

Computing variances and probabilities by conditioning

Computing variances by conditioning

- ❖ $Var(X) = E[X^2] - (E[X])^2$
- ❖ (Example 3.18) Variance of the Geometric random variable with p . Let N be the trial number of the first success.
 - Let $Y = \begin{cases} 1, & \text{if the first trial results in a success} \\ 0, & \text{otherwise} \end{cases}$
 - $E[N] = \frac{1}{p}$ (shown in Example 3.11)
 - $E[N^2|Y = 1] = 1, E[N^2|Y = 0] = E[(N + 1)^2]$
 - $E[N^2] = E[E[N^2|Y]] = E[N^2|Y = 1]P\{Y = 1\} + E[N^2|Y = 0]P\{Y = 0\} = p + (1 - p)E[(N + 1)^2]$
 - Therefore, $E[N^2] = \frac{2-p}{p^2}$ and $Var[N] = E[N^2] - (E[N])^2 = \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}$

Computing probabilities by conditioning

- ❖ $P(E) = \sum_y P(E|Y = y)P(Y = y)$, if Y is discrete
- ❖ $P(E) = \int_{-\infty}^{\infty} P(E|Y = y) f_Y(y) dy$, if Y is continuous
- ❖ (Example 3.22) Let X denote the number of accidents that a randomly chosen policyholder has next year. Let Y denote the Poisson mean number of accidents for this policyholder.
 - $P[X = n|Y = \lambda] = e^{-\lambda} \frac{\lambda^n}{n!}$
 - $g(\lambda) = \lambda e^{-\lambda}, \lambda \geq 0$
 - $P[X = n] = \int_{-\infty}^{\infty} P(X = n|Y = \lambda) g(\lambda) d\lambda = \int_0^{\infty} e^{-\lambda} \frac{\lambda^n}{n!} \lambda e^{-\lambda} d\lambda = \frac{1}{n!} \int_0^{\infty} \lambda^{n+1} e^{-2\lambda} d\lambda$
 - $h(\lambda) = \frac{2e^{-2\lambda}(2\lambda)^{n+1}}{(n+1)!}, \lambda > 0 \Leftrightarrow \lambda \sim \text{gamma}(n+2, 2)$
 - $\int_0^{\infty} \frac{2e^{-2\lambda}(2\lambda)^{n+1}}{(n+1)!} d\lambda = 1 \rightarrow \int_0^{\infty} \lambda^{n+1} e^{-2\lambda} d\lambda = \frac{(n+1)!}{2^{n+2}}$
 - $P[X = n] = \frac{1}{n!} \int_0^{\infty} \lambda^{n+1} e^{-2\lambda} d\lambda = \frac{n+1}{2^{n+2}}$

The Ballot problem

- ❖ $P(E) = \sum_y P(E|Y = y)P(Y = y)$, if Y is discrete
- ❖ (Example 3.27) In an election, candidate A receives n votes and candidate B m votes where $n > m$. Assuming that all orderings are equally likely, prove the probability that A is always ahead in the count of votes is $\frac{n-m}{n+m}$ (Let $P_{n,m}$ denote this probability)?
 - (Solution) Condition on which candidate receives the last vote counted
 - Then, $P_{n,m} = P\{\text{A always ahead} \mid \text{A receives last vote}\} \frac{n}{n+m} + P\{\text{A always ahead} \mid \text{B receives last vote}\} \frac{m}{n+m}$

\swarrow
 $P\{\text{A receives a vote}\}$
 - $P_{n,m} = \frac{n}{n+m} P_{n-1,m} + \frac{m}{n+m} P_{n,m-1}$ (Equation 1)
 - It is true when $n + m = 1 \Rightarrow P_{1,0} = 1$
 - Assume it is true on $n + m - 1$, i.e. $P_{n-1,m} = \frac{n-1-m}{n-1+m}$ and $P_{n,m-1} = \frac{n-(m-1)}{n+m-1}$
 - By induction on $n + m$ and (Equation 1), $P_{n,m} = \frac{n}{n+m} \frac{n-1-m}{n-1+m} + \frac{m}{n+m} \frac{n-m+1}{n+m-1} = \frac{n-m}{n+m}$