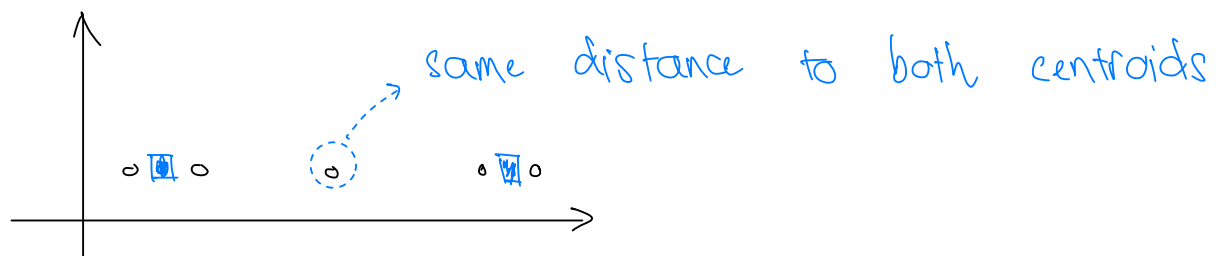


# EXPECTATION MAXIMIZATION

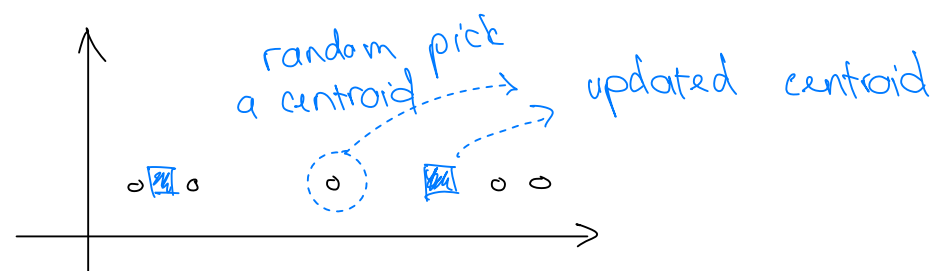
## Main idea:

- Expectation Minimization (EM) a more generalized version of K-means
- Problem with K-means:

Iteration  $t$ :



Iteration  $t+1$ :

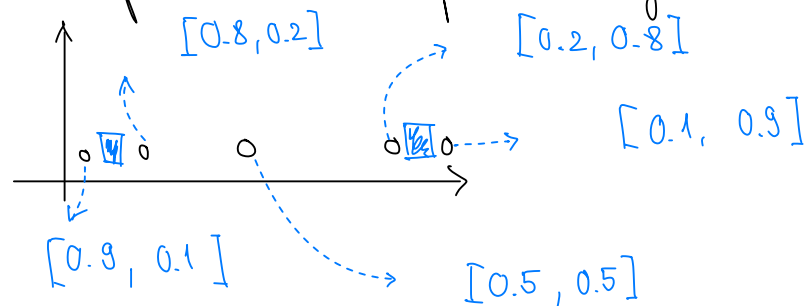


$\Rightarrow$  Not ideal. The centroids are not symmetric even though the data is symmetric

How EM solve that problem:

- In EM, each data point is a probability

Iteration  $t$ :

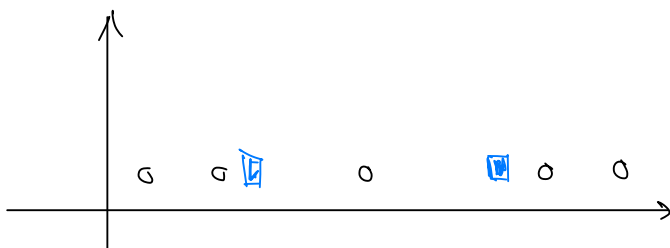


where:

$$\begin{matrix} [0.9 & , & 0.1] \\ \swarrow & & \searrow \\ P(z_i=1) & & P(z_i=2) \end{matrix}$$

- Since  $z_i$  is a probability,  $\mu$  will be more represent the data

$$\left\{ \begin{array}{l} \mu_1 = \frac{0.9x_1 + 0.8x_2 + 0.5x_3 + 0.2x_4 + 0.1x_5}{0.9 + 0.8 + 0.5 + 0.2 + 0.1} \\ \mu_2 = \frac{0.1x_1 + 0.2x_2 + 0.5x_3 + 0.8x_4 + 0.9x_5}{0.1 + 0.2 + 0.5 + 0.8 + 0.9} \end{array} \right.$$



$\Rightarrow$  Look much better! (more symmetrical)

## Expectation Maximization vs K-means

### K-means (Deterministic approach)

- Assignment step:

For each data point  $x_i$ :

$$z_i = \underset{k=1, \dots, K}{\operatorname{argmin}} \|x_i - \mu_k\|^2$$

- Centroids step:

For each centroid:

$$\begin{aligned}\mu_k &= \frac{\sum_{i \in S_k} x_i}{|S_k|} \quad (S_k = \{i: z_i = k\}) \\ &= \frac{\sum_{i=1}^n I(z_i = k) x_i}{\sum_{i=1}^n I(z_i = k)}\end{aligned}$$

### EM (Probabilistic approach)

- E-step (Evaluation)

For each data point  $x_i$ :

$$\begin{cases} P(z_i = 1) \\ \vdots \\ P(z_i = K) \end{cases}$$

$$\Rightarrow P(z_i = k) \text{ for each } k$$

Example: sigmoid func distance

$$P(z_i = k) = \frac{\exp\left(-\frac{1}{\lambda} \|x_i - \mu_k\|^2\right)}{\sum_{k=1}^K \exp\left(-\frac{1}{\lambda} \|x_i - \mu_k\|^2\right)}$$

- M-step: (Maximization)

For each centroid:

$$\mu_k = \frac{\sum_{i=1}^n P(z_i = k) x_i}{\sum_{i=1}^n P(z_i = k)}$$

### How to choose correct $\lambda$ ?

To find  $\theta$ , we use this procedure (Probabilistic Approach to Clustering)

### Probabilistic Approach to Clustering

- Assume some "hidden" joint distribution  $p(x, z | \theta)$  that generates the data  $x$  and the labels  $z$ .

- The goal is to find that distribution by estimating  $\theta$

- To estimate  $\theta$ , there are 2 scenarios:

i)  $x$  and  $z$  are known (complete information). Estimate  $\theta$  by maximizing the joint probability

$$\max_{\theta} \sum_{i=1}^n \log p(x_i, z_i | \theta)$$

ii) Only  $x$  are known (incomplete information). Estimate  $\theta$  by maximizing the marginal probability

$$\max_{\theta} \sum_{i=1}^n \log p(x_i | \theta)$$

## How does this help in clustering?

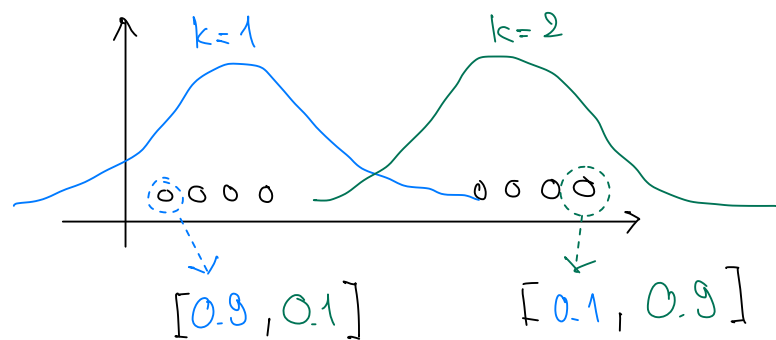
- Remember that the goal of clustering is to find the best possible assignment / label for each data point.
- Once we know the joint distribution  $p(x, z | \theta)$  by estimating  $\theta$  we can infer the label of each data point.

$$z \sim p(z | x, \theta)$$

observed      estimated earlier

## Example with Gaussian Distribution

Visual:



Write down the joint distribution  $p(x, z)$

$$p(x, z) = p(z) \cdot p(x | z)$$

joint      marginal      conditional  
distribution      distribution      distribution

For a specific  $k$  and data point  $x_i$ :

$$p(x_i, z=k) = p(z=k) \cdot p(x_i | z=k)$$

joint probability data and  $z=k$       =       $\pi_k$        $N(x_i | \mu_k, \sigma_k^2)$

probability  $z=k$       probability seeing  $x$  given  $z=k$  (likelihood)

Trick:

Distribution is big if-else where each statement is a probability

The goal is to estimate  $\theta$ , so that we can infer the labels

$$\theta = \{ \pi_k, \mu_k, \sigma_k^2 \}_{k=1}^K$$

The nitty gritty of how to estimate  $\theta$

## Scenario 1: Complete Information

◦ We observed both data and labels (complete information)

◦ Estimate  $\theta$  by maximizing joint probability:

$$\hat{\theta} = \operatorname{argmax}_{\theta} \sum_{i=1}^n \log p(x_i, z_i | \theta)$$

◦ Detail calculations:

i) Write down log-joint probability

For a specific data point and specific  $k$ , the joint probability is:

$$p(x_i, z_i = k) = \pi_k \cdot N(x_i | \mu_k, \sigma_k^2)$$

Then for all data points and all values of  $k$ :

$$\begin{aligned} \sum_{i=1}^n \sum_{k=1}^K I(z_i = k) p(x_i, z_i = k) \\ = \sum_{i=1}^n \sum_{k=1}^K I(z_i = k) [\pi_k \cdot N(x_i | \mu_k, \sigma_k^2)] \end{aligned}$$

Finally, arrive at log-joint probability:

$$\sum_{i=1}^n \sum_{k=1}^K I(z_i = k) \log [\pi_k \cdot N(x_i | \mu_k, \sigma_k^2)]$$

ii) Estimate  $\theta$

$$\theta = \{ \pi_k, \mu_k, \sigma_k^2 \}_{k=1}^K$$

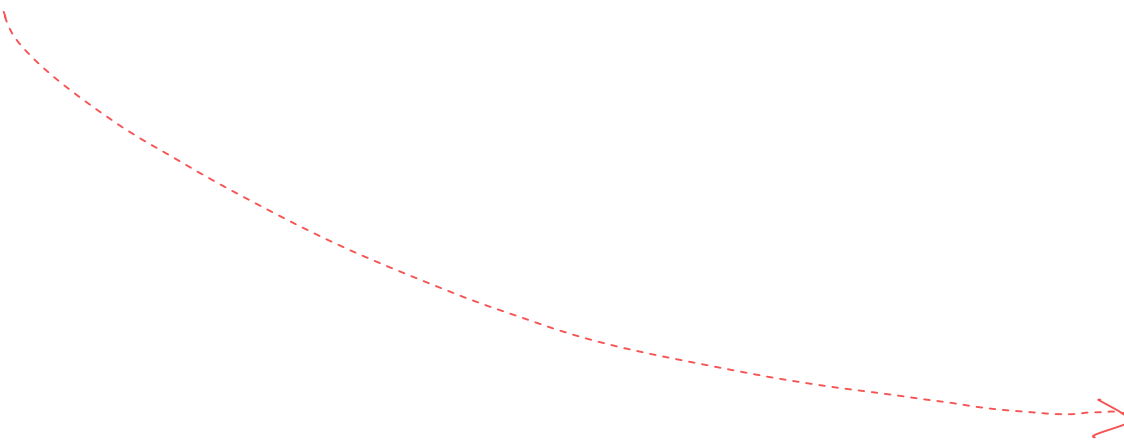
◦ Taking derivative w.r.t  $\pi_k, \mu_k, \sigma_k$ , we arrive at these results:

$$\left\{ \begin{array}{l} \pi_k = \frac{\sum_{i=1}^n I(z_i = k)}{n} \\ \mu_k = \frac{\sum_{i=1}^n I(z_i = k) x_i}{\sum_{i=1}^n I(z_i = k)} \end{array} \right.$$

$$\sigma_k = \operatorname{var}(\{x_i | z_i = k\})$$

→ variance of all data points where  $z = k$

Obviously, in practice we won't always have labels, that is what we will analyze next



## Scenario 2: Incomplete information

- Can only observe data points, labels are hidden
- Estimating  $\theta$  by maximizing marginal probability
$$\hat{\theta} = \underset{\theta}{\operatorname{argmax}} \sum_{i=1}^n \log p(x_i | \theta)$$

Detail calculations:

i) Write down log-marginal probability

For a specific data point:

$$\begin{aligned} p(x_i | \theta) &= \sum_{k=1}^K p(x_i, z_i = k | \theta) \\ &= \sum_{k=1}^K \pi_k \cdot N(x_i | \mu_k, \sigma_k^2) \end{aligned}$$

For all data points:

$$\sum_{i=1}^n p(x_i | \theta) = \sum_{i=1}^n \sum_{k=1}^K \pi_k \cdot N(x_i | \mu_k, \sigma_k^2)$$

Add the log:

$$\sum_{i=1}^n \log p(x_i | \theta) = \sum_{i=1}^n \log \left( \sum_{k=1}^K \pi_k \cdot N(x_i | \mu_k, \sigma_k^2) \right)$$

ii) Estimate  $\theta$ :

$$\theta = \{ \pi_k, \mu_k, \sigma_k^2 \}$$

- There is **no closed form solution** to maximizing the log marginal probability

Why? Lets analyze the marginal probability

$$\begin{aligned} &\sum_{i=1}^n \log \left( \sum_{k=1}^K \pi_k \cdot N(x_i | \mu_k, \sigma_k^2) \right) \\ &= \sum_{i=1}^n \log \left( \sum_{k=1}^K f(\pi_k, x_i, \mu_k, \sigma_k^2) \right) \end{aligned}$$

many distinct functions

$\Leftrightarrow$  many mountains

•  $\log(\text{mountain}) \Leftrightarrow$  steps to that mountain

•  $\log(\text{mountains}) \Leftrightarrow \begin{cases} \text{steps to mountain 1} \\ \text{steps to mountain 2} \\ \dots \end{cases}$

- So no closed form method, luckily we can still use **Expectation Maximization (EM)** to estimate  $\theta$ .

# Expectation Maximization Algorithm

## ◦ Pseudo code:

◦ Initially  $\theta$

◦ For  $t = 1, 2, 3, \dots$

i) E - step (Evaluation):

Fill the hidden value  $z_i$  by drawn i.i.d. from  $p(z | x, \theta)$

$$z_i \stackrel{\text{i.i.d.}}{\sim} p(z | x, \theta)$$

For specific data point  $x_i$  and value  $k$ :

(posterior)  $p(z_i = k | x_i, \theta^t) = \frac{p(x_i, z_i = k | \theta^t)}{\sum_{k=1}^K p(x_i, z_i = k | \theta^t)}$

$\rightarrow$  prior  $\cdot$  likelihood  $= p(z_i = k) \cdot p(x_i | z_i = k)$

$\rightarrow$  marginal  $p(x) = \sum_{k=1}^K p(x, z)$

$\rightarrow$  conditional probability label of data point  $x_i$  is  $k$

ii) M - step (Maximization):

◦ Update  $\theta$  by maximizing the expected log joint probability

For specific data point and value  $k$ :

$$\log p(x_i, z_i = k | \theta^t)$$

For all data points and all values of  $k$ :

$$\sum_{i=1}^n \sum_{k=1}^K p(z_i = k | x_i, \theta^t) \log p(x_i, z_i = k | \theta^t)$$

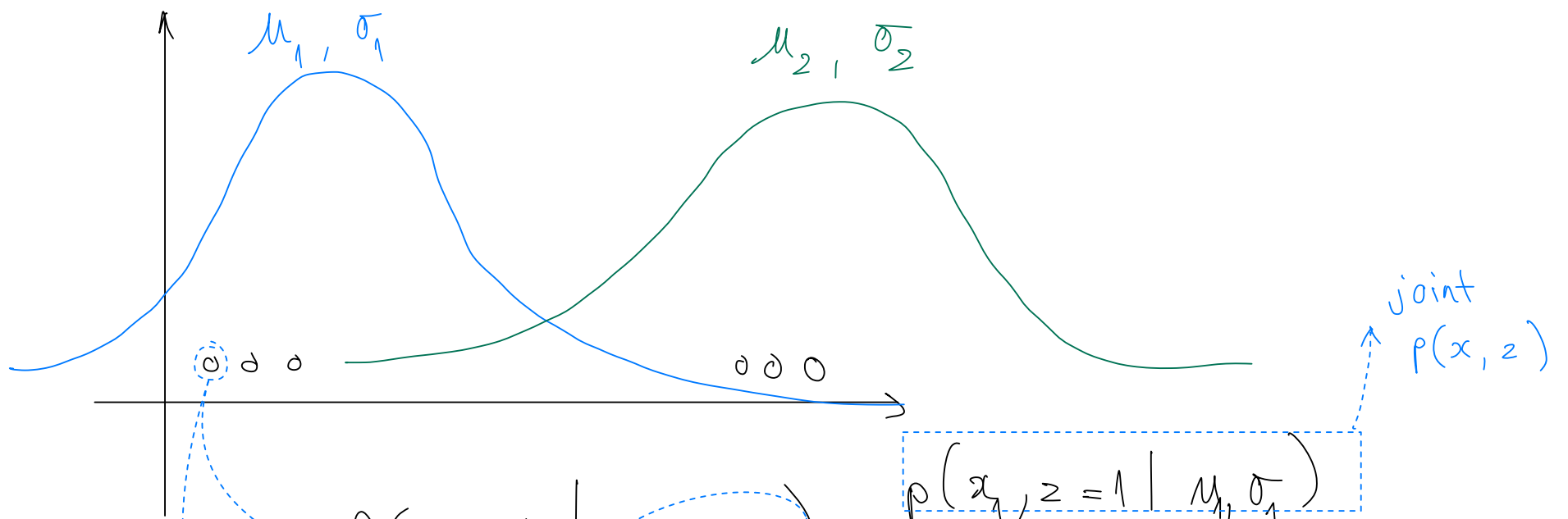
$\rightarrow I(z_i = k)$  in complete information case

$$= \sum_{i=1}^n \sum_{k=1}^K p(z_i = k | x_i, \theta^t) \log [\pi_k \cdot N(x_i | \mu_k, \sigma_k^2)]$$

◦ Taking derivative w.r.t.  $\pi_k, \mu_k, \sigma_k$ , we arrive at these results:

$$\left\{ \begin{array}{l} \pi_k^{t+1} = \frac{\sum_{i=1}^n p(z_i = k | x_i, \theta^t)}{n} \\ \mu_k^{t+1} = \frac{\sum_{i=1}^n p(z_i = k | x_i, \theta^t) \cdot x_i}{\sum_{i=1}^n p(z_i = k | x_i, \theta^t)} \\ \sigma_k^{t+1} = \text{Var}(\{x_i | z_i = k\}) \end{array} \right.$$

$\rightarrow$  variance of data points with same label



Assignments:

$$P(z=1 | x_1, \mu_1, \sigma_1) = \frac{p(x_1, z=1 | \mu_1, \sigma_1)}{\sum_k p(x_1, z=k | \mu_1, \sigma_1)}$$

$\underbrace{\quad}_{\gamma_{11}} \quad \underbrace{\quad}_{\text{observed}} \quad \underbrace{\quad}_{\text{marginal } p(x)}$

$\underbrace{\quad}_{\gamma_{12}} \quad P(z=2 | x_1, \mu_2, \sigma_2)$

Do the same for other data points, then:  
 we have the distribution of  $z_i$   
 $z_i \sim p(z | x, \mu, \sigma)$

Centroids  $(\mu, \sigma)$ :  $\{\mu, \sigma\}_{new} = \operatorname{argmax} \sum E_{z_i \sim p(z|x, \mu, \sigma)} [\log p(x_i, z_i | \mu_1, \sigma_1)]$

joint likelihood  
 randomly drawn from