2020年5月23日 星期六 下午8:23

> T.S. dostases - Process - model - pact > p Regression = 20- plot

no-numeric measurement, - 122 ARIX) precess Simulation V order Vocet - arima. Sim - generate data

* * coet - estimate - arima) - (SSE

* * coet - estimate - arima) + (AIC

715a8285 70906ee...

> Thistleton and Sadigov Akaike Information Criterion and Model Quality Week 4

Measuring the Quality of a Model

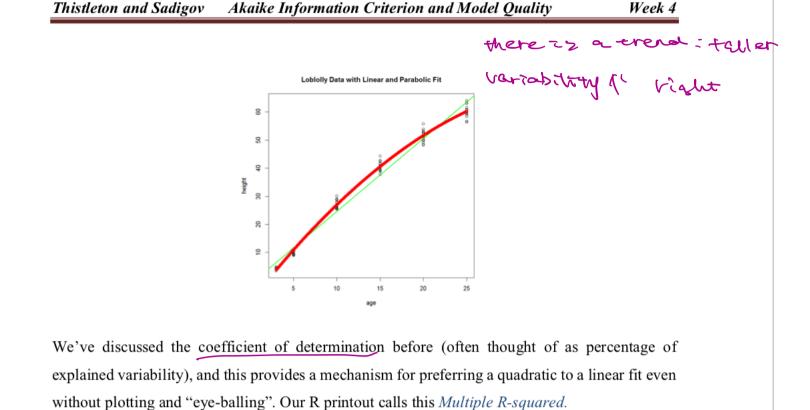
At this point we can do a lot! We understand Autoregressive Processes and Moving Average Processes. We can even, for example using Yule-Walker, estimate the coefficients in an Autoregressive Process. To be honest, we'll usually rely on software for estimating coefficients in models, but it's nice to know how to obtain these techniques anyway because it deepens our general understanding. Moving forward, an important issue in modeling is deciding *which* model we believe describes the

process which generated our data set. This is a more subtle question than it appears at first glance. Time Series data collected in the field doesn't come labelled with the generating process! Unless we have deep subject-matter knowledge, we have to work to determine a good model for our data. As a concrete example, suppose you have time series data and you believe an AR(p) stochastic

process is a good model. We've used the Partial Autocorrelation Function (PACF) to make a guess at the best value for p, the order of the process. Unless the series is quite long and therefore we have a lot of data from which to estimate the PACF coefficients, this is a pretty subjective way to proceed. Also, we will soon look at processes with both AR(p) and MA(q) components (as well as trend and seasonality), called ARMA models. We will even consider data sets exhibiting a trend, bringing us to integrated moving average, or ARIMA models, and finally data sets with a periodic component or seasonality, called SARIMA models. A quick glance at the ACF or PACF won't be enough to judge the quality of a model or to determine the proper values for p and q and the other estimated quantities. Our conundrum is similar to what we find in Regression, so let's first explore these ideas on

model (green) we have a decent fit, but a parabola (red) looks even better. (There is some heteroscedasticity in that the groups of data points develop increasing spread over time, but we'll ignore that for the moment.) We don't always have the luxury of plotting and then choosing (think of a model with several input variables), so we would like to develop a numerical measure of quality. This will help us choose between candidate models. **Practical Time Series Analysis** Page 1

familiar territory. Take a look at our old friend, the Coblolly Pine data. If we fit a straight line



Linear Fit:

Quadratic Fit:

Multiple R-squared: 0.9934, Adjusted R-squarea.

Recall that the multiple R² value gives the proportion of explained variance, while the corresponding adjusted term R² makes each variable "pay a tax" to enter the model. There are of the model of the

Measuring the Quality of a Time Series Model

There are a variety of ways to judge the quality of a time series model. We'll discuss two common ways: SSE and AIC. And, rather than deal with the complexities of a natural process, which can be quite messy to model, let's develop our ideas by generating some data from an AR(p) process so that we already know the order of the model. Then we work to understand how our measures behave.

We'll use the trusty arima.sim() which, of course, simulates autoregressive integrated moving

 $X_t = Z_t + .7X_{t-1} - .2X_{t-2}$

process: list(order = c(2,0,0), ar = c(0.7, -.2)). Generally speaking, as we will explore in greater

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Our code will therefore be:

settle in).

rm(list=ls(all=TRUE))

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average models.

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5

0.7141

-0.2027

1988.660

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0.7137

-0.2016

1996.590

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We start by creating a "list" to specify the parameters as well as include the coefficients of the

We'll start with something simple and simulate the process

detail later this week, by order we mean that we describe the process in terms of its autoregressive order (2), the order of any differencing (0), and then the order of the moving average part (0).

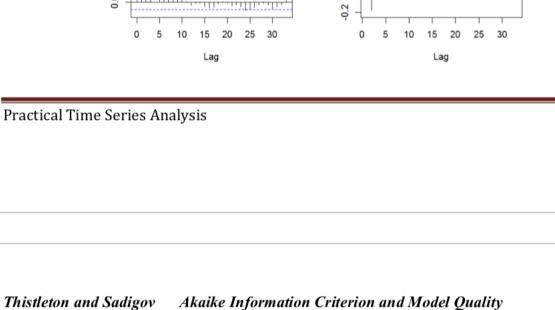
#Roman conquest of Britain set.seed(43) $data = arima.sim(list(order = c(2,0,0), ar = c(0.7, -.2)), \underline{n = 2000})$ Now plot our ACF and PACF (you can plot the time series data itself as well): par(mfrow=c(1,2))

Look at the PACF and feel satisfied that this is consistent with a second order model (some spikes will rise past our dotted line by chance...you can run a much longer time series to see the spikes

ACF of AR Data of Second Order PACF of Time Series

acf(data, main="ACF of AR Data of Second Order") acf(data, type="partial", main="PACF of Time Series")

0.8 0.2



as you tell it the order of the process. For example, I ran this on our data and found: arima(data, order=c(2,0,0), include.mean=FALSE)quess produe

ALL-educated guess Coefficients:

ar2

arima(data, order=c(p, 0, 0), include.mean=F)

0.5970

2072.832

arep

-0.1912

ar1

0.7111

Order of AR process, p

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coefficients

SSE

It may feel a little artificial to estimate the coefficients from the time series data (since we already

know what they are), but we are exploring the behavior of our estimation process and, in general,

Rather than estimate the coefficients by writing our own routines, we can invoke a very useful

command called *arima()*. This command will estimate the parameters of the model for you as long

we won't know the coefficients of a process and will need to estimate them from the data alone.

0.0219s.e. 0.0220 $sigma^2$ estimated as 0.9985: log likelihood = -2836.64, aic = 5679.27These estimates compare quite favorably to our established values $\phi_1 = .7$, $\phi_2 = -.2$.

Again, unless you already happen to know that your AR(p) process is of a certain order, it is hard to look at the ACF and PACF and really say with assurance what the order of the process is. I ran the command

for various values of the order p and produced the following table. So should you! The AIC value

is part of the standard print out. I computed the SSE values as $sum(resid(m)^2)$. More on that later!

0.7111

-0.1912

0.7136

-0.2003

1996.678

0.0128 0.0176 0.0302 -0.0515 -0.0066 0.0629 **AIC** 5680.945 5682.857 5676.917 5751.732 *5*679.274

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If you'd like some code,	here it is!	
SSE=NULL AIC=NULL		
for $(p \text{ in } 1:5) $ { $m = arima$	c(data, order=c(p,0,0), include.mean=FALSE)	

 $X_t = Z_t + \phi_1 X_{t-1} + \phi_2 X_{t-2} + \dots + \phi_p X_{t-p}$

attempts to predict the value of the time series at time t, call it x_t , by using the values at the p

the sum of the squares of the errors (SSE) for each. When I plot the residual squared terms against

preceding time steps. A naïve suggestion would be to try an order, e.g. p=3, estimate the coefficients for this order, and then determine the residual sum of squares for our candidate model. Fita model We could then pick whichever model gives us the lowest aggregate residual sum of squares. We have our 5 models for p=1 through p=5. We can extract the residuals from each, and then look at

the order I obtain the graph on the left.

Using the Residual Sum of Squares The model we are trying to develop

 $SSE[p] = sum(resid(m)^2)$

print(paste(m\$aic, sum(resid(m)^2)))

 $\overline{AIC[p]} = m$ \$aic

print(m\$coef)

SSE plotted on Order of AR(p) Model

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Thistleton and Sadigov Akaike Information Criterion and Model Quality Week 4 par(mfrow=c(1,2))order = c(1,2,3,4,5)plot(SSE~order, main="SSE plotted on Order of AR(p) Model", ylim=c(1800, 2100)) $plot(AIC \sim order, main = "AIC plotted on Order of AR(p) Model", ylim = c(5500, 5800))$ Adding more terms reduces the SSE, but the really big drop occurs when we move from p=1 to

AIC plotted on Order of AR(p) Model

Using the AIC The AIC tries to help you assess the relative quality of several competing models, just like adjusted

 R^2 in linear regression, by giving credit for models which reduce the error sum of squares and at the same time by building in a *penalty for models which bring in too many parameters*. Obviously there's a fair amount of theory behind the exact form of the AIC. This statistic is expressed somewhat differently in various software packages, depending on how they adjust the formula (which is based on maximum likelihood estimation). The version appearing in Akaike's 1974

p=2. After that, the enhancements are really much more modest. There seems to be a tradeoff

between added complexity and a meager diminution of the SSE. On the right I've also plotted the

AIC against the order, as discussed below. Based on these plots we would select a 2nd order model.

 $AIC = \underbrace{\log(\hat{\sigma}^2)}_{\text{5.16}} + \underbrace{\frac{n+2\cdot p}{n}}_{\text{7.16}}, \quad \text{where} \quad \hat{\sigma}_{\text{1}}^{2} = \underbrace{\frac{SSE}{n}}_{\text{7.16}}$ While different authors and software packages use varying forms of the AIC, often eliminating constant terms that are the same for various candidate models. In the end, we are looking to compare the AIC for a variety of candidate models and as long as we can make relative

comparisons we are OK. We prefer a model with a lower AIC. In the current case, that means

we'd believe a two parameter model is the best (though only narrowly) and we would model our

Here is a simple version of a formula for the AIC of a given model with p terms:

 $X_t = 0.7111X_{t-1} + -0.1912 \quad X_{t-2} + Z_t$

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paper is

time series as

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