



Adaptive Randomization with Arm-Elimination

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Adaptive Randomization with Arm-Elimination

GLR Statistics and test

Multi-Arm
Bandit with
Side
Information

Review

treatment effect PDF: $f_{\theta}(x) = e^{\theta x - \psi(\theta)}$

mean: $\psi'(\theta)$ and variance: $\psi''(\theta)$

Special case: Bernoulli Distribution, $\theta = \log(\frac{p}{1-p})$ and $\psi(\theta) = \log(e^{\theta} + 1)$

Kullback-Leibler Information

$$I(\mu, \mu') = \mathbb{E}_{\theta_{\mu}} \left\{ \log \left(\frac{f_{\theta_{\mu}}(X)}{f_{\theta_{\mu'}}(X)} \right) \right\} = (\theta_{\mu} - \theta_{\mu'})\mu - [\psi(\theta_{\mu}) - \psi(\theta_{\mu'})]$$

Generalized Likelihood Ratio (GLR) Statistics

$$l_j^i(k, k') = n_{ijk} \left\{ \hat{\mu}_{ijk} \theta_{\hat{\mu}_{ijk}} - \psi(\theta_{\hat{\mu}_{ijk}}) \right\} + n_{ijk'} \left\{ \hat{\mu}_{ijk'} \theta_{\hat{\mu}_{ijk'}} - \psi(\theta_{\hat{\mu}_{ijk'}}) \right\} - (n_{ijk} + n_{ijk'}) \{ \bar{\mu} \theta_{\bar{\mu}} - \psi(\theta_{\bar{\mu}}) \},$$

where $\bar{\mu} = \frac{n_{ijk} \hat{\mu}_{ijk} + n_{ijk'} \hat{\mu}_{ijk'}}{n_{ijk} + n_{ijk'}}$ of biomarker j and strategy k in interim i . Note that $n_{ij} = \sum_{k=1}^K n_{ijk}$.

Elimination Rule

After trials of interim i , arm(a.k.a strategy) $k \neq \hat{k}_{ij}$ within biomarker j is eliminated if

$$l_j^i(\hat{k}_{ij}, k) \geq a_\alpha$$

where \hat{k}_{ij} refers to the best treatment of biomarker j in interim i

(Note that $l(\lambda, \lambda) = 0$)

a_α and α

- The the best strategy for each biomarker class is not eliminated is guaranteed with the probability $1 - \alpha$
- In my simulation, the best strategy is presumed to be unique
 - An alternative is to choose the best strategy set, with guaranteed probability not being eliminated $1 - \alpha$

Adaptive Randomization with Arm-Elimination

Arm Elimination — Computation of a_α

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○ Restatement of α

$A_j \rightarrow$ the best treatment of biomarker j is eliminated

$$A_j = \{ \max_{k \neq \hat{k}_{ij}} [l_j^i(\hat{k}_{ij}, k) I_{\{\hat{\mu}_{ijk} > \hat{\mu}_{ij\hat{k}_{ij}}\}}] \geq a_\alpha \text{ for some } 1 \leq i \leq I \}$$

$$\text{Thus, } \alpha = P\left(\bigcup_{j=1}^J A_j\right)$$

○ Determine a_α

- P_* is used as the probability measure satisfying $\theta_{j1} = \dots = \theta_{jK}$,
 - Thus α can be computed through the De Morgan's Law
- we can let $1 = \hat{k}_{ij}$, n_{ijk} can be approximated by $\frac{(1+o_p(1))n_{ij}}{K}$

Central Limit Theorem

- $l_j^i(1, k)$ can be approximated by $\left(1 + O_p(1)\right) \frac{n_{ij}}{4K\psi''(\theta_{\mu j1})} (\hat{\mu}_{ijk} - \hat{\mu}_{ij1})^2 = \frac{1}{2} (\Delta_{jk}^i)^2$
 - where $\Delta_{jk}^i = \left(\frac{n_{ij}}{4K\psi''(\theta_{\mu j1})}\right)^{1/2} (\hat{\mu}_{ijk} - \hat{\mu}_{ij1})$
- By applying Central Limit Theorem, $\Delta_{jk}^i = \frac{\hat{\mu}_{ijk} - \hat{\mu}_{ij1}}{\sqrt{2 \frac{\psi''(\theta_{\mu j1})}{\frac{n_{ij}}{K}}}} \sim \mathcal{N}(0, 1)$

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Multivariate Markov Chain

$i \in \{1,2,3,4\}$, given $(\Delta_{j2}^i, \dots, \Delta_{jK}^i)$,

$$\begin{aligned}\Delta_{jk}^{i+1} &= \frac{\hat{\mu}_{i+1,jk} - \hat{\mu}_{i+1,j1}}{\sqrt{2 \frac{\psi''(\theta_{\mu j1})}{\frac{n_{i+1,j}}{K}}}} = \sqrt{\frac{\frac{n_{i+1,j}}{K}}{2\psi''(\theta_{\mu j1})}} \frac{\hat{S}_{i+1,jk} - \hat{S}_{i+1,j1}}{\frac{n_{i+1,j}}{K}} \\ &= \sqrt{\frac{n_{ij}}{n_{i+1,j}}} \sqrt{\frac{\frac{n_{ij}}{K}}{2\psi''(\theta_{\mu j1})}} \left(\frac{\hat{S}_{ijk}}{\frac{n_{i,j}}{K}} - \frac{\hat{S}_{i,j1}}{\frac{n_{i,j}}{K}} \right) + \sqrt{\frac{n_{i+1,j} - n_{i+1,j}}{n_{i+1,j}}} \sqrt{\frac{\frac{n_{i+1,j} - n_{i+1,j}}{K}}{2\psi''(\theta_{\mu j1})}} \left[\frac{(\hat{S}_{i+1,jk} - \hat{S}_{ijk})}{\frac{n_{i+1,j} - n_{i,j}}{K}} - \frac{(\hat{S}_{i+1,j1} - \hat{S}_{ij1})}{\frac{n_{i+1,j} - n_{i,j}}{K}} \right] \\ &= \sqrt{\frac{n_{ij}}{n_{i+1,j}}} \Delta_{jk}^i + \sqrt{\frac{n_{i+1,j} - n_{i+1,j}}{n_{i+1,j}}} \sqrt{\frac{\frac{n_{i+1,j} - n_{i+1,j}}{K}}{2\psi''(\theta_{\mu j1})}} \left[\frac{(\hat{S}_{i+1,jk} - \hat{S}_{ijk})}{\frac{n_{i+1,j} - n_{i,j}}{K}} - \frac{(\hat{S}_{i+1,j1} - \hat{S}_{ij1})}{\frac{n_{i+1,j} - n_{i,j}}{K}} \right]\end{aligned}$$

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Arm Elimination — Computation of a_α

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$$\bullet \quad \Delta_{jk}^{i+1} = \sqrt{\frac{n_{ij}}{n_{i+1,j}}} \Delta_{jk}^i + \sqrt{\frac{n_{i+1,j} - n_{i+1,j}}{n_{i+1,j}}} \sqrt{\frac{\frac{n_{i+1,j} - n_{i+1,j}}{K}}{2\psi''(\theta_{\mu j1})}} \left[\frac{(\hat{S}_{i+1,jk} - \hat{S}_{ijk})}{\frac{n_{i+1,j} - n_{i,j}}{K}} - \frac{(\hat{S}_{i+1,j1} - \hat{S}_{ij1})}{\frac{n_{i+1,j} - n_{i,j}}{K}} \right]$$

$$\text{Where } \sqrt{\frac{\frac{n_{i+1,j} - n_{i+1,j}}{K}}{2\psi''(\theta_{\mu j1})}} \left[\frac{(\hat{S}_{i+1,jk} - \hat{S}_{ijk})}{\frac{n_{i+1,j} - n_{i,j}}{K}} - \frac{(\hat{S}_{i+1,j1} - \hat{S}_{ij1})}{\frac{n_{i+1,j} - n_{i,j}}{K}} \right] \sim \mathcal{N}(0, 1)$$

Adaptive Randomization with Arm-Elimination

Arm Elimination — Computation of a_α

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- Moreover, let $(\frac{n_{ij}}{4K\psi''(\theta_{\mu j1})})^{1/2} = \sigma_{\mu j1}$. For $k \neq k'$, $\text{cov}(\Delta_{jk}^i, \Delta_{jk'}^i) = (\sigma_{\mu j1})^2 \text{cov}(\hat{\mu}_{ijk} - \hat{\mu}_{ij1}, \hat{\mu}_{ijk'} - \hat{\mu}_{ij1})$

$$= (\sigma_{\mu j1})^2 \text{Var}(\hat{\mu}_{ij1}) = \frac{1}{2} \text{Var}(\Delta_{jk}^i)$$
- Therefore, given $(\Delta_{j2}^i, \dots, \Delta_{jK}^i)$, the distribution of $(\Delta_{j2}^{i+1}, \dots, \Delta_{jK}^{i+1})$ is

$$\mathcal{N}\left(\sqrt{\frac{n_{ij}}{n_{i+1,j}}} \begin{bmatrix} \Delta_{j2}^i \\ \vdots \\ \Delta_{jK}^i \end{bmatrix}, \frac{n_{i+1,j} - n_{i,j}}{n_{i+1,j}} \begin{bmatrix} 1 & \cdots & \frac{1}{2} \\ \vdots & \ddots & \vdots \\ \frac{1}{2} & \cdots & 1 \end{bmatrix}\right)$$

Monte Carlo Method

- Obviously, $P_*(A_j)$ is non-increasing in a_α .
- It is also easy to prove that $P_*(\cup_{j=1}^J A_j)$ is non-increasing in a_α .
- Monte Carlo Simulation($I = 5, K = 3, J = 3, t = 100,000$):
 - (1) Generate $k-1$ -dimensional multivariate normal Markov Chain with period $t = 5$, store the maximum value.
 - (2) Repeat (1) t times for J biomarker respectively.
 - (3) Use Bisection Method to find a_α within a prescribed error.
- Result:
`a_alpha: 2.592155653417598`
- Alternative Method in Computing a_α : Recursive Numerical Integration.

Adaptive Randomization with Arm-Elimination Simulation

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Recap of AR (with elimination)

Compute a_α

For $i \in \{1, \dots, I\}$ do

For $j \in \{1, \dots, J\}, t \in \{1, \dots, n_{ij} - n_{i-1,j}\}$ do

if $i = 1$ then $\phi_{jt} \leftarrow \text{random} \{1, \dots, K\}$

else

$$\phi_{jt} \begin{cases} \mathcal{B}_j & \text{with probability } \frac{(1 - |\mathcal{H}_j \setminus \mathcal{B}_j| \epsilon)}{|\mathcal{B}_j|} \\ \mathcal{H}_j \setminus \mathcal{B}_j & \text{with probability } \epsilon \end{cases}$$

end if

if $l_j^i(\hat{k}_j, k_j) \geq a_\alpha$ then eliminate k_j end if

Adaptive Randomization with Arm-Elimination Simulation

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| | | | | | |
|-----------|------------|------------|-----------|-------|---------------|
| 0.625064 | | | | | |
| Biomarker | | | | | |
| | 1 | Strategy | 2 | 3 | Type I error |
| | | | | | Type II error |
| 1 | 0.7 | 0.2 | 0.2 | 0.0 % | 0.0 % |
| | (301.9316) | (6.8882) | (6.9347) | | |
| | (431.1339) | (34.4193) | (34.4468) | | |
| 2 | 0.2 | 0.7 | 0.2 | 0.0 % | 0.0 % |
| | (5.7076) | (239.9474) | (5.7506) | | |
| | (28.5797) | (342.7496) | (28.6707) | | |
| 3 | 0.2 | 0.2 | 0.7 | 0.0 % | 11.36 % |
| | (2.4125) | (2.4148) | (53.0766) | | |
| | (12.0667) | (12.0738) | (75.8595) | | |

- Without Elimination: 0.586

Adaptive Randomization with Arm-Elimination Simulation

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| | | | | | |
|--------------------|------------------------------------|-------------------------------------|-----------------------------------|--------------|---------------|
| 0.6406417307599259 | | | | | |
| Biomarker | | Strategy | | Type I error | Type II error |
| | 1 | 2 | 3 | | |
| 1 | 0.7 (259.975) (371.5823) | 0.5 (47.1925) (94.3107) | 0.2 (6.8121) (3.4107) | 0.0 % | 7.1 % |
| 2 | 0.2 (5.5921) (28.0188) | 0.7 (199.2923) (284.8465) | 0.5 (43.5972) (87.1347) | 0.0 % | 11.88 % |
| 3 | 0.5 (14.1288) (28.2274) | 0.2 (2.3334) (11.6848) | 0.7 (42.053) (60.0878) | 0.0 % | 72.22 % |

- Without Elimination: 0.592

Adaptive Randomization with Arm-Elimination Simulation

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| | | | | | |
|-----------|------------|------------|------------|--------|---------------|
| 0.6931809 | | | | | |
| Biomarker | | | | | |
| | 1 | Strategy | 2 | 3 | Type I error |
| 1 | 0.7 | 0.69 | 0.69 | 1.02 % | Type II error |
| | (120.5078) | (112.9669) | (113.1364) | | 0.0 % |
| | (172.3093) | (163.7081) | (163.9826) | | |
| 2 | 0.69 | 0.7 | 0.69 | 1.24 % | 0.0 % |
| | (90.4203) | (96.6089) | (90.1753) | | |
| | (131.1514) | (138.049) | (130.7996) | | |
| 3 | 0.69 | 0.69 | 0.7 | 1.55 % | 0.0 % |
| | (22.7364) | (22.6029) | (24.026) | | |
| | (32.9345) | (32.739) | (34.3265) | | |

Non-Parametric Method — Simulation & Comparison

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Histogram Method with Local Linear Regression

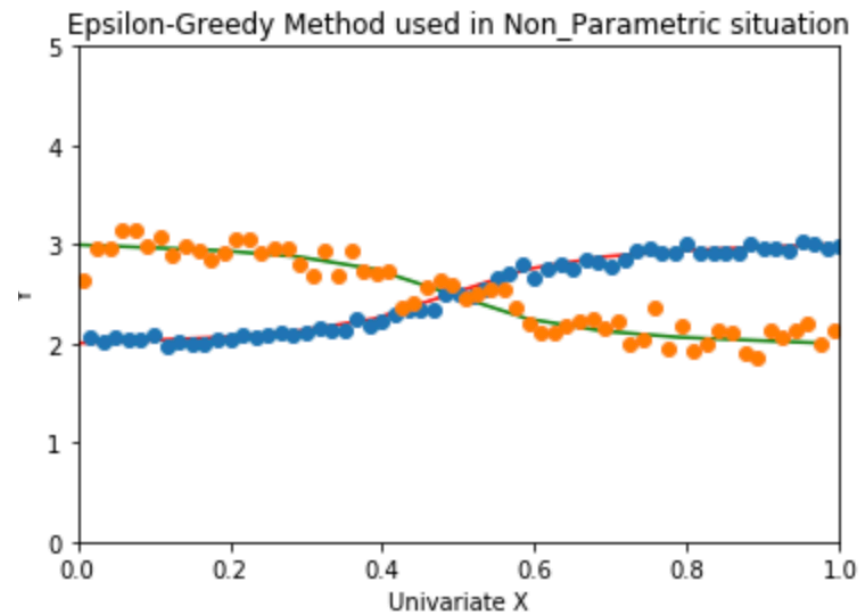
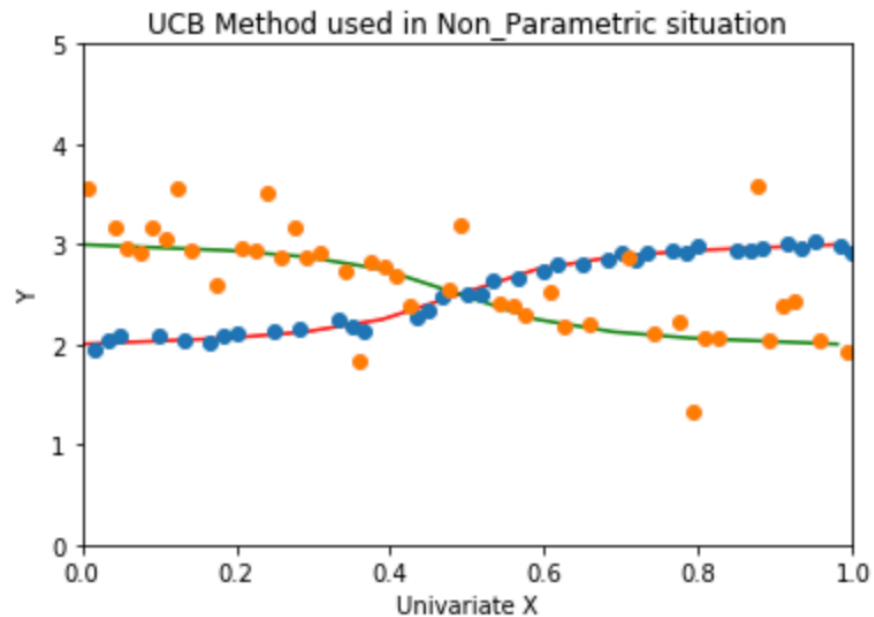
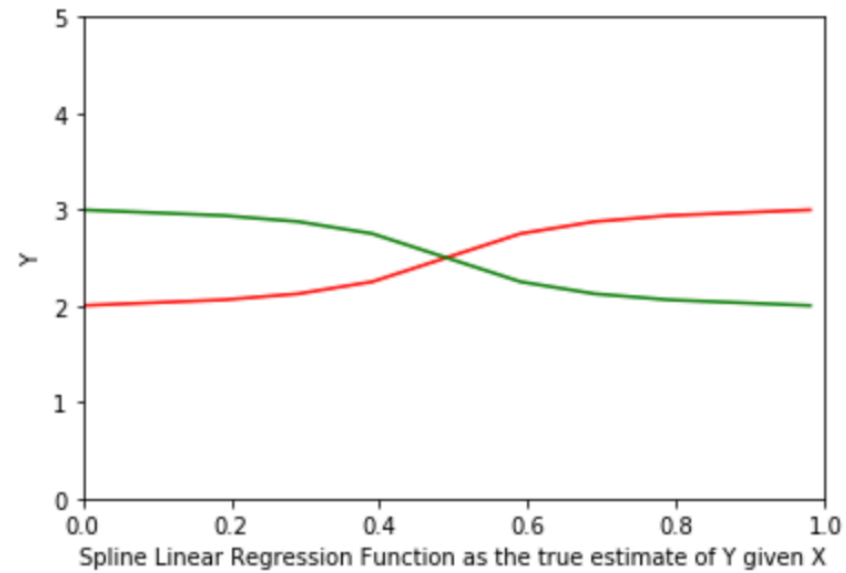
- The basic non-parametric method uses bins to split side information, then compute mean for each bin
- Yet for regions where choosing strategy is complex, this method engenders problems.
- A novel method:
 - Choose a consecutive set bins whose proportion of surviving strategies exceeds δ
(δ is close to 1)
- Arm elimination criterion: At stage t , assume that $X_t \in B_i, N_r \in a^r$ for $a^r > 1$.

$$l_{j,t}^2(i) > g^2(n_{ij}(t)/N_r)$$

$$\text{where } l_{j,t}^2(i) = (\bar{Y}_{jn_{i\hat{j}}(t)} - \bar{Y}_{jn_{ij}(t)}) / \sqrt{\frac{s_{j\hat{j}}^2 n_{i\hat{j}}(t)}{n_{i\hat{j}}(t)} + \frac{s_{jn_{ij}}^2}{n_{ij}(t)}}$$

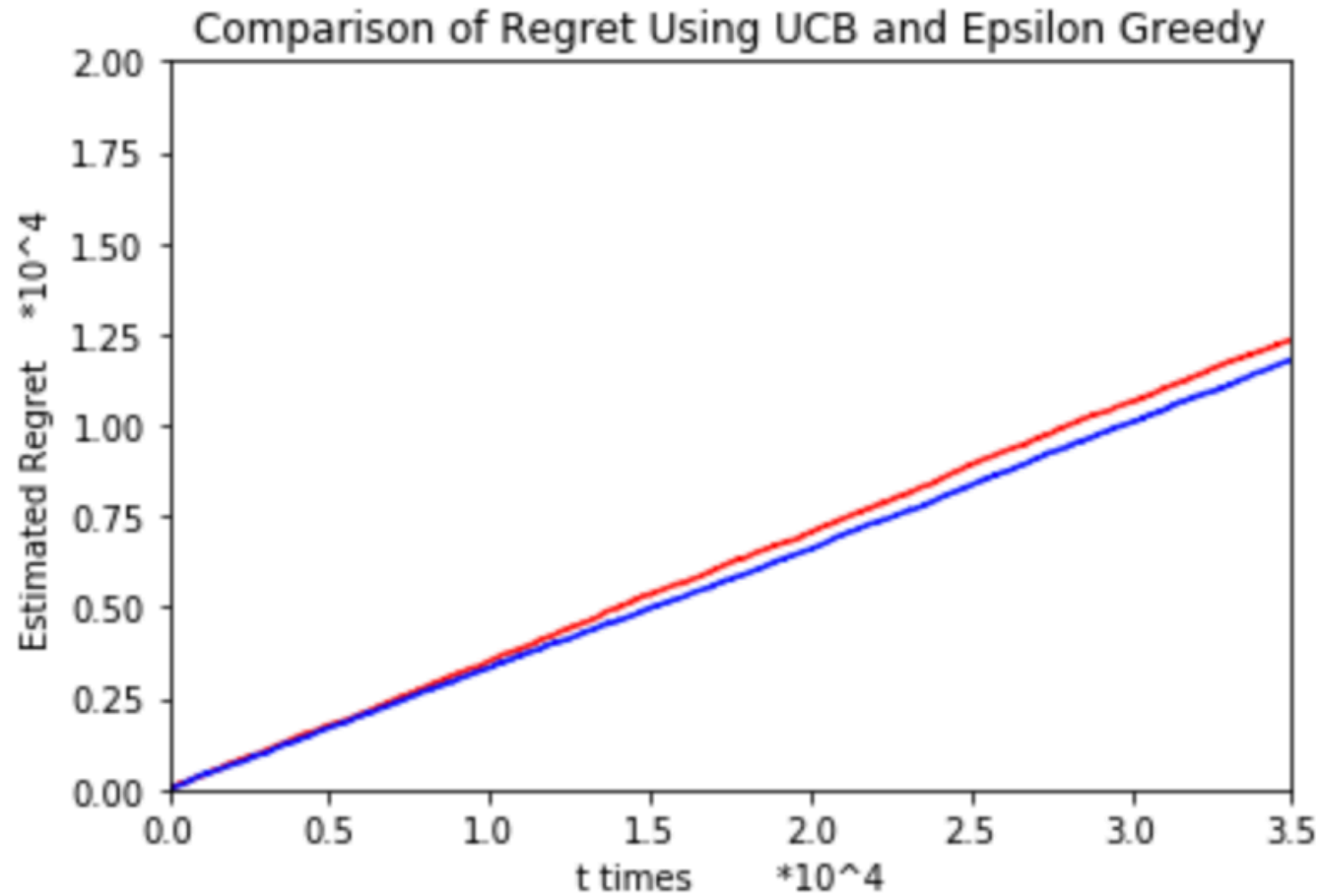
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Adaptive Randomization with Arm-Elimination

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Red: Epsilon-Greedy

Blue: UCB



Future Plan

1

Theory

2

Application





Gnereralized Likelihood Ratio: Distribution and Approximation



A more effective AR method with elimination as a non-parametric method



A more efficient threshold for eliminating inferior strategies

Business Insight : How will a new MoBike allocate their bikes



- **The Emergence of “Sharing Economy”**
- **How can you find the nearest bike or charger?**
- **Cost-effective strategy**
- **Gaming between the juggernauts?**

Appendix

A Snapshot of Codes

Computing a-alpha By Xiaocheng Li

**Multi-Arm
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```
import numpy as np
from numpy.linalg import cholesky
```

```
I = 5
K = 3
J = 3
alpha = 0.1
SampleNo = 100000
n_ij = [0,100,200,300,400,500]
coef_mu = []
coef_cov = []
max_list_final = []
regret = [0,0]
for i in range(len(n_ij)-1):
    coef_mu.append(np.sqrt(n_ij[i]/n_ij[i+1]))
    coef_cov.append(0.5-n_ij[i]/(2*n_ij[i+1]))
```

```
for i in range(J):
    max_list_final.append(0)
```

```
for j in range(J):
    MC_list = []
    for n in range(SampleNo):
        MC_list.append(0)
```

```
for n in range(SampleNo):
```

```
    K_list = []
    for a in range(K-1):
        K_list.append(0)
```

```
    L_list = []
    for b in range(I):
        L_list.append(0)
```

```
for i in range(I):
    if i == 0:
        mu = np.array([K_list])
```

```
    else:
        mu = s
        for c in range(len(s[0])):
            s[0][c] = s[0][c] + coef_mu[i]
```

```
Sigma_list = []
for a in range(K-1):
    Sigma_list.append(0)
for b in range(K-1):
    if a == b:
        Sigma_list[a].append(2*coef_cov[i])
    else:
        Sigma_list[a].append(1*coef_cov[i])
Sigma = np.array(Sigma_list)
```

```
R = cholesky(Sigma)
s = np.dot(np.random.randn(1, 2), R) + mu
for k in s[0]:
    L_list[i].append(float(k))
```

```
    MC_list[n].append(L_list)
```

```
max_list = []
for n in range(SampleNo):
    max_ln = []
    for i in range(I):
        max_ln.append(max(MC_list[n][0][i]))
    max_list.append(max(max_ln))
max_list.sort()
print(len(max_list))
max_list_final[i] = max_list
```

```
alpha_max = max_list_final[0][97999]
alpha_min = max_list_final[2][94999]
print(alpha_max)
print(alpha_min)
```

```
p_max = [0,0,0]
p_min = [0,0,0]
error = 0.0000000001
```

```
for j in range(J):
    for n in range(SampleNo):
        if max_list_final[j][n] > alpha_max:
            p_max[j] += 1
        if max_list_final[j][n] > alpha_min:
            p_min[j] += 1
```

```
p_max[j] = p_max[j] / SampleNo
p_min[j] = p_min[j] / SampleNo
prob_max = sum(p_max)-p_max[0]*p_max[1]-p_max[0]*p_max[2]-
p_max[2]*p_max[1]+p_max[0]*p_max[1]*p_max[2]
prob_min = sum(p_min)-p_min[0]*p_min[1]-p_min[0]*p_min[2]-
p_min[2]*p_min[1]+p_min[0]*p_min[1]*p_min[2]
prob_max
prob_min
```

A Snapshot of Codes

AR with Elimination By Xiaocheng Li

Multi-Arm Bandit with Side Information

```
#准备工作
import numpy as np
import re
J = 3
K = 3
n_interval = 200
l = 5
epsilon = 0.1
a_alpha = 2.592155653417598
```

```
#计算GLR的中间步骤
def theta(mu_st,n):
    import numpy as np
    theta = 0
    theta = mu_st * (np.log(mu_st)-np.log(1-mu_st)) + np.log(1-mu_st)
    theta = theta * n
    return theta
```

```
#计算GLR
def GLR(success_1, trials_1, J, K):
    list1 = []
    none_final = 0
    import numpy as np
    for j in range(J):
        list1.append([])
        for k in range(K):
            if j == k:
                list1[j].append(0)
            else:
                if trials_1[j][k] == 0:
                    list1[j].append(none_final)
                    break
                none_1 = []
                none_2 = []
                none_3 = []
                none_1 = [success_1[j][k], trials_1[j][k]]
                none_2 = [max(success_1[j], trials_1[j])(np.argmax(success_1[j]))]
                none_3 = [none_1[0] + none_2[0], none_1[1] + none_2[1]]
                none_final = theta(none_1[0]/none_1[1], none_1[1]) + theta(none_2[0]/none_2[1],
none_2[1])
                none_final = none_final - theta(none_3[0]/none_3[1], none_3[1])
                list1[j].append(none_final)
    return list1
```

```
def
AR_elimination_loop(stage,J,K,eps,Trials_Begin,Acc_Trials,Success,Mu,a_alpha_index,elimination)
:
```

```
suc = 0
tot = 0

trials = Trials_Begin
# Create Patients Data
Patients = {}
for j in range(J):
    Patients.setdefault(j+1, {})
    for k in range(K):
        Patients[j+1].setdefault(k+1, [])
```

```
# Update Patients Data and other Matrixes
for j in Patients:
    if j == 1:
        trials_j = int((trials * 0.5)
    elif j ==2:
        trials_j = int((trials * 0.4)
    else:
        trials_j = int((trials * 0.1)
```

```
acc_trials_j = []
for k in range(K):
    acc_trials_j.append([Acc_Trials[j-1][k]])
```

```
acc_success_j = []
for k in range(K):
    acc_success_j.append([Success[j-1][k]])
```

```
n_inferior = 0
superior_choice = []
inferior_choice = []
for k in range(K):
    if elimination[j-1][k] == 'superior':
        superior_choice.append(k)
    elif elimination[j-1][k] == 'inferior':
        n_inferior += 1
        inferior_choice.append(k)
```

```
#Begin the trial
for n in range(trials_j):
```

```
prob = 0
prob = np.random.uniform(0,1)
choose_k = 0
```

```
if stage == 0:
#注明： 这里默认只有3个treatment
    choice_k = int(np.random.choice([1,2,3]))
else:
    prob_2 = 0
    prob_2 = np.random.uniform(0,1)
    if prob_2 <= n_inferior * eps:
        choice_k = int(np.random.choice(inferior_choice))+1
    else:
        choice_k = int(np.random.choice(superior_choice))+1
```

```
if prob < p_jk[j-1][choice_k-1]:
    Patients[j][choice_k].append(1)
    acc_success_j[choice_k-1] +=1
    Success[j-1][choice_k-1] +=1
else:
    Patients[j][choice_k].append(0)
```

```
acc_trials_j[choice_k-1] +=1
Acc_Trials[j-1][choice_k-1] +=1
```

```
# Update mu matrix
for k in range(len(Mu[j-1])):
    if acc_trials_j[k] == 0:
        Mu[j-1][k] = 0
    else:
```

```
Mu[j-1][k] = (acc_success_j[k])/(acc_trials_j[k])
for j in Patients:
    # Update mu matrix, decide elimination
    GLR_stats = 0
    GLR_stats = GLR(Success,Acc_Trials,J,K)
    for k in range(K):
        if GLR_stats[j-1][k] >= a_alpha_index:
            elimination[j-1][k] = 'none'

        for k in range(K):
            if elimination[j-1][k] != 'none':
                if abs(Mu[j-1][k]-max(Mu[j-1])) <= (sum(acc_trials_j))**(-0.4):
                    elimination[j-1][k] = 'superior'
                else:
                    elimination[j-1][k] = 'inferior'

        suc += (sum(acc_success_j))
        tot += (sum(acc_trials_j))

return(Patients,Acc_Trials,Success,Mu, suc/tot, elimination)
```

```
#试10000次
count = 0
for simulation in range(400):
    J = 3
    K = 3
    n_interval = 200
    l = 5
    epsilon = 0.1
    a_alpha = 2.592155653417598
    p_jk = [[0.7,0.69,0.69], [0.69,0.7,0.69],[0.69,0.69,0.7]]
```

```
success_jk = []
for j in range(J):
    success_jk.append([])
    for k in range(K):
        success_jk[j].append(0)
```

```
trials_jk = []
for j in range(J):
    trials_jk.append([])
    for k in range(K):
        trials_jk[j].append(0)
```

```
mu_jk = []
for j in range(J):
    mu_jk.append([])
    for k in range(K):
        mu_jk[j].append(0)
```

```
eli = []
for j in range(J):
    eli.append([])
    for k in range(K):
        eli[j].append('start')
```

```
for i in range(l):
    Patients,trials_jk,success_jk,mu_jk,rate,eli =
AR_elimination_loop(i,J,K,epsilon,n_interval,trials_jk,success_jk,mu_jk,a_alpha,eli)
```

```
for j in range(J):
    for k in range(K):
        cul_suc[j][k] += success_jk[j][k]
        cul_trial[j][k] += trials_jk[j][k]
for j in range(J):
    if eli[j][0] == 'none':
        type_eli[j] += 1
count +=1
```

```
print(ave_suc)
print(ave_trial)
sum(ave_trial[0])+sum(ave_trial[1])+sum(ave_trial[2])
```

A Snapshot of Codes

Non-Parametric UCB By Xiaocheng Li

**Multi-Arm
Bandit with
Side
Information**

```
def Y_ARM1(x):
    Y_A1 = 0
    if x > 0.8 and x < 1:
        Y_A1 = (5/16) * (x-1) + 3
    elif x > 0.7 and x <= 0.8:
        Y_A1 = (5/8) * (x-0.9) + 3
    elif x > 0.6 and x <= 0.7:
        Y_A1 = (5/4) * (x-0.8) + 3
    elif x >= 0.5 and x <= 0.6:
        Y_A1 = (5/2) * (x-0.7) + 3
    elif x >= 0 and x < 0.5:
        Y_A1 = 5 - Y_ARM1(1-x)
    return Y_A1

def Y_ARM2(x):
    Y_A2 = 5 - Y_ARM1(x)
    return Y_A2

N_1 = 35000
N0 = 300
x = -1
no_bins = 60
ARM_1 = []
ARM_2 = []
Interim = 0
while N0*(2**(Interim) < N_1:
    Interim += 1
regret = [], []

ARM_1 = []
ARM_2 = []
B_i = []
Interval = 1/no_bins
for k in range(2):
    B_i.append([])
    for i in range(no_bins):
        B_i[k].append([])

def UCB(list1, list2, no_1, no_2):
    import numpy as np

    f1 = sum(list1)
    f2 = sum(list2)
    if list1 != []:
        f1 = f1/len(list1)
    if list2 != []:
        f2 = f2/len(list2)

    u1 = np.sqrt(np.log(no_1 + no_2)/no_1)
    u2 = np.sqrt(np.log(no_1 + no_2)/no_2)
    strategy = np.argmax([f1+u1, f2+u2]) + 1
    return strategy

#一个坐标系上绘制多个图 Plotting more than one plot on the same set of axes
#依次作图即可
import numpy as np
import pylab as pl
i = 0
x1 = []
y1 = []

while i < 0.99:
    x1.append(i)
    i += 0.01
    y1.append(Y_ARM1(i))
i = 0
x2 = []
y2 = []
while i < 0.99:
    x2.append(i)
    i += 0.01
    y2.append(Y_ARM2(i))

x01 = []
y01 = []
for i in range(len(B_i[0])):
    x01.append(1/60 + (1/60)*i)
    y01.append(sum(B_i[0][i])/len(B_i[0][i]))

x02 = []
y02 = []
for i in range(len(B_i[0])):
    x02.append(1/120 + (1/60)*i)
    y02.append(sum(B_i[1][i])/len(B_i[1][i]))

pl.plot(x1, y1, 'red')# use pylab to plot x and y
pl.plot(x2, y2, 'green')

pl.title("")# give plot a title
pl.xlabel('Spline Linear Regression Function as the true estimate of Y given X')# make axis labels
pl.ylabel('Y')

pl.xlim(0.0, 1)# set axis limits
pl.ylim(0, 5)

pl.show()# show the plot on the screen
```


A Snapshot of Codes

Non-Parametric epsilon-greedy By Xiaocheng Li

Multi-Arm Bandit with Side Information

```
def Y_ARM1(x):
    Y_A1 = 0
    if x > 0.8 and x < 1:
        Y_A1 = (5/16) * (x-1) + 3
    elif x > 0.7 and x <= 0.8:
        Y_A1 = (5/8) * (x-0.9) + 3
    elif x > 0.6 and x <= 0.7:
        Y_A1 = (5/4) * (x-0.8) + 3
    elif x >= 0.5 and x <= 0.6:
        Y_A1 = (5/2) * (x-0.7) + 3
    elif x >= 0 and x < 0.5:
        Y_A1 = 5 - Y_ARM1(1-x)
    return Y_A1
```

```
def Y_ARM2(x):
    Y_A2 = 5 - Y_ARM1(x)
    return Y_A2
```

```
N_1 = 35000
N0 = 300
x = -1
no_bins = 60
epsilon = 0.1
ARM_1 = []
ARM_2 = []
regret1 = [0,0]
Interim = 0
while N0*(2**Interim) < N_1:
    Interim += 1
```

```
ARM_1 = []
ARM_2 = []
B_i = []
Interval = 1/no_bins
for k in range(2):
    B_i.append([])
    for i in range(no_bins):
        B_i[k].append(0)
```

```
def greedy(list1, list2):
    import numpy as np

    f1 = sum(list1)
    f2 = sum(list2)
    if list1 != []:
        f1 = f1/len(list1)
    if list2 != []:
        f2 = f2/len(list2)

    strategy = np.argmax([f1,f2]) + 1
    return strategy
```

```
import numpy as np
no_1 = 0
no_2 = 0
```

```
for i in range(N_1):
    x = -1
    y = 0
    choice = 0
```

```
x = np.random.uniform(0,1)
```

```
s = 0
while (Interval * s) < x:
    s += 1

if i < 2:
    choice = i + 1
else:
    advantage = 0
    advantage = greedy(B_i[0][s-1], B_i[1][s-1])
    prob = 0
    prob = np.random.uniform(0,1)
    if prob > epsilon:
        choice = advantage
    else:
        choice = 3 - advantage
```

```
if choice == 1:
    no_1 += 1
    y = np.random.normal(Y_ARM1(x),1)
    ARM_1.append(y)
    B_i[0][s-1].append(y)
else:
    no_2 += 1
    y = np.random.normal(Y_ARM2(x),1)
    ARM_2.append(y)
    B_i[1][s-1].append(y)
regret1[1].append(max([Y_ARM1(x),Y_ARM2(x)]) - y)
regret1[0].append(i)
if i != 0:
    regret1[1][i] += regret1[1][i-1]
len(B_i[1])
for i in range(len(regret1[0])):
    regret1[0][i] = regret1[0][i]/10000
    regret1[1][i] = regret1[1][i]/10000
regret1[1][34999]
```

#一个坐标系上绘制多个图 Plotting more than one plot on the same set of axes
#依次作图即可

```
import numpy as np
import pylab as pl
i = 0
x1 = []
y1 = []
while i < 0.99:
    x1.append(i)
    i += 0.01
    y1.append(Y_ARM1(i))
i = 0
x2 = []
y2 = []
while i < 0.99:
    x2.append(i)
    i += 0.01
    y2.append(Y_ARM2(i))
```

```
x01 = []
y01 = []
```

```
for i in range(len(B_i[0])):
    x01.append(1/60 +(1/60)*i)
    if len(B_i[0][i]) == 0:
        y01.append(0)
    else:
        y01.append(sum(B_i[0][i])/len(B_i[0][i]))
```

```
x02 = []
y02 = []
for i in range(len(B_i[0])):
    x02.append(1/120 +(1/60)*i)
    y02.append(sum(B_i[1][i])/len(B_i[1][i]))
```

```
pl.plot(x1, y1, 'red')# use pylab to plot x and y
pl.plot(x2, y2, 'green')
pl.plot(x01, y01, 'o')
pl.plot(x02, y02, 'o')
```

```
pl.title('Epsilon-Greedy Method used in Non_Parametric situation')# give plot a title
pl.xlabel('Univariate X')# make axis labels
pl.ylabel('Y')
```

```
pl.xlim(0.0, 1)# set axis limits
pl.ylim(0, 5)
```

```
pl.show()# show the plot on the screen
```

A Snapshot of Codes

Non-Paremetric AR Method By Xiaocheng Li (Working)

**Multi-Arm
Bandit with
Side
Information**

```
def Y_ARM1(x):
    Y_A1 = 0
    if x > 0.8 and x <= 1:
        Y_A1 = (5/16) * (x-1) + 3
    elif x > 0.7 and x <= 0.8:
        Y_A1 = (5/8) * (x-0.9) + 3
    elif x > 0.6 and x <= 0.7:
        Y_A1 = (5/4) * (x-0.8) + 3
    elif x >= 0.5 and x <= 0.6:
        Y_A1 = (5/2) * (x-0.7) + 3
    elif x > 0 and x < 0.5:
        Y_A1 = 5 - Y_ARM1(1-x)
    return Y_A1

def Y_ARM2(x):
    Y_A2 = 5 - Y_ARM1(x)
    return Y_A2

def Y_ARM_MAX(x):
    Y_A_MAX = max(Y_A1, Y_A2)

def g(t):
    import numpy as np

    if t > 0.86 and t <= 1:
        result = (1/t-1)**(0.5)*(0.63883-0.40258(1/t-1))
    elif t > 0.28 and t <= 0.86:
        result = -0.5759*t**2 + 0.2987*t + 0.4034
    elif t > 0.01 and t <= 0.28:
        result = -1.58137*t + 1.53343*t**(0.5) + 0.073271
    elif t > 0 and t <= 0.01:
        result = (t*(-2*np.log(t) - np.log(-1*np.log(t))-np.log(16*np.pi)
+0.99232*np.exp(-0.03612*t**0.5)))*0.5
    result = (result**2) / (2*t)
    return result

N_1 = 35000
N0 = 300
x = -1
no_bins = 35
ARM_1 = []
ARM_2 = []
Interim = 0
while N0*(2**(Interim) < N_1):
    Interim += 1
    N_1 = N0*2**(Interim-1)

ARM_1 = []
ARM_2 = []

B_i = []
B_elimination = []
Interval = 1/no_bins
for k in range(2):
    B_i.append([])
    for i in range(no_bins):
        B_i[k].append([])
for i in range(no_bins):
    B_elimination.append([])
    for k in range(2):
```

```
        B_elimination[i].append('start')
B_elimination[-6]

import numpy as np
no_1 = 0
no_2 = 0
for r in range(Interim+1):
    if r == 7:
        trials = N_1 - N0*2**(Interim-1)
    else:
        trials = N0*2**r

for t in range(trials):
    x = -1
    y = 0
    choice = 0

    x = np.random.uniform(0,1)

    s = 0
    while (Interval * s) < x:
        s += 1

    if B_elimination[s-1] == ['start', 'start']:
        choice = np.random.choice([1,2])
    elif:

    if choice == 1:
        no_1 += 1
        y = np.random.normal(Y_ARM1(x),1)
        ARM_1.append([x,y])
        B_i[0][s-1].append([x,y])
    else:
        no_2 += 1
        y = np.random.normal(Y_ARM2(x),1)
        ARM_2.append([x,y])
        B_i[1][s-1].append([x,y])
print(ARM_1)
```