

MATH 60604A

Statistical modelling

§ 6a - Group effects

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Inclusion of group effects

- So far, we have only accounted for group structure by modelling the within-group correlation.
- We may also want to include a **group effect** in the mean model, i.e., a different intercept for each group.
- This is done by adding the categorical group variable g as explanatory variable in the mean model, which translates into $m - 1$ indicator variables $1_{g=i}$ for $i = 1, \dots, m - 1$ if there are m groups.

Equation for group effect

- Suppose that we only include the categorical variable g representing groups,

$$Y_{ij} = \beta_0 + \sum_{i=1}^{m-1} \beta_i \mathbf{1}_{g=i} + \varepsilon_{ij},$$

- for the baseline (group m), the intercept is β_0 ,
- the group effect for $g = i$ is β_i ($i = 1, \dots, m - 1$), and the overall group-specific “intercept” is $\beta_0 + \beta_i$.

Linear model with a group effect for the revenge data

We consider a regression model for revenge with a group effect to illustrate the challenges.

- The idea here is to model the fact that desire for revenge can vary between subjects.
- In the current example, there are only five observations per person to estimate the group effect.
- The model will ignore the within-person correlation for now.

Model with group effects for subject

SAS code to fit a linear model via REML

```
proc mixed data=revenge method=reml;  
class id;  
model revenge = id sex age vc wom t / solution;  
run;
```

In addition to the categorical variable `id`, the model includes the same explanatory variables as before. Each person has his/her own “intercept” parameter (`id=80` is the baseline category).

Estimates of fixed effects

Solution for Fixed Effects						
Effect	id	Estimate	Standard Error	DF	t Value	Pr > t
Intercept		6.7425	0.2290	319	29.44	<.0001
id	1	-1.6400	0.3152	319	-5.20	<.0001
id	2	-3.8400	0.3152	319	-12.18	<.0001
id	3	-1.3200	0.3152	319	-4.19	<.0001
id	4	0.2000	0.3152	319	0.63	0.5262
			⋮			
id	79	-0.6000	0.3152	319	-1.90	0.0578
id	80	0
sex		0
age		0
vc		0
wom		0
t		-0.5675	0.01762	319	-32.21	<.0001

Parameter significance

Type 3 Tests of Fixed Effects				
Effect	Num DF	Den DF	F Value	Pr > F
id	75	319	3.77	<.0001
sex	0	.	.	.
age	0	.	.	.
vc	0	.	.	.
wom	0	.	.	.
t	1	319	1037.49	<.0001

There are **no** parameters estimates or tests for the variables *sex*, *age*, *vc* or *wom*, but there is for the time variable *t*. Because some covariates are fixed over time, their effect are not uniquely estimable (perfect collinearity). If we remove *id* from the model, we can however estimate their effects (hence 75 df rather than 79 in the *F*-table).

Collinearity

- Once we've included a fixed effect for each person, it is impossible to include any variable that does not vary in time for a single person.
- The variables `sex`, `age`, `vc` and `wom` are fixed in time for each person (`vc` and `wom` were only measured once, at time 1).
- These variables are already implicitly included in the individual effect. There is perfect collinearity between a variable fixed in time, and the `id` variable.
- This means that we can perfectly predict the value of `sex` (and the three others) by only looking at the `id` variable.
- Therefore, we cannot have a fixed effect for each individual while simultaneously including variables that are fixed in time for each subject.

Challenges arising from the inclusion of a group effect

- Group is a categorical variable: we need enough observations in each group to reliably estimate the group effects.
- If the number of groups m is large relative to the overall sample size, there may also be too many parameters in the model.
- We cannot estimate the effect of variables that do not vary within group if we add group effects.

Model with group effect and correlation structure

The model fitted next includes only `id` and the time variable `t` as explanatory variables in the mean model, but we specify in addition an AR(1) correlation structure within-individual for the errors ε .

SAS code to include a group effect with AR(1) correlation

```
proc mixed data=revenge method=reml;  
class id tcat;  
model revenge = id t / solution;  
repeated tcat / subject=id type=ar(1);  
run;
```

The effect of the AR(1) correlation parameter is significant (likelihood ratio test statistic of 21.68, negligible p -value under χ^2_1). The estimate of the time effect is -0.5684 , very close to that we got in the model including `sex`, `age`, `vc` and `wom`, and the AR(1) structure model in the previous chapter.

Remark on model comparison

- We have to be careful **not to use** the AIC and BIC reported in the output to compare this model to the earlier one including sex, age, vc and wom, since we used the REML estimation method (the default).
- AIC and BIC obtained through REML, are **not comparable** if the “mean” parts of the models (fixed effects) are not the same.
- If we want to compare these models, we must use the maximum likelihood estimator (option `method=ml` when calling `proc mixed`).

Model Information	
Data Set	WORK.REVENGE
Dependent Variable	revenge
Covariance Structure	Autoregressive
Subject Effect	id
Estimation Method	ML

Remark on model comparison

We fit both models with an AR(1) structure for the errors using maximum likelihood.

Model	AIC	BIC	$\hat{\rho}$ (p -value)
sex, age, vc, wom, t	666.1	685.1	0.48 (10^{-20})
id, t	653.4	851.1	-0.013 (0.83)

- The preferred model according to AIC includes id, but AIC tends to select complicated models.
- The preferred model according to BIC includes sex, age, vc and wom and throws away the id variable.
- Once we include an individual effect for group, the correlation structure seems to be unnecessary — the estimated coefficient is even negative, which is counter-intuitive and suggests the model is over-parametrized.

Remark on model comparison

- The choice of covariates depends on the type of study. If we're interested in studying the effects of one or more of the variables `sex`, `age`, `vc` or `wom`, then we don't have any choice: we must choose a model that contains all of them.
- If we're only interested in the time effect, then the two models will come to the same conclusion either way.
- Often, the optimization routine fails — we cannot estimate both the β and the covariance matrix parameters.
- It is possible to include variables that are fixed within group (within person in our example) **and** group effects (`id` in our example) at the same time by using **random effects**.