

Learning from MOM's principles

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Test of robustness of the LASSO

$$Y = \langle X, t^* \rangle + \mathcal{N}(0, 1) \text{ and } (X_1, Y_1), \dots, (X_N, Y_N) \stackrel{i.i.d.}{\sim} (X, Y)$$

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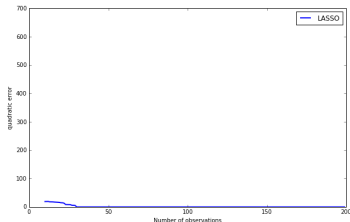
| | |
|-----------|----------------|
| Y_1 | X_1^\top |
| Y_2 | X_2^\top |
| \dots | \dots |
| Y_{100} | X_{100}^\top |
| \dots | \dots |
| Y_n | X_n^\top |
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$$\left\} \operatorname{argmin}_{t \in \mathbb{R}^d} \left(\frac{1}{n} \sum_{i=1}^n (Y_i - \langle X_i, t \rangle)^2 + \sqrt{\frac{2 \log d}{N}} \|t\|_1 \right)$$

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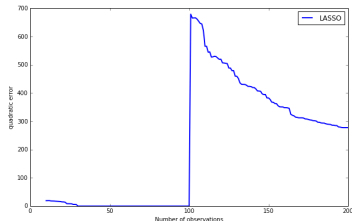
$$\left\{ \begin{array}{c|c} \begin{array}{c} Y_1 \\ Y_2 \\ \dots \\ \tilde{Y}_{100} = 1\bar{M} \\ \dots \\ Y_n \\ \dots \\ Y_N \end{array} & \begin{array}{c} X_1^\top \\ X_2^\top \\ \dots \\ \tilde{X}_{100}^\top = (1)_1^d \\ \dots \\ X_n^\top \\ \dots \\ X_N^\top \end{array} \end{array} \right\} \operatorname{argmin}_{t \in \mathbb{R}^d} \left(\frac{1}{n} \sum_{i=1}^n (Y_i - \langle X_i, t \rangle)^2 + \sqrt{\frac{2 \log d}{N}} \|t\|_1 \right)$$

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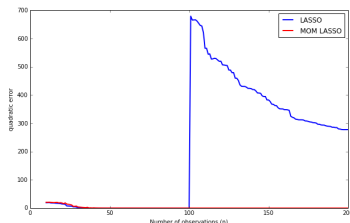


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Test of robustness in classification

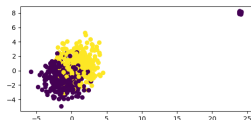
Dataset made of :

- ▶ 600 informative data $\stackrel{i.i.d.}{\sim} (Y, X)$ s.t. $\mathcal{L}(X|Y = 1) = \mathcal{N}((1, 1), 1.4I)$, $\mathcal{L}(X|Y = -1) = \mathcal{N}((-1, -1), 1.4I)$ and $\mathbb{P}(Y = 1) = \mathbb{P}(Y = -1)$.

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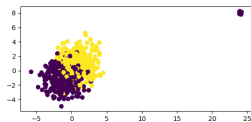
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- ▶ 30 outliers data in the top corner: $Y = -1$ and $X \sim \mathcal{N}((24, 8), 0.1)$



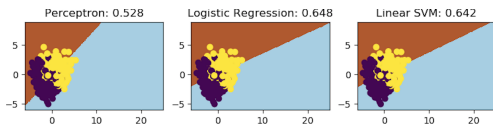
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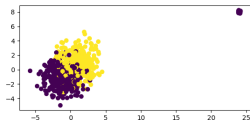
Classical procedures (Perceptron, Logistic regression, SVM):



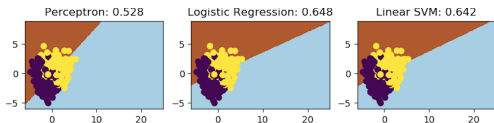
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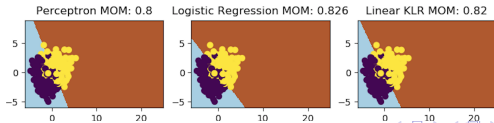
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Their MOM (Median Of Means) version:



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- ▶ Huge datasets are likely to be corrupted by outliers
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[Huber and Ronchetti, "Robust Statistics"]:

"..we can act as if the X_i 's are free of gross error"

The leverage point problem \leadsto preprocessing

Construct procedures robust to outliers in the X_i 's

A benchmark result: Let $(X_i, Y_i)_{i=1}^N$ be

► i.i.d. $\sim (X, Y)$

► $Y = \langle X, t^* \rangle + \zeta$ where $X \sim \mathcal{N}(0, I_{d \times d})$ and $\zeta \sim \mathcal{N}(0, \sigma^2)$ ind. of X ,

then OLS $\hat{t} \in \operatorname{argmin}_{t \in \mathbb{R}^d} \sum_{i=1}^N (Y_i - \langle X_i, t \rangle)^2$ satisfies with probability at least $1 - c_0 \exp(-c_1 d)$,

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Question

Is it possible to construct an estimator satisfying the very same result when 1) the dataset is corrupted by **outliers** and 2) under **weak moment assumption**

From i.i.d. to the $\mathcal{O} \cup \mathcal{I}$ framework

Aim: (X, Y) a r.v., estimate $t^* \in \operatorname{argmin}_{t \in \mathbb{R}^d} \mathbb{E}(Y - \langle X, t \rangle)^2$.

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- ▶ $\forall t \in \mathbb{R}^d$, $\|\langle X_i, t \rangle\|_{L_2} \leq \theta_1 \|\langle X_i, t \rangle\|_{L_1}$
(small ball assumption from [Koltchinskii & Mendelson], [van de Geer & Muro], [Oliveira])

Result for the MOM OLS

In the $\mathcal{O} \cup \mathcal{I}$ framework, the MOM OLS \tilde{t}_d with number of blocks $K = d$ where

$$\tilde{t}_d \in \operatorname{argmin}_{t \in \mathbb{R}^d} \sup_{t' \in \mathbb{R}^d} \operatorname{MOM}_{K=d}(\ell_t - \ell_{t'})$$

is such that with probability at least $1 - c_0 \exp(-c_1 d)$,

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Conclusion

It is possible to recover the same result in the $\mathcal{O} \cup \mathcal{I}$ framework as in the i.i.d. Gaussian with independent noise framework.

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Construction of MOM estimators II: MOM's principle

Refs:

- * [Nemirovsky, Yudin. 1983]
- * [Jerrum, Valiant, Vazirani. 1986]
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Key idea: $MOM_K(Z)$ is a subgaussian estimator of $\mathbb{E}Z$ under a L_2 -moment assumption: if $\|Z\|_{L_2} < \infty$ then with probability at least $1 - c_0 \exp(-c_1 K)$,

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Adaptation to K via Lepski's method:

- ▶ $\hat{l}_K = \left[MOM_K(Z) - \sigma \sqrt{K/N}, MOM_K(Z) + \sigma \sqrt{K/N} \right]$
- ▶ $\hat{K} = \min \left(K : \cap_{k=K}^N \hat{l}_k \neq \emptyset \right)$
- ▶ $\tilde{\mu} \in \cap_{k=\hat{K}}^N \hat{l}_k$

Construction of MOM estimators III: MOM's principle

Aim: We are given:

- ▶ (X, Y) , F and $f^* \in \operatorname{argmin}_{f \in F} R(f)$ where $R(f) = \mathbb{E} \ell_f(X, Y)$ like
$$\ell_f(x, y) = (y - f(x))^2, \log(1 + e^{-yf(x)}), (1 - yf(x))_+, \rho_\kappa(y - f(x))$$
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We want to

- ▶ Estimate f^* : w.h.p. $\left\| \hat{f} - f^* \right\|_{L_2}^2 \leq \text{rate}$
- ▶ Predict Y : w.h.p. $R(\hat{f}) \leq \inf_{f \in F} R(f) + \text{residue}$

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Classical approach via ERM: $\hat{f} \in \operatorname{argmin}_{f \in F} R_N(f)$ where

$$R_N(f) = P_N \ell_f = \frac{1}{N} \sum_{i=1}^N \ell_f(X_i, Y_i)$$

Construction of MOM estimators III: MOM's principle

Aim: We are given:

- ▶ (X, Y) , F and $f^* \in \operatorname{argmin}_{f \in F} R(f)$ where $R(f) = \mathbb{E} \ell_f(X, Y)$ like
$$\ell_f(x, y) = (y - f(x))^2, \log(1 + e^{-yf(x)}), (1 - yf(x))_+, \rho_\kappa(y - f(x))$$
- ▶ $(X_1, Y_1), \dots, (X_N, Y_N)$ some data.

We want to

- ▶ Estimate f^* : w.h.p. $\|\hat{f} - f^*\|_{L_2}^2 \leq \text{rate}$
- ▶ Predict Y : w.h.p. $R(\hat{f}) \leq \inf_{f \in F} R(f) + \text{residue}$

Classical approach via ERM: $\hat{f} \in \operatorname{argmin}_{f \in F} R_N(f)$ where

$$R_N(f) = P_N \ell_f = \frac{1}{N} \sum_{i=1}^N \ell_f(X_i, Y_i)$$

Main idea

Replace the (non-robust) empirical mean $P_N \ell_f$ by a MOM $MOM_K(\ell_f)$ to estimate $R(f) = P \ell_f$

Construction of MOM estimators IV: MOM's principle

1) **MOM minimizer:** $\bar{f} \in \operatorname{argmin}_{f \in F} MOM_K(\ell_f)$ where

$$MOM_K(\ell_f) = \operatorname{Median}(P_{B_1}\ell_f, \dots, P_{B_K}\ell_f).$$

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Minmax MOM estimator. Statistical properties I

Aims: (X, Y) , estimate $f^* \in \operatorname{argmin}_{f \in F} \mathbb{E}(Y - f(X))^2$ and predict Y

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$$\begin{aligned} 2. \forall i \in \mathcal{I}, f \in F \quad & \|f(X_i) - f^*(X_i)\|_{L_2} = \|f(X) - f^*(X)\|_{L_2} \\ & \|Y_i - f(X_i)\|_{L_2} = \|Y - f(X)\|_{L_2} \end{aligned}$$

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 $\|Y_i - f(X_i)\|_{L_2} = \|Y - f(X)\|_{L_2}$
- ▶ $\zeta_i := Y_i - f^*(X_i)$, we assume that for all $f \in F$

$$\operatorname{var}(\zeta_i(f(X_i) - f^*(X_i))) \leq \sigma^2 \mathbb{E}(f(X_i) - f^*(X_i))^2$$

- ▶ $\forall f \in F, \|f(X_i) - f^*(X_i)\|_{L_2} \leq \theta_1 \|f(X_i) - f^*(X_i)\|_{L_1}$ (SBA)

Minmax MOM estimator. Statistical properties II

Two fixed points measuring the complexity of the problem:

$$r_Q(\gamma_Q) = \inf \left\{ r > 0 : \forall J \subset \mathcal{I}, |J| \geq \frac{N}{2}, \mathbb{E} \sup_{\substack{g \in F - f^* \\ \|g\|_{L_p^2} \leq r}} \left| \sum_{i \in J} \epsilon_i g(X_i) \right| \leq \gamma_Q |J| r \right\}$$
$$r_M(\gamma_M) = \inf \left\{ r > 0 : \forall J \subset \mathcal{I}, |J| \geq \frac{N}{2}, \mathbb{E} \sup_{\substack{g \in F - f^* \\ \|g\|_{L_p^2} \leq r}} \left| \sum_{i \in J} \epsilon_i \zeta_i g(X_i) \right| \leq \gamma_M |J| r^2 \right\}$$

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Let

$$r^* = \max\{r_Q(\gamma_Q), r_M(\gamma_M)\}.$$

$(r^*)^2$ is the minimax rate of convergence in the i.i.d. framework with Gaussian design and Gaussian noise independent of the design [L. & Mendelson].

Minmax MOM estimator. Statistical properties III

Theorem

In the $\mathcal{O} \cup \mathcal{I}$ framework. Let $K \in [\max(N(r^*)^2/\sigma^2, |\mathcal{O}|), N]$. With probability at least $1 - c_0 \exp(-c_1 K)$, the minmax MOM estimator

$$\hat{f}_K \in \operatorname{argmin}_{f \in F} \sup_{g \in F} \operatorname{MOM}_K(\ell_f - \ell_g)$$

satisfies

$$\left\| \hat{f}_K - f^* \right\|_{L_2}^2 \leq c_3 \frac{\sigma^2 K}{N} \text{ and } R(\hat{f}_K) \leq \inf_{f \in F} R(f) + \frac{c_4 \sigma^2 K}{N}.$$

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In particular, for $K = \max(N(r^*)^2, |\mathcal{O}|)$,

$$\left\| \hat{f}_K - f^* \right\|_{L_2}^2 \leq R(\hat{f}_K) - \inf_{f \in F} R(f) \leq c_4 \max \left((r^*)^2, \frac{\sigma^2 |\mathcal{O}|}{N} \right)$$

$= c_4 (r^*)^2$ (the minimax rate) when $\sigma^2 |\mathcal{O}| \leq N (r^*)^2$.

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(then, adaptation to K via Lepski's method).

Regularized minmax MOM estimators

$$\hat{f}_K \in \operatorname{argmin}_{f \in F} \sup_{g \in F} \operatorname{MOM}_K(\ell_f - \ell_g) + \lambda(\|f\| - \|g\|)$$

General results:

- ▶ sparsity oracle inequalities and sparse estimation rates (when $\|\cdot\|$ has some sparsity inducing power)

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Example: **MOM version of the LASSO:**

$$\hat{t}_K \in \operatorname{argmin}_{t \in \mathbb{R}^d} \sup_{t' \in \mathbb{R}^d} \operatorname{MOM}_K(\ell_t - \ell_{t'}) + \lambda_K(\|t\|_1 - \|t'\|_1)$$

where $\ell_t(x, y) = (y - \langle x, t \rangle)^2$ and

$$\lambda_K \sim \sigma \sqrt{\frac{1}{N} \log \left(\frac{\sigma^2 d}{K} \right)}$$

MOM version of the LASSO

Aim: Estimate $t^* \in \operatorname{argmin}_{t \in \mathbb{R}^d} \mathbb{E}(Y - \langle X, t \rangle)^2$ w.r.t. $s = \|t^*\|_0$.

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- ▶ No assumption on $|\mathcal{O}|$ observations s.t. $|\mathcal{O}| \leq N/10$

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 4. $\forall t \in \mathbb{R}^d, \operatorname{var}(\zeta \langle X, t \rangle) \leq \sigma^2 \|\langle X, t \rangle\|_{L^2}^2, \|\langle X, t \rangle\|_{L^2} \leq \theta_0 \|\langle X, t \rangle\|_{L^1}$

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for $K = \max(s \log(d/s), |\mathcal{O}|)$. (adaptation via Lepski's method)

Algorithms

Descent methods for the MOM minimizer I

Problem: $u \in \mathbb{R}^d \rightarrow MOM_K(\ell_u)$ is not convex (in general) where

$$MOM_K(\ell_u) = Median \left(\frac{1}{|B_1|} \sum_{i \in B_1} (Y_i - \langle X_i, u \rangle)^2, \dots, \frac{1}{|B_K|} \sum_{i \in B_K} (Y_i - \langle X_i, u \rangle)^2 \right)$$

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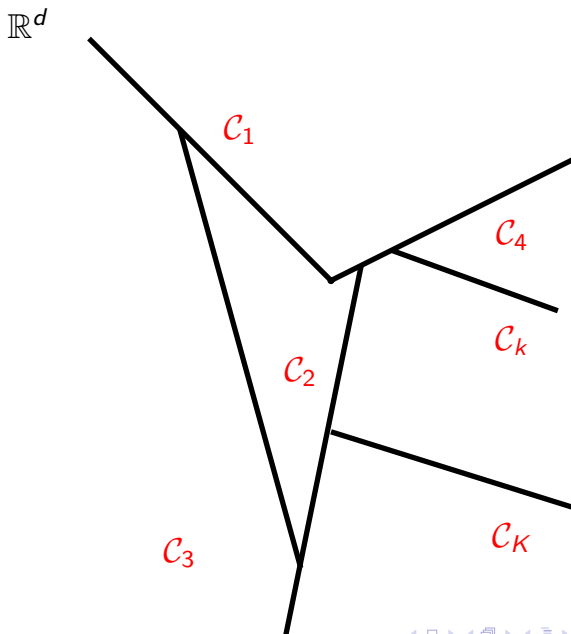
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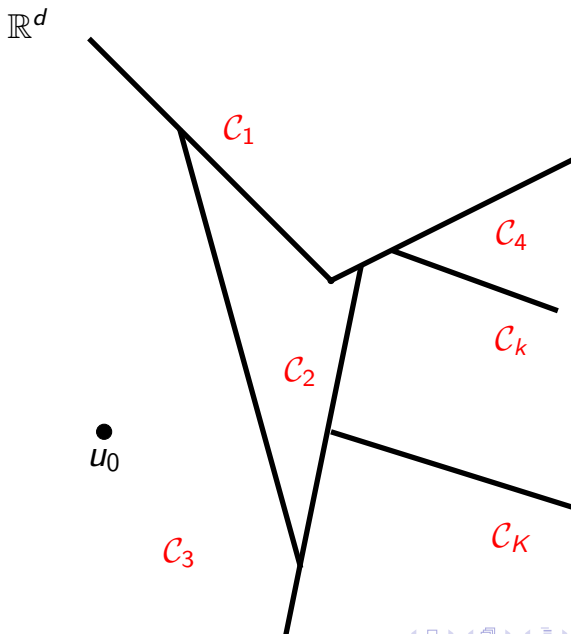
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3. $u_{t+1} = u_t - \eta_t \nabla_t$

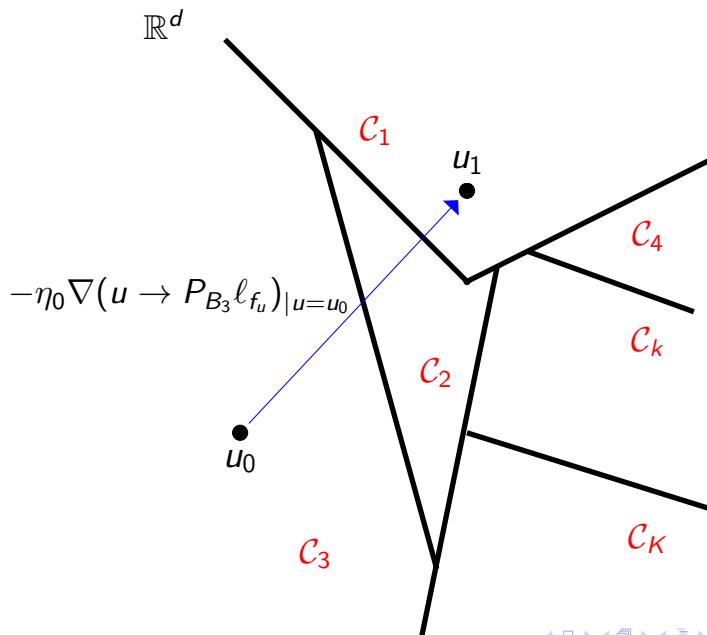
Descent methods for the MOM minimizer II



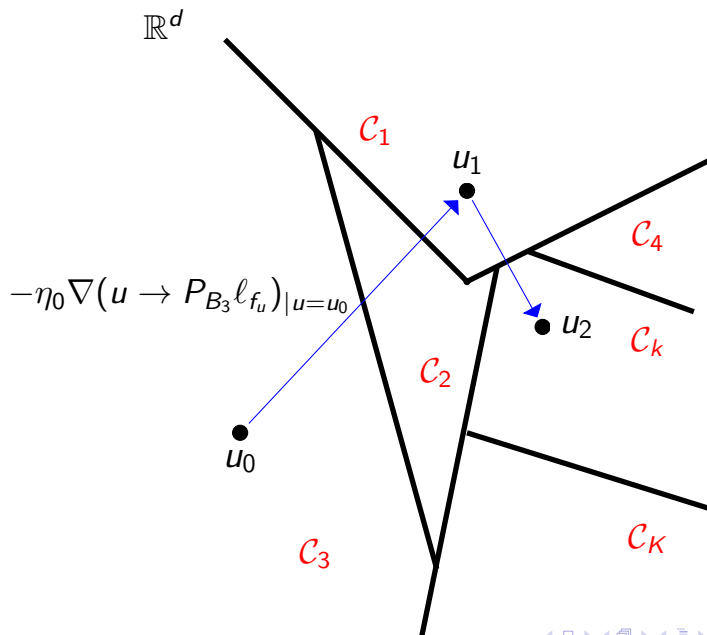
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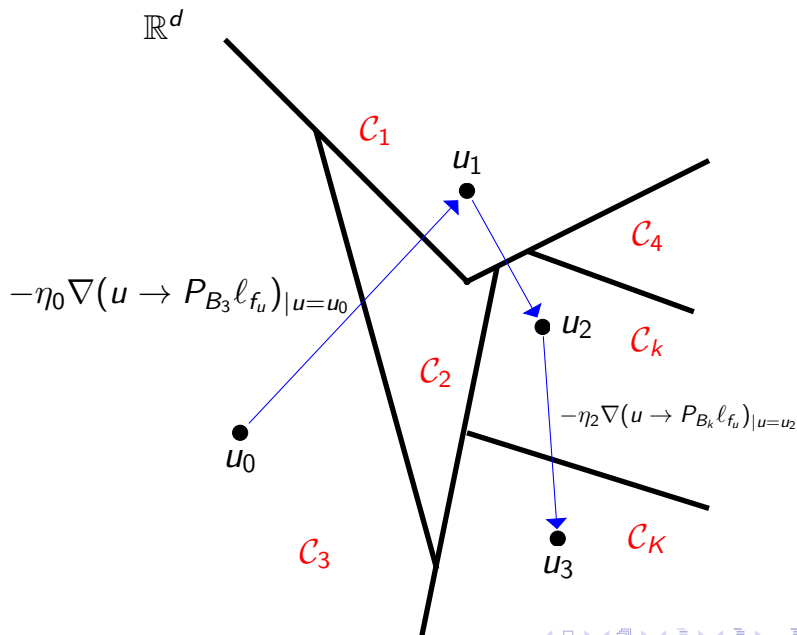
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Descent methods for the MOM minimizer II



MOM GD = BGD

$$\left. \begin{array}{l} (X_1, Y_1) \\ (X_2, Y_2) \\ \vdots \\ (X_{N/K}, Y_{N/K}) \\ (X_{N/K+1}, Y_{N/K+1}) \\ \vdots \\ (X_{N-1}, Y_{N-1}) \\ (X_N, Y_N) \end{array} \right\} P_{B_1} \ell_{u_t}$$
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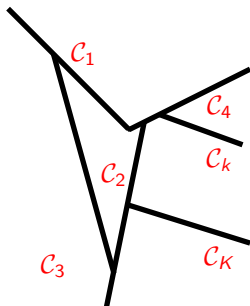
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Idea: Choose the descent block according to its **centrality** via the median operator \leadsto "remove outliers" and closer to $\mathbb{E} \ell_{u_t}$

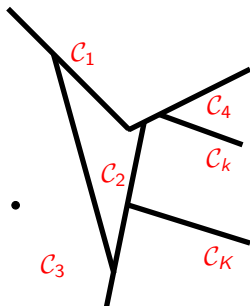
Pb of local minima \Rightarrow Random blocks



Local minima if a cell C_k
contains a minimum from

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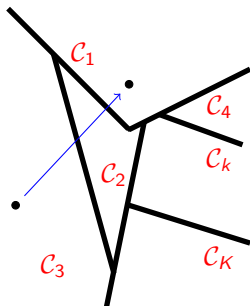
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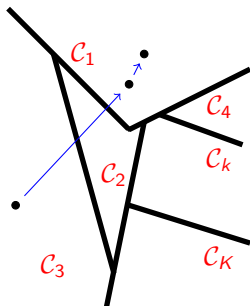
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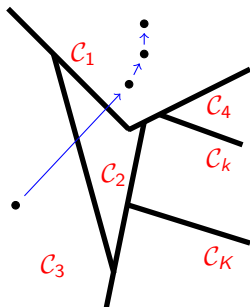
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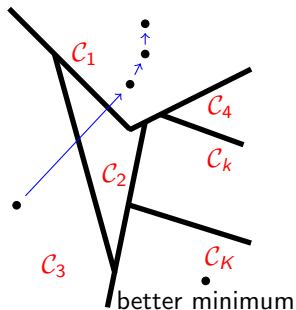
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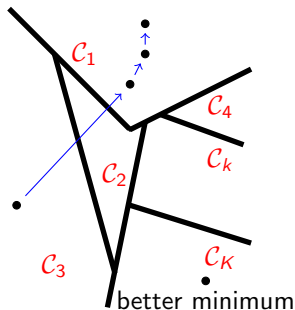
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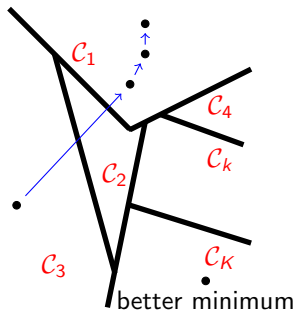
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Solution: choose the blocks of data at **random** at every step:

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MOM GD with random blocks = BSGD with a particular choice of the descent blocks

Convergence of the MOM GD with random blocks

Theorem

Let $\mathcal{D}_N = \{(X_i, Y_i)_{i=1}^N\}$. Assume that

1. $\|\nabla_u \ell_u(x, y)\|_2 \leq L$
2. $\hat{u} \in \operatorname{argmin}_{u \in \mathbb{R}^d} \mathbb{E}_{B_1 \cup \dots \cup B_K} [MOM_K(\ell_u) | \mathcal{D}_N]$ is such that $\forall \epsilon > 0$,

$$\inf_{\|\hat{u} - u\|_2 \geq \epsilon} \langle \hat{u} - u, \mathbb{E}[\nabla_u \ell_u(x, y) | \mathcal{D}_N] \rangle > 0$$

3. $\sum_t \eta_t^2 < \infty$ and $\sum_t \eta_t = \infty$
4. for λ_d -almost all $u \in \mathbb{R}^d$, there exists an open set B such that $u \in B$ and for all partition $B_1 \cup \dots \cup B_K$ and $v \in B$, ℓ_u and ℓ_v have the same median block.

Then, for almost all dataset \mathcal{D}_N ,

$$\|u_T - \hat{u}\|_2 \xrightarrow[T \rightarrow \infty]{a.s.} 0$$

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Idea: Alternate between ascent (for the max) and descent (for the min).

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Example for the **minmax MOM version of the LASSO**:

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where $\ell_u(x, y) = (y - \langle x, u \rangle)^2$ and $\lambda_K \sim \sigma \sqrt{(1/N) \log(\sigma^2 d/K)}$.

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At iteration $(u_t, u_{t'})$ we do:

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Simulations: effect of random blocks on local minima

$N = 200$ i.i.d. copies of (X, Y) where

$$Y = \langle X, t^* \rangle + \zeta, \quad X \sim \mathcal{N}(0, I_{d \times d}) \quad \zeta \sim \mathcal{N}(0, 1) \text{ ind. of } X$$

where $d = 500$ and $\|t^*\|_0 = 20$.

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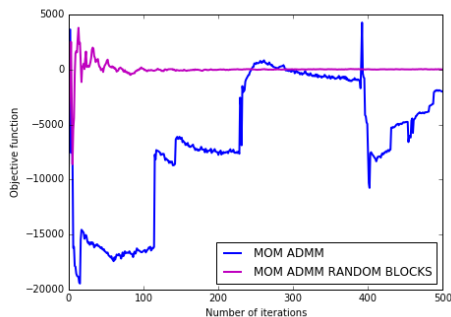
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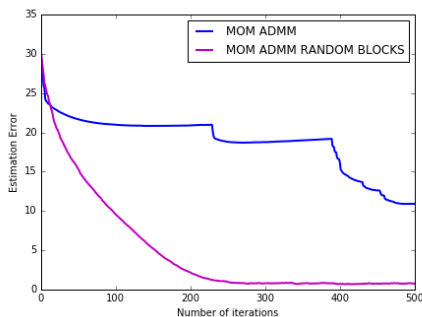
Objective function

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$$\|\hat{t} - t^*\|_2$$



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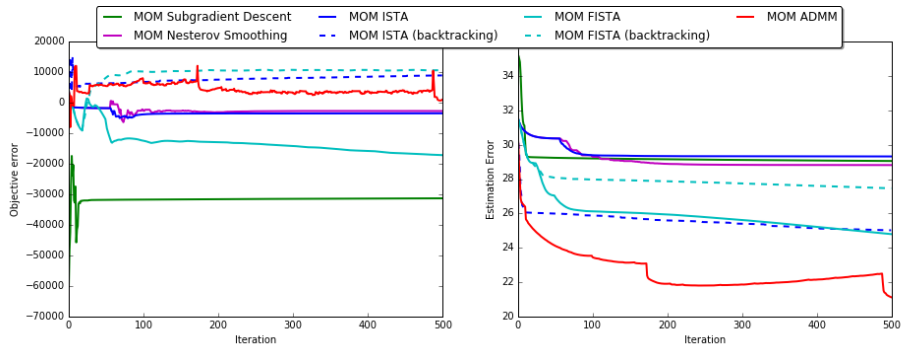
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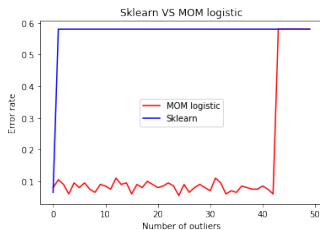
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(non random blocks)

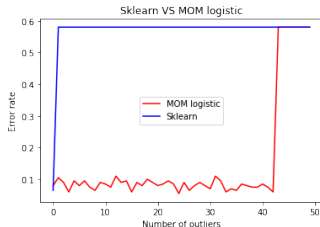
Test of robustness of minmax MOM estimators

Logistic Vs MOM logistic $N = 1000$, $d = 50$, $K = 100$

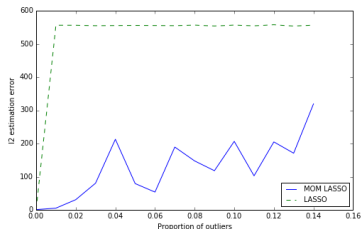


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LASSO Vs MOM LASSO $N = 200$, $d = 500$, $s = 10$, adaptive choice of K and λ



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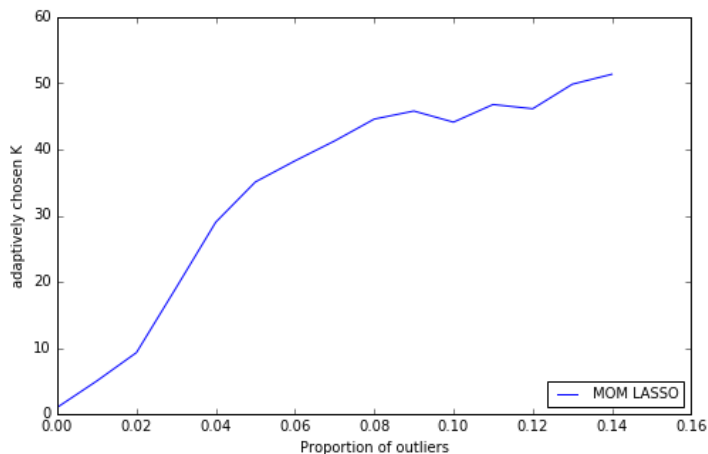
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7. return $\hat{f}_{\hat{K}, \hat{\lambda}}$.

Adaptively chosen number of blocks K



\hat{K} increases with $|\mathcal{O}|/N$ because we need at least $K \geq 2|\mathcal{O}|$ to make MOM estimators working.

An outliers detection algorithm (random blocks)

Idea: Outliers should not be selected in the median blocks along the iterations.

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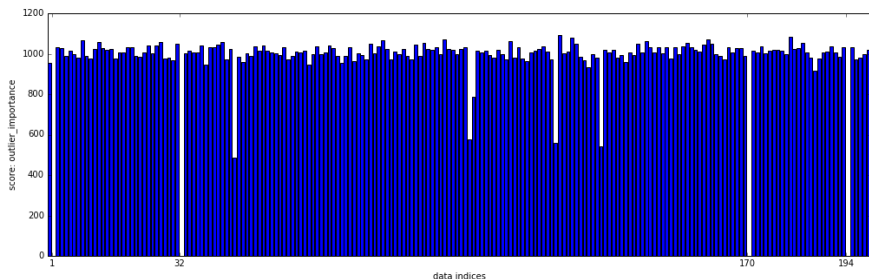
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outliers are data number 1, 32, 170, 194.

Thanks!

Alternating sub-gradient descent

```
input :  $(t_0, t'_0) \in \mathbb{R}^d \times \mathbb{R}^d$  : initial point  
         $(\eta_p)_p, (\beta_p)_p$ : two step size sequences  
output: approximated solution to the min-max problem  
1 for  $t = 1, \dots, T$  do  
2   find  $k \in [K]$  such that  $MOM_K(\ell_{t_p} - \ell_{t'_p}) = P_{B_k}(\ell_{t_p} - \ell_{t'_p})$   
3  
   
$$t_{p+1} = t_p + 2\eta_p \mathbb{X}_k^\top (\mathbb{Y}_k - \mathbb{X}_k t_p) - \lambda \eta_p \text{sign}(t_p)$$
  
4   find  $k \in [K]$  such that  $MOM_K(\ell_{t_{p+1}} - \ell_{t'_p}) = P_{B_k}(\ell_{t_{p+1}} - \ell_{t'_p})$   
5  
   
$$t'_{p+1} = t'_p + 2\beta_p \mathbb{X}_k^\top (\mathbb{Y}_k - \mathbb{X}_k t'_p) - \lambda \beta_p \text{sign}(t'_p)$$
  
6 end  
7 Return  $(t_p, t'_p)$ 
```

Alternating proximal gradient descent

input : $(t_0, t'_0) \in \mathbb{R}^d \times \mathbb{R}^d$: initial point

$(\eta_k)_k, (\beta_k)_k$: two step size sequences

output: approximated solution to the min-max problem

1 **for** $t = 1, \dots, T$ **do**

2 find $k \in [K]$ such that $MOM_K(\ell_{t_p} - \ell_{t'_p}) = P_{B_k}(\ell_{t_p} - \ell_{t'_p})$

$$t_{p+1} = \text{prox}_{\lambda \|\cdot\|_1} (t_p + 2\eta_k \mathbb{X}_k^\top (\mathbb{Y}_k - \mathbb{X}_k t_p))$$

3 find $k \in [K]$ such that $MOM_K(\ell_{t_{p+1}} - \ell_{t'_p}) = P_{B_k}(\ell_{t_{p+1}} - \ell_{t'_p})$

$$t'_{p+1} = \text{prox}_{\lambda \|\cdot\|_1} (t'_p + 2\beta_k \mathbb{X}_k^\top (\mathbb{Y}_k - \mathbb{X}_k t'_p))$$

4 **end**

MOM ADMM

input : $(t_0, t'_0) \in \mathbb{R}^d \times \mathbb{R}^d$: initial point. ρ : a parameter

output: approximated solution to the min-max problem

1 **for** $t = 1, \dots, T$ **do**

2 find $k \in [K]$ such that $MOM_K(\ell_{t_p} - \ell_{t'_p}) = P_{B_k}(\ell_{t_p} - \ell_{t'_p})$

$$t_{p+1} = (\mathbb{X}_k^\top \mathbb{X}_k + \rho I_{d \times d})^{-1} (\mathbb{X}_k^\top \mathbb{Y}_k + \rho z_p - u_p)$$

$$z_{p+1} = \text{prox}_{\lambda \|\cdot\|_1}(t_{p+1} + u_p / \rho)$$

$$u_{p+1} = u_p + \rho(t_{p+1} - z_{p+1})$$

3 find $k \in [K]$ such that $MOM_K(\ell_{t_{p+1}} - \ell_{t'_{p+1}}) = P_{B_k}(\ell_{t_{p+1}} - \ell_{t'_{p+1}})$

$$t'_{p+1} = (\mathbb{X}_k^\top \mathbb{X}_k + \rho I_{d \times d})^{-1} (\mathbb{X}_k^\top \mathbb{Y}_k + \rho z'_p - u'_p)$$

$$z'_{p+1} = \text{prox}_{\lambda \|\cdot\|_1}(t'_{p+1} + u'_p / \rho)$$

$$u'_{p+1} = u'_p + \rho(t'_{p+1} - z'_{p+1})$$

4 **end**

5 **Return** (t_p, t'_p)