

Detecting and Tracking Sports Players with Random Forests and Context-Conditioned Motion Models

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Abstract Player movements in team sports are often complex and highly correlated with both nearby and distant players. A single motion model would require many degrees of freedom to represent the full motion diversity of each player and could be difficult to use in practice. Instead, we introduce a set of *Game Context Features* extracted from noisy detection data to describe the current state of the match, such as how the players are spatially distributed. Our assumption is that players react to the current game situation in only a finite number of ways. As a result, we are able to select an appropriate simplified motion model for each player and at each time instant using a random decision forest which examines characteristics of individual trajectories and broad game context features derived from all current trajectories. Our context-conditioned motion models implicitly incorporate complex inter-object correlations while remaining tractable. We demonstrate significant performance improvements over existing multi-target tracking algorithms on basketball and field hockey sequences of several minutes in duration containing ten and twenty players respectively.

1 Introduction

Multi-target tracking has been a difficult problem of broad interest for years in computer vision. Surveillance is perhaps the most common scenario for multi-target tracking, but team sports is another popular domain that has a wide range of applications in strategy analysis, automated broadcasting, and content-based retrieval.

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Recent work in pedestrian tracking has demonstrated promising results by formulating multi-target tracking in terms of data association [1, 4, 7, 20, 25, 28, 30, 32]: a set of potential target locations are estimated in each frame using an object detector, and target trajectories are inferred by linking similar detections (or tracklets) across frames. However, if complex inter-tracklet affinity models are used, the association problem becomes NP-hard.

Tracking players in team sports has three significant differences compared to pedestrians in surveillance. First, the appearance features of detections are less discriminative because players on the same team will be visually similar. As a result, the distinguishing characteristics of detections in team sports are primarily position and velocity. Second, sports players move in more erratic fashions, whereas pedestrians tend to move along straight lines at constant speed. Third, although pedestrians deviate to avoid colliding with each other, the motions between pedestrians are rarely correlated in complex ways (some scenarios, like sidewalks, may contain a finite number of common global motions). The movements of sports players, on the other hand, are strongly correlated both locally and globally. For example, opposing players may exhibit strong local correlations when ‘marking’ each other (such as one-on-one defensive assignments). Similarly, players who are far away from each other move in globally correlated ways because they are reacting to the same ball.

Simple, independent motion models have been popular for pedestrian tracking because they limit the complexity of the underlying inference problem [7]. However, the models may not always characterize the motion affinity between a pair of tracklets accurately. Brendel *et al.* [4] modeled inter-target correlations between pedestrians using *context*, which consisted of additional terms in the data association affinity measure based on the spatiotemporal properties of tracklet pairs. Following this convention, we will describe correlations between player movements in terms of *game context*. Much like the differences between the individual target motions in surveillance and team sports, game context is more complex and dynamic compared to context in surveillance. For example, teams will frequently gain and lose possession of the ball, and the motions of all players will change drastically at each turnover.

Because a player’s movement is influenced by multiple factors, the traditional multi-target tracking formulation using a set of independent autoregressive motion models is a poor representation of how sports players actually move. However, motion affinity models conditioned on multiple targets (and that do not decompose into a product of pairwise terms) make the data association problem NP-hard [7]. In this work, we show how data association is an effective solution for sports player tracking by devising an accurate model of player movements that remains tractable by conditioning on features describing the current state of the game, such as which team has possession of the ball. One of our key contributions is a new set of broad *game context features* (GCF) for team sports, and a methodology to estimate them from noisy detections. Using game context features, we can better assess the affinity between trajectory segments by implicitly modeling complex interactions through a random decision forest involving a combination of kinematic and game context

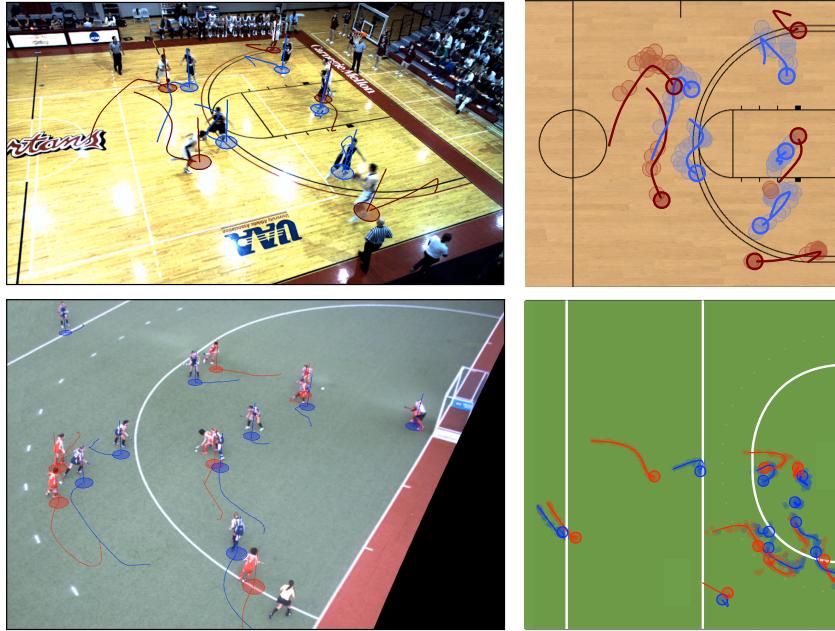


Fig. 1 Motion Models. A player’s future motion is contingent on the current game situation. The global distribution of players often indicates which team is attacking, and local distributions denote when opposing players are closely following each other. We use contextual information such as this to create a more accurate motion affinity model for tracking players. The overhead views of basketball and field hockey show the input detections and corresponding ground truth annotations. Player trajectories are strongly correlated with both nearby and distant players.

features. We demonstrate the ability to track 20 players in over 30 minutes of international field hockey matches, and 10 players in 5 minutes of college basketball.

2 Related Work

Recent success in pedestrian tracking has posed multi-target tracking as hierarchical data association: long object trajectories are found by linking together a series of detections or short tracklets. The problem of associating tracklets across time has been investigated using a variety of methods, such as the Hungarian algorithm [9, 21], linear programming [10], cost-flow networks [30], maximum weight independent sets [4], continuous-discrete optimization [3] and higher-order motion models [7]. Data association is often formulated as a linear assignment problem where the cost of linking one tracklet to another is some function of extracted features (typically motion and appearance). More recent work (discussed shortly) considers more complex association costs.

Crowds are an extreme case of pedestrian tracking where it is often not possible to see each individual in their entirety. Because of congestion, pedestrian motions are often quite similar, and crowd tracking algorithms typically estimate a finite set of global motions. Often, the affinity for linking two tracklets together depends on how well the hypothesized motion agrees with one of the global motions. [1, 32] solve tracking in crowded structured scenes with floor fields estimation and Motion Structure Tracker, respectively. [23] uses a Correlated Topic Model for crowded, unstructured scenes.

Team sports is another relevant domain for multi-target tracking [24], with algorithms based on particle filters being extremely popular [5, 8, 14, 16, 18, 27]. However, results are quite often demonstrated only on short sequences (typically less than two minutes). In contrast, only a small amount of work has investigated long-term sports player tracking. Nillius *et al.* [19] generated a Bayes network of splitting and merging tracklets for a long ten minute soccer sequence, and found the most probable assignment of player identities using max-margin message passing. Kumar and Vleeschouer proposed discriminative label propagation [11] and Shitrit *et al.* use multi-commodity network flow for tracking multiple people [26].

In both pedestrian and player tracking, object motions are often assumed to be independent and modeled as zero displacement (for erratic motion) and/or constant velocity (for smooth motion governed by inertia). In reality, the locations and motions of sports players are strongly correlated. Pairwise repulsive forces have been used in multi-target tracking to enforce separability between objects [2–4, 12, 29]. Recently, multi-object motion models have been used in pedestrian tracking to anticipate how people will change their trajectories to avoid collisions [20], or for estimating whether a pair of trajectories have correlated motions [4]. In team sports, Kim *et al.* [13] estimated motion fields using the velocities of tracklets to anticipate how the play would evolve, but did not use the motion fields to track players over long sequences. Zhang *et al.* [31] augmented the standard independent autoregressive motion model with a database of *a priori* trajectories manually annotated from other games.

3 Hierarchical Data Association Tracking

Our tracking formulation (see Fig. 2) employs an object detector to generate a set \mathcal{O} of hypothesized sports player locations through the duration of the video. Each detection $\mathcal{O}_i = [t_i, \mathbf{x}_i, \mathbf{a}_i]$ contains a time stamp, the player’s location on the ground plane, and the player’s appearance information respectively. The goal is to find the most probable set $\mathcal{T}^* = \{\mathcal{T}_1, \mathcal{T}_2, \dots, \mathcal{T}_N\}$ of player trajectories where each trajectory is a temporal sequence of detections $\mathcal{T}_n = \{\mathcal{O}_a, \mathcal{O}_b, \dots\}$

$$\mathcal{T}^* = \arg \max_{\mathcal{T}} P(\mathcal{O}|\mathcal{T})P(\mathcal{T}). \quad (1)$$

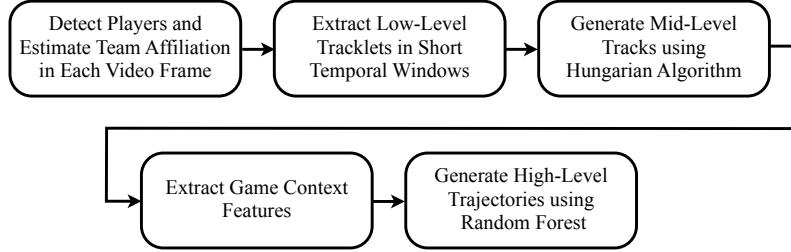


Fig. 2 Algorithm. Our tracking algorithm has five distinct phases: detection, low-level tracklets, mid-level tracks, game context features, and high-level trajectories.

The likelihood $P(\mathcal{O}|\mathcal{T})$ indicates how well a set of trajectories \mathcal{T} matches the observations, and the prior $P(\mathcal{T})$ describes, in the case of sports tracking, how realistic the set of estimated player trajectories \mathcal{T} is. In multi-target tracking, the prior is often simplified to consider each trajectory as an independent Markov chain

$$P(\mathcal{T}) \sim \prod_n P(\mathcal{T}_n) \quad (2)$$

$$\sim \prod_n \prod_t P(\mathcal{T}_n^t | \mathcal{T}_n^{t-1}), \quad (3)$$

where \mathcal{T}_n^t indicates the trajectory of the n th player at time interval t .

In team sports, the prior is a highly complex function and is not well approximated by a series of independent trajectory assessments. We maintain the formulation of conditional independence between trajectories, but condition each individual trajectory prior on a set of game context features θ which describe the current state of the match

$$P(\mathcal{T}) \stackrel{\text{def}}{=} \prod_{n,t} P(\mathcal{T}_n^{t-1} \rightarrow \mathcal{T}_n^t | \theta). \quad (4)$$

Conditioning the individual motion models on game context implicitly encodes higher-order inter-trajectory relationships and long-term intra-trajectory information without sacrificing tractability.

3.1 Detection

We use the method of Carr *et al.* [6] to generate possible (x,y) positions of players in all video frames. The technique requires a calibrated camera, and uses background subtraction to generate foreground masks. Player locations are estimated by evaluating how well a set of hypothesized 0.5m wide cylinders 1.8m tall can explain each observed foreground mask. The method is tuned to only detect high confidence situations, such as when a player is fully visible and well separated from other players. For each detected position, a rectified patch is extracted from the image and a coarse

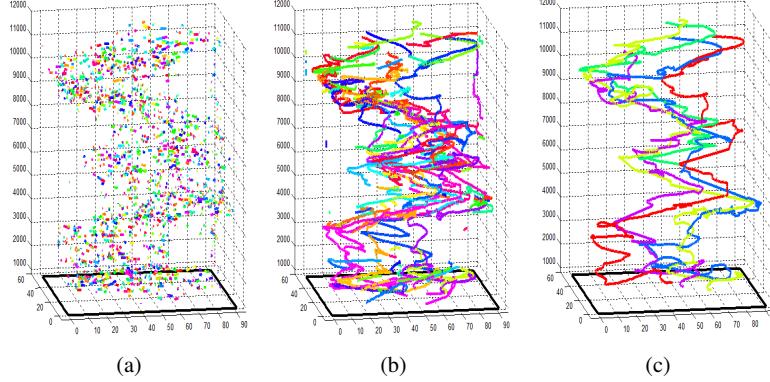


Fig. 3 Hierarchical Data Association. (a) low-level tracklets Υ from noisy detections; (b) mid-level tracks Γ obtained via the Hungarian algorithm [9]; (c) N high-level player trajectories \mathcal{T} via a cost flow network [30].

RGB histogram (4 bins per channel) is computed as an appearance feature. The first few seconds of video are accompanied by user supplied labels: each detection is assigned to one of four categories $\{\text{home}, \text{away}, \text{referee}, \text{none}\}$. A random forest is constructed from the training data, and at test time, outputs the four-class probability histogram for each detection.

3.2 Hierarchical Association

Because the solution space of data association grows exponentially with the number of frames, we adopt a hierarchical approach to handle sequences that are several minutes long (see Fig. 3).

3.2.1 Low-Level Tracklets

A set Υ of low-level tracklets is extracted from the detections by fitting constant velocity models to clusters of detections in 0.5s long temporal windows using RANSAC. Each Υ_i represents an estimate of a player's instantaneous position and velocity (see Fig. 4).

3.2.2 Mid-Level Tracks

Similar to [9], the Hungarian algorithm is used to combine subsequent low-level trajectories into a set Γ of mid-level tracks up to 60s in duration. The method auto-

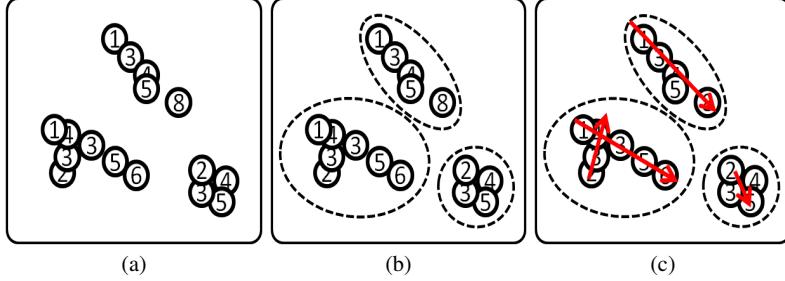


Fig. 4 Low-level Tracklet Extraction. Each detection is represented as a circle with a frame number. (a) detection responses within a local spatial-temporal volume; (b) identified clusters; (c) RANSAC fitted constant velocity models (red).

matically determines the appropriate number of mid-level tracks, but is tuned to prefer shorter, more reliable tracks. Generally, mid-level tracks terminate when abrupt motions occur or when a player is not detected for more than two seconds.

3.2.3 High-Level Trajectories

MAP association is equivalent to the minimum cost flow in a network [30] where a vertex i is defined for each mid-level track Γ_i and edge weights reflect the likelihood and prior in (4). Unlike the Hungarian algorithm, it is possible to constrain solutions to have exactly N trajectories by pushing N units of flow between special source s and sink t vertices (see Fig. 5). The complete trajectory \mathcal{T}_n of each player corresponds to the minimum cost path for one unit of flow from s to t . The cost c_{ij} per unit flow from i to j indicates the negative affinity, or negative log likelihood that Γ_j is the immediate successor of Γ_i , which we decompose into probabilities in continuity of appearance, time and motion

$$c_{ij} = -\log P(\mathcal{O}|\Gamma_i \rightarrow \Gamma_j)P(\Gamma_i \rightarrow \Gamma_j|\theta) \quad (5)$$

$$= -\log (P_a \cdot P_\tau \cdot P_m). \quad (6)$$

The probability that Γ_i and Γ_j belong to the same team is

$$P_a(\Gamma_i \rightarrow \Gamma_j) = a_i \cdot a_j + (1 - a_i) \cdot (1 - a_j) \quad (7)$$

where a_i and $1 - a_i$ are the confidence scores of the mid-level track belonging to team A and B respectively.

Let t_{i0} and t_{i1} denote the start and end times of Γ_i respectively. If Γ_j is the immediate successor of Γ_i , any non-zero time gap implies that missed detections must have occurred. Therefore, the probability based on temporal continuity is defined as

$$P_\tau(\Gamma_i \rightarrow \Gamma_j) = \exp(-\lambda(t_{j0} - t_{i1})). \quad (8)$$

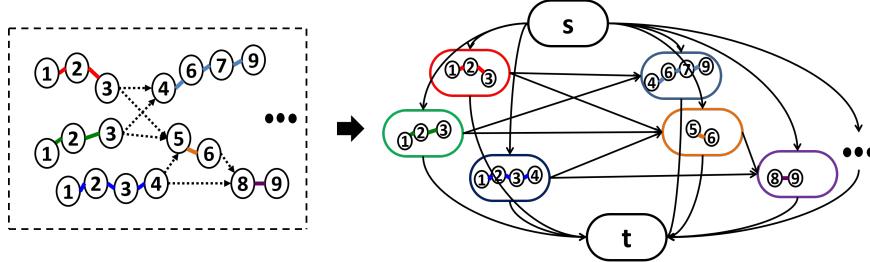


Fig. 5 Cost Flow Network from Mid-Level Tracks. Each small circle with a number inside indicates a detection and its time stamp. Colored edges on the left indicate mid-level track associations. The corresponding Cost Flow Network is shown on the right. Each mid-level track Γ_i , as well as the global source s and sink t forms a vertex, and each directed edge in black from vertex a to b has a cost indicating the negative affinity of associating Γ_a to Γ_b .

Each mid-level trajectory Γ_i has ‘miss-from-the-start’ and ‘miss-until-the-end’ costs on edges (s, i) and (i, t) respectively. The weights are computed using 8 for temporal gaps (T_0, t_{i0}) and (t_{j1}, T_1) , where T_0 and T_1 are the global start and end times of the sequence.

Before describing the form of $P_m(\Gamma_i \rightarrow \Gamma_j | \theta)$ in more detail, we first discuss how to extract a set of game context features θ from noisy detections \mathcal{O} .

4 Game Context Features

In team sports, players assess the current situation and react accordingly. As a result, a significant amount of contextual information is implicitly encoded in player locations. In practice, the set of detected player positions in each frame contains errors, including both missed detections and false detections. We introduce four game context features (two based on absolute position and two based on relative position) for describing the current game situation with respect to a pair of mid-level tracks that can be extracted from a varying number of noisy detected player locations \mathcal{O} .

4.1 Absolute Occupancy Map

We describe the distribution of players during a time interval using an occupancy map, which is a spatial quantization of the number of detected players, so that we get a description vector of constant length regardless of missed and false detections. We also apply a temporal averaging filter of 1s on the occupancy map to reduce the noise from detections. The underline assumption is that players may exhibit different motion patterns under different spatial distributions. For example, a concentrated

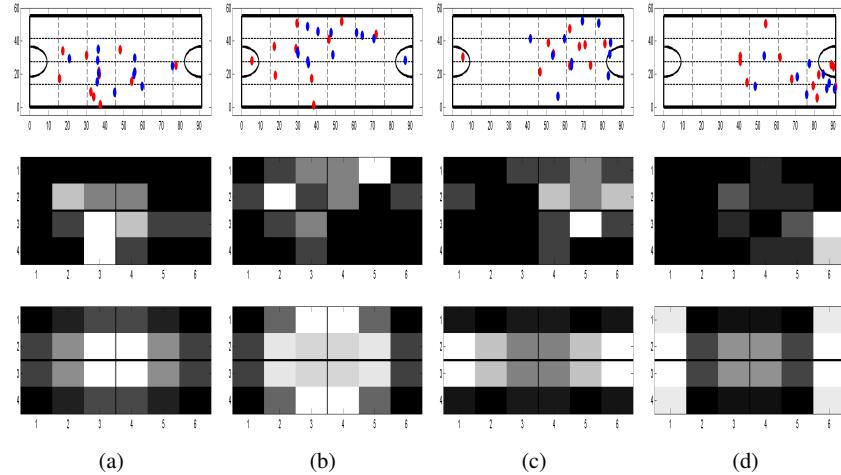


Fig. 6 Absolute Occupancy Map. Four clusters are automatically obtained via K-means: (a) center-concentrated, (b) center-diffuse, (c) goal, (d) corner. The rows show: noisy detections (top), estimated occupancy map (middle), and the corresponding cluster center (bottom), which is symmetric horizontally and vertically.

distribution may indicate a higher likelihood of abrupt motion changes, and smooth motions are more likely to happen during player transitions with a spread-out distribution.

We compute a time-averaged player count for each quantized area. We assume the same distribution could arise regardless of which team is attacking, implying a 180° symmetry in the data. Similarly, we assume a left/right symmetry for each team, resulting in a four-fold compression of the feature space.

Similar to visual words, we use K-means clustering to identify four common distributions (see Fig. 6) roughly characterized as: center concentrated, center diffuse, goal, and corner.

When evaluating the affinity for $\Gamma_i \rightarrow \Gamma_j$, we average the occupancy vector over the time window (t_{i1}, t_{j0}) and the nearest cluster ID is taken as the context feature of absolute occupancy $\theta_{ij}^{(A)} = k \in \{1, \dots, K\}$.

The spatial quantization scheme may be tuned for a specific sport, and does not necessarily have to be a grid.

4.2 Relative Occupancy Map

The relative distribution of players is often indicative of identity [19] or *role* [17]. For example, a forward on the right side typically remains in front and to the right of teammates regardless of whether the team is defending in the back-court or attacking

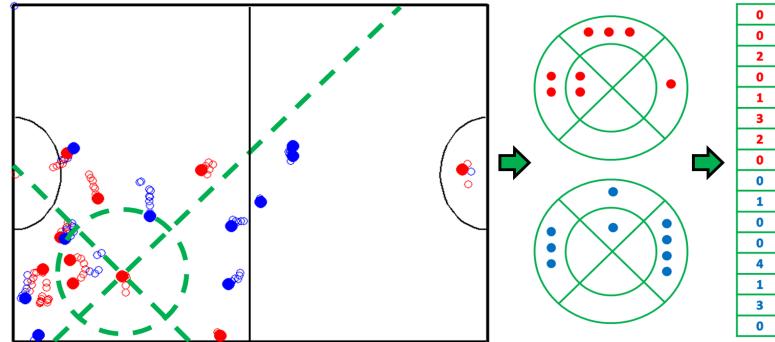


Fig. 7 Relative Occupancy Map. The quantization scheme is centered on a particular low-level tracklet Γ_i at time t . The same-team distribution and opponent distribution are counted separately.

in the front-court. Additionally, the motion of a player is often influenced by nearby players.

Therefore, we define a relative occupancy map specific to each low-level tracklet Γ_i which quantizes space similarly to the shape context representation: distance is divided into two levels, with a threshold of 4 meters, and direction into four bins (see Fig. 7). The per-team occupancy count is then normalized to sum to one for both the inner circle and outer ring. Like absolute occupancy maps, we cluster the 16 bin relative occupancy counts (first 8 bins describing same-team distribution, last 8 bins describing opponent distribution) into a finite set of roles using K-means.

For each pair of possible successive $\Gamma_i \rightarrow \Gamma_j$ mid-level tracks, we extract the occupancy vector \mathbf{v}_i and \mathbf{v}_j , with cluster ID k_i, k_j , from the end tracklet of Γ_i and the beginning tracklet of Γ_j . We also compute the Euclidian distance of $d_{ij} = |\mathbf{v}_i - \mathbf{v}_j|_2$. Intuitively, a smaller d_{ij} indicates higher likelihood that Γ_j is the continuation of Γ_i . The context feature of relative occupancy is the concatenation of $\theta_{ij}^{(R)} = (d_{ij}, k_i, k_j)$.

4.3 Focus Area

In team sports such as soccer or basketball, there is often a local region with relatively high player density that moves smoothly in time and may indicate the current or future location of the ball [13, 22]. The movement of the focus area in absolute coordinates also strongly correlates to high-level events such as turnovers. We assume the movement of individual players should correlate with the focus area over long time periods, thus this feature is useful for associations $\Gamma_i \rightarrow \Gamma_j$ with large temporal gaps (when the motion prediction is also less reliable). For example, mid-level trajectory Γ_i in Fig. 8 is more likely to be matched to Γ_{j1} with a constant velocity



Fig. 8 Focus Area. Kinematic constraints are less reliable across larger time windows. Because player motions are globally correlated, the affinity of two mid-level tracks over large windows should agree with the overall movement trend of the focus area.

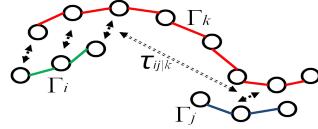


Fig. 9 Chasing. If Γ_i and Γ_j both correlate to a nearby track Γ_k , there is a higher likelihood that Γ_j is the continuation of Γ_i .

motion model. However, if the trajectory of the focus area is provided as in Fig. 8, it is reasonable to assume $\Gamma_i \rightarrow \Gamma_{j2}$ has a higher affinity than $\Gamma_i \rightarrow \Gamma_{j1}$.

We estimate the location and movement of the focus area by applying meanshift mode-seeking to track the local center of mass of the noisy player detections. Given a pair of mid-level tracks with hypothesized continuity $\Gamma_i \rightarrow \Gamma_j$, we interpolate the trajectory within the temporal window (t_{i1}, t_{j0}) and calculate the variance of its relative distance to the trajectory of the focus area σ_{ij} . We also extract the average speed of the focus area v_f during the time window, which describes the momentum of the global motion. The focus area context feature is thus set as $\theta_{ij}^{(F)} = (\sigma_{ij}, v_f)$.

4.4 Chasing Detection

Individual players are often instructed to follow or *mark* a particular opposition player. Basketball, for example, commonly uses a one-on-one defense system where a defending player is assigned to follow a corresponding attacking player. We introduce chasing (close-interaction) links to detect when one player is marking another. If trajectories Γ_i and Γ_j both appear to be following a nearby reference trajectory Γ_k , there is a strong possibility that Γ_j is the continuation of Γ_i (assuming the mid-level track of the reference player is continuous during the gap between Γ_i and Γ_j , see Fig. 9).

We identify chasing links by searching for pairs of low-level tracklets (γ_i, γ_k) that are less than 2 meters apart and moving along similar directions (We use the angular threshold of 45° during the experiment). Let $\tau_{ij|k}$ be the temporal gap between γ_i 's last link with γ_k and γ_j 's first link with γ_k , and $\tau_{ij|k} = \infty$ when there are no links

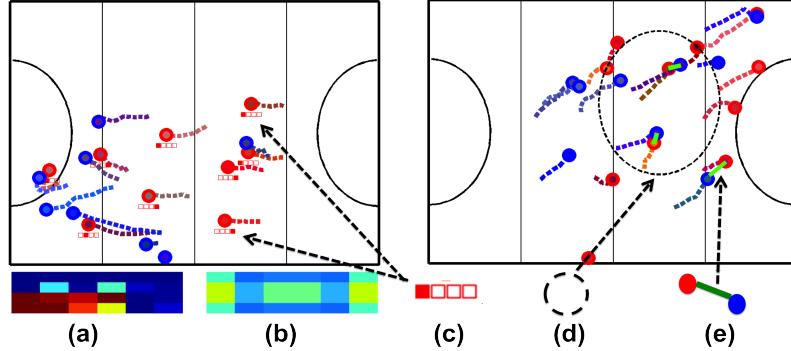


Fig. 10 Visualization of all game context features, where detections of different teams are plot as blue and red dots. (a) the computed occupancy map based on noisy player detection; (b) the corresponding cluster center of the global occupancy map; (c) indicator of the role categories of each player based on its local relative occupancy map; (d) the focus area; (e) chasing between players from different teams.

between either Γ_i or Γ_j and Γ_k . The chasing continuity feature $\theta_{ij}^{(C)}$ that measures whether trajectories Γ_i and Γ_j are marking the same player is given by

$$\theta_{ij}^{(C)} = \min_{k=1, \dots, /i,j} \{\tau_{ij|k}\}. \quad (9)$$

Intuitively, the likelihood that (Γ_i, Γ_j) correspond to the same player increases as $\theta_{ij}^{(C)}$ decreases.

A visualization of all game context features is shown in Fig. 10. Features based on position appear on the left. The computed instantaneous occupancy map (a) and its corresponding cluster center (b) are shown underneath. Each mid-level track is assigned to one of four role categories based on its local relative occupancy map as indicated in (c). Features based on motion appear on the right. The focus area (d) is shown as a dashed ellipse, and (e) detected correlated movements between opposition players is illustrated using green lines.

5 Context-Conditioned Motion Models

Although we have introduced a set of context features $\theta = \{\theta^{(A)}, \theta^{(R)}, \theta^{(F)}, \theta^{(C)}\}$, it is nontrivial to design a single fusion method for generating the final motion likelihood score. For instance, each game context feature may have varying importance between different sports. For example, the chasing-based feature is less important in sports where one-on-one defense is less common. To make our framework general across different sports, we use a purely data-driven approach to learn a motion

Table 1 Our context-conditioned motion models employ traditional kinematic features [9], such as the position extrapolation error when joining two tracks across a temporal gap.

Feature ID	Symbol	Meaning
1	t_g	temporal gap duration
2	e_0	const-position extrapolation error
3	e_1	const-velocity extrapolation error
4	e_2	const-acceleration extrapolation error
5	$\Delta\mathbf{v}$	change in velocity
6	σ_{ij}	motion correlation with focus area
7	v_f	velocity of focus area
8	d_{ij}	consistency of local relative player distribution
9-12	k_i, k_j	local occupancy cluster encoded into binary digit
13-14	$\theta_{ij}^{(A)}$	global occupancy cluster encoded into binary digit
15	$\theta_{ij}^{(C)}$	chasing gap

model which uses both the traditional kinematic features as well as our game context features.

The kinematic features $\mathbf{K} = \{t_g, e_0, e_1, e_2, \Delta\mathbf{v}\}$ describe the motion smoothness between two successive mid-level tracks $\Gamma_i \rightarrow \Gamma_j$, and is based on the distance error with extrapolate constant position, constant velocity and constant acceleration models respectively. Additionally, the velocity change in velocity is also included (see Tab. 1).

We generate training data by extracting kinematic features $f_{ij}^{(K)}$ and game context features θ_{ij} for all pairs of mid-level tracks (Γ_i, Γ_j) that have a temporal gap $t_{g|i \rightarrow j} \in [0, 60)$ seconds. Using ground truth tracking data, we assign binary labels $y_{ij} \in \{1, 0\}$ indicating whether the association is correct or not (two trajectories belonging to the same ground truth player identity). However the total number of incorrect associations are usually much more than the correct ones. To avoid severely unbalanced training data, we only select a subset of hard negative examples that has high motion smoothness.

A random forest containing 500 decision trees is then trained to learn the mapping $C(f_{ij}^{(k)}, \theta_{ij}) \rightarrow y_{ij}$. A random forest is robust against the overfitting that might occur when using limited training data via bootstrapping, especially when the data is not easily separable due to association ambiguity in the real world. More importantly, by recursively splitting the data with random subsets of features, the random forest model automatically optimizes local adaptivity, *i.e.*, decisions for linking pairs of tracks having small or large temporal gaps may be split at different levels and handled with different feature sets. As confirmed in our experiments (see Sec.6), the occupancy-feature is more effective at handling short-term association (when feature t_g is small) and the chasing-feature is more important in connecting trajectories with long temporal gaps (t_g is big).

During the testing stage, the average classification score across all trees provides a continuous affinity score to approximate $P(\Gamma_i \rightarrow \Gamma_j | \theta) = C(f_{ij}^{(K)}, \theta_{ij})$ in Eqn. 5.

6 Experiments

We validate our framework on two sports: field hockey with 20 players and basketball with 10 players. Player detection is transformed from multiple calibrated views using the method in [6] with frame rates of 30 (hockey) and 25 (basketball), respectively. We use simple RGB-based color histogram classifiers to estimate the confidence score $a_i \in [0, 1]$ of track i belonging to home team or the away team. We also discard tracks likely to correspond to the referees and goalies.

6.1 Baseline Models and Evaluation metrics

To verify the contribution of the various GCFs, we construct 5 models for a quantitative comparison. All models apply hierarchical association and start with the same set of mid-level tracks $\{\Gamma\}$. The only difference between the models is the motion affinity used during the final association stage. Model 1 (K) only uses kinematic features ($f^{(K)}$) for training, which is equivalent to the combined algorithm of [9, 15, 30]. Models 2-4 use focus area features (F), chasing related features (C) and occupancy feature ($A + R$), respectively in addition to motion-smoothness features. Model 5 uses all features ($f^{(K)}, \theta$).

We have also examined other features for describing aspects of game context, such as variance of track velocity or team separability. However we found these features to be less effective than the ones described in Sec. 4.

Three errors are commonly evaluated in the multi-target tracking literature: (1) the number of incorrect associations N_{err} , (2) the number of missed detections N_{miss} , and (3) the number of false detections N_{fa} . The Multiple Object Tracking Accuracy measure $MOTA = 1 - (N_{err} + N_{miss} + N_{fa})/N$ combines all three errors with equal weighting. However the equal weighting de-emphasizes N_{err} in a hierarchical association framework with a high frame rate. Therefore, we report the individual error sources and normalize for the situation of a known fixed number of objects: N_{err}^* is an average count of incorrect ID associations per minute per player; P_{miss} and P_{fa} are the proportion of missed and false mid-level trajectory segments of \mathcal{T}_n as compared to the ground truth, ranging from 0 to 1.

In addition to overall tracking performance, we also evaluate in isolation the high-level association stage $\{\Gamma\} \rightarrow \mathcal{T}$, which is the key part of our framework. We report association precision and recall rate, where precision = $N_{TP}/(N_{TP} + N_{FA})$, and N_{TP}, N_{FA} are correct/incorrect number of associations of $\Gamma_i \rightarrow \Gamma_j$. We define recall = $1 - T_{miss}/T_{gap}$, where T_{gap} is the accumulation of temporal gaps t_{gap} between high-level associations, and T_{miss} is the total length of mid-level tracks Γ_i .

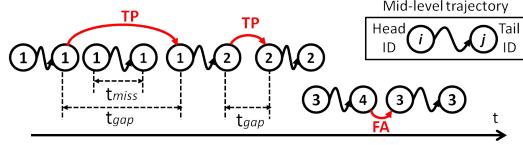


Fig. 11 Demonstration of evaluation metrics for high-level association (red).

being missed. The motivation is to exclude miss-associations in previous stages. An illustration of these metrics is given in Fig. 11. Finally, we also report the statistics of average length temporal gap \bar{t}_{gap} being correctly associated during the high-level association, which reflects the algorithm’s ability to associate trajectories with long-term misses.

6.2 Field Hockey Dataset

We generated and labeled 6 field hockey sequences for a total length of 29 minutes, from 3 games played by different teams . The average player detection miss and false-alarm rates are 14.0% and 10.3%, respectively, or the multi-target detection accuracy MODA = $1 - (N_{miss} + N_{fa})/N = 0.75$. Our first experiment uses as much training data as possible: testing one sequence and using the remaining five for training.

The introduction of each individual *GCF* achieves better performance, and using all *GCFs* generally produces the best performance (see Tab. 2).

According to Tab. 2, all methods are good in terms of low false-alarm rate. Thus the major difference in their performances is reflected in the terms for incorrect association N_{err}^* and miss association P_{err} .

We can also introduce a weighting w_m on motion likelihood relative to the appearance likelihood into the objective function of Eqn. 1, where w_m plays an essential role in the trade-off between miss-associations and false associations:

$$\log P(\mathcal{T}|\mathcal{O}, \theta) = \log P(O|\mathcal{T}) + w_m \cdot \log P(\mathcal{T}|\theta) + c. \quad (10)$$

Instead of the default setting of $w_m = 1$, a lower weight for the motion likelihood ($w_m < 1$) gives higher priority to optimizing the observation likelihood $P(O|\Gamma)$, which prefers to have fewer missing players. On the other hand, a higher weighting $w_m > 1$ encourages smoother motions and results in fewer false alarms but more missed detections. As we vary w_m from 0.2 to 3, the trade-off curves are plotted in Fig. 12(a).

We also conduct an experiment studying the cross-game-generalization of the *GCFs*. Instead of testing 1 sequence trained on the other 5, we perform all pairwise combinations (30 in total) of 1 sequence training with 1 other sequence testing.

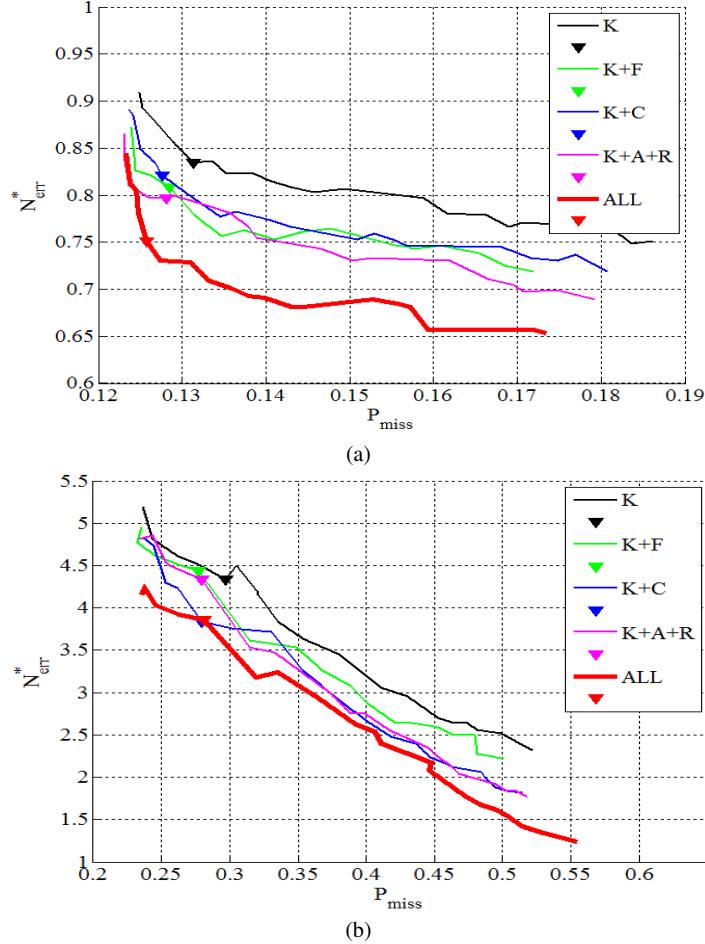


Fig. 12 Trade-off curve between P_{miss} and N_{err}^* for (a) field hockey sequences and (b) basketball sequences. N_{err}^* is averaged association error per minute per person. The triangle marks indicate the default operating point ($w_m = 1$ in Eqn.10). Our proposed method using all GCFs achieves more than 10% of improvements on both cases.

We then evaluate the resulting statistics for same-game learning and different-game learning respectively, as summarized in Tab. 4.

It can be seen that the introduction of GCFs again improves the result both in the case of same-game and different-game learning, yet this time the amount of training data used is much smaller (4 minutes on average). On the other hand, same-game learning outperforms cross-game learning in terms of generalization, which matches our intuition that the game context features are more similar within the

Table 2 Quantitative evaluations on field hockey dataset

Features	N_{err}^*	P_{miss}	P_{fa}	precision	recall	\bar{t}_{gap} (sec)
Kinematic	.84	.131	.032	.69	.97	3.68
Kinematic + Focus	.81	.129	.032	.71	.97	3.97
Kinematic + Chasing	.82	.128	.032	.70	.97	3.56
Kinematic + Occupancy	.80	.128	.033	.71	.98	3.62
All	.75	.126	.031	.75	.97	3.95

Table 3 Quantitative evaluations on basketball dataset

Features	N_{err}^*	P_{miss}	P_{fa}	precision	recall	\bar{t}_{gap} (sec)
Kinematic	4.33	.30	.027	.65	.99	3.26
Kinematic + Focus	4.43	.280	.031	.67	.99	3.99
Kinematic + Chasing	.380	.280	.024	.71	.99	5.09
Kinematic + Occupancy	4.32	.280	.025	.68	.99	3.60
All	3.81	.281	.018	.71	.99	3.81

Table 4 Comparison of same/cross game learning (Hockey)

Features	Same Game			Different Game		
	N_{err}^*	P_{miss}	P_{fa}	N_{err}^*	P_{miss}	P_{fa}
Kinematic	0.810	0.141	0.034	1.240	0.130	0.036
Kinematic + Focus	0.840	0.133	0.034	1.230	0.125	0.034
Kinematic + Chasing	0.780	0.134	0.034	1.190	0.127	0.035
Kinematic + Occupancy	0.780	0.136	0.034	1.170	0.126	0.034
All	0.770	0.134	0.033	1.140	0.124	0.034

same game with the same players, *e.g.*, the team distribution/tactics and the velocity/acceleration of players are more consistent.

6.3 Basketball Dataset

We also conduct the same evaluation on a basketball dataset of 4 sequences for a total length of more than 5 minutes. The dataset is more challenging due to a higher player density and less training data. Each sequence is tested while using the other 3 sequences for training. The average testing performance is reported in the trade-off curve of Fig. 12(b) and Tab. 2. As can be seen, the chasing feature is much more important for basketball sequences, indicating that one-on-one defensive situations occur more frequently in basketball than field hockey.

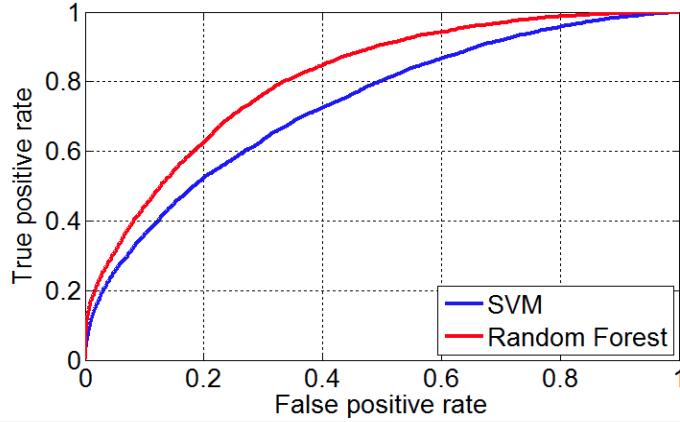


Fig. 13 RoC Curve comparison between random forest (red) and SVM (blue) classifier.

6.4 Classifier

In addition to random forests, we also examined the performance of linear SVMs for representing context-conditioned motion models. We utilize all the training data for a 5-fold cross-validation scheme and evaluate the RoC curves (see Fig. 13). The experiments suggest the underlying motion models are not linear functions of the features because random forests outperform SVMs in all parameter configurations.

6.5 Feature Importance

When training a random forest, we can also calculate the average decrease in Gini-index after removing each feature, which indicates the significance of the feature as shown in Fig. 14. In general, kinematic features¹ have greater significance than context features, which makes sense because kinematic features have direct correlation with the track association affinity. The constant velocity model (feature 3) is most important among kinematic features. However, in hockey, the consistency of the local player distribution (feature 8) is more important than any kinematic feature. Features 9 – 14 have low significance because they are binary features with each bit containing very limited information. Furthermore, game context features are individually more important in the hockey sequence than in the basketball sequence.

¹ refer to Tab. 1 for the meaning of each feature number

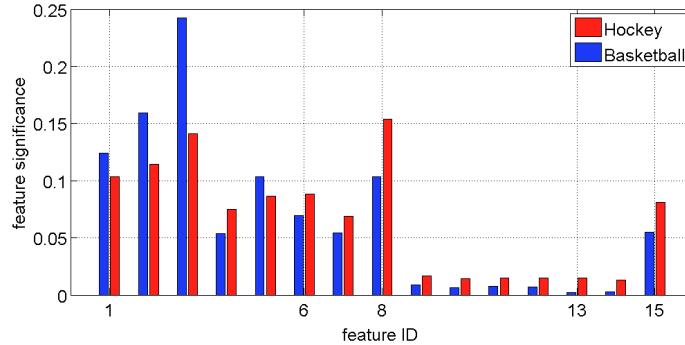


Fig. 14 Normalized feature significance in hockey (red) and basketball (blue) sequences.

7 Summary

In this work, we use hierarchical association to track multiple players in team sports over long periods of time. Although the motions of players are complex and highly correlated with teammates and opponents, the short-term movement of each player is often reactive to the current situation. Using this insight, we define a set of game context features and decompose the motion likelihood of all players into independent per-player models contingent on game state. Higher-order inter-player dependencies are implicitly encoded into a random decision forest based on track and game context features. Because the conditioned model decomposes into pairwise terms, our formulation remains efficiently solvable using cost flow networks. We validate our approach on 30 minutes of international field hockey and 10 minutes of college basketball. In both sports, motion models conditioned on game context features consistently improve tracking results by more than 10%.

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