

## Stat 3022 Homework 3

## &gt; ## Problem 2.25

&gt; library(Stat2Data)

&gt; data(Pines)

&gt; head(Pines)

	Row	Col	Hgt90	Hgt96	Diam96	Grow96	Hgt97	Diam97	Spread.97	Needles97	Deer95
1	1	1	NA	NA	NA	NA	NA	NA	NA	NA	NA
2	1	2	14	284	4.2	96	362	6.6	162	66	0
3	1	3	17	387	7.4	110	442	9.3	250	77	0
4	1	4	NA	NA	NA	NA	NA	NA	NA	NA	NA
5	1	5	24	294	3.9	70	369	7.0	176	72	0
6	1	6	22	310	5.6	84	365	6.9	215	76	0

	Deer97	Cover95	Fert	Spacing
1	NA	0	0	15
2	1	2	0	15
3	0	1	0	15
4	NA	0	0	15
5	0	2	0	15
6	0	1	0	15

&gt; lm1=lm(Hgt97~Hgt90, data=Pines)

&gt; summary(lm1)

Call:

lm(formula = Hgt97 ~ Hgt90, data = Pines)

Residuals:

	Min	1Q	Median	3Q	Max
	-261.886	-44.343	7.308	55.114	196.114

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	307.439	9.841	31.239	< 2e-16 ***
Hgt90	2.322	0.492	4.721	2.77e-06 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 78.79 on 807 degrees of freedom

(191 observations deleted due to missingness)

Multiple R-squared: 0.02687, Adjusted R-squared: 0.02567

F-statistic: 22.28 on 1 and 807 DF, p-value: 2.772e-06

#2.25(a): The regression equation is  $\hat{Hgt97} = 307 + 2.32 \cdot Hgt90$ . From summary, the t value is 4.72 and the p value is  $2.77e-06$  which is very smaller than zero. Thus we accept the hypothesis  $H_1: \beta_1 \neq 0$ .

&gt; summary(lm1)\$adj.r.squared

[1] 0.02566567

&gt; summary(lm1)\$r.squared

[1] 0.02687153

#2.25(b): The R-squared is about 2.7%, which means only 2.7% of variability in Hgt97 can be explained by Hgt90.

&gt; anova(lm1)

Analysis of Variance Table

Response: Hgt97

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Hgt90	1	138344	138344	22.284	2.772e-06 ***
Residuals	807	5010010	6208		

---

Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

#2.25(c): From the anova table shows above.

```
> 138344/(5010010+138344)
[1] 0.0268715
```

#2.25(d): The coefficient value of determination computed from the anova table is about 2.7%.

#2.25(e): No. Because from the r-squared value we know that only small part of the dataset Hgt97 can be explained by this model, it is not that acceptable.

```

> ## Problem 3.13
> data(MathEnrollment)
> head(MathEnrollment)
  Ayear Fall Spring
1  2001  259   246
2  2002  301   206
3  2003  343   288
4  2004  307   215
5  2005  286   230
6  2006  273   247
> mathenroll = data.frame(MathEnrollment)
> newMathenroll = subset(mathenroll, Ayear!=2003)
> lm2=lm(formula = Spring~Fall+Ayear, data=newMathenroll)
> summary(lm2)
Call:
lm(formula = Spring ~ Fall + Ayear, data = newMathenroll)

```

Residuals:

	Min	1Q	Median	3Q	Max
	-16.1945	-9.3982	0.3212	5.8503	18.2036

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-1.172e+04	2.686e+03	-4.361	0.00331 **
Fall	-1.007e+00	2.041e-01	-4.933	0.00169 **
Ayear	6.107e+00	1.337e+00	4.566	0.00258 **

---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

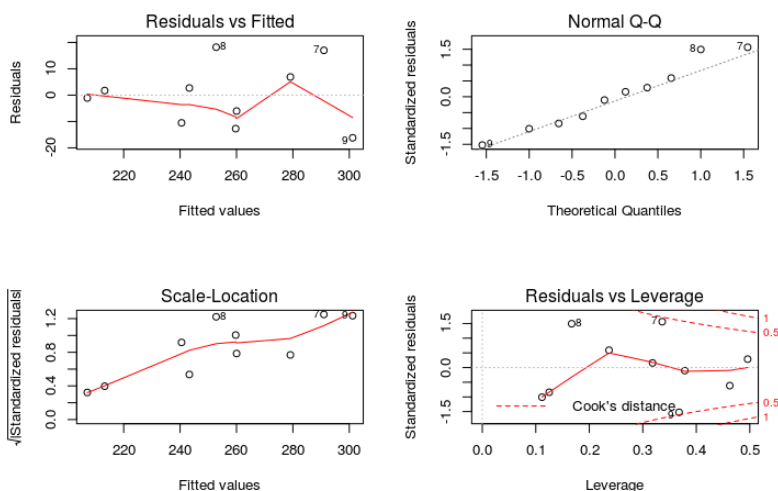
Residual standard error: 13.37 on 7 degrees of freedom  
Multiple R-squared: 0.871, Adjusted R-squared: 0.8342  
F-statistic: 23.64 on 2 and 7 DF, p-value: 0.0007704

**#3.13(a):** The regression equation is  $\hat{\text{Spring}} = -11720 - 1.007 \cdot \text{Fall} + 6.107 \cdot \text{Ayear}$

```

> par(mfrow=c(2,2))
> plot(lm2)

```



**#3.13(b):** The residuals vs fitted value plot shows that the points around the zero line is random and the zero mean assumption holds. The normal Q-Q plot shows a general linear pattern while there is one point shows significant difference from the linear line. Generally, this model is acceptable.

```
> ## Problem 3.14
> summary(lm2)$r.squared
[1] 0.8710292
> summary(lm2)$adj.r.squared
[1] 0.8341804
```

**#3.14(a):** The R-squared value is about 87.1% which means that there is 87.1% of Spring enrollment can be explained by this model.

```
> summary(lm2)$sigma
[1] 13.36684
```

**#3.14(b):** The standard error is about 13.37 which is the size of typical error for this multiple regression.

```
> anova(lm2)
Analysis of Variance Table
```

Response: Spring

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Fall	1	4721.1	4721.1	26.423	0.001338 **
Ayear	1	3725.8	3725.8	20.852	0.002585 **
Residuals	7	1250.7	178.7		

---

Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

**#3.14(c):** From the anova table shows the p-values of F-test for both Fall and Ayear, which are 0.001338 and 0.002585, are smaller than 0.05, thus we say that these two variables are significant.

**#3.14(d):** The hypotheses for Fall are  $H_0: \beta_1 = 0$  and  $H_1: \beta_1 \neq 0$ . The t statistic value is -4.933 and the p-value is 0.00169.

The hypotheses for Ayear are  $H_0: \beta_1 = 0$  and  $H_1: \beta_1 \neq 0$ . The t statistic value is 4.566 and the p-value is 0.00258.

For both Ayear and Fall, the p-value is smaller than 0.05 which means we reject the  $H_0$  and the coefficient for Fall and Ayear are different from zero.

> ## Problem 3.19

> data(Speed)

> head(Speed)

	Year	FatalityRate	StateControl
1	1987	2.41	0
2	1988	2.32	0
3	1989	2.17	0
4	1990	2.08	0
5	1991	1.91	0
6	1992	1.75	0

> lm3=lm(FatalityRate~Year, data=Speed)

> summary(lm3)

Call:

lm(formula = FatalityRate ~ Year, data = Speed)

Residuals:

	Min	1Q	Median	3Q	Max
	-0.18959	-0.07550	-0.02576	0.09346	0.24606

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	91.320887	8.374227	10.9	1.28e-09 ***
Year	-0.044870	0.004193	-10.7	1.75e-09 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1164 on 19 degrees of freedom

Multiple R-squared: 0.8577, Adjusted R-squared: 0.8502

F-statistic: 114.5 on 1 and 19 DF, p-value: 1.75e-09

#3.19(a): The regression equation is  $\hat{\text{FatalityRate}} = 91.32 - 0.045 \cdot \text{Year}$ , and the slope is -0.045 which shows this is a decline line.

> anova(lm3)

Analysis of Variance Table

Response: FatalityRate

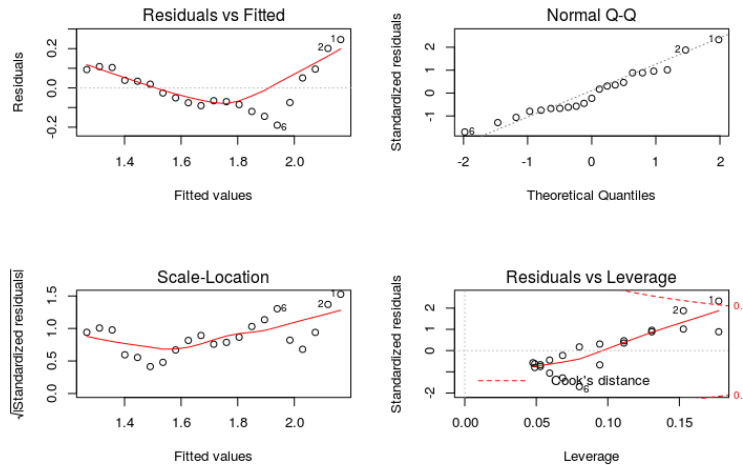
	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Year	1	1.55026	1.55026	114.49	1.75e-09 ***
Residuals	19	0.25726	0.01354		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

> par(mfrow=c(2,2))

> plot(lm3)



**#3.19(b):** The residuals vs fitted value plot is in V shape from the plot which is not that random and doesn't against the zero mean assumption.

```
> lm4=lm(formula = FatalityRate~Year+StateControl+Year*StateControl, data=Speed)
> summary(lm4)
Call:
lm(formula = FatalityRate ~ Year + StateControl + Year * StateControl,
    data = Speed)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.103571	-0.020769	0.004048	0.022473	0.091667

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.162e+02	1.303e+01	16.59	6.19e-12 ***
Year	-1.076e-01	6.548e-03	-16.44	7.19e-12 ***
StateControl	-1.614e+02	1.447e+01	-11.15	3.07e-09 ***
Year:StateControl	8.097e-02	7.264e-03	11.15	3.08e-09 ***

---  
 Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.04243 on 17 degrees of freedom  
 Multiple R-squared: 0.9831, Adjusted R-squared: 0.9801  
 F-statistic: 329 on 3 and 17 DF, p-value: 2.998e-15

```
> anova(lm4)
Analysis of Variance Table
```

Response: FatalityRate

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Year	1	1.55026	1.55026	860.9841	5.288e-16 ***
StateControl	1	0.00292	0.00292	1.6211	0.2201
Year:StateControl	1	0.22373	0.22373	124.2562	3.082e-09 ***
Residuals	17	0.03061	0.00180		

---  
 Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

**#3.19(c):** The regression equation for this model is  $\hat{\text{FatalityRate}} = 216.2 - 0.1076 \cdot \text{Year} - 161.4 \cdot \text{StateControl} + 0.08097 \cdot \text{Year} \cdot \text{StateControl}$ . The p-value for all state control, year and the year\*state control are very small (nearly equal to 0). This shows that all these three factors are very important for this model and the relationship between fatality rate and year is different before and after 1995.

**#3.19(d):** Plug the value 0 and 1 for state control into the regression before, we

get two different equations:

1.hat(FatalityRate) = 216.2-0.1076\*Year for Year<1995

2.hat(FatalityRate) = 54.8-0.02663\*Year for Year>1995

> ## Problem 3.21

> data(BritishUnions)

> head(BritishUnions)

	Date	AgreePct	DisagreePct	NetSupport	Months	Late	Unemployment
1	Oct-75	75	16	-59	2	0	4.9
2	Aug-77	79	17	-62	23	0	5.7
3	Sep-78	82	16	-66	36	0	5.5
4	Sep-79	80	16	-64	48	0	5.4
5	Jul-80	72	19	-53	58	0	6.8
6	Nov-81	70	22	-48	74	0	10.2

> BritishUnions\$Late = as.factor(BritishUnions\$Late)

> lm5=lm(formula = NetSupport~Months+Late+Months\*Late, data=BritishUnions)

> summary(lm5)

Call:

lm(formula = NetSupport ~ Months + Late + Months \* Late, data = BritishUnions)

Residuals:

	Min	1Q	Median	3Q	Max
	-9.6724	-5.3454	-0.1211	3.6432	14.7972

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-66.62827	4.94880	-13.464	5.2e-09 ***
Months	0.21037	0.07392	2.846	0.0138 *
Late1	13.11464	21.57377	0.608	0.5537
Months:Late1	0.17398	0.12761	1.363	0.1959

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7.241 on 13 degrees of freedom

Multiple R-squared: 0.9752, Adjusted R-squared: 0.9695

F-statistic: 170.4 on 3 and 13 DF, p-value: 1.102e-10

> anova(lm5)

Analysis of Variance Table

Response: NetSupport

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Months	1	25734.8	25734.8	490.8652	1.04e-11 ***
Late	1	965.7	965.7	18.4200	0.0008764 ***
Months:Late	1	97.5	97.5	1.8588	0.1959052
Residuals	13	681.6	52.4		

---

Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

#3.21(a): The regression equation is  $\hat{\text{NetSupport}} = -66.628 + 0.2104 \cdot \text{Months} + 13.115 \cdot \text{Late1} + 0.1740 \cdot \text{Months} \cdot \text{Late1}$

#3.21(b): The t statistic value is 1.363 for Months:Late1 and the p-value is 0.1959 which is vary larger than 0.05. So we say that parallel lines are adequate for describing the relationship between NetSupport and Months and the interaction can be dropped from this model. We can't drop Late factor because we don't know what the model will look like when interaction dropped.

```
> lm6=lm(formula = NetSupport~Months, data=BritishUnions)
> anova(lm6,lm5)
```

Analysis of Variance Table

Model 1: NetSupport ~ Months

Model 2: NetSupport ~ Months + Late + Months \* Late

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	15	1744.73				
2	13	681.56	2	1063.2	10.139	0.002221 **

---

Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

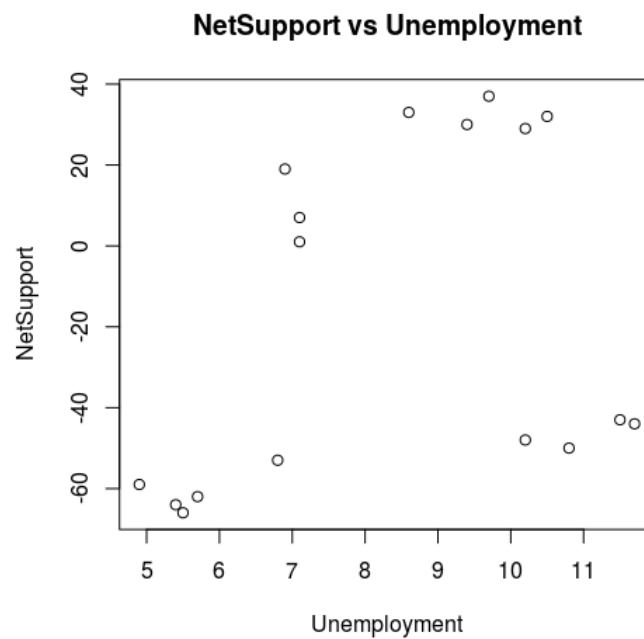
#3.21(c): The F test statistic value is 10.139 from the anova table and the p-value is 0.002221 which is very small. The hypotheses are  $H_0: \beta_2 = \beta_3 = 0$  and  $H_1: \beta_2 \neq 0$  and  $\beta_3 \neq 0$ . Since the p-value is very small, we reject the  $H_0$  and the Late factor is important for the model.



```

> ## Problem 3.22
> par(mfrow=c(1,1))
> plot(NetSupport~Unemployment, data=BritishUnions, main="NetSupport vs Unemployment")

```



**#3.22(a):** The scatter plot shows above. The plot shows that unemployment rate vs net support is random and can't find a specific pattern for it.

```

> lm7=lm(formula = NetSupport~Unemployment, data=BritishUnions)
> summary(lm7)

```

Call:  
 lm(formula = NetSupport ~ Unemployment, data = BritishUnions)

Residuals:

Min	1Q	Median	3Q	Max
-46.93	-31.23	-20.64	36.87	49.23

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-67.660	37.862	-1.787	0.0942 .
Unemployment	5.980	4.379	1.366	0.1921

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

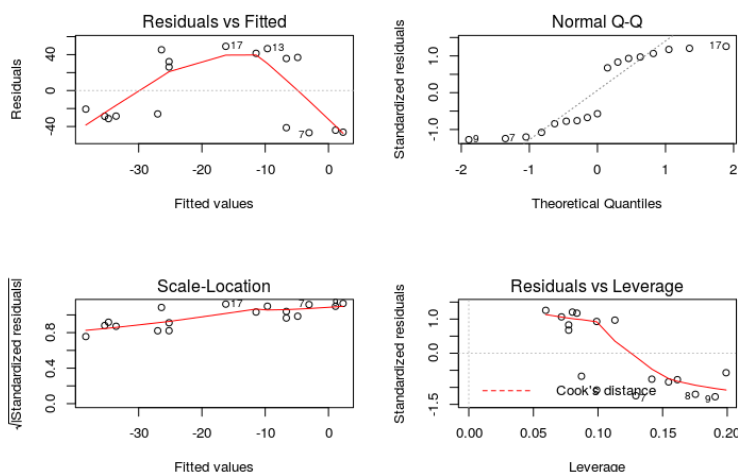
Residual standard error: 40.37 on 15 degrees of freedom

Multiple R-squared: 0.1106, Adjusted R-squared: 0.05132

F-statistic: 1.865 on 1 and 15 DF, p-value: 0.1921

> par(mfrow=c(2,2))

> plot(lm7)



**#3.22(b):** From the residuals vs fitted value plot, we know that the points around the zero line are not very random and the zero mean assumption is againsted, so this model is not that acceptable. The p-value from the summary is 0.192 which is very larger than 0.05 so we can say that unemployment is not that significant for the net support.

> lm8=lm(formula = NetSupport~Unemployment+Months, data=BritishUnions)

> summary(lm8)

Call:

lm(formula = NetSupport ~ Unemployment + Months, data = BritishUnions)

Residuals:

	Min	1Q	Median	3Q	Max
	-11.628	-6.924	-2.717	4.554	19.202

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-65.51220	9.27541	-7.063	5.66e-06 ***
Unemployment	-2.35767	1.20207	-1.961	0.07 .
Months	0.53898	0.03508	15.362	3.71e-10 ***

---

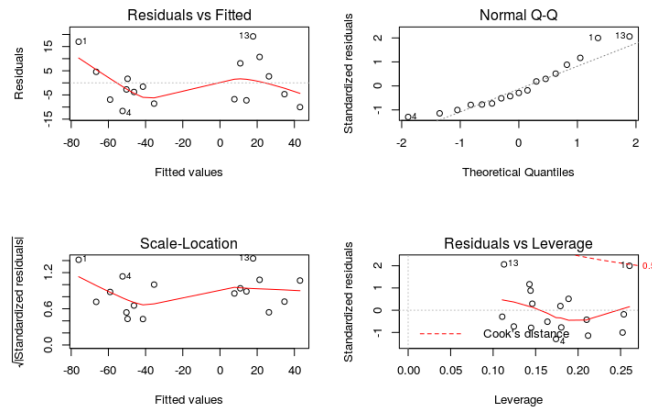
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 9.887 on 14 degrees of freedom

Multiple R-squared: 0.9502, Adjusted R-squared: 0.9431

F-statistic: 133.5 on 2 and 14 DF, p-value: 7.603e-10

> plot(lm8)



**#3.22(c):** From the summary, the p-value for months is very small which indicate that this factor is important for this model and the residuals vs fitted value plot shows that the variances around the zero line is acceptable and doesn't against the zero mean assumption. As compare at the level of 0.10, p-value for both unemployment and months are smaller than 0.10. So the model fitted is acceptable and we can reject the  $H_0: \beta_1=0$ .

**#3.22(d):** The p-value of unemployment is smaller in part c than part b which indicate that its importance becomes higher in the second model. The coefficient is positive in the first model and negative in the second model. Thus, we can conclude that after adjust months, unemployment rate can be associated with net support in better way.

```
> ## Problem 3.30
> data(Pollster08)
> head(Pollster08)
```

	PollTaker	PollDates	MidDate	Days	n	Pop	McCain	Obama	Margin	Charlie
1	Rasmussen	8/28-30/08	8/29	1	3000	LV	46	49	3	0
2	Zogby	8/29-30/08	8/30	2	2020	LV	47	45	-2	0
3	Diageo/Hotline	8/29-31/08	8/30	2	805	RV	39	48	9	0
4	CBS	8/29-31/08	8/30	2	781	RV	40	48	8	0
5	CNN	8/29-31/08	8/30	2	927	RV	48	49	1	0
6	Rasmussen	8/30-9/1/08	8/31	3	3000	LV	45	51	6	0

```
Meltdown
```

1	0
2	0
3	0
4	0
5	0
6	0

```
> lm9=lm(formula = Margin~Days+I(Days^2),data=Pollster08)
> summary(lm9)
```

Call:

```
lm(formula = Margin ~ Days + I(Days^2), data = Pollster08)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-10.7496	-2.0461	-0.1227	1.9297	6.8969

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	4.477958	1.095676	4.087	8.89e-05 ***
Days	-0.604426	0.138598	-4.361	3.18e-05 ***
I(Days^2)	0.021129	0.003776	5.595	1.97e-07 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.014 on 99 degrees of freedom

Multiple R-squared: 0.3495, Adjusted R-squared: 0.3363

F-statistic: 26.59 on 2 and 99 DF, p-value: 5.711e-10

```
> anova(lm9)
```

Analysis of Variance Table

Response: Margin

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Days	1	198.74	198.736	21.879	9.205e-06 ***
I(Days^2)	1	284.34	284.345	31.304	1.966e-07 ***
Residuals	99	899.24	9.083		

---

Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

#3.30(a): The R-squared value is 34.95% and the SSE is 899.24 from summary and the anova table.

```
> Pollster08$Charlie = as.factor(Pollster08$Charlie)
> lm10=lm(formula = Margin~Days+Charlie, data=Pollster08)
> summary(lm10)
Call:
lm(formula = Margin ~ Days + Charlie, data = Pollster08)

Residuals:
    Min       1Q   Median       3Q      Max
-10.7871  -2.1513  -0.1123   1.7988   9.0684

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.31282    0.78692  -0.398   0.6918
Days         0.12222    0.06774   1.804   0.0742 .
Charlie1     0.63640    1.30386   0.488   0.6266
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.454 on 99 degrees of freedom
Multiple R-squared:  0.1458,    Adjusted R-squared:  0.1286
F-statistic: 8.451 on 2 and 99 DF,  p-value: 0.0004089
> anova(lm10)
Analysis of Variance Table
```

```
Response: Margin
      Df Sum Sq Mean Sq F value    Pr(>F)
Days    1  198.74  198.736  16.6630 9.056e-05 ***
Charlie 1    2.84   2.841   0.2382   0.6266
Residuals 99 1180.75  11.927
---
Signif. Codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

#3.30(b): The R-squared value is 14.58% and the SSE is 1180.75 from summary and the anova table.

```
> Pollster08$Meltdown = as.factor(Pollster08$Meltdown)
> lm11=lm(formula = Margin~Days+Meltdown, data=Pollster08)
> summary(lm11)
Call:
lm(formula = Margin ~ Days + Meltdown, data = Pollster08)

Residuals:
    Min       1Q   Median       3Q      Max
-10.7480  -2.5448   0.0408   2.0390   8.2618

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.735333    0.881351   0.834   0.4061
Days        0.001412    0.072015   0.020   0.9844
Meltdown1    3.187183    1.340626   2.377   0.0194 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.363 on 99 degrees of freedom
Multiple R-squared:  0.19,    Adjusted R-squared:  0.1736
F-statistic: 11.61 on 2 and 99 DF,  p-value: 2.949e-05
> anova(lm11)
Analysis of Variance Table
```

Response: Margin

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Days	1	198.74	198.736	17.572	6.024e-05 ***
Meltdown	1	63.92	63.922	5.652	0.01936 *
Residuals	99	1119.67	11.310		

---

Signif. Codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

**#3.30(c):** The R-squared value is 19% and the SSE is 1119.67 from summary and the anova table.

**#3.30(d):** After comparing these three model I would choose model in part a because it has the largest R-squared value which means it can explain the most data for margin on days among these three models and the SSE for the model in part a is the smallest one among all three models which means its predicting error is the smallest in these three.