

# HW 5 M362M, Fall 2024

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## Problem 1

We always have  $Y \geq 0$  and it only takes integral values. Thus  $\mathbb{P}[Y < 0] = 0$ . For  $i = 0, 1, 2, \dots$ , we have

$$\begin{aligned}\mathbb{P}[Y = i] &= \sum_{n=1}^{\infty} \mathbb{P}[\xi_1 + \dots + \xi_N = i \mid N = n] \cdot \mathbb{P}[N = n] \\ &= \sum_{n=1}^{\infty} \mathbb{P}[\xi_1 + \dots + \xi_n = i] \cdot \mathbb{P}[N = n]\end{aligned}$$

Note that  $\xi_1 + \dots + \xi_n \sim \text{Binom}(n, p_B)$ . If  $i = 0$ , then

$$\begin{aligned}\mathbb{P}[Y = 0] &= \sum_{n=1}^{\infty} \binom{n}{0} p_B^0 (1 - p_B)^n \cdot p_g (1 - p_g)^{n-1} \\ &= (1 - p_B) p_g \cdot \sum_{n=0}^{\infty} [(1 - p_B)(1 - p_g)]^n \\ &= \frac{(1 - p_B) p_g}{1 - (1 - p_B)(1 - p_g)}\end{aligned}$$

If  $i = 1, 2, \dots$ ,  $\mathbb{P}[\xi_1 + \dots + \xi_n = i] = 0$  for  $n < i$ , thus

$$\begin{aligned}\mathbb{P}[Y = i] &= \sum_{n=i}^{\infty} \binom{n}{i} p_B^i (1 - p_B)^{n-i} \cdot p_g (1 - p_g)^{n-1} \\ &= p_g (1 - p_g)^{i-1} p_B^i \cdot \sum_{n=i}^{\infty} \binom{n}{i} [(1 - p_B)(1 - p_g)]^{n-i}\end{aligned}$$

Now we prove the identity, for all  $i \in \mathbb{N}_0$ ,  $|x| < 1$ , we have

$$S_i := \sum_{n=i}^{\infty} \binom{n}{i} x^{n-i} = \frac{1}{(1-x)^{i+1}}.$$

We proceed with induction on  $i$ . When  $i = 0$ , it is a geometric series, and the result holds. Suppose it holds when  $i = k - 1$ , then for  $i = k$ ,

$$\begin{aligned}
S_k - xS_k &= 1 + \sum_{n=k+1}^{\infty} \left[ \binom{n}{k} - \binom{n-1}{k} \right] x^{n-k} \\
&= 1 + \sum_{n=k+1}^{\infty} \binom{n-1}{k-1} x^{n-k} \\
&= S_{k-1} \\
&= \frac{1}{(1-x)^k}
\end{aligned}$$

Therefore

$$S_k = \frac{1}{(1-x)^{k+1}}.$$

Thus the statement holds for all  $i \in \mathbb{N}_0$ . Back to the original problem, we have

$$\mathbb{P}[Y = i] = \begin{cases} \frac{p_g(1-p_g)^{i-1}p_B^i}{[1-(1-p_B)(1-p_g)]^{i+1}}, & i = 1, 2, \dots \\ \frac{(1-p_B)p_g}{1-(1-p_B)(1-p_g)}, & i = 0, \\ 0 & , \text{ else.} \end{cases}$$

## Problem 2

### 2.1

$$\begin{aligned}
M_{X_n}(t) &= M_{\sum_{i=1}^n \delta_i}(t) \\
&= \prod_{i=1}^n M_{\delta_i}(t) && (\{\delta_i\}_{i=1}^n \text{ independent}) \\
&= [M_{\delta_1}(t)]^n && (\delta_1, \dots, \delta_n \text{ identically distributed}) \\
&= \left[ \frac{1}{2}(e^t + e^{-t}) \right]^n
\end{aligned}$$

### 2.2

We first calculate the first two derivatives of  $M_{X_n}(t)$ .

$$M'_{X_n}(t) = n \left[ \frac{1}{2}(e^t + e^{-t}) \right]^{n-1} \cdot \frac{1}{2}(e^t - e^{-t}) = \frac{n}{2^n} (e^t + e^{-t})^{n-1} (e^t - e^{-t}).$$

$$M''_{X_n}(t) = \frac{n}{2^n} [(n-1)(e^t + e^{-t})^{n-2} (e^t - e^{-t})^2 + (e^t + e^{-t})^n].$$

We are now in a position to calculate  $\mathbb{E}[X_n]$  and  $\text{Var}(X_n)$ .

$$\begin{aligned}
\mathbb{E}[X_n] &= M'_{X_n}(0) \\
&= 0, && (e^0 - e^{-0} = 0)
\end{aligned}$$

and

$$\begin{aligned}
\text{Var}(X_n) &= \mathbb{E}[X_n^2] - (\mathbb{E}[X_n])^2 \\
&= \mathbb{E}[X_n^2] & (\mathbb{E}[X_n] = 0) \\
&= M''_{X_n}(0) + M'_{X_n}(0) \\
&= n
\end{aligned}$$

Therefore

$$Z_n = \frac{X_n}{\sqrt{n}}$$

### 2.3

$$\begin{aligned}
M_{Z_n}(t) &= \left[ M_{\frac{\delta_1}{\sqrt{n}}}(t) \right]^n \\
&= \left[ \frac{1}{2} \left( e^{\frac{t}{\sqrt{n}}} + e^{\frac{-t}{\sqrt{n}}} \right) \right]^n
\end{aligned}$$

### 2.4

$$\begin{aligned}
M_{Z_\infty}(t) &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cdot e^{tx} dx \\
&= e^{\frac{1}{2}t^2} \cdot \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x-t)^2} dx \\
&= e^{\frac{1}{2}t^2} \cdot \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2} dy & (y = x - t) \\
&= e^{\frac{1}{2}t^2}
\end{aligned}$$

### 2.5

$$\begin{aligned}
\ln \left( \lim_{n \rightarrow \infty} M_{Z_n}(t) \right) &= \lim_{n \rightarrow \infty} \ln M_{Z_n}(t) & (\ln(\cdot) \text{ is continuous}) \\
&= \lim_{u \rightarrow 0^+} \frac{\ln \left( \frac{e^{ut} + e^{-ut}}{2} \right)}{u^2} & (u = \frac{1}{\sqrt{n}}) \\
&= \lim_{u \rightarrow 0^+} \frac{te^{ut} - te^{-ut}}{2u \cdot \underbrace{(e^{ut} + e^{-ut})}_{\text{Converges to 2}}} & (\text{L'Hospital's Rule}) \\
&= \lim_{u \rightarrow 0^+} \frac{e^{ut} - e^{-ut} + 2t^2 e^{-ut}}{4} & (\text{L'Hospital's Rule}) \\
&= \frac{1}{2}t^2.
\end{aligned}$$

Thus  $\lim_{n \rightarrow \infty} M_{Z_n}(t) = e^{\frac{1}{2}t^2} = M_{Z_\infty}(t)$

## Problem 3

### 3.1

$$\begin{aligned}\mathbb{P}[Z_2 = 4] &= \mathbb{P}[Z_{1,1} = 2, Z_{1,2} = 2] \cdot \mathbb{P}[Z_1 = 2] \\ &= 0.3^2 \times 0.3 \\ &= 0.027\end{aligned}$$

$$\begin{aligned}\mathbb{P}[Z_2 = 3] &= \mathbb{P}[Z_{1,1} = 1, Z_{1,2} = 2] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 2, Z_{1,2} = 1] \cdot \mathbb{P}[Z_1 = 2] \\ &= 0.6 \times 0.3 \times 0.3 + 0.3 \times 0.6 \times 0.3 \\ &= 0.108\end{aligned}$$

$$\begin{aligned}\mathbb{P}[Z_2 = 2] &= \mathbb{P}[Z_{1,1} = 1, Z_{1,2} = 1] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 0, Z_{1,2} = 2] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 2, Z_{1,2} = 0] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 2] \cdot \mathbb{P}[Z_1 = 1] \\ &= 0.6^2 \times 0.3 + 0.1 \times 0.3 \times 0.3 \\ &\quad + 0.3 \times 0.1 \times 0.3 + 0.3 \times 0.6 \\ &= 0.306\end{aligned}$$

$$\begin{aligned}\mathbb{P}[Z_2 = 1] &= \mathbb{P}[Z_{1,1} = 1, Z_{1,2} = 0] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 0, Z_{1,2} = 1] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 1] \cdot \mathbb{P}[Z_1 = 1] \\ &= 0.6 \times 0.1 \times 0.3 + 0.1 \times 0.6 \times 0.3 \\ &\quad + 0.6 \times 0.6 \\ &= 0.396\end{aligned}$$

$$\begin{aligned}\mathbb{P}[Z_2 = 0] &= \mathbb{P}[Z_{1,1} = 0, Z_{1,2} = 0] \cdot \mathbb{P}[Z_1 = 2] \\ &\quad + \mathbb{P}[Z_{1,1} = 0] \cdot \mathbb{P}[Z_1 = 1] \\ &\quad + \mathbb{P}[Z_1 = 0] \\ &= 0.1^2 \times 0.3 + 0.1 \times 0.6 + 0.1 \\ &= 0.163\end{aligned}$$

$$\therefore \mathbb{P}[Z_2 = i] = \begin{cases} 0.027, & i = 4 \\ 0.108, & i = 3 \\ 0.306, & i = 2 \\ 0.396, & i = 1 \\ 0.163, & i = 0 \\ 0, & \text{else.} \end{cases}$$

### 3.2

Let  $Z$  be a random variable with the following probability mass function

$$P(Z = i) = \begin{cases} 0.3, & i = 2 \\ 0.6, & i = 1 \\ 0.1, & i = 0 \\ 0, & \text{else.} \end{cases}$$

Define  $G_Z(s) = \mathbb{E}[s^Z]$ ,  $G_{Z_n}(s) = \mathbb{E}[s^{Z_n}]$  for  $0 \leq s \leq 1$ . We claim that

$$G_{Z_n}(s) = \underbrace{G_Z \circ \cdots \circ G_Z}_{n \text{ times}}(s).$$

We proceed by induction on  $n$ . When  $n = 1$ , the results holds as  $Z_1 = Z$ . Suppose the statement is true for some  $n = k_0$ . Then

$$\begin{aligned} G_{Z_{k_0+1}}(s) &= \mathbb{E} \left[ \mathbb{E} \left[ s^{Z_{k_0+1}} \mid Z_{k_0} \right] \right] && \text{(Tower Property)} \\ &= \mathbb{E} \left[ \mathbb{E} \left[ s^{\sum_{j=1}^{Z_{k_0}} Z_{k_0,j}} \mid Z_{k_0} \right] \right] \\ &= \mathbb{E} \left[ \left( \mathbb{E} [s^Z] \right)^{Z_{k_0}} \right] && (\{Z_{k_0,j}\}_{j=1}^{Z_{k_0}} \text{ i.i.d.}) \\ &= G_{Z_{k_0}}(G_Z(s)) \\ &= \underbrace{G_Z \circ \cdots \circ G_Z}_{k_0 \text{ times}}(G_Z(s)) && \text{(Induction Hypothesis)} \\ &= \underbrace{G_Z \circ \cdots \circ G_Z}_{k_0+1 \text{ times}}(s) \end{aligned}$$

Thus the statement is true for all  $n \in \mathbb{N}$ . Note that

$$G_{Z_n}(s) = \sum_{k=0}^{\infty} s^k \cdot \mathbb{P}[Z_n = k],$$

therefore  $G_{Z_n}(0) = \mathbb{P}[Z_n = 0]$ , and we have

$$\mathbb{P}[Z_n = 0] = \underbrace{G_Z \circ \cdots \circ G_Z}_{n \text{ times}}(0).$$

We know that  $G_Z(s) = 0.1 + 0.6s + 0.3s^2$ , and using a computer,  $\mathbb{P}[Z_{10} = 0] = \underbrace{G_Z \circ \cdots \circ G_Z}_{10 \text{ times}}(0) = 0.3109$ .

(Please Turn Over)

### 3.3

```
n_sim = 100000
p_0 = 0.1
p_1 = 0.6
p_2 = 0.3
# This function returns Z_n (n = 1,2,3, ...)
branch = function(n,p_0,p_1,p_2){
  if (n==1){
    # The "1" here represents Z_0 = 1
    return (sample(c(0,1,2), 1, replace=TRUE, prob=c(p_0,p_1,p_2)))
  }
  # This is Z_{n-1}
  prev = branch(n-1,p_0,p_1,p_2)
  if (prev == 0){
    return (0)
  }
  # This is (Z_{n-1,1}, ..., Z_{n-1,Z_{n-1}})
  offsprings = sample(c(0,1,2), prev, replace=TRUE, prob=c(p_0,p_1,p_2))
  # Return Z_n
  return (sum(offsprings))
}
# These are the results of Z_10 from the 100000 simulations
Z_10 = replicate(n_sim, branch(10,p_0,p_1,p_2))

# Probability of extinction (Theoretical)
the_prob = 0.3109

# Probability of extinction (Simulation)
sim_prob = mean(Z_10==0)

cat(" The probability of extinction by the 10th generation", "\n", "\n",
    "Probability (Theoretical) :", round(the_prob, 4), "\n",
    "Probatility (Simulation) :", round(sim_prob, 4))
## The probability of extinction by the 10th generation
##
## Probability (Theoretical) : 0.3109
## Probatility (Simulation) : 0.3101
```

End of Homework