

MATH 60604A

Statistical modelling

§ 6e - Random slope model

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Model formulation

We consider a linear mixed model with a random slope and a random intercept for the revenge data, of the form

$$\begin{aligned} \mathbf{Y}_i \mid \mathcal{B}_i = \mathbf{b}_i &\sim \text{No}_5 \left(\mathbf{X}_i \boldsymbol{\beta} + \mathbf{Z}_i \mathbf{b}_i, \sigma^2 \mathbf{I}_5 \right) \\ \mathcal{B}_i &\sim \text{No}_2(\mathbf{0}_2, \boldsymbol{\Omega}) \end{aligned}$$

where $\mathbf{Z}_i = [\mathbf{1}_5, \text{time}_i]$ is a 5×2 model matrix for the random effects and $\boldsymbol{\Omega} = \begin{pmatrix} \omega_{11} & \omega_{12} \\ \omega_{12} & \omega_{22} \end{pmatrix}$.

The columns of \mathbf{Z}_i typically include as covariates

- time or
- indicators for categorical variables (group effect).

Random effects on the predictor variables

Suppose the matrix $\mathbf{Z}_i = [\mathbf{1}_{n_i}, \mathbf{X}_{1i}]$.

$$Y_{ij} = (\beta_0 + b_{0i}) + (\beta_1 + b_{1i})X_{ij1} + \beta_2 X_{ij2} + \cdots + \beta_p X_{ijp} + \varepsilon_{ij}.$$

- The conditional effect of the variable X_1 for group i is $\beta_1 + b_{1i}$
- The parameter β_1 is the “slope” of X_1 averaged over the entire population.
- $\beta_1 + b_{1i}$ is the effect of X_1 specific to group i .

Covariance of the response

- The covariance matrix of Y_{ij} depends on the predictors in \mathbf{Z}_i which have random effects.
- For example, if $\mathbf{Z}_i = [\mathbf{1}_{n_i}, \mathbf{X}_{1i}]$, the marginal variance of Y_{ij} is

$$\text{Var}(Y_{ij} \mid \mathbf{X}_i) = \omega_{11} + X_{ij1}^2 \omega_{22} + 2X_{ij1} \omega_{12} + \sigma_\varepsilon^2.$$

- With independent errors, the covariance between two observations in the same group is

$$\text{Cov}(Y_{ij}, Y_{ik} \mid \mathbf{X}_i) = \omega_{11} + X_{ij1} X_{ik1} \omega_{22} + (X_{ij1} + X_{ik1}) \omega_{12}.$$

- It may be difficult to estimate parameters if the errors has a complex covariance structure (not to mention computational costs).

SAS code for random slope model

```
proc mixed data=statmod.revenge;  
model revenge = sex age vc wom t  
      / ddfm=kenwardroger solution;  
random intercept t / subject=id type=un v=1 vcorr=1;  
run;
```

The output includes information about the number of covariance parameters, the number of random effects, etc.

Dimensions	
Covariance Parameters	4
Columns in X	6
Columns in Z per Subject	2
Subjects	80
Max Obs per Subject	5

Covariance matrix of response

Covariance Parameter Estimates			Estimated V Matrix for Subject 1					
Cov Parm	Subject	Estimate	Row	Col1	Col2	Col3	Col4	Col5
UN(1,1)	id	0.3064	1	0.4239	0.1830	0.1476	0.1122	0.07682
UN(2,1)	id	-0.05268	2	0.1830	0.3704	0.1468	0.1287	0.1106
UN(2,2)	id	0.01730	3	0.1476	0.1468	0.3515	0.1452	0.1444
Residual	id	0.2055	4	0.1122	0.1287	0.1452	0.3672	0.1782
			5	0.07682	0.1106	0.1444	0.1782	0.4175

- The variance of the random intercept is $\omega_{11} = 0.3064$
- The variance of the random slope is $\omega_{22} = 0.01730$
- The correlation between the random effects is -0.72 .

Testing for correlation between random effects

- We can test whether $\mathcal{H}_0 : \omega_{12} = 0$ versus $\mathcal{H}_a : \omega_{12} \neq 0$ by fitting the model with diagonal covariance and performing a likelihood ratio test (REML, since they have the same fixed effects)
 - in SAS, change type=un to type=vc (default option)
 - the test statistic is $R = 8.98$
 - its null distribution is χ_1^2 (regular problem, covariance can be negative)
 - the p -value is 0.002:
 - the correlation between the random effects is strongly significant.

Model comparison

- We can do similar comparisons with the random intercept-only model,
 - this corresponds to $\mathcal{H}_0 : \omega_{12} = \omega_{22} = 0$
 - the asymptotic null distribution is $\frac{1}{2}\chi_1^2 + \frac{1}{2}\chi_2^2$
 - but this approximation is poor ...probably better to use information criteria.