

# Generalized linear models for longitudinal data with biased sampling designs

A sequential offsetted regression approach

L.S. McDaniel<sup>1</sup>   J.S. Schildcrout<sup>2</sup>   E.F. Schisterman<sup>3</sup>  
P.J. Rathouz<sup>4</sup>

<sup>1</sup>Biostatistics Program, School of Public Health  
LSU Health Sciences Center, New Orleans

<sup>2</sup>Department of Biostatistics, Department of Anesthesiology  
Vanderbilt University School of Medicine

<sup>3</sup>*Eunice Kennedy Shriver* National Institute of Child Health and Human  
Development  
National Institutes of Health

<sup>4</sup>Department of Biostatistics and Medical Informatics  
University of Wisconsin, Madison

- Goal: Identify risk and prognostic factors for ADHD in **early** childhood
- Sampling: 255 subjects, about half cases/controls
  - Cases: Referred by parent or teacher
  - Controls: Matched demographically
- Followed for 15 years (we have 8)
- Analyze: Time course of hyperactivity symptom count

## Subject-level Sampling:

A subject is either in or out of the study

## Observation-level Sampling:

A subject may be sampled at each time point

- $Y_j$ : count or continuous outcome at times  $t_j = 1, \dots, T$
- $x_j$ :  $p$ -vector of covariates at times  $t_j$
- $X = (x_1, \dots, x_T)'$  is a  $T \times p$  matrix of covariates

Marginal population mean model for  $Y_j$ :

$$\mu_{P_j} = E(Y_j|X) = g^{-1}(x_j'\beta)$$

Finally,

- $Z_j$ : subject was referred at time  $t_j$
- $S_j$ : subject was sampled at time  $t_j$

- 1 No interference assumption

$$E(Y_j|X) = E(Y_j|x_j)$$

- 2 Known value for

$$\frac{\Pr(S_j = 1|Z_j = 1)}{\Pr(S_j = 1|Z_j = 0)} = \frac{\pi(1)}{\pi(0)}$$

- 3 Sampling only depends on  $Z_j$ , and possibly baseline covariates

The assumptions allow for three modeling steps:

- 1 Estimate  $\Pr(Z_j = 1 | Y_j, X)$  from sample, for each  $t_j$
- 2 Compute  $\Pr(S_j = 1 | Y_j, X)$ , for each  $t_j$
- 3 Estimate  $E(Y_j | X)$  from sample

## Step 1: Estimate $\Pr(Z_j = 1|Y_j, X)$

Let  $w_j$  be a vector of covariates (possibly overlapping  $x_j$ )

In the population:

$$\begin{aligned}\Pr(Z_j = 1|Y_j, X) &= \lambda_{P_j}(y, X) \\ &= \text{logit}^{-1} \{ w'_{1j}\gamma_1 + h(y)w'_{2j}\gamma_2 \}\end{aligned}$$

Then, in the sample:

$$\begin{aligned}\Pr(Z_j = 1|Y_j, X, S_j) &= \lambda_{S_j}(y, X) \\ &= \text{logit}^{-1} \{ w'_{1j}\gamma_1 + h(y)w'_{2j}\gamma_2 + \log \pi(1)/\pi(0) \}\end{aligned}$$

## Step 2: Compute $\Pr(S_j | Y_j, X)$

$$\begin{aligned}\rho_j(y, X) &= \Pr(S_j = 1 | y, X) \\ &= \pi(0) \{1 - \lambda_{P_j}(y, X)\} + \pi(1) \lambda_{P_j}(y, X)\end{aligned}$$

Gain stability by using

$$\frac{\rho_j(y, X)}{\rho_j(y_0, X)} = \frac{1 - \lambda_{P_j}(y, X) + \{\pi(1)/\pi(0)\} \lambda_{P_j}(y, X)}{1 - \lambda_{P_j}(y_0, X) + \{\pi(1)/\pi(0)\} \lambda_{P_j}(y_0, X)}$$



## Step 3: Estimate $E(Y_j|X)$

In the population, conditional density is **exponential family**:

$$f_P(y|X) = \exp \left\{ \frac{\theta_j y - b(\theta_j)}{\phi} + c(y; \phi) \right\}$$

Use canonical link:

$$g(\mu_{P_j}) = g(E(Y_j|X)) = x_j' \beta = \theta_j$$

In sample:

$$f_S(y|X) \propto \exp \left\{ \frac{\theta_j y - b(\theta_j)}{\phi} + c(y; \phi) + \log \rho_j(y, X) \right\}$$

View the solution as stacked estimating equations:

$$\sum_i \begin{pmatrix} \mathbf{T}_i(\gamma) \\ \mathbf{U}_i(\gamma, \beta) \end{pmatrix} = \mathbf{0}.$$

Use sandwich estimate for SE

Worked out for

- Binary data (binomial)
- Count data (Poisson)
- Continuous data (normal)

For continuous data, need to estimate variance

For each subject,  $i$ ,

$$\log \mu_{ij} = \beta_0 + \beta_{x_1} x_{1i} + \beta_t t_j + \beta_{tx_1} (t_j \times x_{1i})$$

$x_{1i}$  is time-invariant, binary covariate.

Oversample **subjects** with high values for  $Y_{i1}$

Table includes % bias and coverage probability (target of 95%)

Estimation	$\beta_0 = -1.4$	$\beta_{x_1} = 0.4$	$\beta_t = -0.1$	$\beta_{tx_1} = 0.1$
Naive GEE	-42 (0)	-21 (88)	24 (84)	-11 (92)
IPW	0 (94)	-1 (95)	3 (93)	2 (94)
SOR	1 (95)	2 (95)	2 (94)	0 (95)

# Efficiency Relative to Simple Random Sampling

Estimation	$\beta_0 = -1.4$	$\beta_{x_1} = 0.4$	$\beta_t = -0.1$	$\beta_{tx_1} = 0.1$
IPW	1.17	1.17	0.63	0.60
SOR	1.40	1.37	1.26	1.14

For each subject,  $i$ ,

$$Y_{ij} = \beta_0 + \beta_{x_1}x_{1i} + \beta_{x_2}x_{2ij} + \beta_{x_3}x_{3ij} + \epsilon_{ij}$$

$x_{1i}$  is time-invariant,  $x_{2ij}$  and  $x_{3ij}$  vary with time.

Oversample **observations** with high values for  $Y_{ij}$

Estimation	$\beta_0 = 1$	$\beta_{x_1} = 1$	$\beta_{x_2} = 1$	$\beta_{x_3} = 1$
Naive GEE	443 (0)	29 (74)	29 (70)	28 (72)
IPW	-1 (94)	1 (95)	1 (94)	-1 (95)
SOR	-5 (90)	0 (95)	0 (95)	-1 (96)

# Efficiency Relative to Simple Random Sampling

Estimation	$\beta_0 = 1$	$\beta_{x_1} = 1$	$\beta_{x_2} = 1$	$\beta_{x_3} = 1$
IPW	0.34	0.37	0.34	0.33
SOR	0.75	1.33	1.31	1.26

Response is hyperactivity symptom count.  
Coefficients are exponentiated

	Naive		SOR	
Intercept	4.01	(3.32, 4.90)	2.75	(2.23, 3.39)
$t$	0.98	(0.90, 1.05)	1.06	(0.97, 1.15)
$(t - 2)_+$	<b>0.95</b>	<b>(0.88, 1.03)</b>	<b>0.89</b>	<b>(0.80, 0.97)</b>
age	0.87	(0.75, 1.00)	0.87	(0.75, 1.01)
sex	<b>0.80</b>	<b>(0.53, 1.22)</b>	<b>0.62</b>	<b>(0.41, 0.93)</b>
afr	1.58	(1.25, 2.03)	1.67	(1.30, 2.12)
other	1.11	(0.63, 1.95)	1.03	(0.58, 1.84)
sex* $t$	1.04	(0.84, 1.28)	1.07	(0.85, 1.35)
sex* $(t - 2)_+$	0.91	(0.71, 1.16)	0.89	(0.68, 1.15)