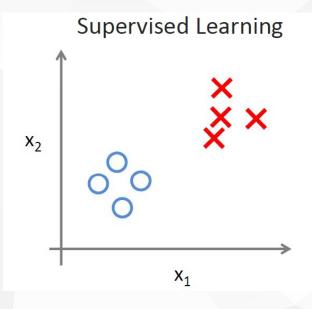
第六章: 有监督学习方法



什么是有监督学习



- ■有监督学习: 从有标记的训练数据中学习推断函数。训练数据包括一个训练样本的集合。
- ■有监督学习算法分析训练数据,产生推断函数。 推断函数能够对新的样本进行预测。
- 最优的情形: 算法能够准确地对没见过的样本进行正确地分类。
- ■目标函数 (target function) :Y=f(X) 或 P(Y|X)

→ 有监督学习方法一: 产生式模型 Generative Model

■首先对联合分布的进行推断:

$$p(x, y) = p(y)p(x \mid y)$$

■然后使用贝叶斯定理来计算条件分布p(y/x)

$$p(y | x) = \frac{p(x, y)}{p(x)} = \frac{p(y)p(x | y)}{\int p(y)p(x | y)dy}$$

- ■利用条件概率密度来预测。
- ■例子:确定某人所说语言的类别 (英语, 法语, 汉语)
 - •产生式模型的做法首先学习每种语言,然后确定所说的语主属于哪种语言。

→ 有监督学习方法二: 判别式模型Discriminative Model

- ■直接估计条件概率分布P(y|x)或条件概率密度函数p(y|x)。
- ■根据估计的函数确定输出。
- ■例子:确定某人所说语言的类别(英语,法语,汉语)?
 - •不学习任何语言的情况下确定语言差异——这是一项容易得多的任务!

为 有监督学习方法三: 判别函数

■寻找一个函数f(x), 将每个输入直接映射到目标输出。

例如: 学习二值 分类器时, f(.)是一个二值函数

•f=1 表示第一类 C_1 , f=0 或 (-1) 表示第二类 C_2

•例如: f(x) = sign(d(x))

- ■概率不起直接作用
 - •不能直接获取后验概率。
 - •f 通常旨在近似条件分布p(y|x)。

回归方法

>> 回归

- ■输入: N i.i.d 训练样本 $(x^i, y^i) \in X \times R$, $i = 1, \dots, N$
- ■目标函数: $f \in \mathcal{F}$
- ■损失函数: $L(f;x,y) = (f(x)-y)^2$
- ■期望风险: $\int (f(x)-y)^2 dP(x,y)$
- ■如果f是关于x的线性函数,最优化问题为: $\min_{w} J(w) = \sum_{i=1}^{N} (w^{T}x^{i} y^{i})$

$$\frac{\partial J(\boldsymbol{w})}{\partial w_{j}} = 2\sum_{i=1}^{N} x_{j}^{i} (\boldsymbol{w}^{T} \boldsymbol{x}^{i} - \boldsymbol{y}^{i})$$
 梯度等于0
$$\boldsymbol{w}^{*} = (\boldsymbol{X}^{T} \boldsymbol{X})^{-1} \boldsymbol{X}^{T} \boldsymbol{y}$$

$$\boldsymbol{X}^{T} = (\boldsymbol{x}^{1}, \boldsymbol{x}^{2}, \dots, \boldsymbol{x}^{N}), \quad \boldsymbol{y}^{T} = (\boldsymbol{y}^{1}, \boldsymbol{y}^{2}, \dots, \boldsymbol{y}^{N}).$$

▶ 最小二乘法 (Least Mean Squares Algorithm)

量优化问题:
$$\min_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^{N} (\mathbf{w}^{T} \mathbf{x}^{i} - \mathbf{y}^{i})^{2}$$

■梯度下降:
$$\frac{\partial J(w)}{\partial w_i} = 2\sum_{i=1}^N x_j^i (w^T x^i - y^i)$$

更新规则:
$$w_i = w_i - 2\alpha$$

$$w_j = w_j - 2\alpha \sum_{i=1}^{N} (\mathbf{w}^T \mathbf{x}^i - y^i) x_j^i, \ \alpha > 0$$

批梯度下降,BGD

Batch Gradient Descent

■优点:

- ●一次迭代是对所有样本进行计算,此时利用矩阵进行操作,实现了并行。
- ●由全数据集确定的方向能够更好地代表样本总体,从而更准确地朝向极值所在的方向。当目标函数为凸函数时,BGD一定能够得到全局最优。

■缺点:

●当样本数目 N很大时,每迭代一步都需要对所有样本计算,训练过程会很慢。

→ 最小二乘法 (Least Mean Squares Algorithm)

量优化问题:
$$\min_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^{N} (\mathbf{w}^{T} \mathbf{x}^{i} - \mathbf{y}^{i})^{2}$$

■梯度下降:
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■更新规则:

$$w_j = w_j - 2\alpha \sum_{i=1}^{N} (\mathbf{w}^T \mathbf{x}^i - y^i) x_j^i, \ \alpha > 0$$

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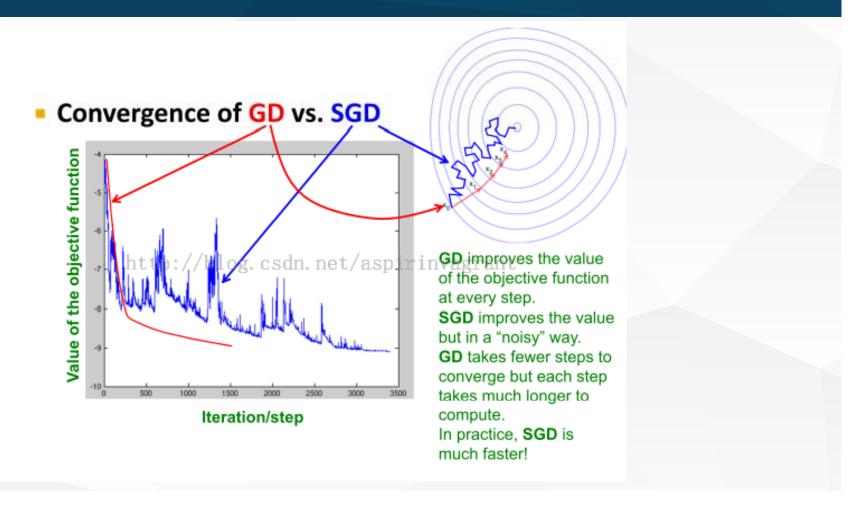
批梯度下降,BGD

Batch Gradient Descent

随机梯度下降,SGD **Stochastic Gradient Descent**

$$\boldsymbol{w} = \boldsymbol{w} - 2\alpha \boldsymbol{X}^T \boldsymbol{b}$$
, where $b_i = \boldsymbol{w}^T \boldsymbol{x}^i - y^i$, $\boldsymbol{b} = (b_1, b_2, \dots, b_N)^T$

>>> SGD 与BGD/GD对比



> 作业(选做)

■给定训练数据集
$$X = \begin{pmatrix} 1 & 2 & 5 & 4 \\ 2 & 5 & 1 & 2 \end{pmatrix}, y = (19 \ 26 \ 19 \ 20)^T$$
, 令 step=0.001, w_0 =[1 1]^T

■编程实现 SGD和GD算法, 求解 w



iteration= 9998, loss=0.4715

iteration= 9999, loss=0.4713

iteration= 10000, loss=0.4712

output_w =

0.6289

2.3083

13.2834

iteration= 12671, loss=0.3017

iteration= 12672, loss=0.3017

iteration= 12673, loss=0.3017

iteration= 12674, loss=0.3016

iteration= 79998, loss=0.3063

iteration= 79999, loss=0.3072

iteration= 80000, loss=0.3017

output_w =

0.5595

2.2090

13.8365

>> 利用非线性基进行线性回归 (广义线性回归)

■对非线性基进行线性组合: $f(\mathbf{w}, \mathbf{x}) = w_0 + \sum_{j=1}^K w_j \phi_j(\mathbf{x})$

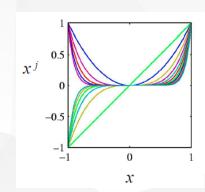
$$\Phi = (1, \phi_1, \cdots, \phi_K)$$

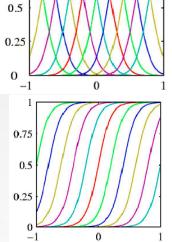
■非线性基函数

• $\phi(\mathbf{x}) = (1, x, x^2, \dots, x^K)$ 多项式基函数

•
$$\phi_j(\mathbf{x}) = \exp\left(-\frac{\left(x - \mu_j\right)^2}{2s^2}\right)$$
高斯函数

•
$$\phi_j(\mathbf{x}) = \sigma\left(\frac{\mathbf{x} - \mu_j}{s}\right), \sigma(a) = \frac{1}{1 + \exp(-a)}$$
 Sigmoid 逐数





>> 广义线性回归的闭式解

量最优化问题:
$$\min_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^{N} (\mathbf{w}^{T} \phi(\mathbf{x}^{i}) - \mathbf{y}^{i})^{2}$$

■最优化问题:
$$\min_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^{N} (\mathbf{w}^{T} \phi(\mathbf{x}^{i}) - \mathbf{y}^{i})^{2}$$
■梯度:
$$\frac{\partial J(\mathbf{w})}{\partial w_{j}} = 2 \sum_{i=1}^{N} \phi_{j}(\mathbf{x}^{i}) (\mathbf{w}^{T} \phi(\mathbf{x}^{i}) - \mathbf{y}^{i})$$

■闭式解:
$$\mathbf{w}^* = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{y}$$

其中
$$\Phi = \begin{pmatrix} \phi_0(\mathbf{x}^1) & \cdots & \phi_K(\mathbf{x}^1) \\ \vdots & \vdots & \vdots \\ \phi_0(\mathbf{x}^N) & \cdots & \phi_K(\mathbf{x}^N) \end{pmatrix}, \mathbf{y} = (\mathbf{y}^1, \dots, \mathbf{y}^N)^T$$

■也可用 SGD求解。

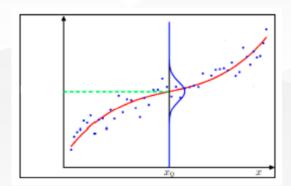
➤ 解释: MLE

- ■在高斯噪声模型下,最小化平方误差与最大似然的解相同。
- ■假设y是具有加性高斯噪声的确定函数f 给出的标量,即 $y = f(x, w) + \varepsilon$, ε 是均值为0,方差为 β^{-1} 的高斯噪声
- ■训练数据:

$$p(y \mid \mathbf{x}, \mathbf{w}, \boldsymbol{\beta}) = N(y \mid f(\mathbf{x}, \mathbf{w}), \boldsymbol{\beta}^{-1})$$

■似然函数: $(x^i, y^i), i = 1, 2, \dots, N$

$$\prod_{i=1}^{N} N(y^{i} \mid \boldsymbol{w}^{T} \boldsymbol{x}^{i}, \boldsymbol{\beta}^{-1})$$



>> 对数似然函数

■对数似然

$$\sum_{i=1}^{N} \ln N(y^{i} \mid \mathbf{w}^{T} \mathbf{x}^{i}, \beta^{-1}) = \frac{N}{2} \ln \beta - \frac{N}{2} \ln 2\pi - \frac{1}{2} \beta J(\mathbf{w})$$

■结论: 最大化似然相当于最小化平方误差之和。

最小二乘法实际上是在假设误差项满足高斯分布且独立同分布情况下, 使似然性最大化。

>> 正则化的 LMS 与 MAP

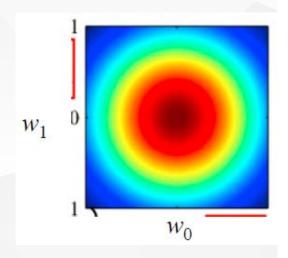
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■闭式解为:
$$w^* = (\Phi^T \Phi + \lambda I)^{-1} \Phi^T y$$

■似然函数 为: $\prod_{i=1}^{N} N(y^{i} | \boldsymbol{w}^{T} \boldsymbol{x}^{i}, \boldsymbol{\beta}^{-1})$

■参数的先验: $p(w) = N(\theta, \lambda^{-1}I)$ 多变量高斯

•w的各分量间的协相关系数为0,各分量方差为 λ^{-1}



>> 参数的后验概率

■贝叶斯定理: $p(\mathbf{w} | \mathbf{y}) = p(\mathbf{y} | \mathbf{w}) p(\mathbf{w}) / p(\mathbf{y})$

•似然函数: $p(\mathbf{y} | \mathbf{X}, \mathbf{w}, \boldsymbol{\beta}) = \prod_{i=1}^{N} N(\mathbf{y}^{i} | \mathbf{w}^{T} \mathbf{x}^{i}, \boldsymbol{\beta}^{-1})$

•先验: $p(\mathbf{w}) = N(\mathbf{0}, \lambda^{-1}\mathbf{I})$

■后验概率依然是高斯分布,对后验取对数:

$$\ln(p(\mathbf{w} \mid \mathbf{y})) = -\beta \sum_{i=1}^{N} (y^{i} - \mathbf{w}^{T} \mathbf{x}^{i})^{2} - \lambda \mathbf{w}^{T} \mathbf{w} + constant$$

■最大化后验等同于最小化带有正则项的平方和误差。

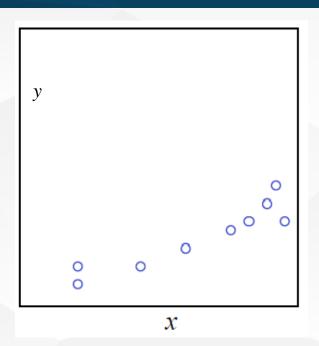
$$\min_{\mathbf{w}} \sum_{i=1}^{N} (\mathbf{w}^{T} \mathbf{x}^{i} - \mathbf{y}^{i})^{2} + \alpha \mathbf{w}^{T} \mathbf{w}, \ \alpha = \frac{\lambda}{\beta}$$

>> 例子: 直线拟合

- ■輸入自变量x, 目标变量 y
- ■目标: 利用直线拟合数据
- ■线性回归的目标: 利用给定的样本,

来学习函数: $f(w,x) = w_0 + w_1 x$

$$\mathbf{w} = (w_0, w_1)^T$$

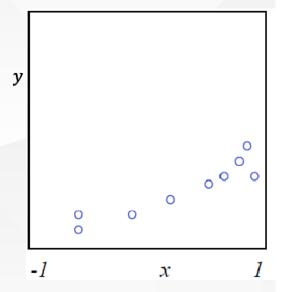


>> 产生数据

- ■利用 w_0 =-0.3, w_1 =0.5的函数 $f(x, \mathbf{w}) = w_0 + w_1 x$ 产生数据
 - 首先从均匀分布U(-1,1)中产生 x^i ,然后利用函数 $f(x^i, w)$ 计算相应的输出。
 - ●添加 $\sigma = 5$ 的高斯噪声来得到目标变量 y^{i} ,

$$\beta = \left(\frac{1}{0.2}\right)^2 = 25$$

•对于 w 的先验,我们选择如下正态分布 $\alpha = 2$ $p(\mathbf{w} \mid \alpha) = N(0, \alpha^{-1}\mathbf{I})$



>> 具体推导

$$p(y_n \mid x_n, \mathbf{w}) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\left(w_0 + w_1 x_n - y_n\right)^2}{2\sigma^2}\right)$$

$$w_0 + w_1 x_n = y_n$$

在权向量空间,等高线是什么形状?

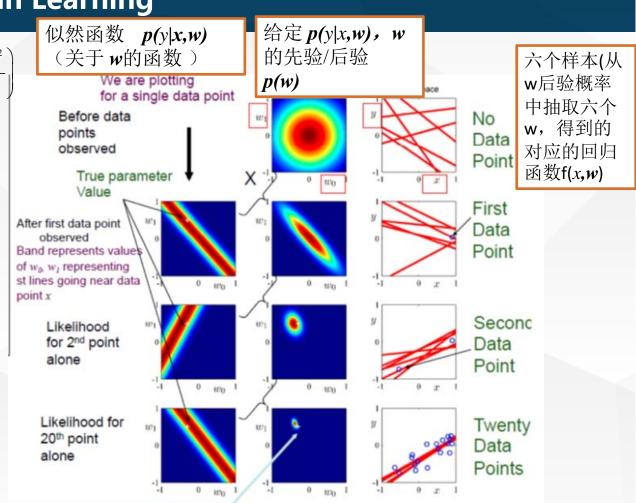
$$\begin{split} p(y_n \mid x_n, \pmb{w}) &= \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\left(w_0 + w_1 x_n - y_n\right)^2}{2\sigma^2}\right) \\ w_0 + w_1 x_n &= y_n \\ \text{在权向量空间,等高线是什么形状?} \\ &= \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta\left(w_0 + w_1 x_1 - y_1\right)^2}{2}\right) \\ &= \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta\left(w_0 + w_1 x_1 - y_1\right)^2}{2}\right) \frac{\sqrt{\alpha}}{\sqrt{2\pi}} \exp\left(-\frac{\alpha\left(w_0^2 + w_1^2\right)}{2}\right) \\ &= \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta\left(w_0 + w_1 x_1 - y_1\right)^2}{2}\right) \frac{\sqrt{\alpha}}{\sqrt{2\pi}} \exp\left(-\frac{\alpha\left(w_0^2 + w_1^2\right)}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}}^* \exp\left(-\frac{\left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T \hat{\pmb{\Sigma}}^{-1} \left(\mathbf{w} - \hat{\pmb{\mu}}\right)^T}{2}\right) \\ &= \underbrace{\mathbb{X}$$

>>> Sequential Bayesian Learning

$$p(y_n \mid x_n, \mathbf{w}) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(w_0 + w_1 x_n - y_n)^2}{2\sigma^2}\right)$$

- 由于只有两个参数 •我们可以在参数空间画出参 数的参验及后验分布
- 对后验概率进行依We look at sequential update of posterior

当有无限的数据点时, 参数的后验分布是以 其真实值为中心的 Delta分布。



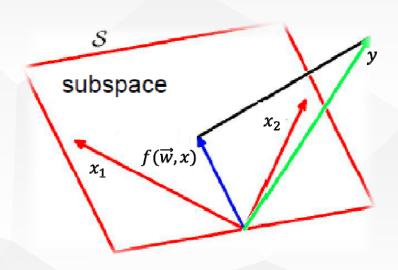
>> MLE 与MAP

- $\blacksquare \mathsf{MLE:} \quad \hat{\theta}_{MLE} = \arg\max_{\theta} P(D \mid \theta)$
- $\blacksquare \mathsf{MAP}: \ \hat{\theta}_{MAP} = \arg\max_{\theta} P(\theta \mid D) = \arg\max_{\theta} P(D \mid \theta) P(\theta)$
- ■MLE 可以看作是使用均匀分布作为先验的 MAP。
- ■MLE是频率学派的想法, 而MAP是贝叶斯学派的。
- 二者之间的比较与 ERM 与 SRM 的比较类似。
- ■更多的数据会使MLE拟合得更好,但是容易过拟合。
- ■MAP能够通过添加参数的先验产生与正则项类似的效果。添加参数先验也可看作是模型选择。

>> LMS-几何解释

- ■如果 w 的维度(m)小于N, 那么样本 x_i s 在m维的子空间 S
- $X = (x^1, x^2, \dots, x^N), 那么 f(w, X) 是$ 个N维向量
- 最优的w 使得 f(w,X) 为子空间 S 中与y最近的向量。

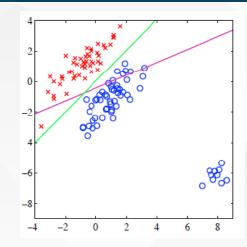
相当于y在子空间S上的投影。



分类方法

>> 分类问题

- ■输入: N i.i.d 训练样本 $(x^i, y^i) \in X \times C$, $i = 1, \dots, N$
- ■目标函数: $f \in \mathcal{F}$
- ■损失函数: $L(f;x,y) = I_{\{f(x)\neq y\}}$
- ■期望风险 (损失): $\int I_{\{f(x)\neq y\}}dP(x,y) = P(f(x)\neq y)$



为什么不直接应用回归的损失,这样就可以用回归问题的求解方法。

异常值问题!

>> 判别函数法

- ■二类问题: y=1 or -1
- 文解 $\mathbf{w}, f(\mathbf{x}, \mathbf{w}) = \mathbf{y}$,例如 $f(\mathbf{x}, \mathbf{w}) = \operatorname{sgn}(\mathbf{w}^T \mathbf{x}) = \begin{cases} 1, \mathbf{w}^T \mathbf{x} > 0 \\ -1, \mathbf{w}^T \mathbf{x} < 0 \end{cases}$
- ■使用 y=1 或 -1, 所有数据需要满足:

$$\mathbf{w}^T \mathbf{x} \mathbf{y} > 0$$

■对于每一个错分的样本, 一些方法试图最小化如下的式子:

$$\hat{E}_p(\mathbf{w}) = -\sum_{i \in \overline{I_M}} \mathbf{w}^T \mathbf{x}^i y^i$$

错分样本的集合

>> 判别式模型: Logistic 回归

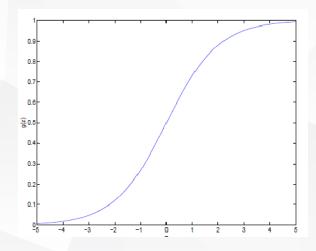
■估计后验概率 p(y|x)

$$P(y=1 | x) = f(x, w) = g(w^{T} x) = \frac{1}{1 + \exp(-w^{T} x)}$$
$$g(z) = \frac{1}{1 + e^{-z}}$$

Logistic function

Sigmoid function

- *g*(*z*)的性质
 - $z \to \infty$ 时 $g(z) \to 1$
 - $z \rightarrow -\infty$ Af, $g(z) \rightarrow 0$
 - g(z)的0-1之间
 - g'(z)=g(z)(1-g(z))



→ 求解:最大似然估计 (Maximum Likelihood Estimator)

• 概率分布:

$$P(y | x, w) = (f(x, w))^{y} (1 - f(x, w))^{1-y}$$

• 似然:

$$L(w) = \prod_{i=1}^{N} P(y^{i} | x^{i}, w) = \prod_{i=1}^{N} (f(x^{i}, w))^{y^{i}} (1 - f(x^{i}, w))^{1-y^{i}}$$

Bernoulli

• 最大化log 似然:

$$l(w) = \log L(w) = \sum_{i=1}^{N} \left(y^{i} \log f(x^{i}, w) + (1 - y^{i}) \log \left(1 - f(x^{i}, w) \right) \right)$$

• 梯度:

$$\frac{\partial l(\mathbf{w})}{\partial w_j} = (y^i - f(\mathbf{x}^i, \mathbf{w}))x_j^i, \forall (\mathbf{x}^i, y^i)$$

• SGD:

$$w_j = w_j + \alpha(y^i - f(\mathbf{x}^i, \mathbf{w}))x_j^i$$

>> 补充材料

•
$$g'(z)=g(z)(1-g(z))$$

$$g(z) = \frac{1}{1 + e^{-z}}$$

$$L(w) = \prod_{i=1}^{N} P(y^{i} \mid x^{i}, w) = \prod_{i=1}^{N} (f(x^{i}, w))^{y^{i}} (1 - f(x^{i}, w))^{1-y^{i}}$$

$$P(y=1|x) = f(x,w) = g(w^{T}x) = \frac{1}{1 + \exp(-w^{T}x)}$$

$$l(w) = \log L(w) = \sum_{i=1}^{N} \left(y^{i} \ln f(x^{i}, w) + (1 - y^{i}) \ln \left(1 - f(x^{i}, w) \right) \right)$$

$$\frac{\partial l(\mathbf{w})}{\partial w_j} = \sum_{i=1}^{N} \left(y^i \frac{1}{f(\mathbf{x}^i, \mathbf{w})} \frac{\partial f(\mathbf{x}^i, \mathbf{w})}{\partial w_j} + (1 - y^i) \frac{1}{\left(1 - f(\mathbf{x}^i, \mathbf{w})\right)} \frac{-\partial f(\mathbf{x}^i, \mathbf{w})}{\partial w_j} \right)$$

$$\frac{\partial f(\mathbf{x}^i, \mathbf{w})}{\partial w_i} = f(\mathbf{x}^i, \mathbf{w})(1 - f(\mathbf{x}^i, \mathbf{w}))x_j^i$$

$$\frac{\partial l(\mathbf{w})}{\partial w_j} = \sum_{i=1}^{N} \left(y^i (1 - f(\mathbf{x}^i, \mathbf{w})) x_j^i + (1 - y^i) \left(-f(\mathbf{x}^i, \mathbf{w}) \right) x_j^i \right) = \sum_{i=1}^{N} \left(y^i - f(\mathbf{x}^i, \mathbf{w}) \right) x_j^i$$

>>> 多类 Logistic回归

- Softmax 函数取代logistic sigmoid: 这里 w_1, w_2, \dots, w_K 是待学习的参数。
- $P(C_k \mid \mathbf{w}, \mathbf{x}) = \frac{\exp(\mathbf{w}_k^T \mathbf{x})}{\sum_{j=1}^K \exp(\mathbf{w}_j^T \mathbf{x})}$
- ■y可以看作是取K值之一的离散变量
 - 将y表示为K维的向量
 - •如果 y = 3, 那么y = (0,0,1,0,...,0) (第三个元素为1, 其它元素为0)

One hot 表示 独热表示

•K维向量满足:

$$\sum_{i=1}^{K} P(y_i \mid \boldsymbol{w}, \boldsymbol{x}) = 1$$

■ 概率分布:

$$\mu_i = P(y_i = 1 \mid \boldsymbol{w}, \boldsymbol{x})$$

$$P(\boldsymbol{y} \mid \boldsymbol{\mu}) = \prod_{i=1}^K \mu_i^{y_i}$$

Generalized Bernoulli 广义伯努里分布

>> 似然函数

$$P(y^{1}, y^{2}, \dots, y^{N} | w_{1}, w_{2}, \dots, w_{K}, x^{1}, \dots, x^{N})$$

$$= \prod_{i=1}^{N} \prod_{k=1}^{K} P(C_{k} | w_{k}, x^{i}) = \prod_{i=1}^{N} \prod_{k=1}^{K} \mu_{ik}^{y_{k}^{i}}$$

■其中

$$\mu_{ik} = \frac{\exp(\boldsymbol{w}_k^T \boldsymbol{x}^i)}{\sum_{j=1}^K \exp(\boldsymbol{w}_j^T \boldsymbol{x}^i)}$$

→ 优化多类LR

■最优化问题

$$\min E = -\ln P(\mathbf{y}^1, \dots, \mathbf{y}^N \mid \mathbf{w}_1, \dots, \mathbf{w}_k, \mathbf{x}^1, \dots, \mathbf{x}^N)$$



$$\min - \sum_{i=1}^{N} \sum_{k=1}^{K} y_k^i \ln \mu_{ik}$$
 Cross Entropy Loss Function 交叉熵损失函数

■梯度:

$$\nabla_{\mathbf{w}_j} E = \sum_{i=1}^N (\mu_{ij} - y_j^i) \mathbf{x}^i$$

■N: 样本数量

>> 补充材料

$$E = -\sum_{i=1}^{N} \sum_{k=1}^{K} y_{k}^{i} \ln \mu_{ik}$$

$$\mu_{ik} = \frac{\exp(\mathbf{w}_{k}^{T} \mathbf{x}^{i})}{\sum_{j=1}^{K} \exp(\mathbf{w}_{j}^{T} \mathbf{x}^{i})}$$

$$\ln \mu_{ik} = \ln \exp(\mathbf{w}_{k}^{T} \mathbf{x}^{i}) - \ln \sum_{j=1}^{K} \exp(\mathbf{w}_{j}^{T} \mathbf{x}^{i})$$

$$if \quad k = j, \frac{\partial \ln \mu_{ij}}{\partial \mathbf{w}_{j}} = \mathbf{x}^{i} - \frac{\mathbf{x}^{i} \exp(\mathbf{w}_{j}^{T} \mathbf{x}^{i})}{\sum_{j=1}^{K} \exp(\mathbf{w}_{j}^{T} \mathbf{x}^{i})} = \mathbf{x}^{i} - \mathbf{x}^{i} \mu_{ij}$$

$$if \quad k \neq j, \frac{\partial \ln \mu_{ik}}{\partial \mathbf{w}_{j}} = \frac{\mathbf{x}^{i} \exp(\mathbf{w}_{j}^{T} \mathbf{x}^{i})}{\sum_{j=1}^{K} \exp(\mathbf{w}_{j}^{T} \mathbf{x}^{i})} = -\mathbf{x}^{i} \mu_{ij}$$

$$\nabla_{\mathbf{w}_{j}} E = -\sum_{i=1}^{N} \left(\sum_{k=1}^{K} (-y_{k}^{i} \mathbf{x}^{i} \mu_{ij}) + y_{j}^{i} \mathbf{x}^{i} \right) = \sum_{i=1}^{N} \left(\sum_{k=1}^{K} (y_{k}^{i} \mathbf{x}^{i} \mu_{ij}) - y_{j}^{i} \mathbf{x}^{i} \right)$$

$$= \sum_{i=1}^{N} \left(\mathbf{x}^{i} \mu_{ij} \sum_{k=1}^{K} (y_{k}^{i}) - y_{j}^{i} \mathbf{x}^{i} \right)$$
由于
$$\sum_{k=1}^{K} (y_{k}^{i}) = 1$$
所以上式=
$$\sum_{i=1}^{N} \left(\mu_{ij} - y_{j}^{i} \right) \mathbf{x}^{i}$$

多类 Logistic 回归(另一种形式)

■ 只学习K-1的向量 w_1, w_2, \dots, w_{K-1}

$$P(C_k \mid \mathbf{w}, \mathbf{x}) = \frac{\exp(\mathbf{w}_k^T \mathbf{x})}{1 + \sum_{j=1}^{K-1} \exp(\mathbf{w}_j^T \mathbf{x})} \qquad P(C_K \mid \mathbf{w}, \mathbf{x}) = \frac{1}{1 + \sum_{j=1}^{K-1} \exp(\mathbf{w}_j^T \mathbf{x})}$$

■交叉熵损失:

$$\min - \sum_{i=1}^{N} \sum_{k=1}^{K} y_k^i \ln \mu_{ik}$$

$$\mu_{ik} = \frac{\exp(\mathbf{w}_k^T \mathbf{x}^i)}{1 + \sum_{j=1}^{K-1} \exp(\mathbf{w}_j^T \mathbf{x}^i)}, \mu_{iK} = \frac{1}{1 + \sum_{j=1}^{K-1} \exp(\mathbf{w}_j^T \mathbf{x}^i)}$$

■梯度:

$$\nabla_{\mathbf{w}_j} E = \sum_{i=1}^N (\mu_{ij} - y_j^i) \mathbf{x}^i$$

▶ 生成式模型: 高斯判别分析Gaussian Discriminative Analysis

■伯努利分布: 单个二进制变量 $x \in \{0,1\}$ 的分布由单个连续参数 $\in [0,1]$ 控制。 $P(x \mid \beta) = \beta^x (1-\beta)^{1-x}$

■二项式分布: 给出N个服从伯努利分布的样本中观察到m次x=1的的概

本: $P(m \mid N, \beta) = {N \choose m} \beta^m (1 - \beta)^{N-m}$

■多项式分布 (Multinomial Distribution) 是二项式分布的推广。变量可以取K个状态,第k个状态被观测到了 m_k 次的概率为:

$$P(m_1, \dots, m_K \mid N, \beta) = \binom{N}{m_1, \dots, m_K} \prod_{k=1}^K \beta_k^{m_k}$$

>> 多变量正态分布

