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Geophysical Solutions



Gaussian Mixture ESMDA

Preliminary conclusions

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Summary

- The idea of generalizing ESMDA for a Gaussian Mixture prior models
- Advantages:
 - Partly overcome the Gaussian unimodal assumption
 - In the particular case where the modes are interpreted as different facies, the facies estimation becomes an integral part of the solution
 - It might be useful to test different assumptions of the reservoir, e.g. depositional trends
- However, it may rise numerical and computation issues such as
 - Evaluation of PDF with high dimension
 - The number of models in the ensemble is probably higher



- Only one reference was found in the literature by searching for keywords in google scholar



Contents lists available at [ScienceDirect](#)

Advances in Water Resources

journal homepage: www.elsevier.com/locate/advwatres

Iterative ensemble smoothers in the annealed importance sampling framework

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- However, the context of the use of Gaussian mixture differs from our proposal



- It is introduced in the context of “importance sampling”
- The mixture is used to weigh each model of the ensemble

Set $j = 0$
Draw X_0^i from the prior $p(\cdot)$ for $i = 1, \dots, N$
Propagate the sample to obtain $\mathcal{H}(X_0^i)$ for $i = 1, \dots, N$
Draw X_1^i from X_0^i using $K_1(\cdot | X_0^i)$, $i = 1, \dots, N$
 Evaluate the weights

$$\{w_1(X_1^i)\} = \frac{p(X_1^i) \ell(X_1^i)^{\alpha_1} T_1(X_0^i | X_1^i)}{p(X_0^i) K_1(X_1^i | X_0^i)}$$

Normalize the weights to obtain $\{w_1^i\}_{i=1}^N$

while -- Sampling -- **do**
 Set $j = j + 1$
 Draw X_j^i from $K_j(\cdot | X_{j-1}^i)$, for $i = 1, \dots, N$
 Calculate $\{w_j(X_j^i)\}$ from Eq. (2.19), which is the weight correction term for the transition from X_{j-1}^i to X_j^i
 if -- Resampling -- **then**
 Sample $\{i^k\}_{k=1}^N$ with replacement from $(1, \dots, N)$ with $P(i^k = \ell) = w_j^\ell$
 Set $X_j^k = X_j^{i^k}$
 Set $w(X_j^i) = N^{-1}$, $i = 1, \dots, N$
 end
end

Algorithm 1: Weighted ESMDA.

Set $j = 0$
Draw X_0^i from the prior $p(\cdot)$ for $i = 1, \dots, N$
 Propagate the sample to obtain $\mathcal{H}(X_0^i)$ for $i = 1, \dots, N$
Draw X_1^i from X_0^i defined by Eq. (4.3) Evaluate the weights $\{w_1(X_1^i)\} = \Phi(y | \mathcal{H}(X_{j-1}), h^2 \text{Cov}\{\mathcal{H}(X_{j-1})\} + \alpha_j^{-1} \mathbb{R})$
 Normalize the weights to obtain $\{w_1^i\}_{i=1}^N$

while -- Sampling -- **do**
 Set $j = j + 1$
 Draw X_j^i from Eq. (4.3), for $i = 1, \dots, N$
 Calculate $\{w_j(X_j^i)\} \propto w_{j-1}^i \Phi(y | \mathcal{H}(X_{j-1}), h^2 \text{Cov}\{\mathcal{H}(X_{j-1})\} + \alpha_j^{-1} \mathbb{R})$
 Calculate τ from Eq. (4.2)
 Set $w_j = \tau w_j + (1 - \tau) N^{-1}$
 if -- Resampling -- **then**
 Sample $\{i^k\}_{k=1}^N$ with replacement from $(1, \dots, N)$ with $P(i^k = \ell) = w_j^\ell$
 Set $X_j^k = X_j^{i^k}$
 Set $w_j(X_j^i) = N^{-1}$, $i = 1, \dots, N$
 end
end

end

Algorithm 2: Gaussian Mixture MDA.

Posterior distribution:

$$\hat{p}_j(x|y) \propto \sum_{i=1}^N w(X_j^i) \Phi(x | X_j^i, \mathbb{P}_j),$$



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 Evaluate the weights

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Normalize the weights to obtain $\{w_1^i\}_{i=1}^N$

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Draw X_j^i from $K_j(\cdot | X_{j-1}^i)$, for $i = 1, \dots, N$

Calculate $\{w_j(X_j^i)\}$ from Eq. (2.19), which is the weight correction term for the transition from X_{j-1}^i to X_j^i

if -- Resampling -- **then**

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 Normalize the weights to obtain $\{w_1^i\}_{i=1}^N$

while -- Sampling -- **do**

Set $j = j + 1$

Draw X_j^i from Eq. (4.3), for $i = 1, \dots, N$

Calculate

$\{w_j(X_j^i)\} \propto w_{j-1}^i \Phi(y | \mathcal{H}(X_{j-1}), h^2 \text{Cov}\{\mathcal{H}(X_{j-1})\} + \alpha_j^{-1} \mathbb{R})$

Calculate τ from Eq. (4.2)

Set $w_j = \tau w_j + (1 - \tau) N^{-1}$

if -- Resampling -- **then**

Sample $\{i^k\}_{k=1}^N$ with replacement from $(1, \dots, N)$ with $P(i^k = \ell) = w_j^\ell$

Set $X_j^k = X_j^{i^k}$

Set $w_j(X_j^i) = N^{-1}$, $i = 1, \dots, N$

end

end

Algorithm 2: Gaussian Mixture MDA.

Posterior distribution:

$$\hat{p}_j(x|y) \propto \sum_{i=1}^N w(X_j^i) \Phi(x | X_j^i, \mathbb{P}_j),$$

“weights according to (2.19) an adaptive weight shrinkage is applied”

$$w_{\text{new}} = \tau w_{\text{old}} + (1 - \tau) N^{-1},$$

where

$$\tau = \left(\sum_{j=1}^N (w_{\text{old}}^i)^2 \right)^{-1}.$$



Gaussian Mixture BLI

- Forward model with Gaussian error:

$$\mathbf{d} = \mathbf{G}\mathbf{m} + \mathbf{e}$$

- Gaussian mixture prior:

$$\mathbf{m} \sim \sum_{k=1}^F \lambda_k N(\mathbf{m}; \boldsymbol{\mu}_{m|k}, \mathbf{C}_{m|k})$$

- The posterior distribution is also a Gaussian mixture and it is analytically obtained:

$$\mathbf{m}|\mathbf{d} \sim \sum_{k=1}^F \lambda_{k|d} N(\mathbf{m}; \boldsymbol{\mu}_{m|d,k}, \mathbf{C}_{m|d,k})$$

$$\boldsymbol{\mu}_{r|m,k} = \boldsymbol{\mu}_{r|k} + \boldsymbol{\Sigma}_{r|k} \mathbf{G}^T (\mathbf{G} \boldsymbol{\Sigma}_{r|k} \mathbf{G}^T + \boldsymbol{\Sigma}_e)^{-1} (\mathbf{m} - \mathbf{G} \boldsymbol{\mu}_{r|k})$$

$$\boldsymbol{\Sigma}_{r|m,k} = \boldsymbol{\Sigma}_{r|k} - \boldsymbol{\Sigma}_{r|k} \mathbf{G}^T (\mathbf{G} \boldsymbol{\Sigma}_{r|k} \mathbf{G}^T + \boldsymbol{\Sigma}_e)^{-1} \mathbf{G} \boldsymbol{\Sigma}_{r|k}$$

$$\lambda_{k|d} \propto \lambda_k N(\mathbf{d}; \mathbf{G} \boldsymbol{\mu}_{m|k}, \mathbf{G} \mathbf{C}_{m|k} \mathbf{G}^T + \mathbf{C}_e)$$

Diagram illustrating the components of the posterior distribution:

- A blue arrow points from $\boldsymbol{\mu}_{m|k}$ in the Gaussian term to $g(\boldsymbol{\mu}_{r|k})$.
- A blue arrow points from $\mathbf{C}_{m|k}$ in the Gaussian term to $\mathbf{C}_{dd|k} + \mathbf{C}_e$.



General comments

- Since the distributions in the ESMDA are indirectly given by the multiple realizations in the ensemble, in our implementation we define a Gaussian mixture prior by drawing multiple realizations from each Gaussian component
- In practice, it is like we have multiple ensembles
- The posterior is then computed from the ensemble after the updates



GM-ESMDA - First approach

- Straightforward application of the Gaussian Mixture BLI

- Prior considering 2 modes:

$$\mathbf{m} \sim \sum_{k=1}^2 \omega_k N(\mathbf{m}; \boldsymbol{\mu}_{m|k}, \mathbf{C}_{m|k})$$

- GM-ESMDA:

$$\mathbf{m}_1 \sim N(\mathbf{m}; \boldsymbol{\mu}_{m|1}, \mathbf{C}_{m|1})$$

$$\mathbf{m}_2 \sim N(\mathbf{m}; \boldsymbol{\mu}_{m|2}, \mathbf{C}_{m|2})$$

$$\mathbf{m}_1^{i+1} = \mathbf{m}_1^i + \mathbf{C}_{md|1}^i \left(\mathbf{C}_{dd|1}^i + \alpha \mathbf{C}_e \right)^{-1} (\mathbf{d}_{obs}^i - g(\mathbf{m}_1^i)) \quad \mathbf{m}_2^{i+1} = \mathbf{m}_2^i + \mathbf{C}_{md|2}^i \left(\mathbf{C}_{dd|2}^i + \alpha \mathbf{C}_e \right)^{-1} (\mathbf{d}_{obs}^i - g(\mathbf{m}_2^i))$$

$$\lambda_k^{i+1} \propto \lambda_k^i N(\mathbf{d}_{obs}; g(\boldsymbol{\mu}_{m_k}), \mathbf{C}_{dd|k} + \mathbf{C}_d)$$

$$p(\mathbf{m}|\mathbf{d}) \approx \lambda_1 kdensity(\mathbf{m}_1) + \lambda_2 kdensity(\mathbf{m}_2)$$



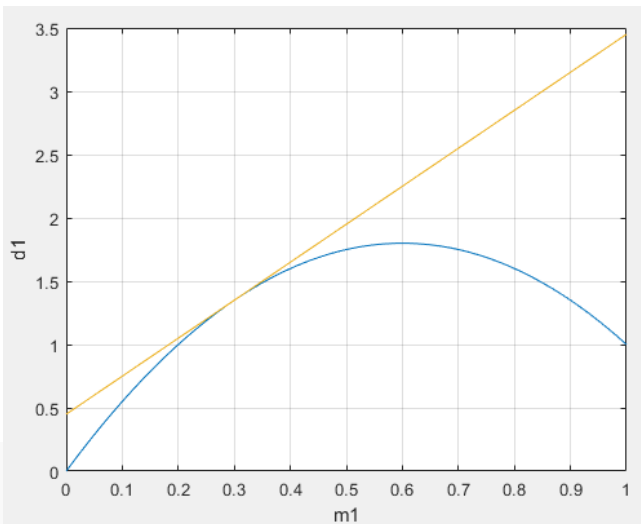
Ultra Toy Problem

- Bivariate problem with a 2nd order polynomial forward model

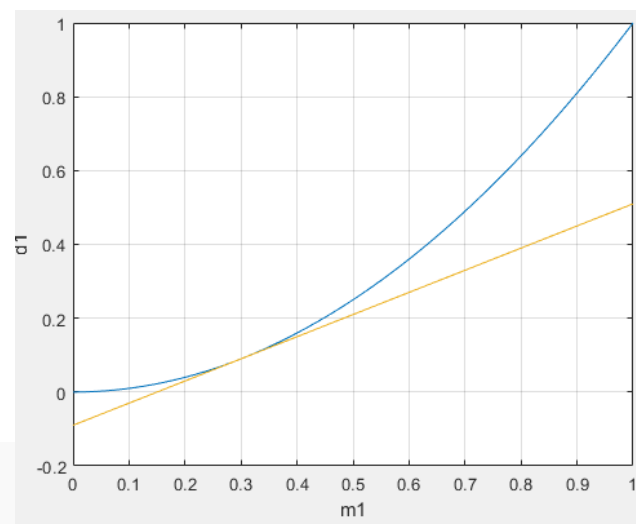
$$\begin{pmatrix} d_1 \\ d_2 \end{pmatrix} = \begin{pmatrix} g(m_1) \\ g(m_2) \end{pmatrix}$$

- It allows us to compare the algorithm with Gaussian Mix BLI
- Additional parameter to control de non-linearity

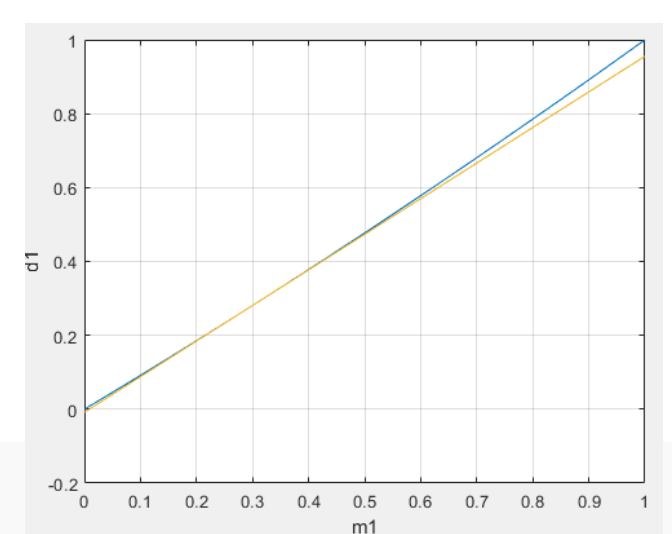
Linear factor = - 0.6



Linear factor = 0



Linear factor = 5

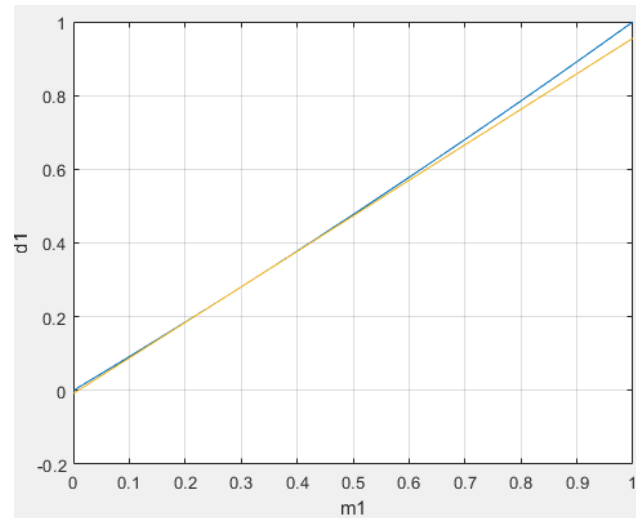


Experiment 1

Bivariate linear case

Priors with the same mean and variance
but different correlations

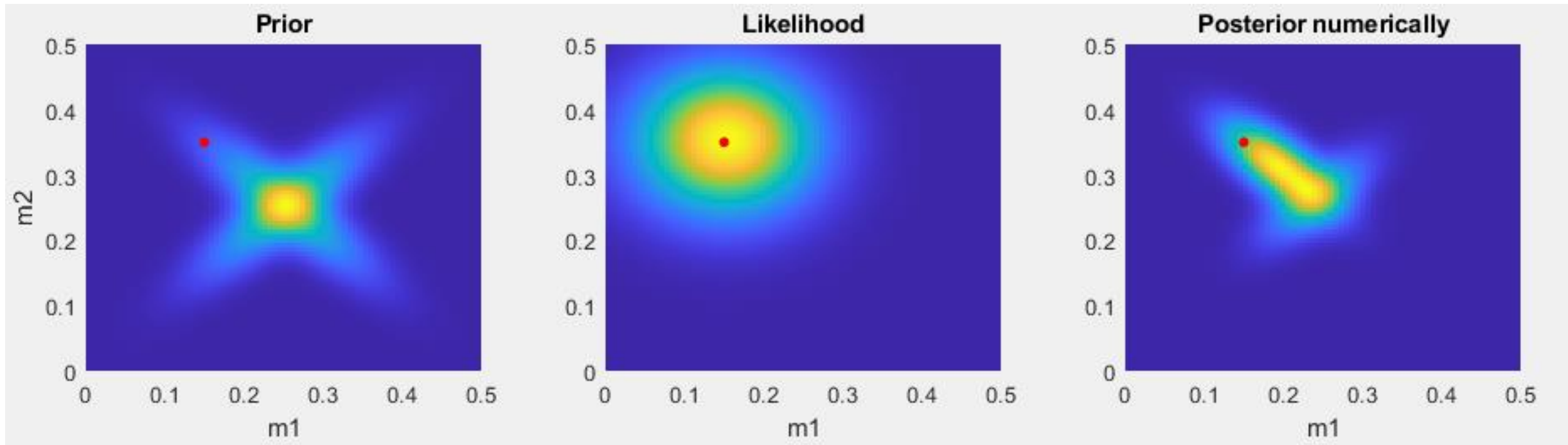
Linear factor = 10





Reference solution – Linear case

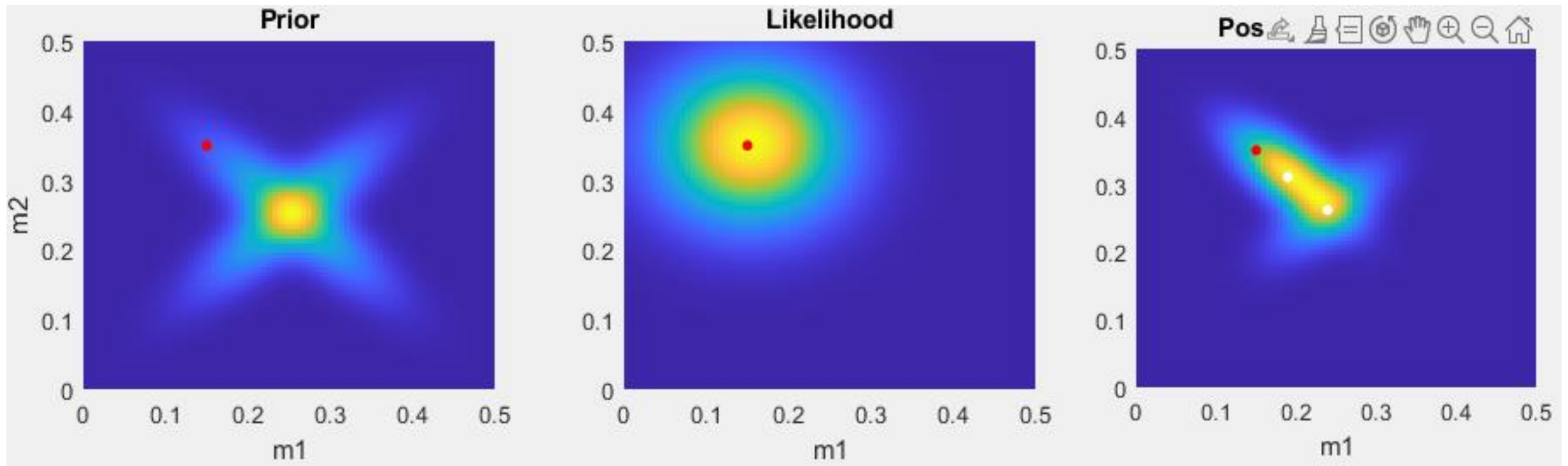
- Extensively numerical evaluation of the model space





Reference solution – Linear case

- Analytical Gaussian mixture BLI

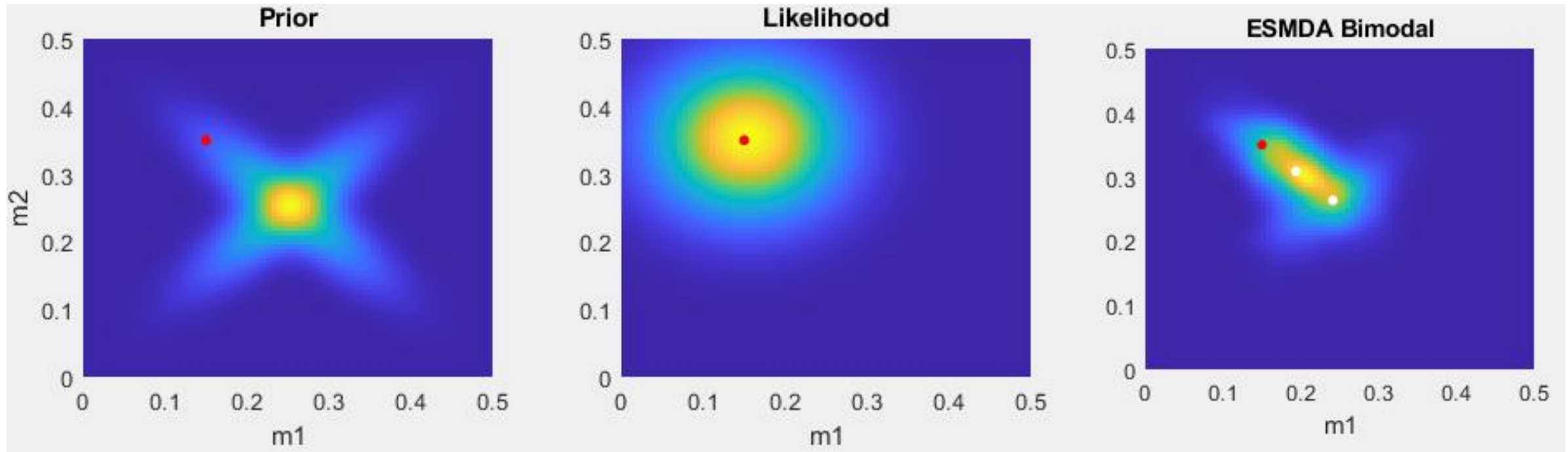


$$\lambda_1 = 0.3 \quad \lambda_2 = 0.7$$



Gaussian Mix ESM DA - Linear case

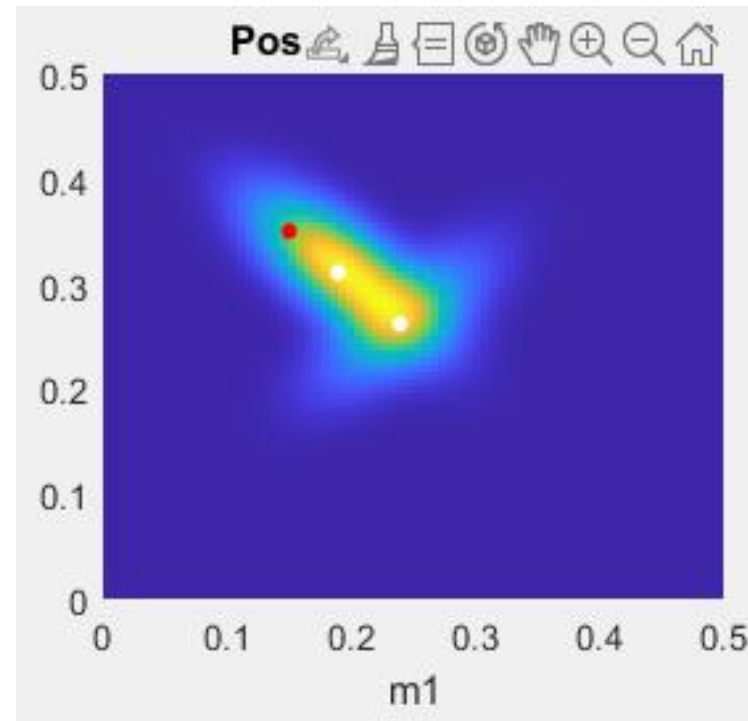
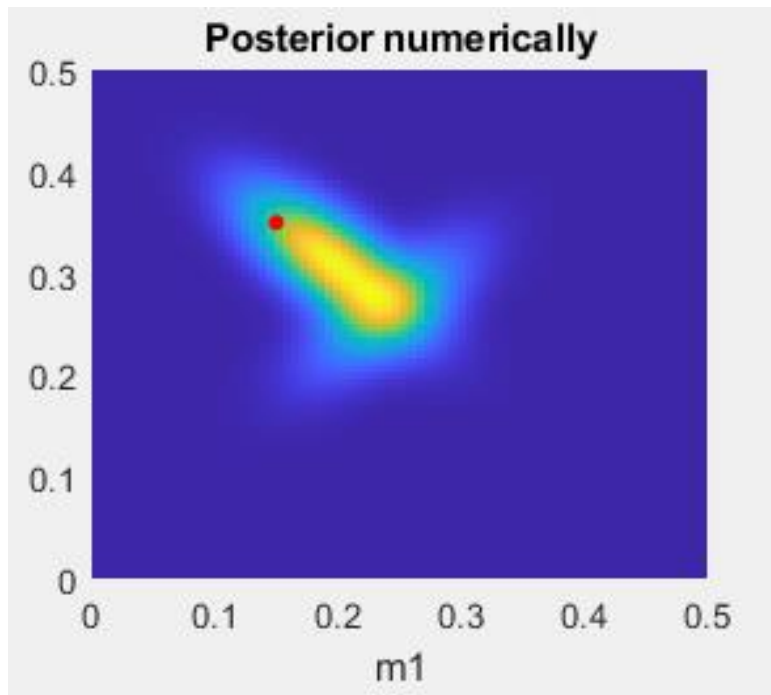
- First approach



$$\lambda_1 = 0.26 \quad \lambda_2 = 0.74$$

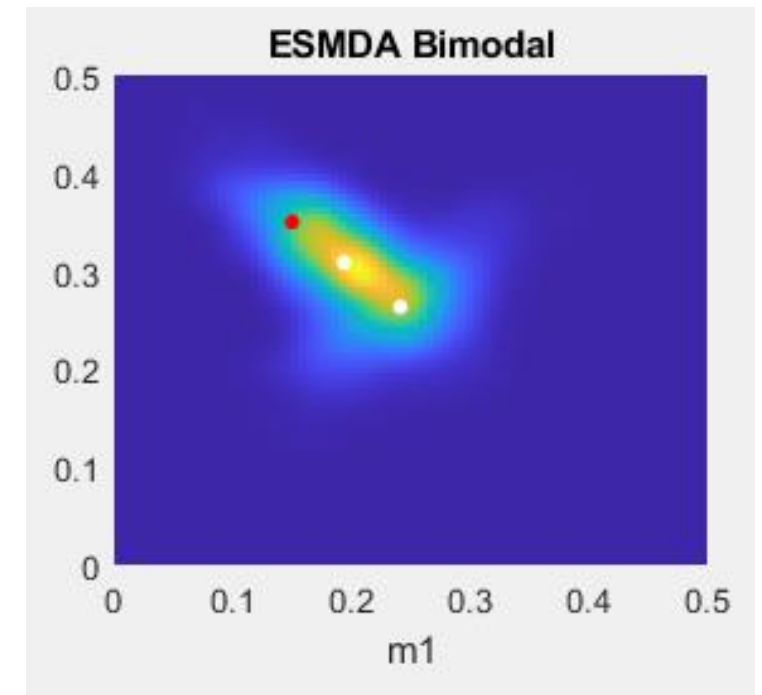


Posterior estimates – Linear case



$$\lambda_1 = 0.3$$

$$\lambda_2 = 0.7$$



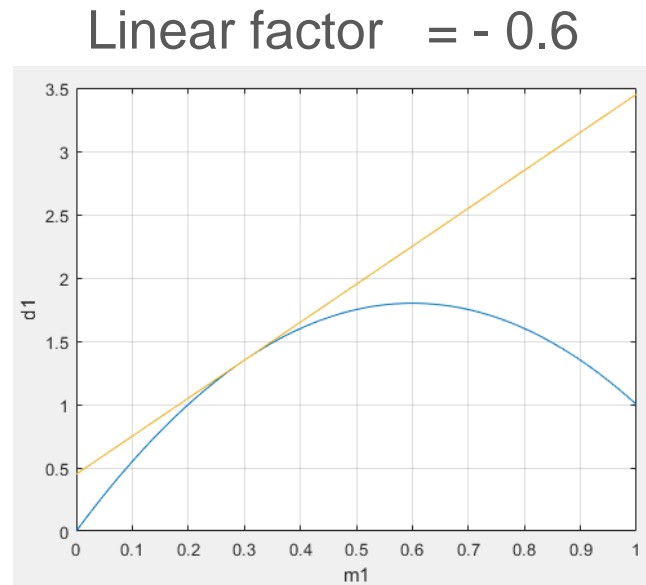
$$\lambda_1 = 0.26$$

$$\lambda_2 = 0.74$$

Experiment 2

Bivariate NON linear case

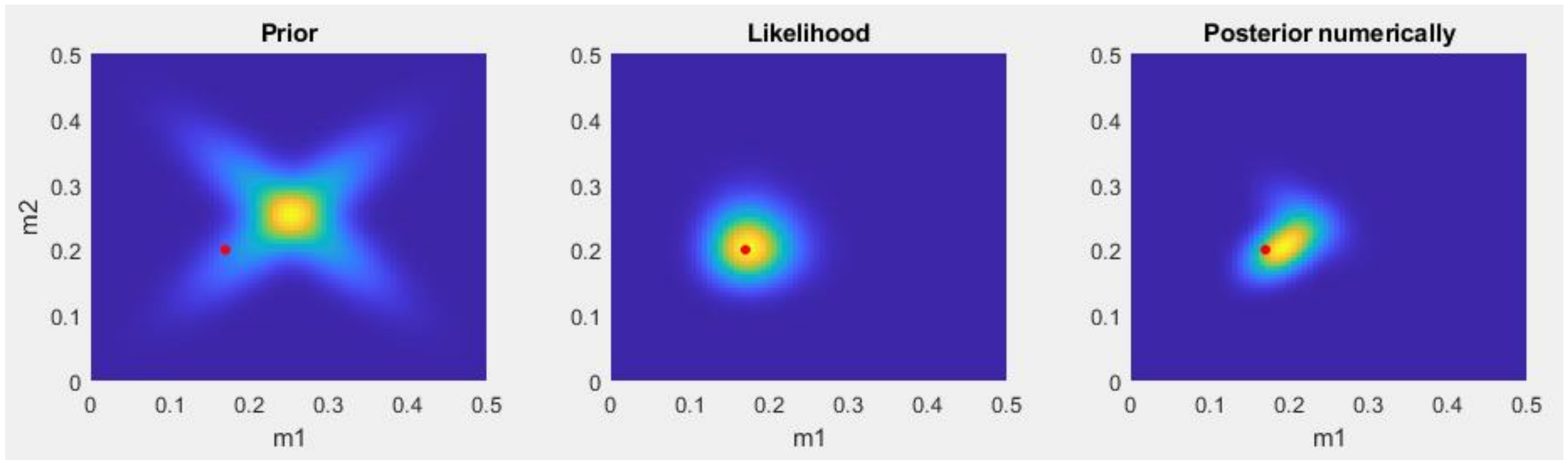
Priors with the same mean and variance
but different correlations





Reference solution – Non-Linear case

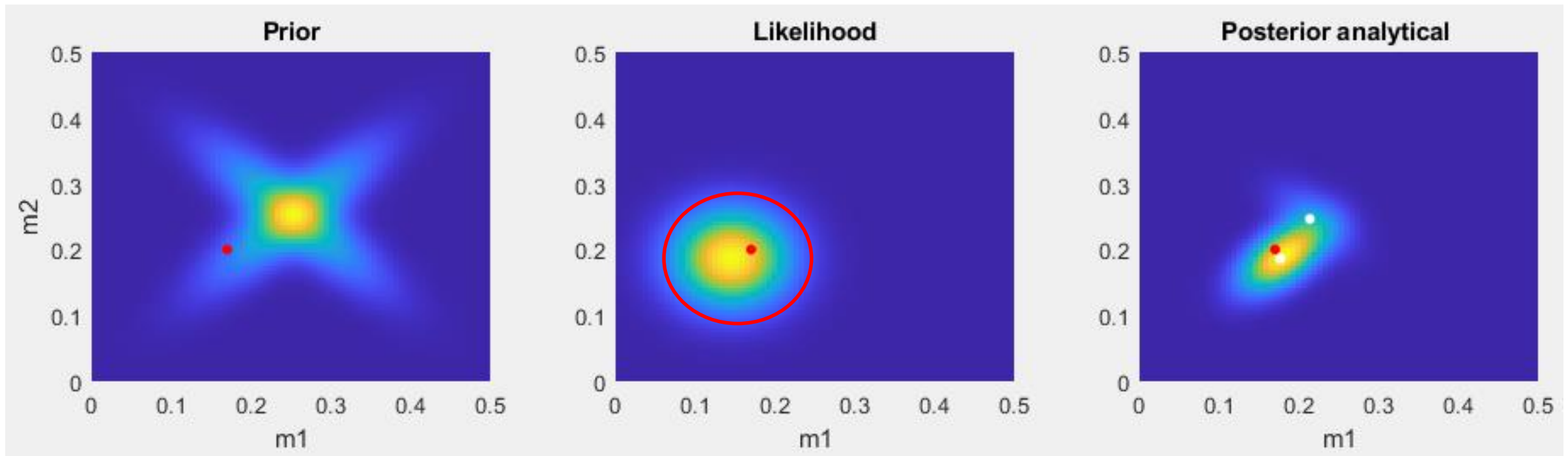
- Extensively numerically evaluation of the model space





Reference solution – Non-Linear case

- Analytical Gaussian mixture BLI

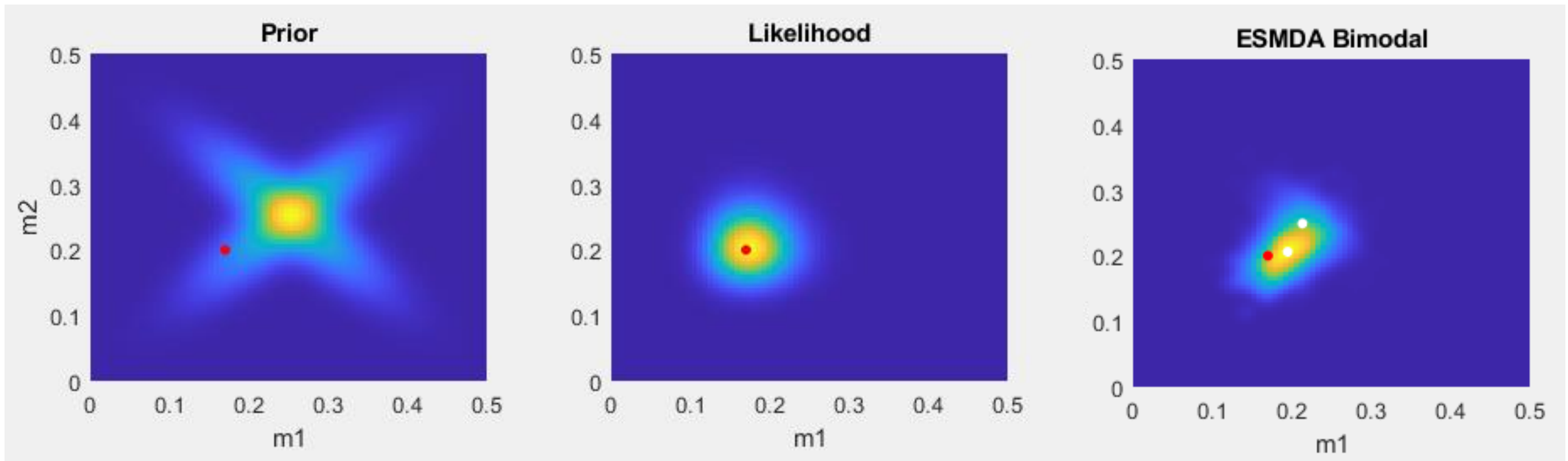


$$\lambda_1 = 0.83 \quad \lambda_2 = 0.17$$



Gaussian Mix ESM DA – Non-Linear case

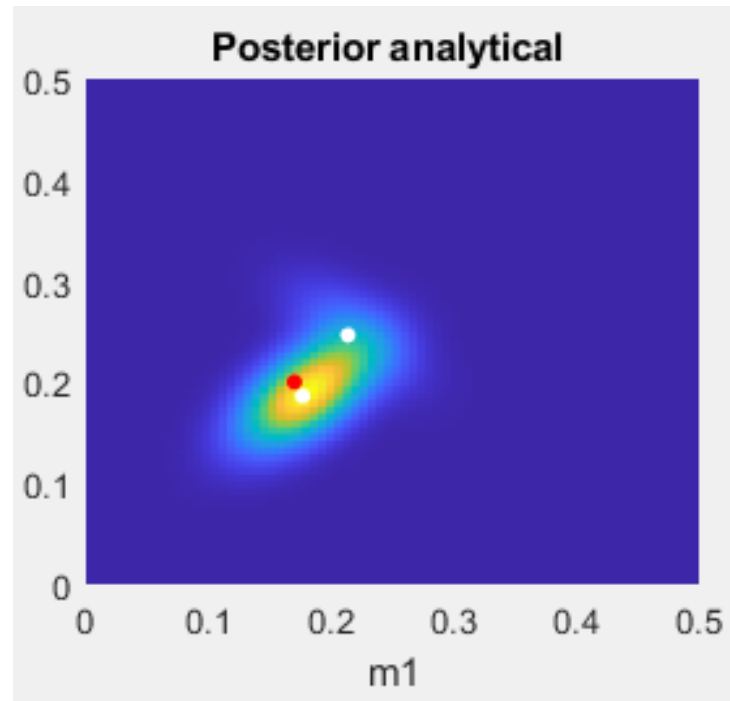
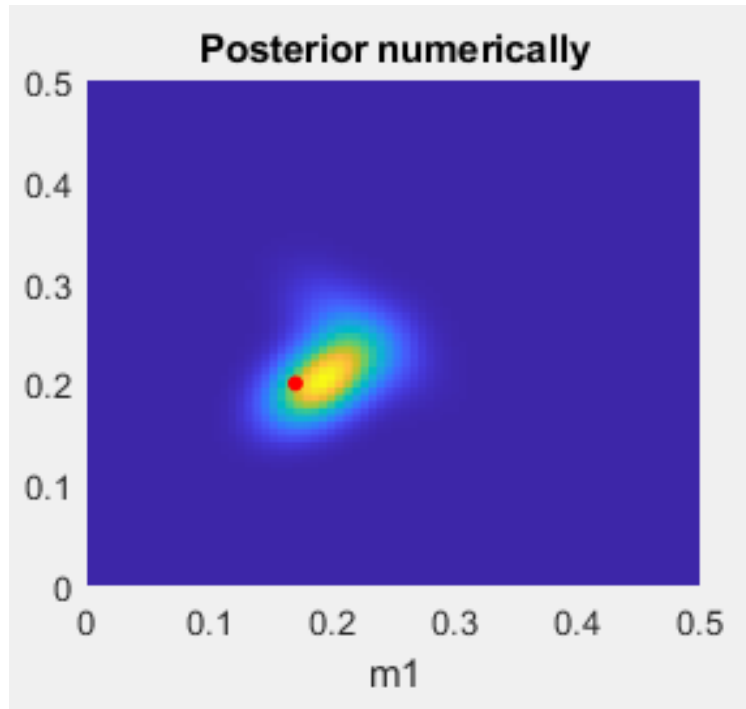
- First approach



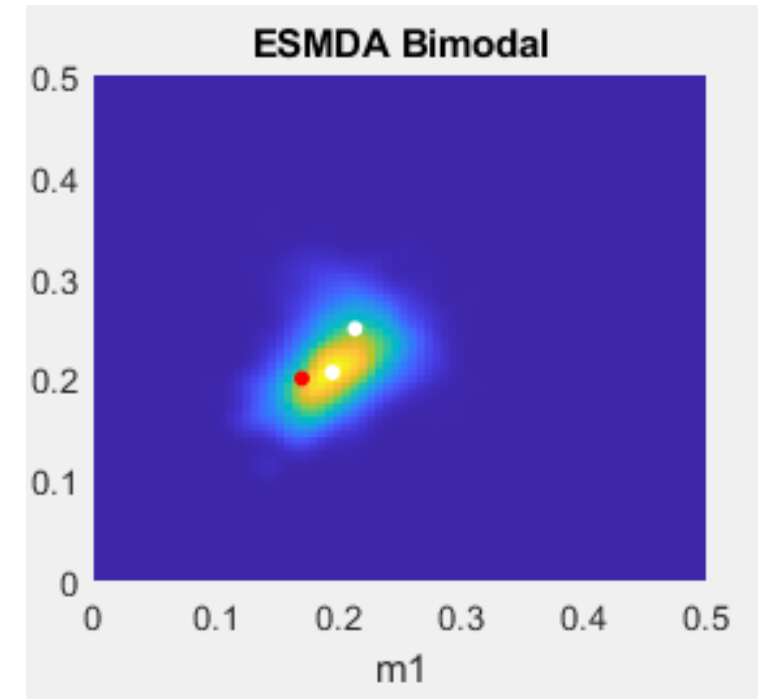
$$\lambda_1 = 0.72 \quad \lambda_2 = 0.28$$



Posterior estimates – Non linear case



$$\lambda_1 = 0.83 \quad \lambda_2 = 0.17$$



$$\lambda_1 = 0.72 \quad \lambda_2 = 0.28$$



Conclusions

- For the linear model, the proposed GM-ESMDA shows similar results to the GM-BLI and numerical results
- For the non-linear model, the proposed GM-ESMDA shows similar results to the numerical results and it differs from the GM-BLI, which is expected since the forward model was linearized.



Future work

- Experiments with different prior means
- Compute the statistical moments of the posterior distributions for a quantitative comparison
- Investigate further Gibbs algorithm approach
- Compute Kullback–Leibler divergence between the reference posterior and the estimated posterior of the methods
- Application in a pointwise petrophysical inversion from elastic properties and non-linear rock-physics data – Well log of Volve data set as reference petrophysical data
 - It is still a bi/tri-variate problem
 - “Easy” and fast publication
- Meanwhile, we think in a way for handling the lambda evaluation for high-dimensional problems



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Thank you

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