered voters in a polling district, if no one was treated and the mean_age was 0. The median value for local + SMS is our best estimate for beta1, which represents the change in percent of registered voters treated counties. The median value for mean_age is the best estimate for beta2, which represents the change in the percent of registered voters for every increase of one year in the polling district's mean_age. For all of these estimates, we can discern our confidence in them using the MAD SD; we are 95% confident that the true value lies within two MAD SDs of the median.

Written 3

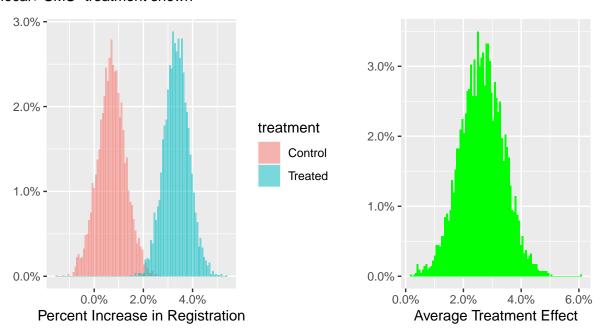
```
## elpd_diff se_diff
## fit_2 0.0 0.0
## fit_1 -3.6 5.4
```

From the results, our second model, fit_2, has a slightly lower elpd and is therefore more accurate. However, the median value for the difference between fit_1 and fit_2 is less than 4, a value so small it becomes hard to

distinguish it from noise and so to determine which of the two values is truly better. This median estimate is really only our best guess of the difference between the two. We, however, can be 95% confident that the actual value of the difference is within two standard errors of the mean. With that taken into account, fit_1 could either be more appreciably worse of an estimate than fit_2 or appreciably greater than fit_2. This only gives more reason to not consider one of these models better than the other.

Posterior Distribution of the Expected Value of the Change in Proportion of Registered Voters

Survey of a 2013 Kenyan Voter Registration Experiment "local+ SMS" treatment shown



Source: Kenyan Voter Registration Experiment

Written 4

The first plot above shows the posterior for the expected value of the change in the number of voters in a district whose mean age is 42 under both treatment and control. While this same value was negative when age was 0, now that the age has been given a much more reasonable value, draws from the same fitted object show much more reasonable answers. The second graph is a posterior for the estimated average treatment effect. It was obtained by subtracting each of the 4000 draws which comprise the first plot's control posterior from one of the 4000 draws which comprise the first plot's treatment posterior. In doing so, we have calculated an estimate for the Average Treatment Effect, which we cannot know exactly due to the Fundamental Problem of Causal Inference.

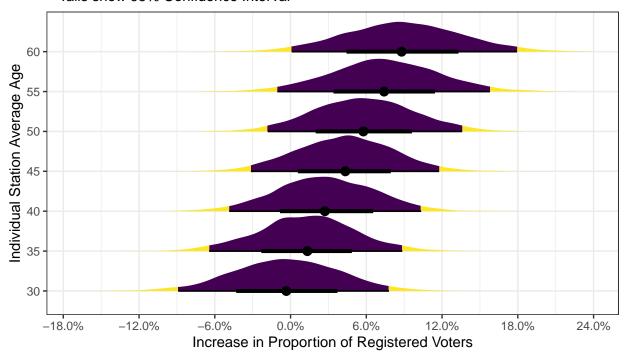
```
## Warning: 'as.tibble()' was deprecated in tibble 2.0.0.
```

^{##} Please use 'as_tibble()' instead.

^{##} The signature and semantics have changed, see '?as_tibble'.

Predicted New Voter Registration by Station Average Age

Treated Distrcts Polled before the Upcoming Election Tails show 95% Confidence Interval



Source: Kenya Voter Experiment

Written 5

For each age on the y axis, a posterior distribution of newly registered voters for a treated districts of that average age is plotted. As average age increases, the amount of individuals who registered increase. The same is true in the U.S. that old people are more likely to vote. For districts with an average age of 60, the average proportion of newly registered voters is approximately 9%. We would be surprised if any district of an average age of 60 reported an increase greater than 19% or fewer than 0%, since these are outside of our 95% confidence interval.

Code

```
knitr::opts_chunk$set(echo = FALSE)
library(primer.data)
library(tidyverse)
library(rstanarm)
library(patchwork)
library(ggdist)
# I did not know you could put it set_seed this way.
set.seed(2319)
object_1 <- kenya %>%
  filter(treatment %in% c("local + SMS", "control")) %>%
  drop_na() %>%
  rename(perc_reg = reg_byrv13) %>%
# Without a vector, you need a minus for every column name
  select(-rv13, -block, -poll_station) %>%
#It seems n is required for slice_sample. This is not what I remembered
  slice sample(n = 100)
fit_1 <- stan_glm(data = object_1,</pre>
         formula = perc_reg ~ treatment + mean_age,
         seed = 54,
         refresh = 0)
print(fit_1, digits = 4)
\#\ I\ don't\ think\ I\ like\ mean\_age\ on\ its\ own\ line,\ seems\ inconsistent. Is this the
# best way to format this?
fit_2 <- stan_glm(data = object_1,</pre>
         formula = perc_reg ~ treatment + poverty + distance + pop_density +
           mean_age,
         seed = 47,
         refresh = 0
# I wish I knew more about the loo function. Preceptir's explanation was
# somewhat helpful but the error (and I had to get rid of another manually)
# reminds how little I really know of what is going on.
loo_1 <- loo(fit_1, k_threshold = 0.7)</pre>
loo_2 <- loo(fit_2, k_threshold = 0.7)</pre>
object_2 <- loo_compare(loo_1, loo_2)</pre>
print(object_2)
# I have gotten used to making newobs which is good.
newobs <- tibble(mean_age = 42, treatment = c("local + SMS", "control"))</pre>
object_3 <- posterior_epred(object = fit_1,</pre>
                  newdata = newobs) %>%
  as_tibble() %>%
  mutate all(as.numeric) %>%
  rowwise() %>%
  rename(Treated = `1`,
```

```
Control = `2`) %>%
  mutate(ATE = Treated - Control)
# This pivot longer was instinctual which was good. I at first made a posterior
# epred instead of posterior predict. The weird values on the graph helped me
# realize that I did not read the question closely.
p1 <- object_3 %>%
  pivot longer(cols = Treated:Control,
                names_to = "treatment",
               values to = "results") %>%
  ggplot(aes(x = results, fill = treatment)) +
  geom_histogram(bins = 100,
                 position = "identity",
                 alpha = 0.5,
                 aes(y = after_stat(count) / sum(count))) +
  scale_x_continuous(labels = scales::percent_format()) +
  scale_y_continuous(labels = scales::percent_format()) +
  labs(x = "Percent Increase in Registration",
       y = "")
# I know we have done colors without hex codes but I cannot remember. Is this
# the best way to do a null y?
p2 <- object_3 %>%
  ggplot(aes(x = ATE)) +
  geom_histogram(bins = 100,
                 fill = "#00FF00".
                 aes(y = after_stat(count) / sum(count))) +
  scale_x_continuous(labels = scales::percent_format()) +
  scale_y_continuous(labels = scales::percent_format()) +
  labs(x = "Average Treatment Effect",
       y = "")
# Very cool function. I feel like the way I had to balance labels was somewhat
# haphazard, but it works well.
p1 + p2 +
  plot_annotation(title = "Posterior Distribution of the Expected Value of the Change in
Proportion of Registered Voters",
         subtitle = 'Survey of a 2013 Kenyan Voter Registration Experiment
"local+ SMS" treatment shown',
caption = "Source: Kenyan Voter Registration Experiment")
# This new obs was also pretty instictual, but the numbers need to be numeric
# and not characters as I did before.
newobs1 \leftarrow tibble(mean_age = c(30, 35, 40, 45, 50, 55,
                          60), treatment = "local + SMS")
posterior_predict(fit_1,
                  newobs1) %>%
  as.tibble() %>%
  janitor::clean_names() %>%
  mutate_all(as.numeric) %>%
# I see why Preceptor recommneded janitor here, which didn't work when I tried
```

```
# to use it before as_tibble. The renaming only seemed to work with the numbers
# referred to in strings.
 rename("30" = x1,
        "35" = x2,
        "40" = x3,
         "45" = x4,
        "50" = x5,
         "55" = x6.
         "60" = x7) \%
 pivot_longer(cols = "30":"60",
              names_to = "age",
              values_to = "results") %>%
# At first I made these numbers numeric, which the halfeye did not seem to like.
# Looking at Nana's code in slack helped me realize my mistake.
  ggplot(aes(x = results, y = age)) +
  stat_halfeye(aes(fill = stat(cut_cdf_qi(cdf, .width = c(0.95, 1))))) +
 labs(title = "Predicted New Voter Registration by Station Average Age",
       subtitle = "Treated Distrcts Polled before the Upcoming Election
  Tails show 95% Confidence Interval",
  y = "Individual Station Average Age",
  x = "Increase in Proportion of Registered Voters",
   caption = "Source: Kenya Voter Experiment") +
scale_x_continuous(breaks = (seq(-.24, .24, .06)),
                   labels = scales::percent_format()) +
 theme_bw() +
 theme(legend.position = "none")
```