Mathematics for Computer Science: Homework 7

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Problem 1

1. We let $y_i = x_i - c_1, z_i = x'_i - c_2$, then

$$\begin{split} \Pr(|y_i| \leq 1 - \varepsilon) &= (1 - \varepsilon)^d \leq e^{-\varepsilon d}. \\ \Pr\left(|y_i y_j| \geq \frac{c}{\sqrt{d-1}}\right) \leq \frac{2}{c} e^{-\frac{c^2}{2}}. \\ \Pr\left(|y_i (c_1 - c_2)| \geq \delta \frac{c}{\sqrt{d-1}}\right) \leq \frac{2}{c} e^{-\frac{c^2}{2}}. \end{split}$$

It's similar for z_i .

Let $\varepsilon = \frac{\log n \ln n}{d}$, $c = \sqrt{\log n \ln n}$, then

$$\begin{split} \Pr(|y_i| \leq 1 - \varepsilon) \leq e^{-\varepsilon d} &= n^{-\log n} = n^{-\Omega(\log n)}. \\ \Pr\Big(|y_i y_j| \geq \frac{c}{\sqrt{d-1}}\Big) \leq \frac{2}{\sqrt{\log n \ln n}} e^{-\frac{\log n \ln n}{2}} &= \frac{2}{\sqrt{\log n \ln n}} n^{-\frac{\log n}{2}} = n^{-\Omega(\log n)} \\ \Pr\Big(|y_i (c_1 - c_2)| \geq \delta \frac{c}{\sqrt{d-1}}\Big) \leq \frac{2}{\sqrt{\log n \ln n}} e^{-\frac{\log n \ln n}{2}} &= \frac{2}{\sqrt{\log n \ln n}} n^{-\frac{\log n}{2}} = n^{-\Omega(\log n)}. \end{split}$$

Back to the original problem, let $L^2(d,n) \geq 2 + 2\frac{c}{\sqrt{d-1}}$

$$\begin{split} \Pr \left(\left| x_i - x_j \right| & \leq L(d,n) \right) = \Pr \left(\left| y_i - y_j \right|^2 \leq L^2(d,n) \right) \\ & \geq \Pr \left(y_i^2 + y_j^2 - 2 y_i y_j \leq L^2(d,n) \right) \\ & \geq \Pr (y_i^2 \leq 1) \Pr (y_j^2 \leq 1) \Pr \left(\left| y_i y_j \right| \leq \frac{c}{\sqrt{d-1}} \right) \\ & \geq 1 - n^{-\Omega(\log n)}. \end{split}$$

It's similar for x_i' .

Then, let
$$L^2(d,n) \leq 2(1-\varepsilon)^2 + \delta^2 - 2\frac{c}{\sqrt{d-1}} - 4\delta\frac{c}{\sqrt{d-1}}$$

$$\Pr(\left|x_i - x_j'\right| \geq L(d,n)) = \Pr(\left|y_i - c_1 - z_j + c_2\right|^2 \geq L^2(d,n))$$

$$= \Pr(y_i^2 + z_j^2 + (c_1 - c_2)^2 - 2y_i z_j + 2(c_1 - c_2)y_i - 2(c_1 - c_2)z_j \geq L^2(d,n))$$

$$\geq \Pr(y_i^2 + z_j^2 + (c_1 - c_2)^2 - 2y_i z_j + 2(c_1 - c_2)y_i - 2(c_1 - c_2)z_j \geq L^2(d,n))$$

$$\geq \Pr(y_i^2 \geq (1-\varepsilon)^2) \Pr(z_j^2 \geq (1-\varepsilon)^2) \Pr((c_1 - c_2)^2 = \delta^2)$$

$$\Pr(\left|y_i z_j\right| \leq \frac{c}{\sqrt{d-1}}) \Pr(\left|(c_1 - c_2)y_i\right| \leq \delta\frac{c}{\sqrt{d-1}}) \Pr(\left|(c_1 - c_2)z_j\right| \leq \delta\frac{c}{\sqrt{d-1}})$$

$$\geq (1 - n^{-\Omega(\log n)})^5 = 1 - n^{-\Omega(\log n)}.$$

so, we only need to have

$$2+2\frac{c}{\sqrt{d-1}} \leq L^2(d,n) \leq 2(1-\varepsilon)^2 + \delta^2 - 2\frac{c}{\sqrt{d-1}} - 4\delta\frac{c}{\sqrt{d-1}}.$$

There exists L(d, n) if we have

$$\begin{split} 2 + 4 \left(1 + \frac{(\log n)^2}{d^{\frac{1}{4}}}\right) \frac{\sqrt{\log n \ln n}}{\sqrt{d - 1}} &\leq 2 \left(1 - \frac{\log n \ln n}{d}\right)^2 + \frac{(\log n)^4}{\sqrt{d}} \\ &\Leftarrow 2 + 4 \left(1 + \frac{(\log n)^2}{d^{\frac{1}{4}}}\right) \frac{\sqrt{\log n \ln n}}{\sqrt{d - 1}} &\leq 2 + 2 \left(\frac{\log n \ln n}{d}\right)^2 - 4 \frac{\log n \ln n}{d} + \frac{(\log n)^4}{\sqrt{d}} \\ &\iff 4 \frac{1 + \frac{(\log n)^2}{d^{\frac{1}{4}}}}{\sqrt{d - 1}} &\leq 2 \frac{\log n \ln^2 n}{d^2} - 4 \frac{\ln n}{d} + \frac{(\log n)^3}{\sqrt{d}} \\ &\iff 4 \frac{1 + \frac{(\log d^3)^2}{d^{\frac{1}{4}}}}{\sqrt{d - 1}} + 4 \frac{\ln d^3}{d} &\leq 2 \frac{\log d^3 \ln^2 d^3}{d^2} + \frac{(\log d^3)^3}{\sqrt{d}}. \end{split}$$

It's true for large enough d. Therefore, there exists L(d,n) and we can just let $L(d,n) = \sqrt{2 + 2\sqrt{\frac{\log n \ln n}{d-1}}}$.

Problem 2.5

Answer: $1.\Pr(x \ge 3) \le \frac{E[x]}{3} = \frac{\int_0^4 x f(x) dx}{3} = \frac{\int_0^4 \frac{x}{4} dx}{3} = \frac{2}{3}$

 $2.\Pr(|x| \ge a) = \Pr(x^2 \ge a^2) \le \frac{E[x^2]}{a^2}$. For a = 3, we have

$$\Pr(|x| \ge 3) \le \frac{E[x^2]}{9} = \frac{\int_0^4 \frac{x^2}{4} \, \mathrm{d}x}{9} = \frac{16}{27}$$

3. For a = 3,

$$\Pr(|x| \ge 3) = \Pr(x^r \ge 3^r) \le \frac{E[x^r]}{3^r} = \frac{\int_0^4 \frac{x^r}{4} \, \mathrm{dx}}{3^r} = \frac{4^r}{3^r(r+1)}.$$

Problem 2.12

Answer: We know that $V(d) = \frac{2\pi^{\frac{d}{2}}}{d\Gamma(\frac{d}{2})}$, so we have

$$\frac{V(d)}{V(d-1)} = \frac{2\pi^{\frac{d}{2}}}{d\Gamma\left(\frac{d}{2}\right)} \frac{(d-1)\Gamma\left(\frac{d-1}{2}\right)}{2\pi^{\frac{d-1}{2}}}.$$

If d = 2k and $k \ge 3$, then

$$\begin{split} \Gamma\bigg(\frac{d}{2}\bigg) &= \Gamma(k) = (k-1)! \\ \Gamma\bigg(\frac{d-1}{2}\bigg) &= \Gamma\bigg(k-\frac{1}{2}\bigg) = \frac{(2k-3)!!}{2^{k-1}}\sqrt{\pi} \\ \frac{V(d)}{V(d-1)} &= \pi\frac{(2k-1)!!}{2^k k!} = \frac{(2k-1)...(5\cdot 3\cdot 1\cdot \pi)}{(2k)...(6\cdot 4\cdot 2)} < 1. \end{split}$$

If d = 2k + 1 and $k \ge 3$, then

$$\begin{split} \Gamma\bigg(\frac{d}{2}\bigg) &= \Gamma\bigg(k+\frac{1}{2}\bigg) = \frac{(2k-1)!!}{2^k}\sqrt{\pi} \\ \Gamma\bigg(\frac{d-1}{2}\bigg) &= \Gamma(k) = (k-1)! \\ \frac{V(d)}{V(d-1)} &= \frac{2^{k+1}(k)!}{(2k+1)!!} = \frac{(2k)...(6\cdot 4\cdot 2\cdot 2)}{(2k+1)...(7\cdot 5\cdot 3)} < 1. \end{split}$$

Therefore, the volume of a sphere decreases as the dimension increases above 5. Since

$$V(2) = \pi < V(3) = \frac{4}{3}\pi < V(4) = \frac{\pi^2}{2} < V(5) = \frac{8}{15}\pi^2.$$

Therefore, the volume of a d-dimensional unit ball take on its maximum when d=5.

Problem 2.19

Answer:

1. We integral x_1 . For $x_1=t$, then volume of the remaining ball is $V(d-1)\big(1-t^2\big)^{\frac{d-1}{2}}$.

Therefore, we have
$$V(d) = \int_{-1}^{1} (1-t^2)^{\frac{d-1}{2}} V(d-1) dt$$
.
2. $V(1) = 2$, $V(2) = 2 \int_{-1}^{1} \sqrt{1-t^2} dt = 2 \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \cos(t) d(\sin t) = \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} (\cos(2t) + 1) dt = \frac{\sin(2t)}{2} + t|_{-\frac{\pi}{2}}^{\frac{\pi}{2}} = \pi \text{ and } V(3) = \int_{-1}^{1} (1-t^2) \pi dt = \pi \left(t - \frac{t^3}{3}\right)|_{-1}^{1} = \frac{4}{3}\pi$.

Problem 2.22

Answer: We assume the height is h, then the face of the hypercube is a ball of dimension d – 1 and radius $\sqrt{1-\left(\frac{h}{2}\right)^2}$, so the volume of the cylinder is $h\left(1-\left(\frac{h}{2}\right)^2\right)^{\frac{d-1}{2}}V(d-1)$.

$$\begin{split} f'(h) &= \left(1 - \frac{h^2}{4}\right)^{\frac{d-1}{2}} + h \frac{d-1}{2} \left(-\frac{h}{2}\right) \left(1 - \frac{h^2}{4}\right)^{\frac{d-3}{2}} \\ &= \left(1 - \frac{h^2}{4}\right)^{\frac{d-1}{2}} - h^2 \frac{d-1}{4} \left(1 - \frac{h^2}{4}\right)^{\frac{d-3}{2}} = \left(1 - h^2\right)^{\frac{d-3}{2}} \left(1 - \frac{dh^2}{4}\right). \end{split}$$

Since $h \le 1$, we have $f'(h) \ge 0$ for $h \le \frac{2}{\sqrt{d}}$ and $f'(h) \le 0$ for $h \ge \frac{2}{\sqrt{d}}$. Therefore, the maximum volume is achieved when $h = \frac{2}{\sqrt{d}}$ and the volume is $h^d = \left(\frac{2}{\sqrt{d}}\right)^d$.

Problem 2.43

Answer: 1.We assume there are n samples, then $\|\mu - m_s\|_{\infty} = \max(|\mu_1 - m_{s1}|, |\mu_2 - m_{s2}|)$ $\begin{array}{l} m_{s,2}\big|,...,|\mu_d-m_{sd}|\big). \text{ Union bound gives us } \Pr(\|\mu-m_s\|_{\infty}\geq\varepsilon)\leq d\Pr(|\mu_i-m_{si}|\geq\varepsilon). \text{ We assume that } m_{si}=\frac{x_1+x_2+...+x_n}{n}, \text{ then } \Pr(|\mu_i-m_{si}|\geq\varepsilon)=\Pr\left(\sum_{i=1}^n(x_i-\mu_i)\geq n\varepsilon\right). \text{ Let } y_i=x_i-\mu_i, \text{ then } \Pr\left(\sum_{i=1}^n(x_i-\mu_i)\geq n\varepsilon\right)=\Pr\left(\sum_{i=1}^ny_i\geq n\varepsilon\right). \end{array}$

 y_i has normal distribution with mean 0 and variance 1, so we have $|E[y_i^{2r+1}]| = 0, |E[y_i^{2r}]| =$ $E(\left|y_i^{2r}\right|) = \frac{(2r)!}{2^r r!} \le (2r)!$. And we also know that $n\varepsilon \in \left[0, \sqrt{2}n\right]$, thus

$$\begin{split} \Pr\!\left(\sum_{i=1}^n y_i \geq n\varepsilon\right) \leq 3e^{-\frac{n^2\varepsilon^2}{12n}} &= 3e^{-n\frac{\varepsilon^2}{12}}.\\ \Pr(\|\mu - m_s\|_\infty \geq \varepsilon) \leq 3de^{-n\frac{\varepsilon^2}{12}} \end{split}$$

Take $n>\frac{1200\ln d}{\varepsilon^2}$ and we can see that $\Pr(\|\mu-m_s\|_\infty\geq\varepsilon)<3e^{-100}<0.01,$ thus $\Pr(\|\mu-m_s\|_\infty\geq\varepsilon)>99\%.$

 $\begin{aligned} &2.\text{Let }z_i=\mu_i-m_{si} \text{ and }Z\coloneqq \sum_{i=1}^d z_i^2. \text{ Then, since }z_i=\frac{1}{n}\sum_{i=1}^n y_i \text{ and }y_i\sim N(0,1), \text{ we know that }\\ &z_i\sim N\left(0,\frac{1}{n}\right). \text{ Let }t_i=\sqrt{n}z_i \text{ thus }t_i\sim N(0,1). \text{ Then, we have }Z=\sum_{i=1}^d z_i^2=\frac{1}{n}\sum_{i=1}^d t_i^2. \text{ so }nZ\sim \chi_d^2.\\ &E\left[\chi_d^2\right]=d(\text{var}[t_i]+E[t_i])=d \text{ therefore }E[Z]=\frac{d}{n}. \end{aligned}$

$$\Pr(\|\mu - m_s\|_2 \geq \varepsilon) = \Pr\!\left(\sqrt{\sum_{i=1}^d \left(\mu_i - m_{si}\right)^2} \geq \varepsilon\right) = \Pr\!\left(\sqrt{\sum_{i=1}^d z_i^2} \geq \varepsilon\right) = \Pr\!\left(Z \geq \varepsilon^2\right) \leq \frac{E[Z]}{\varepsilon^2} = \frac{d}{n\varepsilon^2}.$$

When $n>100\frac{d}{\varepsilon^2}$, we get $\Pr(\|\mu-m_s\|_2\geq\varepsilon)<0.01, \Pr(\|\mu-m_s\|_2\leq\varepsilon)>99\%$. So we can have $n=O\left(\frac{d}{\varepsilon^2}\right)$.