

Stat 207 HW8

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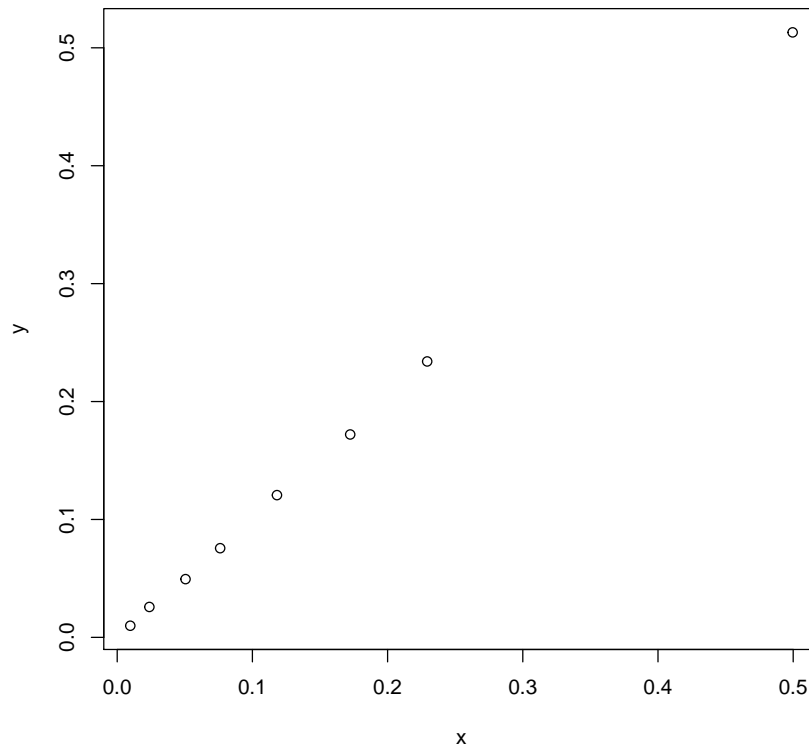
March 9, 2015

1 14.28

```
(a) dat = read.table("CH14PR14.txt")
names(dat) = c("Y", "X1", "X2", "X3")
logit = glm(Y ~ X1 + X2, data = dat, family = "binomial")
logitv = logit$fitted.values
dat = dat[order(logitv), ]
a = rep(1:8, each = 20)
a = a[-1]
b = split(dat, a)
Oj1 = sapply(b, function(x){sum(x[[1]])})
Ej1 = sapply(split(sort(logitv), a), sum)
Oj0 = sapply(b, function(x){length(x[[1]])-sum(x[[1]])})
Ej0 = sapply(b, function(x){length(x[[1]])})-Ej1
rbind(Oj1, Ej1, Oj0, Ej0)

##           1           2           3           4           5           6           7
## Oj1  0.000000  1.000000  0.000000  2.000000  1.000000  8.000000  2.000000
## Ej1  0.187472  0.5159059  0.9878718  1.512501  2.412695  3.44151  4.680125
## Oj0 19.000000 19.000000 20.000000 18.000000 19.000000 12.00000 18.000000
## Ej0 18.812528 19.4840941 19.0121282 18.487499 17.587305 16.55849 15.319875
##           8
## Oj1 10.000000
## Ej1 10.261919
## Oj0 10.000000
## Ej0  9.738081

x = sapply(split(sort(logitv), a), median)
y = Ej1/sapply(b, function(x){length(x[[1]])})
plot(x, y)
```



The plot seems to be linear, it's consistent with a response function of monotonic sigmoidal shape.

```
(b) X.squ = sum((rbind(Oj1, Oj0)-rbind(Ej1, Ej0))^2/rbind(Ej1, Ej0));X.squ
## [1] 12.11578
```

$$H_0: E(Y) = [1 + \exp(-\beta_0 - \beta_1 X_1 - \beta_2 X_2)]^{-1}$$

$$\text{VS. } H_1: E(Y) \neq [1 + \exp(-\beta_0 - \beta_1 X_1 - \beta_2 X_2)]^{-1}$$

$$X^2 = \sum_j \sum_k \frac{(O_{jk} - E_{jk})^2}{E_{jk}} = 12.11578$$

we can reject H_0 if $X^2 > \chi^2(0.95, 8 - 3) = 11.0705$, otherwise reject H_1

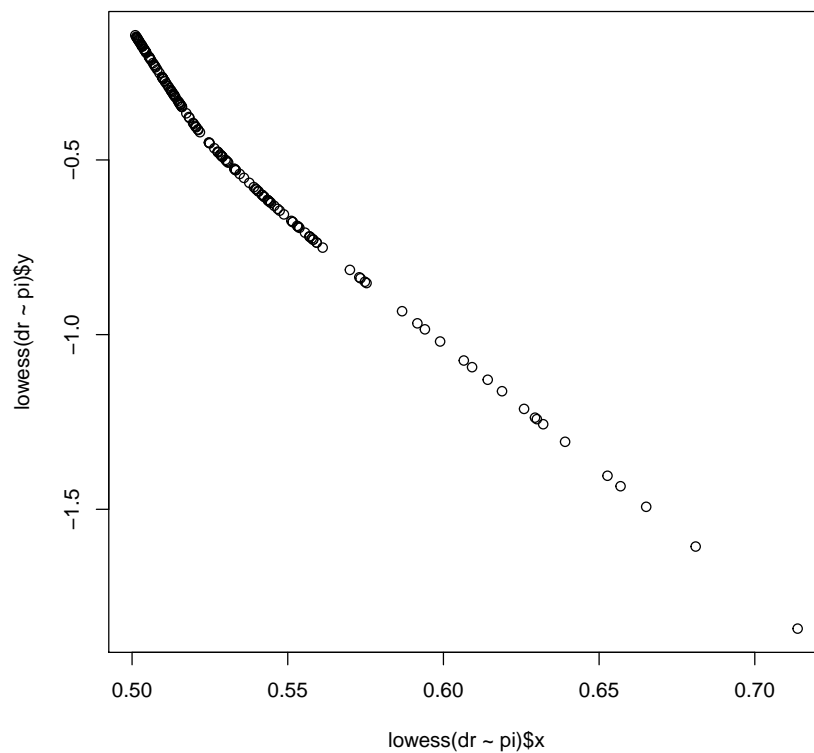
so that reject H_0 because $X^2 > 11.0705$, Pvalue is 0.03323531.

```
(c) p = summary(logit)
     dr = p$deviance.resid;dr
```

##	1	2	3	4	5	6
##	-0.54602312	-0.51373259	1.15260237	-0.17517944	-0.18924467	-0.69185916
##	7	8	9	10	11	12
##	-0.18367445	-0.73777334	-0.47288089	-0.35855218	-0.62785882	-0.33244658
##	13	14	15	16	17	18
##	-0.49058859	-0.23345211	-0.11467682	-0.59485941	-0.85058862	-0.16214474
##	19	20	21	22	23	24
##	-0.22270681	-0.13939895	-0.34822874	-0.28168748	-0.24470112	-0.09084127
##	25	26	27	28	29	30
##	-0.50595282	-0.10798927	-0.69416101	-0.26781191	-0.83451488	-0.68109529
##	31	32	33	34	35	36
##	-0.32519247	-0.45851859	-0.92893094	-0.53211206	-0.33897782	-0.16073331
##	37	38	39	40	41	42
##	0.74365705	1.38971984	-0.61999873	-0.41310085	1.89079347	-0.35679774
##	43	44	45	46	47	48
##	1.71627562	-0.33244658	-0.72647043	-0.32961191	2.84304938	-0.69729005
##	49	50	51	52	53	54
##	-0.10990097	-0.18854516	-0.33530428	-0.21617318	-0.72171075	-0.59079687
##	55	56	57	58	59	60
##	2.23462510	-0.24806553	-0.42012390	-0.28029236	-1.30166964	2.00689554
##	61	62	63	64	65	66
##	1.10849910	-0.15933400	-0.68109529	-0.32124458	-0.10387671	-1.01249619
##	67	68	69	70	71	72
##	-0.12087236	-0.36424502	-0.62497861	1.86302691	-0.23960963	-0.48311783
##	73	74	75	76	77	78
##	-0.38006335	1.82187884	-1.44789883	-0.84697959	-0.83715068	-0.35248880
##	79	80	81	82	83	84
##	-0.58403320	-0.45631940	-0.35855218	1.94481807	-0.72970910	-0.53462959
##	85	86	87	88	89	90
##	-0.36293048	-1.15124676	-0.62286031	-0.65995923	2.34647993	-0.81345988
##	91	92	93	94	95	96
##	1.12311780	-0.17737218	1.27437628	-0.55762351	-0.49703364	-0.30029670
##	97	98	99	100	101	102
##	-1.06493877	-0.71059309	0.95982139	-0.33201751	-0.67802259	1.90510957
##	103	104	105	106	107	108
##	-0.45631940	-0.42162878	-0.12900059	-0.17648324	-0.19257415	-0.14876060
##	109	110	111	112	113	114
##	1.82969794	-0.32283373	-0.53091804	-0.13409993	-0.61013597	-0.72086770
##	115	116	117	118	119	120
##	-0.50887859	-0.61087534	-0.16356841	-0.16073331	-0.12675915	-0.49293406
##	121	122	123	124	125	126
##	-0.14931436	-0.21351074	1.84295971	-0.40421564	1.40875849	-0.30814469

```
##          127          128          129          130          131          132
## -0.40618549 -0.49528940 -0.51130056 -0.73777334 -0.28412789 -0.58877426
##          133          134          135          136          137          138
## -0.49703364 -1.22831729 -0.31850052  1.89079347 -0.29409388 -0.69416101
##          139          140          141          142          143          144
## -0.20982689 -0.48482421 -0.61925036 -0.28658859 -0.39176366 -0.48311783
##          145          146          147          148          149          150
## -1.11866815 -0.31307905 -0.35422338 -0.28065744 -0.72970910 -0.20982689
##          151          152          153          154          155          156
## -0.41110024 -0.42725579 -0.28554134 -0.38936035 -0.29263987 -0.25647527
##          157          158          159
##  0.42476809  0.86785144  1.67453806

pi = exp(logitv)/(1+exp(logitv))
plot(lowess(dr ~ pi))
```



2 14.39

```
(a) dat = read.table("CH14PR39.txt")
names(dat) = c("Y", "X1", "X2", "X3", "X4")
poi = glm(Y ~ ., data = dat, family = "poisson")
summary(poi)

##
## Call:
## glm(formula = Y ~ ., family = "poisson", data = dat)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.1854  -0.7819  -0.2564   0.5449   2.3626
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.489467   0.336869   1.453  0.14623
## X1          -1.069403   0.133154  -8.031 9.64e-16 ***
## X2           -0.046606   0.119970  -0.388  0.69766
## X3            0.009470   0.002953   3.207  0.00134 **
## X4            0.008566   0.004312   1.986  0.04698 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##      Null deviance: 199.19  on 99  degrees of freedom
## Residual deviance: 108.79  on 95  degrees of freedom
## AIC: 377.29
##
## Number of Fisher Scoring iterations: 5
```

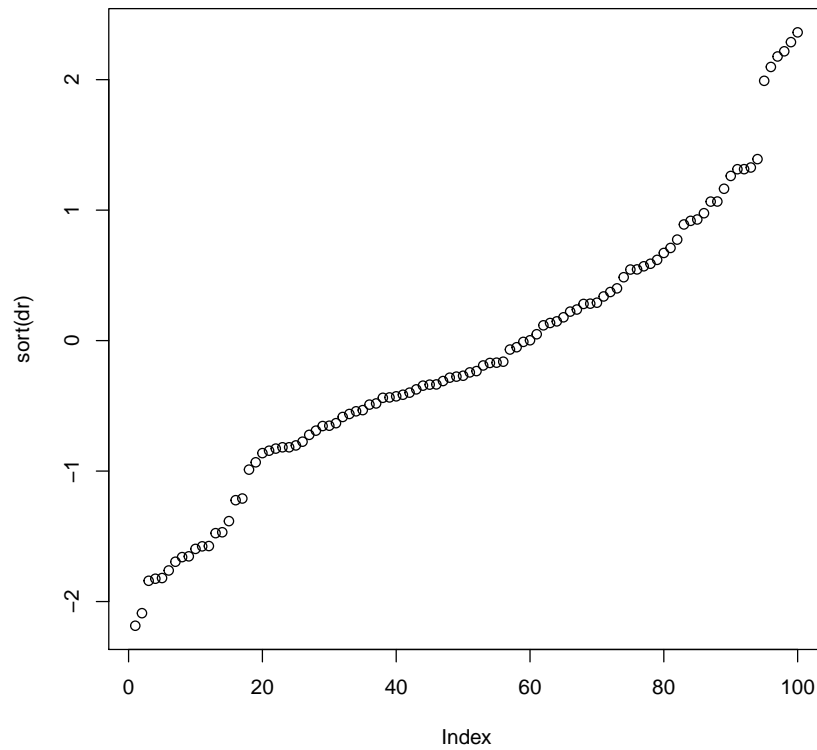
$$\begin{aligned}
 b_0 &= 0.489467 & s(b_0) &= 0.336869 \\
 b_1 &= -1.069403 & s(b_1) &= 0.133154 \\
 b_2 &= -0.046606 & s(b_2) &= 0.119970 \\
 b_3 &= 0.009470 & s(b_3) &= 0.002953 \\
 b_4 &= 0.008566 & s(b_4) &= 0.004312
 \end{aligned}$$

$$\mu = \exp(0.489467165 - 1.069402551X_1 - 0.046606063X_2 + 0.009469987X_3 + 0.008565829X_4)$$

```
(b) p = summary(poi)
     dr = p$deviance.resid;dr

##           1           2           3           4           5
## -0.481563003 -0.632820229  0.485684782 -1.819828480  0.238302497
##           6           7           8           9          10
## -0.427206484 -1.574566470 -1.694831446 -0.190494237  0.372414202
##          11          12          13          14          15
##  0.290568448  0.917894609  0.618945304  0.710218879 -0.169117683
##          16          17          18          19          20
##  0.774212615  1.313966589  0.976533023 -0.284161510  0.281078819
##          21          22          23          24          25
##  0.671826155 -0.309823439 -0.585274974 -1.659501048 -1.653549229
##          26          27          28          29          30
##  0.545695411 -2.089197070 -1.825162331  0.283158353  1.066089357
##          31          32          33          34          35
##  0.338495236 -0.414912576 -0.276062096  2.097482057 -0.373770021
##          36          37          38          39          40
## -2.185378873  2.217373787 -0.269196968 -1.468917177 -0.490568471
##          41          42          43          44          45
## -0.243689445  2.287659410 -0.435460637 -0.171363655 -1.596370920
##          46          47          48          49          50
##  0.178588604 -1.840819930  1.313615961 -0.233624875 -1.576829772
##          51          52          53          54          55
## -0.652282150  2.177759158 -0.862317963 -1.223172388 -0.161565009
##          56          57          58          59          60
## -0.438457085 -0.817656201  0.002451800 -0.690070466 -0.932653968
##          61          62          63          64          65
## -0.532725622  0.400431092  1.326821516  0.569505372  0.589162479
##          66          67          68          69          70
##  0.222202292  1.991163395  0.134986579  1.163783804  0.890153115
##          71          72          73          74          75
## -0.398888037 -0.335860573 -0.843868576  1.065308951 -0.068529260
##          76          77          78          79          80
## -1.210403791  0.928316209 -0.803424449 -0.050707876 -0.817762385
##          81          82          83          84          85
##  0.544674549 -1.761754891 -0.562223893 -0.541853422  0.147418906
##          86          87          88          89          90
## -0.009840529 -0.337422359 -0.774688780 -1.383658148 -0.654757856
##          91          92          93          94          95
## -1.475851025 -0.722343228  2.362545161  1.391144309  1.262066676
##          96          97          98          99         100
##  0.117099736 -0.827717570 -0.345172951  0.048822205 -0.988857456

plot(sort(dr))
```

The points whose devaiance residual near 100 seem to be outliers.

```
(c) logLik(poi)
## 'log Lik.' -183.6439 (df=5)

poiR = glm(Y ~ .-X2 , data = dat, family = "poisson")
logLik(poiR)
## 'log Lik.' -183.7194 (df=4)

qchisq(1-0.05, 5-4)
## [1] 3.841459

pchisq(0.151, 1, lower.tail = FALSE)
## [1] 0.6975815
```

$H_0: \beta_2 = 0$
 VS. $H_1: \beta_2 \neq 0$
 The full model: $\mu = \exp(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4)$
 $\ln(L(F)) = -183.6439$
 The reduced model: $\mu = \exp(\beta_0 + \beta_1 X_1 + \beta_3 X_3 + \beta_4 X_4)$
 $\ln(L(R)) = -183.7194$
 $G^2 = -2(\ln(L(R)) - \ln(L(F))) = 0.151$
 we can reject H_0 if $G^2 > \chi^2(1 - 0.05, 5 - 4) = 3.8415$, otherwise reject H_1
 so that reject H_1 because $G^2 < 3.8415$,
 therefore, X_2 can be dropped from the regression model, and the P-value
 is 0.6975815. And the result is the same as the result we get in (b).

(d) `summary(poiR)`

```
##
## Call:
## glm(formula = Y ~ . - X2, family = "poisson", data = dat)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.2152  -0.7512  -0.2594   0.5830   2.2893
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.443890   0.317289   1.399  0.16181
## X1          -1.077770   0.131415  -8.201 2.38e-16 ***
## X3           0.009471   0.002957   3.203  0.00136 **
## X4           0.008979   0.004190   2.143  0.03209 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##      Null deviance: 199.19  on 99  degrees of freedom
## Residual deviance: 108.94  on 96  degrees of freedom
## AIC: 375.44
##
## Number of Fisher Scoring iterations: 5

b1 = -1.077770
s1 = 0.131415
z = qnorm(1-0.05/2)
c(b1-s1*z, b1+s1*z)
```

```
## [1] -1.3353387 -0.8202013
```

From summary(poiR), we get $s(b_1) = 0.131415$, $b_1 = -1.077770$, based on $b_k \pm z(1 - \alpha/2)sb_k$, we conclude that we are 95 % confident that β_1 is between -1.3353387 and -0.8202013. Because the confidence interval is smaller than 0, aerobic exercise reduce the frequency of falls when controlling for balance and strength.

3 14.44

4 14.45

5 14.46

$$E(Y) = [1 + \exp(-\beta_0 - \beta_1 X_1 - \beta_2 X_2 - \beta_3 X_1 X_2)]^{-1}$$

$$\pi'(X_1) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_1 X_2$$

$$\pi'(X_1 + 1) = \beta_0 + \beta_1(X_1 + 1) + \beta_2 X_2 + \beta_3(X_1 + 1)X_2$$

$$\pi'(X_1 + 1) - \pi'(X_1) = \ln(\text{oddsratio}) = \beta_1 X_1 + \beta_3 X_1 X_2$$

Hence the odds ratio for X_1 is $\exp(\beta_1 X_1 + \beta_3 X_1 X_2)$ therefore, they are different.

6 14.47