410-57, Assignment #3 Anamitra Bhattacharyya

#### INTRODUCTION

This assignment specifically deals with building linear regression models (LRMs), applying and comparing transformations on variables, to improve fit, in addition to identifying and removing select outliers. The overall goal of the assignment and project as a whole is to identify suitable predictors of house sale price in Ames, Iowa. The variables are derived from an observational data set from the Ames Assessor's Office used in obtaining values for individual residential properties sold in Ames, Iowa between 2006 and 2010. The transformed LRMs and multiple regression models (MRMs) are characterized separately and compared, as well as evaluated for goodness-of-fi. Furthermore, identification and removal of outliers in the models of house sales price is performed and evaluated using analysis of variance (ANOVA) metrics, such as the F-statistic and p-value, to determine whether they can be further improved.

#### **RESULTS**

Part A: Transformations – Comparison of Y versus Log(Y)

#### 1. Transformations of X and Y

For the purposes of this section, various transformations of SalePrice (e.g. log and square root) and GrLivArea (e.g. Log) were performed. The transformations were appended to the Ames housing data set as shown in the table below.

Obs	BldgType	OverallQual	BsmtFinSF1	CentralAir	GrLivArea	FullBath	BedroomAbvGr	KitchenQual	GarageCars	WoodDeckSF	SalePrice	TotalFlrSF	houseage	LogSalePrice	SqrtSalePrice	LogGrLivArea
1	1Fam	6	639	Y	1656	1	3	TA	2	210	215000	1656	50	12.2784	463.681	7.41216
2	1Fam	5	468	Y	896	1	2	TA	1	140	105000	896	49	11.5617	324.037	6.79794
3	1Fam	6	923	Y	1329	1	3	Gd	1	393	172000	1329	52	12.0552	414.729	7.19218
4	1Fam	7	1065	Y	2110	2	3	Ex	2	0	244000	2110	42	12.4049	493.964	7.65444
5	1Fam	5	791	Y	1629	2	3	TA	2	212	189900	1629	13	12.1543	435.775	7.39572
6	1Fam	6	602	Y	1604	2	3	Gd	2	360	195500	1604	12	12.1833	442.154	7.38026
7	TwnhsE	8	616	Y	1338	2	2	Gd	2	0	213500	1338	9	12.2714	462.061	7.19893
8	TwnhsE	8	263	Y	1280	2	2	Gd	2	0	191500	1280	18	12.1626	437.607	7.15462
9	TwnhsE	8	1180	Y	1616	2	2	Gd	2	237	236500	1616	15	12.3737	486.313	7.38771
10	1Fam	7	0	Υ	1804	2	3	Gd	2	140	189000	1804	11	12.1495	434.741	7.49776

*Conclusion:* Transformations of SalePrice and GrLivArea were created and will be evaluated and used in the following sections of this assignment.

#### 2. Model fitting transformations

Four models were fitted using SalePrice and GrLivArea and the various transformations on predictor and response variables. These models comprised:

- a) GrLivArea (X) versus SalePrice (untransformed, starting point)
- b) GrLivArea (X) versus LogSalePrice
- c) LogGrLivArea (X) versus SalePrice
- d) LogGrLivArea (X) versus LogSalePrice

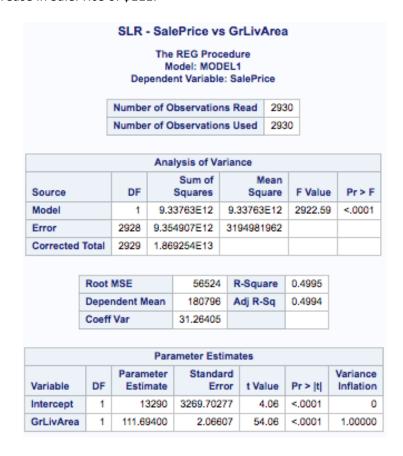
Each of these models were systematically analyzed for goodness-of-fit (GOF) and adequacy (see data below).

#### (a) GrLivArea (X) versus SalePrice (Y)

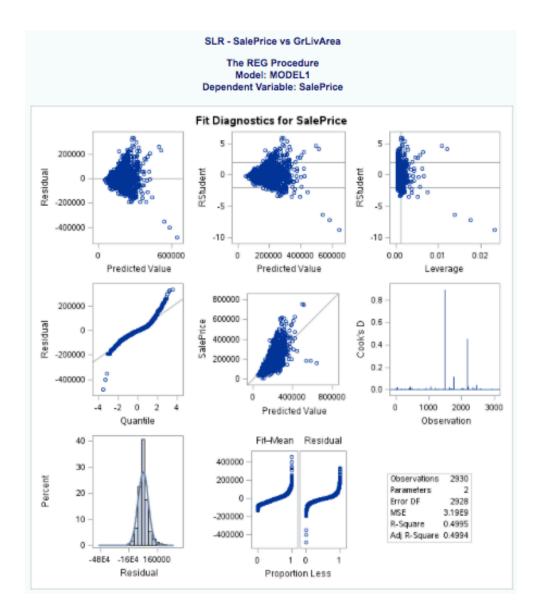
For the purposes of this section, Y= SalePrice X=GrLivArea ( $R^2$ =0.4995). The R-squared for the fitted LRM was approximately 50% (see table below), which is good and suggests that 50% of the variance in SalePrice is explained by GrLivArea. The equation of the fitted, single LRM is described below. In the format y=b0 + b1x1 + e, where x=GrLivArea, y=SalePrice and e = error term.

SalePrice = 13290 + 111GrLivArea, that is, y = 13290 + 111x

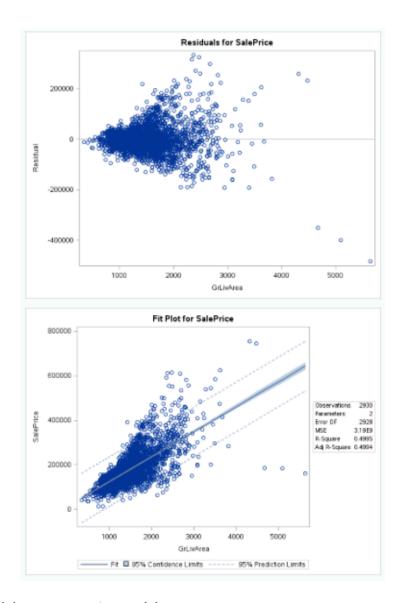
The table of parameter estimates shown below allows one to derive the formula of the linear regression model for SalePrice and GrLivArea (noted above). From the formula slope for the fitted model (above), we interpret this as: for every 1 sq. foot increase in a property GrLivArea there is an increase in SalePrice of \$111.



From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between SalePrice and GrLivArea. The R-squared value is approximately 0.5, indicating that approximately 50% of the variance in SalePrice is explained by GrLivArea. Reviewing the diagnostics for goodness-of-fit (see graphic below), the quantile-quantile plot of residuals (QQplot; central left plot), shows deviation away from the ideal line, that is predicted to be obtained if it were a normal distribution. Instead there is a deviation of points, especially at the upper end away from normality.



Looking at the residuals (top left panel above) rather than having a random distribution points around the zero level, there is a funnel-shape distribution of residuals. The QQ plot shows (center left panel above) indicating significant deviation away from the normality distribution line, especially at the higher quartile range. The SalePrice against predicted value plot (middle center above) shows a line through the data points, which does not align with the best-fit 45° line.



#### (b) GrLivArea (X) versus LogSalePrice (Y)

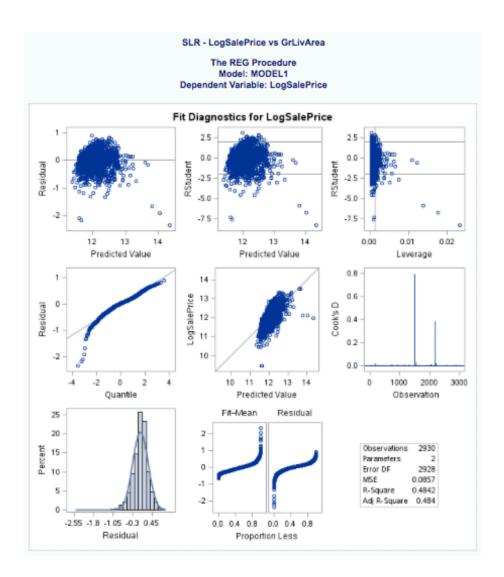
For the purposes of this section, Y= LogSalePrice X=GrLivArea. The R-squared ( $R^2$ =0.4842) for the fitted LRM was approximately 48% (see table below), which has decreased from the untransformed data see (a) above (approx. 0.5). The equation of the fitted, single LRM is described below. In the format y=b0 + b1x1 + e, where x=GrLivArea, y=LogSalePrice and e = error term.

LogSalePrice = 11 + 0.00056107GrLivArea, that is, y = 11 + 0.00056107x

The table of parameter estimates shown below allows one to derive the formula of the linear regression model for LogSalePrice and GrLivArea (noted above). From the formula slope for the fitted model (above), we interpret this as: for every 1 sq. foot increase in a property GrLivArea there is an increase in LogSalePrice of \$0.00056107.

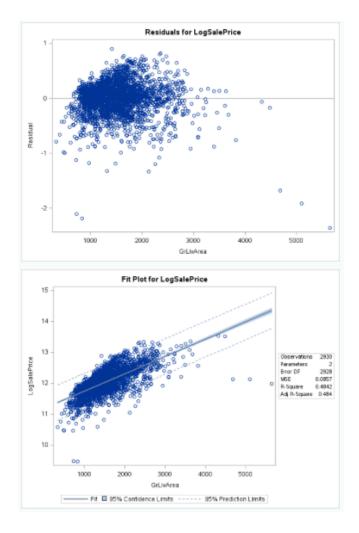
	SLR - LogSalePrice vs GrLivArea  The REG Procedure  Model: MODEL1  Dependent Variable: LogSalePrice									
			_				290 290	-		
		A	na	lysis of V	aria	ince				
Source	Source		Sum of Squares			Mean Square		F Value	Pr > F	
Model		1	235.61694		23	235.61694		2748.89	<.0001	
Error		2928	2	250.96931		0.08571				
Correcte	ed Tot	al 2929	486.58626							
	Root	MSE		0.29277	,	R-Squ	are	0.4842		
	Depe	ndent Mean		12.02097	,	Adj R-	Sq	0.4840		
	Coef	f Var		2.43548	3					
			_	meter Es						
Variable			Parameter Estimate			ard ror t Value		Pr >  t	Variance Inflation	
Intercept	1	11.1795	4	0.01	694	660	.12	<.0001		
GrLivArea	1	0.0005610	7	0.00001	070	52	.43	<.0001	1.00000	

From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between LogSalePrice and GrLivArea. The R-squared value is approximately 0.48, indicating that approximately 48% of the variance in LogSalePrice is explained by GrLivArea. Reviewing the diagnostics for goodness-of-fit (see graphic below), the quantile-quantile plot of residuals (QQplot; central left plot), has improved at the upper end of the distribution but still shows deviation away from the ideal line at the lower range, that is predicted to be obtained if it were a normal distribution.



Looking at the residuals (top left panel above) the funnel-shaped distribution has diminished so there is a more random distribution of points around the zero level. The QQ plot shows (center left panel above) indicating some deviation away from the normality distribution line, especially at the lower quartile range. The LogSalePrice against predicted value plot (middle center above) shows a line through the data points, which has improved its alignment with the best-fit 45° line, but still there is room for improvement.

Inspection of the scatter-plot of LogSalePrice *versus* GrLivArea, there is a greater proportion of data points within the prediction range (see below), though there continues to numerous points outside that range as well.



#### (c) LogGrLivArea (X) versus SalePrice

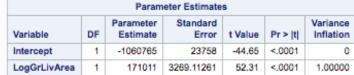
For the purposes of this section, Y=SalePrice X=LogGrLivArea. The R-squared ( $R^2$ =0.4831) for the fitted LRM was approximately 48% (see table below), which has decreased from the untransformed data see (a) above (approx. 0.5) and is the same as the  $R^2$ value for (b) LogSalePrice vs GrLivArea. The equation of the fitted, single LRM is described below. In the format y=b0 + b1x1 + e, where x=LogGrLivArea, y=SalePrice and e = error term.

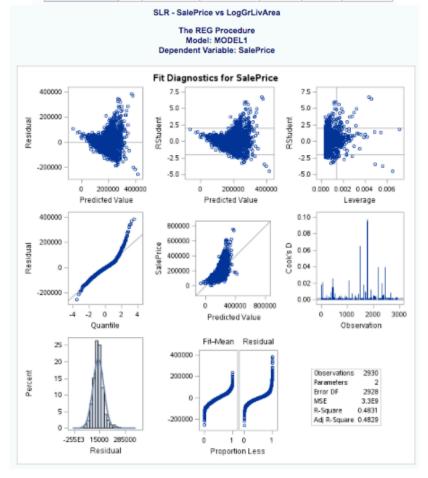
SalePrice = -1060765 + 171011LogGrLivArea, that is, y = -1060765 + 171011x

The table of parameter estimates shown below allows one to derive the formula of the linear regression model for SalePrice and LogGrLivArea (noted above). From the formula slope for the fitted model (above), we interpret this as: for every 1 unit increase in LogGrLivArea there is an increase in SalePrice of \$171,011.

From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between SalePrice and LogGrLivArea. The R-squared value is approximately 0.48, indicating that approximately 48% of the variance in LogSalePrice is explained by LogGrLivArea. Reviewing the diagnostics for goodness-of-fit (see graphic below),

#### SLR - SalePrice vs LogGrLivArea The REG Procedure Model: MODEL1 Dependent Variable: SalePrice Number of Observations Read 2930 Number of Observations Used Analysis of Variance Sum of Mean Source DF Squares Square F Value Pr > F Model <.0001 9.030218E12 9.030218E12 2736.45 9.662319E12 3299972433 Error 2928 Corrected Total 2929 1.869254E13 Root MSE 0.4831 57445 R-Square 180796 0.4829 Dependent Mean Adj R-Sq Coeff Var 31.77358 Parameter Estimates

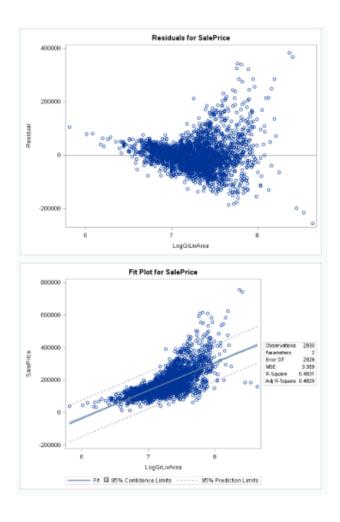




the quantile-quantile plot of residuals (QQplot; central left plot), deviates away from the normal distribution ideal line in the upper range.

Looking at the residuals (top left panel above and top panel below) there is an unusual U-shaped distribution unlike the more random distribution we saw in (b) above with LogSalePrice *versus* GrLivArea. This suggests a non-random distribution of points around the zero level. The SalePrice against predicted value plot (middle center above) shows a line through the data points, which does not align with the best-fit 45° line.

Inspection of the scatter-plot of SalePrice *versus* LogGrLivArea, there is a funnel shape distribution observed and a greater proportion of data points outside the prediction range (see below), compared to the same plot in (b) above, especially in the upper range of LogGrLivArea.



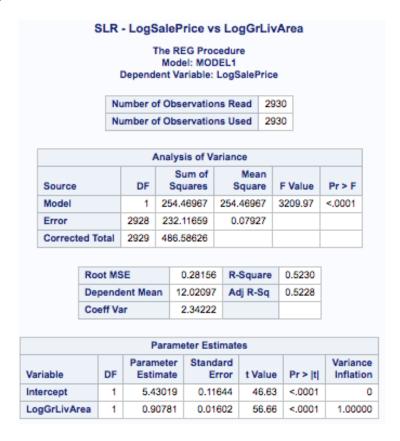
#### (d) LogGrLivArea (X) versus LogSalePrice

For the purposes of this section, Y=LogSalePrice X=LogGrLivArea. The R-squared ( $R^2$ =0.5230) for the fitted LRM was approximately 52% (see table below), which has increased from the untransformed data see (a) above (approx. 0.5). The equation of the fitted, single LRM is

described below. In the format y=b0 + b1x1 + e, where x=LogGrLivArea, y=LogSalePrice and e=error term.

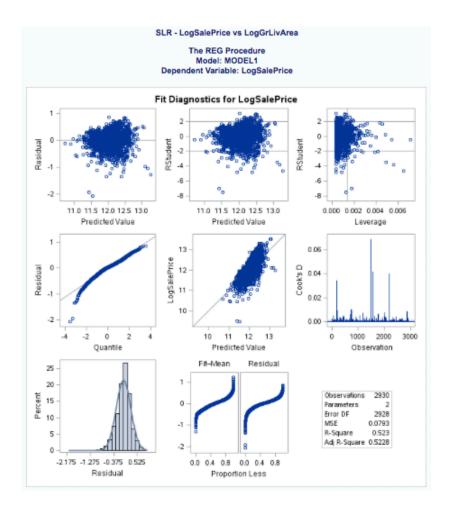
LogSalePrice = 5 + 0.9LogGrLivArea, that is, y = 5 + 0.9x

The table of parameter estimates shown below allows one to derive the formula of the linear regression model for SalePrice and LogGrLivArea (noted above). From the formula slope for the fitted model (above), we interpret this as: for every 1 unit increase in LogGrLivArea there is an increase in LogSalePrice of \$0.9.



From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between LogSalePrice and LogGrLivArea. The R-squared value is approximately 0.52, indicating that approximately 52% of the variance in LogSalePrice is explained by LogGrLivArea. Reviewing the diagnostics for goodness-of-fit (see graphic below), the quantile-quantile plot of residuals (QQplot; central left plot), there is a good fit but there continues to be deviation away from the normal distribution ideal line in the lower range.

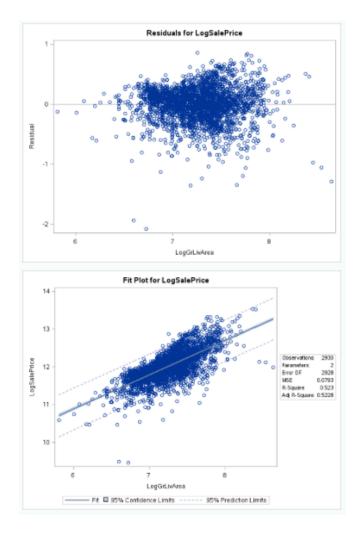
Looking at the residuals (top left panel below) there is no unusual shaped distribution unlike in (a) and (c) above, indicating a more random distribution like we saw in (b) above with LogSalePrice versus GrLivArea. This suggests a non-random distribution of residuals around the zero level. The SalePrice against predicted value plot (middle center above) shows a line through the data points, which is much more in alignment with the best-fit 45° line as in (b), LogSalePrice versus GrLivArea.



Looking at the residuals in the plot below versus LogGrLivArea shows a relatively random distribution of points either side of the zero point. Inspection of the scatter-plot of LogSalePrice *versus* LogGrLivArea, there is a good distribution of data points within the prediction range (see below), better than (a) and (c), although it is slightly worse than (b) which appeared to have a greater portion of observations with the prediction range compared to (d).

Summary
A summary table of the four models created is shown below for comparison.

Model	R <sup>2</sup> coefficient metric
a) GrLivArea (X) versus SalePrice	0.4995
b) GrLivArea (X) versus LogSalePrice	0.4842
c) LogGrLivArea (X) versus SalePrice	0.4831
d) LogGrLivArea (X) versus LogSalePrice	0.5230



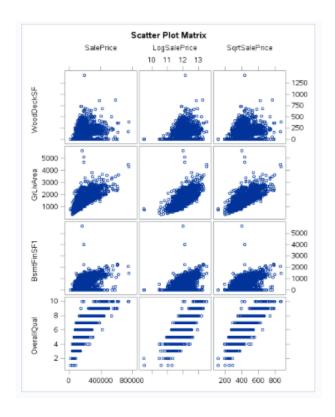
Conclusion: Model (d) based on the R-squared metrics scores highest (0.52) of the four models, and is the best fitting model. Based on the scatter plot of residuals, models (b) and (d) are best, though the only reservation with (d) is that there appears to be more observations outside the prediction range in (d) compared to (b), where the observations are more tightly clustered at the lower range. One of the issues with using a transformed variable(s) is they have been transformed and are thus more difficult to interpret back to reality, they are an abstraction from reality. So for instance, while the original untransformed model in (a) is unaltered it is simple to interpret the effect of a change in the predictor variable on the real-world response variable -sale price. In contrast, when using the Log of the SalePrice one always has to remember to reverse translate the log value (log Y) to Y to derive a meaningful result about SalePrice.

#### 3. Best Predictor with Transformations of SalePrice

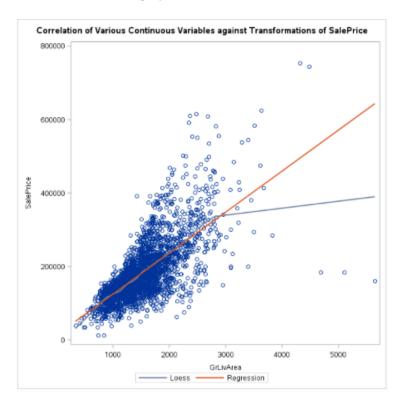
The <u>continuous</u> variables used in this part of the assignment were WoodDeckSF, GrLivArea, BsmtFinSF1 and OverallQual. These variables were use to develop correlations with transformations of the response variable – SalePrice, specifically: SalePrice, LogSalePrice and SqrtSalePrice. The results of this analysis are shown below.

				The	COR	R Proc	edure			
	4 Witi	Vith Variables: WoodDeckSF GrLivArea BsmtFinSF1 OverallQua					al			
	3 Vari	ables:		SalePr	ice Lo	gSalePi	rice SqrtS	SaleP	rice	
				S	imple	Statist	tics			
Variable		N		Mean	S	td Dev	S	um	Minimum	Maximum
WoodDe	ckSF	2930	93.	75188	126	36156	274	693	0	1424
GrLivAre	а	2930		1500	505	50889	4394	093	334.00000	5642
BsmtFin9	SF1	2929	442.	62957	455	59084	1296	462	0	5644
OverallQ	ual	2930	6.	09488	1	.41103	17	858	1.00000	10.00000
SalePrice	•	2930	1	80796		79887	529732	456	12789	755000
LogSaleF	Price	2930	12.	02097	0.	40759	35	221	9.45634	13.53447
SqrtSalel	Price	2930	416.26208		86	74391	1219	648	113.08846	868.90736
			P	earson Correlation Coefficients Prob >  r  under H0: Rho=0 Number of Observations						
				SaleP	rice	LogSalePrice		Sqr	rtSalePrice	
	V	VoodDe	kSF		714	(	0.33332		0.33566	
					0001 2930		<.0001		<.0001 2930	
	G	irLivAre	а	_	678		0.69586		0.71240	
					0001		<.0001		<.0001	
		smtFin9	254	_	930		2930 0.41080		2930	
	В	smtrina	<b>5</b> F1		3291 3001	'	<.0001		0.42719 <.0001	
				2	929		2929		2929	
	0	verallQ	ual		926		0.82564		0.82460 <.0001	
				J 40	1001		<.0001			

For each of the predictor variables (rows) in the table above, the R-squared value stays approximately the same irrespective of the SalePrice or LogSalePrice (transformed). The variables GrLivArea and OverallQual score most highly in terms of R-squared. For GrLivArea the R-squared using LogSalePrice goes down fractionally from 0.70678 (untransformed) to 0.69586 (LogSalePrice). For OverallQual the R-squared value increases with the log transformation 0.79926 (untransformed) to 0.82564. A scatter plot matrix for all these variables is shown below; note that the GrLivArea scatter shows a strong correlation and the scatter is less pronounced compared to WoodDeckSF and BsmtFinSF1.

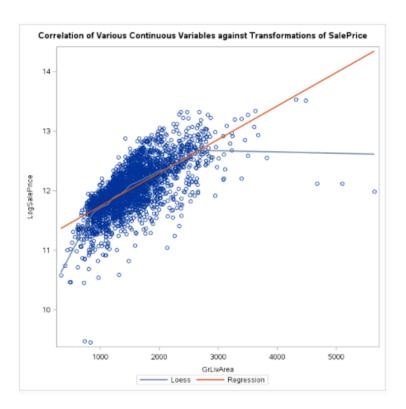


A scatterplot matrix using one of the highly scoring continuous variables (e.g. GrLivArea) and transformations of SalePrice (e.g. SalesPrice, LogSalePrice, SqrtSalePrice) were employed for this analysis. The results are shown in the graphics below.



For the untransformed plot above there is a good correlation of variables until approximately x=2,600 for GrLivArea, with the fitted line approximately passing through the origin. The linear model (brown line) and LOESS (blue line) are collinear below x=2,600. Above x=2,600 there are outliers that act to distort the best fit line as shown by the LOESS line (blue) diverging at an acute angle.

For the correlation plot with LogSalePrice *versus* GrLivArea there is a similar effect, except that the y-intercept is pushed up further and the point of divergence of the linear fitted model and LOESS is at approximately x=2,400. The observations are also more tightly clustered with the linear fitted model (brown line). Also it should be noted that the LOESS line diverges not only at the higher ranges of x but also at the lower end.



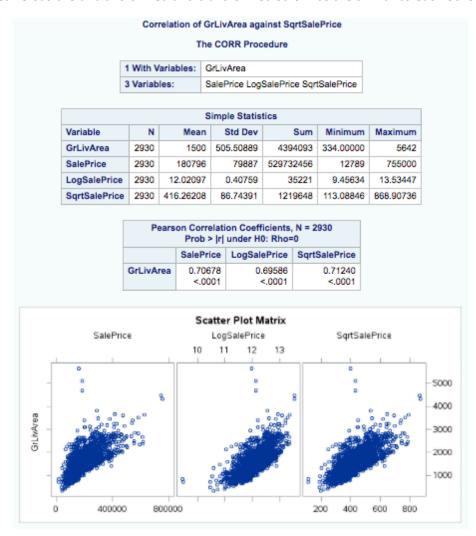
Conclusion: The GrLivArea is an example of a continuous variable that correlates strongly with both SalePrice and LogSalePrice in terms of R-squared specifically 0.71 and 0.69, respectively. The plot of the LogSalePrice versus GrLivArea produces a much more clustered set of observations, though above x=2,400 there is a divergence in the LOESS line due to the presence of potential outlier observations at the higher ranges of GrLivArea, that can affect the trajectory of the fitted model.

#### 4. Other transformations

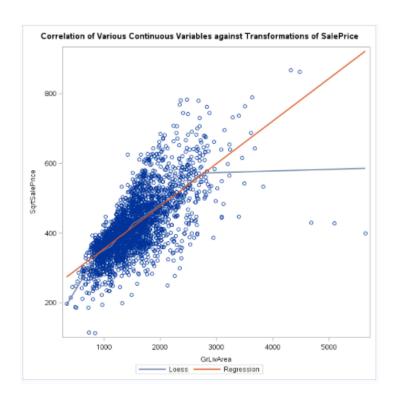
From the analysis performed in part 3 above, the log transformation improves response variable (e.g. SalePrice) when there is a lot of variability or scatter, and acts to tighten the spread of the observations. Also from the analysis performed, we only know that the fit is improved in a

simple linear regression model (LRM), we have not evaluated it in a multiple regression model (MRM) yet.

Another transformation was chosen to transform SalePrice specifically, square root of SalePrice (SqrtSalePrice). The correlation with GrLivArea and SalePrice, LogSalePrice and SqrtSalePrice is shown and compared below. The R-squared value improves slightly for SqrtSalePrice (0.71240) over LogSalePrice (0.69586) and SalePrice (0.70678). The scatter plots of the observations of GrLivArea *versus* the untransformed and transformed SalePrice are similar to each other.



The scatter plot of SqrtSalePrice against GrLivArea has a strong correlation (R<sup>2</sup>=0.71240), see plot below. The scatter below illustrates that the LRM fitted line (brown) and the LOESS (blue) lines diverge at a value of GrLivArea of approximately 2,600 sq. ft. due to outliers.



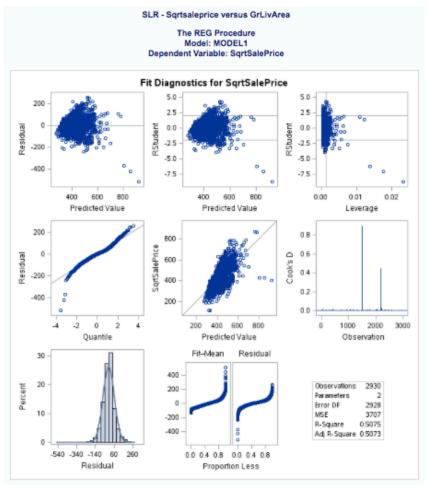
The equation of the fitted, single LRM is described below. In the format y=b0 + b1x1 + e, where x=GrLivArea, y=SqrtSalePrice and e=error term.

SqrtSalePrice = 232 + 0.1GrLivArea, that is,

Sqrt(y) = 232 + 0.1x

		SLF	R - Sqr	tsa	leprice v	ers	us GrLi	vAre	ea .	
		De		M	REG Pro odel: MC Variable	DE	L1	Price	,	
		Nun	Number of Observations Read					293	10	
		Nun	nber o	f O	bservati	ons	Used	293	10	
			Α	nal	ysis of \	/aria	ance			
Source			DF	Sum of Squares			Mean Square		F Value	Pr > F
Model	Model			11	185123	,	11185123		3017.28	<.0001
Error			2928	10	0854156		07.0205	0		
Correct	ed Tot	al :	2929	22	2039279					
	Root	MSE		60.8853		31	R-Square		0.5075	5
	Depe	nder	nt Mear	n	416.262	80	Adj R-	Sq	0.5073	3
	Coeff	Var			14.626	68				
			P	ara	meter E	stim	nates			
Variable	DI		arame Estima			ard ror	t Valu	ie I	Pr >  t	Variance Inflation
Intercept	1	1 2	32.932	09	3.52	198	66.1	4 -	<.0001	(
GrLivAre	a f	1	0.122	25	0.002	223	54.9	3 4	<.0001	1.00000

The table of parameter estimates shown above allows one to derive the formula of the linear regression model for SqrtSalePrice and GrLivArea (noted above). From the formula slope for the fitted model (above), we interpret this as: for every 1 unit increase in GrLivArea there is an increase in SqrtSalePrice of \$0.1. From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between SqrtSalePrice and GrLivArea. The R-squared value is approximately 0.5075, indicating that approximately 51% of the variance in SqrtSalePrice is explained by GrLivArea. Reviewing the diagnostics for goodness-of-fit (see graphic below), the quantile-quantile plot of residuals (QQplot; central left plot), there is a good fit but there continues to be deviation away from the normal distribution ideal line in the upper and especially lower ranges.



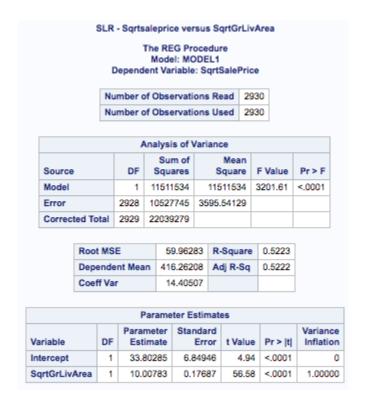
Looking at the residuals (top left panel above) there is a little funnel-shaped distribution indicating less than random distribution The SqrtSalePrice against predicted value plot (middle center above) shows a line through the data points, which is not alignment with the best-fit 45° line.

Another regression model was fitted using SqrtSalePrice and SqrtGrLivArea. The equation of the fitted, single LRM is described below. In the format y=b0 + b1x1 + e, where x=SqrtGrLivArea, y=SqrtSalePrice and e=error term.

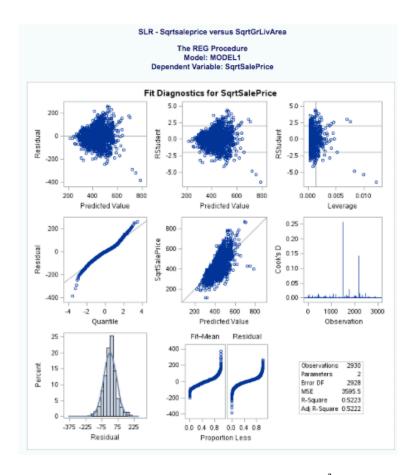
SqrtSalePrice = 33 + 10SqrtGrLivArea, that is,

#### Sqrt(y) = 232 + 0.1Sqrt(x)

The table of parameter estimates shown above allows one to derive the formula of the linear regression model for SqrtSalePrice and SqrtGrLivArea (noted above). From the formula slope for the fitted model (above), we interpret this as: for every 1 unit increase in SqrtGrLivArea there is an increase in SqrtSalePrice of \$10. From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between SqrtSalePrice and SqrtGrLivArea. The R-squared value is 0.5223, indicating that approximately 52% of the variance in SqrtSalePrice is explained by SqrtGrLivArea.



Reviewing the diagnostics for goodness-of-fit (see graphic below), the QQplot of residuals (central left plot), there is a reasonable fit but there continues to be deviation away from the normal distribution ideal line in the upper and lower ranges. Looking at the residuals (top left panel below) there is a little funnel-shaped distribution indicating less than random distribution The SqrtSalePrice against predicted value plot (middle center above) shows a line through the data points, which is not aligned with the best-fit 45° line through the origin.



Conclusion: The fitted models of SqrtSalePrice versus GrLivArea (R<sup>2</sup>=0.5075) and SqrtSalePrice versus SqrtGrLivArea (R<sup>2</sup>=0.5223) are comparable, though the latter is a slightly better correlation. Both the latter models have slight funnel-shaped distribution of residuals, which is absent in the model for LogSalePrice versus LogGrLivArea (R<sup>2</sup>=0.5228), suggesting the log-log model is perhaps a better model from several perspectives.

#### Part B: Outliers

#### 5. Identifying outliers

I used a the 2 standard deviation rule to filter out outliers, which is that 95% of the data under a normal distribution curve resides under (-  $2\sigma$  from the mean)) from the mean. The mean for the SalePrice is \$180,000. Since sigma is ~\$80,000, then  $2\sigma$  value is \$160,000. The lower range would be properties less than \$20,000 (- $2\sigma$  from the mean) while upper range would be above \$340,000 (+ $2\sigma$  from the mean). The counts of outlier by this definition is noted below:

Criteria	Range	Count
SalePrice < \$20,000	- 2σ	2 (removed)
SalePrice > \$340,000	+ 2σ	134 (removed)
SalePrice within range		2794

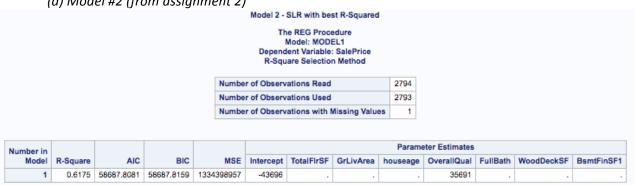
These criteria for outliers were also corroborated by the EDA that was performed in Assignment 1 earlier.

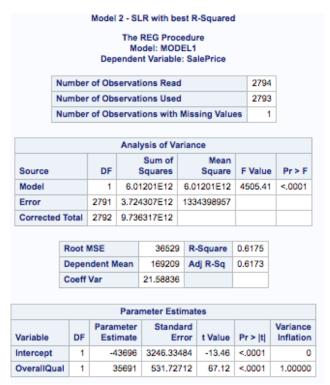
Conclusion: Using the  $2\sigma$  rule, it is possible to filter out outliers from the lower and upper SalePrice ranges, such that we are left with 2794 properties for consideration in further analysis. This results in 136 observations being removed for further analysis.

#### 6. Removing outliers

Using the cleaned data to re-fit models 2, 5 and 6 (from assignment 2), again OverallQual is the best fit as shown below. Also the table confirms that number of observations in SalePrice has been reduced to 2794 from 2930.

## (a) Model #2 (from assignment 2)

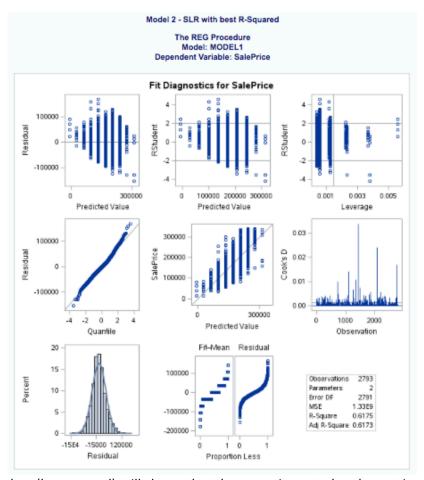




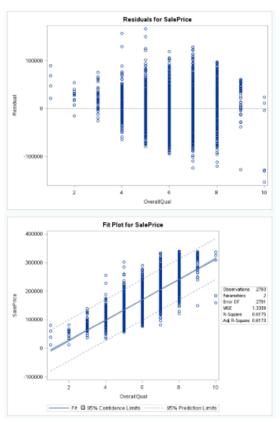
From the parameter estimates table above the equation for the cleaned fitted model is: SalePrice(Y)= -43696 + 35691OverallQual(X)

y = -43696 + 35691x

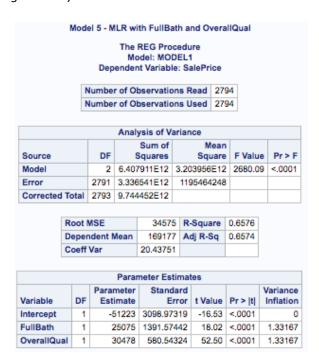
So for every unit increase in OverallQual, the SalePrice increases by \$35,691. From the analysis of variance (ANOVA) table, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between SalePrice and OverallQual with the cleaned data. The R-squared value is 0.6175, indicating that approximately 61% of the variance in SalePrice is explained by OverallQual. Reviewing the diagnostics for goodness-of-fit (see graphic below), the QQplot of residuals (central left plot), there is a reasonable fit but there continues to be deviation away from the normal distribution ideal line in the upper and lower ranges but less so than previously. Looking at the residuals (top left panel below) there is a little funnel-shaped distribution indicating less than random distribution The cleaned SalePrice against predicted value plot (middle center above) shows a line through the data points, which is quite well aligned with the best-fit 45° line through the origin.



The fit plot below (lower panel) still shows that there continues to be observations outside the prediction limits. Interestingly, there continue to be outliers in the middle ranges especially.

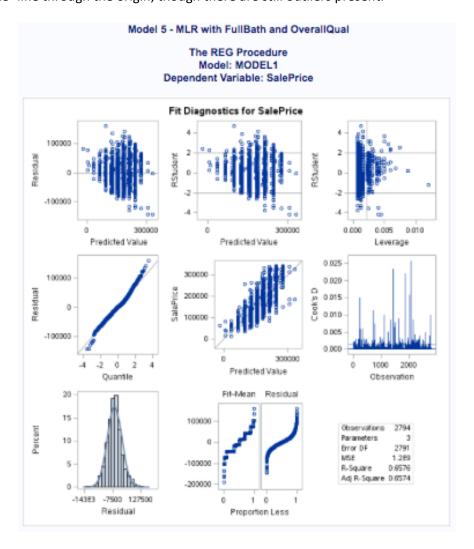


### (b) Model #5 (from assignment 2)

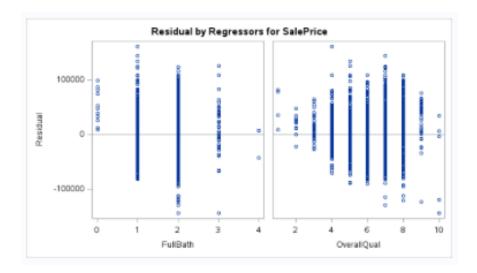


From the parameter estimates table above the equation for the cleaned fitted model is: SalePrice(Y)= -51223 + 25075FullBath(x1) + 30478OverallQual(x2) y = -51223 + 25075x1 + 30478x2

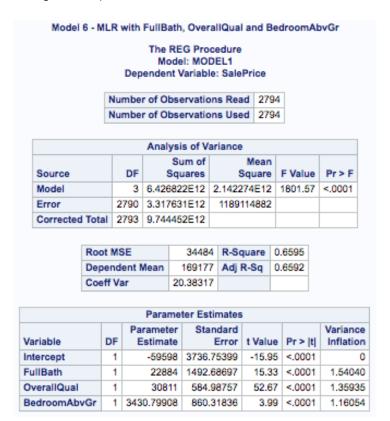
So for every unit increase in FullBath and OverallQual, the SalePrice increases by \$25,075 and \$30,478, respectively. From the analysis of variance (ANOVA) table above, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is low, indicating to reject the null hypothesis. This indicates that there is a correlation between SalePrice and OverallQual with the cleaned data. The R-squared value is 0.6576, FullBath and OverallQual explain approximately 65% of the variance in SalePrice. Reviewing the diagnostics for goodness-of-fit (see graphic below), the QQplot of residuals (central left plot), there is a reasonable fit but there continues to be deviation away from the normal distribution ideal line in the upper and lower ranges. Looking at the residuals (top left panel below) there is a little funnel-shaped distribution indicating less than random distribution The cleaned SalePrice against predicted value plot (middle center above) shows a line through the data points, which is quite well aligned with the best-fit 45° line through the origin, though there are still outliers present.



The fit plot below (lower panel) for residuals still shows a funnel-shape for the OverallQual suggesting some non-randomness, and non-random distribution of residuals at the upper and lower ranges for FullBath.



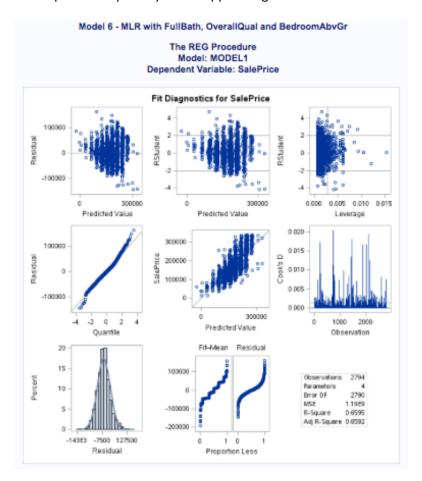
#### (c) Model #6 (from assignment 2)



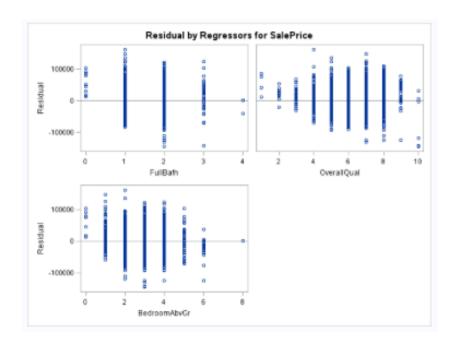
From the parameter estimates table above the equation for the cleaned fitted model is: SalePrice(Y)= -59598 + 22884FullBath(x1) + 30811OverallQual(x2) + 3430BedroonAbvGr(x3) y = -59598 + 22884x1)+ 30811x2 + 3430x3

So for every unit increase in FullBath, OverallQual and BedroomAbvGr the SalePrice increases by \$22,884, \$30,811 and \$3,430, respectively. From the analysis of variance (ANOVA) table above, to assess the quality of the fitted model, the F-statistic is significantly high and the p-value is

low, indicating to reject the null hypothesis. This indicates that there is a correlation between SalePrice and OverallQual with the cleaned data. The R-squared value is 0.6595, FullBath, OverallQual and BedroomAbvGr explain approximately 65% of the variance in SalePrice. Reviewing the diagnostics for goodness-of-fit (see graphic below), the QQplot of residuals (central left plot), there is a reasonable fit but there continues to be deviation away from the normal distribution ideal line in the upper and lower ranges. Looking at the residuals (top left panel below) there is a little funnel-shaped distribution indicating less than random distribution. The cleaned SalePrice against predicted value plot (middle center above) shows a line through the data points, which is quite well aligned with the best-fit 45° line through the origin, though there are still outliers present especially at the upper range.



The fit plot below (lower panel) for residuals shows an elliptical shape for the OverallQual suggesting some non-randomness, and non-random distribution of residuals at the upper and lower ranges for FullBath and similarly for BedroonAbvGr.



#### Summary

Using the cleaned data for SalePrice that has been refit to the models in assignment 2 a summative table of the R-squared coefficients is shown below for comparison.

Model	R <sup>2</sup> coefficient				
Wiodei	Assignment 2	Assignment 3			
(2) LRM Best variable (GrLivArea)	0.6386	0.6175			
(5) MLR with FullBath and OverallQual	0.6614	0.6576			
(6) MLR with FullBath OverallQual BedroomAbvGr	0.6629	0.6595			

Conclusion: While the R-squared coefficients for the models using the cleaned data are comparable (though fractionally less than from assignment 2, there are improvement in the distribution of the QQ plots using the cleaned versus not cleaned data. It is likely that additional cleaning of the SalePrice needs to be performed, but this is also likely for the predictor variables as well.

#### 7. Model-based outliers (Influential outliers)

After removing 'outliers' for SalePrice, there are still unusually large residuals observed in the residual plots earlier. This is in part since we only removed some outliers in SalePrice but we did not remove any with the other predictor variables. These are due to the 'influential' points, which exert a disproportionate affect on the model coefficients. These points are identified by several statistics such as DFFITS. The DFFITS statistic was used to identify and remove these influential points and re-fit the model. The results from before and after refitting the model are shown below.

A total of 147 observations were removed after cleaning further with DFFITS.

# Before refitting (2794 observations):

# Three variable Regression Model with log transformation The REG Procedure Model: MODEL1 Dependent Variable: LogSalePrice

Number of Observations Used

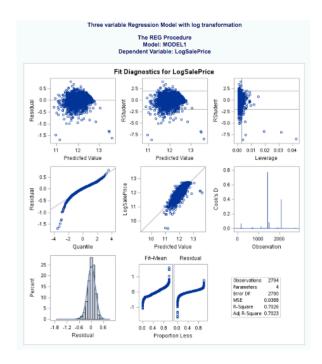
Number of Observations Read 2794

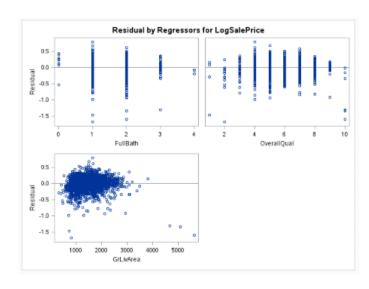
2794

Analysis of Variance										
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F					
Model	3	256.08257	85.36086	2197.24	<.0001					
Error	2790	108.38892	0.03885							
Corrected Total	2793	364.47149								

Root MSE	0.19710	R-Square	0.7026	
Dependent Mean	11.97661	Adj R-Sq	0.7023	
Coeff Var	1.64572			

	Parameter Estimates										
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variance Inflation					
Intercept	1	10.56716	0.01791	590.13	<.0001	0					
FullBath	1	0.06018	0.00904	6.65	<.0001	1.73082					
OverallQual	1	0.16768	0.00346	48.53	<.0001	1.45179					
GrLivArea	1	0.00021758	0.00001068	20.37	<.0001	1.73783					





## After refitting (2647 observations):

Three variable Regression Model with log transformation-Outliers Removed

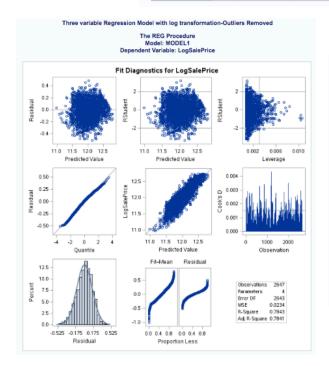
The REG Procedure Model: MODEL1 Dependent Variable: LogSalePrice

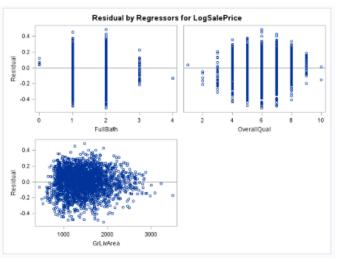
Number of Observations Read 2647 Number of Observations Used 2647

Analysis of Variance										
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F					
Model	3	225.07280	75.02427	3203.45	<.0001					
Error	2643	61.89853	0.02342							
Corrected Total	2646	286.97134								

Root MSE	0.15304	R-Square	0.7843	
Dependent Mean	11.99478	Adj R-Sq	0.7841	
Coeff Var	1.27585			

Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variance Inflation
Intercept	1	10.58913	0.01491	710.37	<.0001	0
FullBath	1	0.05447	0.00765	7.12	<.0001	1.86390
OverallQual	1	0.15725	0.00291	54.10	<.0001	1.49829
GrLivArea	1	0.00026124	0.00000932	28.02	<.0001	1.81235





Conclusion: The before and after removal of the influential points using DFFITS the R-squared value for the multiple linear regression model (MRM) were 0.7026 and 0.7843, respectively. A total of 147 observations were removed during this process. Though only 147 data points were deleted there was a substantial improvement in the fitted model afterwards suggesting that some of these 'model-based' outliers had a disproportionate effect on the overall fit of the model. Furthermore comparing before and after models, the residuals improved so that the plot of residuals versus predicted value after cleaning showed no funnel-shaped distribution, indicating a random disposition of residuals. Furthermore, the QQplot after cleaning showed a much tighter fit to the ideal normality line than before the cleaning. A similar effect was observed in the plot of LogSalePrice versus Predicted value (center middle panel above), which displayed a line through the data points, which is very well aligned with the best-fit 45° line through the origin. We conclude that the removal of influential points did improve the fit of the model based on the R-squared comparison and the disposition of the residuals.

#### **CONCLUSIONS**

In what ways do variable transformation and outlier deletion impact the modeling process and the results? Are these analytical activities a benefit or do they create additional difficulties? The use of variable transformation and outlier detection and removal, separately and in combination, helped to drastically improve the fitted models. This helped not only in tightening the observations such that there was less scatter and variability but also the distribution of residuals was improved and become more random (ideal assumption), so that no unusual shapes (e.g. funnel-shape) observed in residual plots. While the fits to the linear models and their goodness-of-fit improved, producing a 'better' model, there were observations removed from the analysis. From the observations removed from the SalePrice, most were deleted from the upper price range. While absent it is unclear how useful these observations are, though that depends on the nature of the question we are asking. Specifically, if we wanted to know what makes these expensive houses more valuable than most residences in Ames, then we are no longer in a position to address that.

What do you consider to be next steps in the modeling process?

Since we have not yet been able to incorporate categorical variables into any fitted MRM model, this remains another step. In addition further iterations of model adequacy and validation may be required in order to identify and remove additional outliers and influential points in the predictor variable data set, in addition to SalePrice. Once we are completed with these steps and we are satisfied with the model, we can set about using the model – the original purpose.