

Unit 2: Conditional Probability and Bayes' Rule

Conditioning Probability

If A and B are events with $P(B) > 0$, then the conditional probability of A given B , $P(A|B)$ is:

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Intuition:



Theorem 2.3.1

$$P(A \cap B) = P(B) P(A|B) = P(A) P(B|A)$$

Applying theorem 2.3.1 repeatedly, we can generalize to the intersection of n events

Theorem 2.3.2

$$P(A_1, A_2, \dots, A_n) = P(A_1) P(A_2 | A_1) P(A_3 | A_1, A_2) \dots P(A_n | A_1, \dots, A_{n-1})$$

We are now ready to introduce the 2 main theorems about conditional probability: Bayes' rule and Law of Total Prob Probability

Theorem 2.3.3 (Bayes' rule)

$$P(A|B) = \frac{P(B|A) P(A)}{P(B)}$$

Theorem 2.3.4 (Law of Total Probability LOTP)

Let A_1, A_2, \dots, A_n be a partition of the sample space S (i.e., the A_i are disjoint events and their union is S), with $P(A_i) > 0$ for all i . Then:

$$P(B) = \sum_{i=1}^n P(B|A_i) P(A_i)$$

Example 2.3.7 2.3.6

You have 1 fair coin and 1 biased coin with heads $3/4$. You pick 1 coin and flip it 3 times. What is the probability that the coin you picked is the fair one?

- A is events of the coin lands heads 3 times
- F is event that the chosen coin is fair one
- We need to find $P(F|A)$, but it's easier to find $P(A|F)$ and $P(A|F^c)$, therefore:

$$P(F|A) = \frac{P(A|F) P(F)}{P(A)} = \frac{P(A|F) P(F)}{P(A|F) P(F) + P(A|F^c) P(F^c)} \approx 0.23$$

Example 2.3.7 (Testing for a rare disease)

Jimmy being tested for a disease, which affects 1% of population, the test result is positive. Let D be the event that Jimmy has the disease and T be the event that he tests positive.

The test is 95% accurate, which means $P(T|D)$ and $P(T^c|D^c)$ is 0.95

Find the probability that Jimmy has the disease, given the evidence provided by the test result.

$$P(D|T) = \frac{P(T|D) P(D)}{P(T)} = \frac{P(T|D) P(D)}{P(T|D) P(D) + P(T|D^c) P(D^c)} \approx 0.16$$

Note: $P(D|T)$ is a balance between $P(D)$ and $P(T)$

With extra conditioning

Theorem 2.4.1 (Bayes' rule with extra conditioning)

$$P(A|B, E) = \frac{P(B|A, E) P(A|E)}{P(B|E)}$$

Theorem 2.4.2 (LOTP with extra conditioning)

$$P(B|E) = \sum_{i=1}^n P(B|A_i, E) P(A_i|E)$$

Independent events

Events A and B are independent if

$$P(A \cap B) = P(A) P(B)$$

If $P(A) > 0$ and $P(B) > 0$, then:

$$P(A|B) = P(A) \text{ and } P(B|A) = P(B)$$

Independence vs disjoint

- Independence events give no information about each other, they can happen at the same time. $P(A \cap B) = P(A) P(B)$
- Disjoint event give information about each other, they cannot happen at the same time. Since $P(A \cap B) = 0$

Pairwise independence and (complete) independence

$$\text{Complete } \left\{ \begin{array}{l} P(A \cap B) = P(A) P(B) \\ P(A \cap C) = P(A) P(C) \\ P(B \cap C) = P(B) P(C) \\ P(A \cap B \cap C) = P(A) P(B) P(C) \end{array} \right\} \text{ pairwise}$$

Example 2.5.4 (pairwise independence doesn't imply independence)

Consider 2 coin tosses, A is event that the first is Head, B is event that the second is Head, C is event that both has same results

- Knowing what happen with A or B separately doesn't give any information about C
- Knowing what happen with both A and B give us information about C . A and B imply C by definition

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Example 2.5.6 (conditional independence doesn't imply independence)

Return to example 2.5.6, suppose we have chosen a coin but not sure if it is a fair or a biased one. And we flip the coin a number of times. Conditional on choosing the fair coin, the coin tosses are independent. Conditional on choosing the biased coin, the tosses are also independent.

However, the coin tosses are not unconditionally independent. Because if we don't know which coin we have chosen, then observing the sequence of tosses give us information about the coin, and in turn predict future tosses.

Example 2.5.7 (independence doesn't imply conditional independence)

Suppose A and B are the only two ppl who ever call me. A and B are clearly independent.

However, conditional on the phone ringing (R), if A is calling meaning B is not calling and vice versa. $P(A|R) < 1 = P(A)$
 $= P(A|B^c, R)$

Therefore, A and B are not conditionally independent.

Conditioning as a problem-solving tool

Conditioning allow us to split a problem into sub-problems.

We can say, "if condition on E and E^c , then combine them using LOTP"

Strategy 1: Condition on what you wish you know

Example 2.6.1

There are 3 doors, 1 of which has a car behind it, Monty always reveal the first door without a car. The contestant is given the option to switch door before open, should he?

Condition on wish that we know where the car is.

Let C_i be the event that the car behind door $i = 1, 2, 3$:

$$P(\text{get car}) = P(\text{get car} | C_1) \frac{1}{3} + P(\text{get car} | C_2) \frac{1}{3} + P(\text{get car} | C_3) \frac{1}{3}$$

Suppose we switch, then if car is behind door 1 ^{tail} and we switch to door 2, given Monty has already opened door 3 without a car, then switching will succeed:

$$P(\text{get car}) = 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} = \frac{2}{3}$$

So switching works $\frac{2}{3}$ of the time.

Strategy 2: Condition on the first step

Example 2.6.3

A single Bobo, after 1 minute can either die, stay the same, or duplicate. What is the probability that Bobo's population will eventually die out?

Let D be the event that the population die out

Let B_i be the event that Bobo turns into i Bobo

$$\begin{aligned} \text{Then: } P(D) &= P(D|B_0) \frac{1}{3} + P(D|B_1) \frac{1}{3} + P(D|B_2) \frac{1}{3} \\ &= 1 \cdot \frac{1}{3} + P(D) \frac{1}{3} + P(D) \cdot \frac{1}{3} \end{aligned}$$

\Rightarrow Solve the above equation, $P(D) = 1$. So the population will eventually die out.

Example 2.6.4

2 gambler A and B make sequence of \$1 bets. Each bet, A has p chance of winning, B has $q = 1 - p$ chance of winning. A starts with i dollars and B starts with $N - i$ dollars. When A gains, B loses. What is the probability that A wins the game?

Let p_i be event that A wins the game, so p_0^0 means A

loses and $p_N = 1$ means A wins

Condition on outcome of the first round, we have:

$$P_i = P(W|A \text{ starts at } i; \text{win})p + P(W|A \text{ starts at } i; \text{lose})q$$

$$= P(W|A \text{ starts at } i+1)p + P(W|A \text{ starts at } i-1)q$$

$$= P_{i+1} \cdot p + P_{i-1} \cdot q$$

$$\Rightarrow P_i = \begin{cases} \frac{1 - (\frac{q}{p})^i}{1 - (\frac{q}{p})^M} & \text{if } p \neq \frac{1}{2} \\ \frac{i}{M} & \text{if } p = \frac{1}{2} \end{cases}$$

Example 2.6.5 (Simpson's paradox)

Dr. Hilbert			Dr. Nick		
	Heart	Band-aid		Heart	Band-aid
Success	70	10	Success	2	81
Failure	20	0	Failure	8	9

Simpson's paradox happens when

$$P(A|B, C) < P(A|B^c, C)$$

$$P(A|B, C^c) < P(A|B^c, C^c)$$

but

$$P(A|B) > P(A|B^c)$$

In this example, A is success surgery, B is Dr. Nick perform, C is heart surgery. Here even tho Dr. Nick perform worse than Dr. Hilbert in both types of surgery, his overall stats are still better!

LOTP can tell us why why:

$$P(A|B) = P(A|C, B)P(C|B) + P(A|C^c, B)P(C^c|B)$$

$$P(A|B^c) = P(A|C, B^c)P(C|B^c) + P(A|C^c, B^c)P(C^c|B^c)$$

Although: $P(A|C, B) < P(A|C, B^c)$

$$P(A|C^c, B) < P(A|C^c, B^c)$$

The weights $P(C|B)$ and $P(C^c|B)$ can flip the balance.

Since Dr. Nick perform much more Band-aid $P(C|B) \uparrow$, lead to increase overall performance $P(A|B) \uparrow$

Practice Problems

Problem 1

There is an email spam filter. Suppose that 80% of email is spam. In 10% of spam emails, the phrase "free money" is used, whereas this phrase is only used in 1% of non-spam emails. A new email with "free money" arrived, what probability is it a spam?

Let F be event that email has "free money"

S be event that email is Spam

$$P(S|F) = \frac{P(F|S)P(S)}{P(F)} = \frac{P(F|S)P(S)}{P(F|S)P(S) + P(F|S^c)P(S^c)}$$

$$= \frac{0.1 \times 0.8}{0.1 \times 0.8 + 0.01 \times 0.2} = 0.975$$

Problem 2

The screen for phone is manufactured by 3 companies A, B, C. With the proportions of 0.5, 0.3, 0.2 respectively. Their defective probabilities are 0.01, 0.02, 0.03 respectively. Given the screen on the phone is defective, what probability that company A did it?

Let D be event of Defective

A, B, C be events that company manufactured

$$P(A|D) = \frac{P(D|A)P(A)}{P(D)} = \frac{P(D|A)P(A)}{P(D|A)P(A) + P(D|B)P(B) + P(D|C)P(C)}$$

$$= \frac{0.01 \times 0.5}{0.01 \times 0.5 + 0.02 \times 0.3 + 0.03 \times 0.2}$$

$$= 0.294$$

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Problem 3

A family has 3 children, named A, B, and C

a) Is the event "A is older than B" independent of "A is older than C"?

No they are not independent events, because if we consider the sequence of birth; given $A > B$:

$ABC \rightarrow A \text{ older than } C$
 $ACB \rightarrow A \text{ older than } C$
 $CAB \rightarrow A \text{ younger than } C$

So knowing $A > B$ (older), gives us information about whether A is older than C.

To make this more intuitive, think of extreme case with 100 children A_1, \dots, A_{100} . Given $A_1 > A_2, \dots, A_{99}$, then the only way for A_{100} to be older than A_1 is $A_{100} A_1, \dots, A_{99}$ sequence.

b) Find Probability that A is older than B, given A older than C

$$P(A > B | A > C) = \frac{P(A > B, A > C)}{P(A > C)} = \frac{1/3}{1/2} = \frac{2}{3}$$

Problem 4

Consider the Monty Hall problem, except that Monty enjoys open door 2 more than door 3, if he has a choice, he would open door 2 with prob $p = 3/4$

a) Find the probability that the strategy of always switching succeeds

$$\begin{aligned}
 P(W) &= P(W|C_1)P(C_1) + P(W|C_2)P(C_2) + P(W|C_3)P(C_3) \\
 &= 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} \\
 &= \frac{2}{3}
 \end{aligned}$$

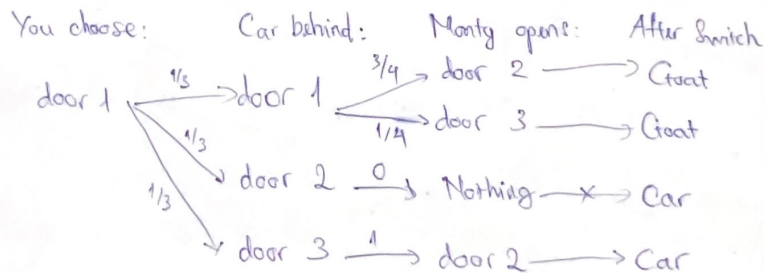
b) Find the probability that the strategy of always switching succeeds, given Monty opens door 2

Let D_i be event that Monty opens Door i

We are looking for $P(W | D_2)$, which is the same as $P(C_3 | D_2)$

$$\begin{aligned}
 P(C_3 | D_2) &= \frac{P(D_2 | C_3) \cdot P(C_3)}{P(D_2)} \\
 &= \frac{P(D_2 | C_3) P(C_3)}{P(D_2 | C_1) P(C_1) + P(D_2 | C_2) P(C_2) + P(D_2 | C_3) P(C_3)} \\
 &= \frac{1 \cdot \frac{1}{3}}{\frac{4}{3} \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}} = \frac{4}{7}
 \end{aligned}$$

Visualizing this:



c) Find the probability that the strategy of always switching succeeds, given Monty opens door 3

$$\begin{aligned}
 P(C_2 | D_3) &= \frac{P(D_3 | C_2) \cdot P(C_2)}{P(D_3)} = \frac{P(D_3 | C_2) \cdot P(C_2)}{P(D_3 | C_1) P(C_1) + P(D_3 | C_2) P(C_2) + P(D_3 | C_3) P(C_3)} \\
 &= 0.8
 \end{aligned}$$

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Homework Problems

Problem 1

Fred is answering a multiple choice problem with n options.

Let K be event that Fred know the answer

R be event that Fred gets the problem right

Suppose that if he knows the answer then he'll get the problem right. But if he doesn't know the answer then he'll guess randomly.

Let $P(K) = p$.

a) Find $P(K|R)$.
$$P(K|R) = \frac{P(R|K) \cdot P(K)}{P(R)}$$

$$= \frac{P(R|K) \cdot P(K)}{P(R|K) \cdot P(K) + P(R|K^c) \cdot P(K^c)} = \frac{p}{1 \cdot p + \frac{1}{n} \cdot (1-p)}$$

b) When (if ever) does $P(K|R) = p$?
$$= \frac{p}{p + \frac{(1-p)}{n}}$$

Given that $P(K|R) = \frac{p}{p + \frac{(1-p)}{n}}$

\Rightarrow With these extreme cases: $n=1, p=0, p=1$.

Problem 2

A hat contains 100 coins, where 99 are fair and 1 is always head. A coin is chosen, flipped 7 times and lands head all 7 times. What is the probability that the chosen coin is double-headed (always heads)?

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Let $\begin{cases} F \text{ is event fair coin chosen} \\ F^c \text{ is event double-headed coin chosen} \\ H_i \text{ is event that lands head } i \text{ times} \end{cases}$

$$P(F^c | H_7) = \frac{P(H_7 | F^c) \cdot P(F^c)}{P(H_7 | F^c) \cdot P(F^c) + P(H_7 | F) \cdot P(F)} = \frac{1 \cdot 0.01}{1 \cdot 0.01 + 0.5^7 \cdot 0.99} = 0.564$$

Problem 3

Jimmy takes a series of n tests. Let D be event that he has the disease, $p = P(D)$ be the prior probability that he has the disease, and $q = 1 - p$. Let T_j be the event that he tests positive on the j th test.

a) Assume the test results are conditionally independent given Jimmy's disease status. Let $a = P(T_j | D)$, $b = P(T_j | D^c)$. Find the posterior probability that Jimmy has the disease, given that he tests positive on exactly k out of n tests.

$$P(D | X=k) = \frac{P(X=k | D) \cdot P(D)}{P(X=k)} = \frac{a^k (1-a)^{n-k} \cdot p}{a^k (1-a)^{n-k} \cdot p + b^k (1-b)^{n-k} \cdot q}$$

b) There is a gene G that makes all the tests results positive. Assume $P(G) = \frac{1}{2}$, D and G are independent. Let $a_0 = P(T_j | D, G^c)$, $b_0 = P(T_j | D^c, G^c)$. Find probability that Jimmy has the disease, given that he tests positive all n times?

Let T be event that Jimmy tests positive all n times.

$$P(D | T) = \frac{P(T | D) \cdot P(D)}{P(T)} = \frac{P(T | D) \cdot p}{P(T | D) \cdot p + P(T | D^c) \cdot q}$$

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Add extra conditioning on if Jimmy has the gene G :

$$P(D) = P(D|G) \cdot P(G) + P(D|G^c) \cdot P(G^c)$$

$$= \frac{1}{2} + \frac{a_0^n}{2}$$

$$P(D^c) = P(D^c|G) \cdot P(G) + P(D^c|G^c) \cdot P(G^c)$$

$$= \frac{1}{2} + \frac{b_0^n}{2}$$

$$\text{Thus: } P(D|T) = \frac{p(1+a_0^n)}{p(1+a_0^n) + q(1+b_0^n)}$$

Problem 4

An election with 2 candidates A and B . Every voter is invited to do a poll. Let A be event that voter voted A , W be event that the voter willing to participate the poll. We know that $P(W|A) = 0.7$, $P(W|A^c) = 0.3$. The final poll said that 60% of respondents say they voted for A . Find $P(A)$, true proportion of ppl who voted for A .

$$P(A|W) = \frac{P(W|A) \cdot P(A)}{P(W|A) \cdot P(A) + P(W|A^c) \cdot P(A^c)}$$

$$\Leftrightarrow 0.6 = \frac{0.7 P(A)}{0.7 P(A) + 0.3 (1 - P(A))}$$

$$\Rightarrow P(A) = 0.39$$