

Lesson 20

José M. Corcuera.
University of Barcelona.

Definition

A zero-coupon bond with *default* is a contract with maturity time T and payoff

$$X = \mathbf{1}_{\{\tau > T\}},$$

where τ is the (random) *default time*.

Then the arbitrage price of this bond at time t will be given by

$$D(t, T) := \mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_t^T r_s ds} \mathbf{1}_{\{\tau > T\}} \middle| \mathcal{G}_t \right),$$

where $(\mathcal{G}_t)_{0 \leq t \leq T}$ represents the flow of the *total* information we have in the market. So far we use the letter \mathcal{F} to indicate the available information in the market but now we will consider different kind of information and we will use $(\mathcal{F}_t)_{0 \leq t \leq T}$ for *the default free market information* that includes the short rate process. Obviously $D(t, T)$ will depend on the model for τ . There are different approaches.

Merton's approach to pricing defaultable bonds

In the Merton approach there is a firm's value process $(V_t)_{t \geq 0}$, that evolves as a geometric Brownian motion under the risk-neutral martingale measure \mathbb{P}^* , specifically

$$dV_t = V_t((r - \kappa)dt + \sigma_V dW_t^*).$$

where r and κ are the interest and dividend rate respectively. That is

$$V_t = V_0 \exp \left(\left(r - \kappa - \frac{1}{2} \sigma_V^2 \right) t + \sigma_V W_t^* \right).$$

Therefore, we have that

$$V_T = V_t \exp \left(\left(r - \kappa - \frac{1}{2} \sigma_V^2 \right) (T - t) + \sigma_V (W_T^* - W_t^*) \right),$$

Then if at the maturity T the total value V_T of the firm's assets is less than the total notional value L of the firm's debt, the firm defaults. Otherwise, the firm does not default, and the debt, that is the L zero-coupon bonds, is paid. Consequently, the default time τ is defined as

$$\tau := T \mathbf{1}_{\{V_T < L\}} + \infty \mathbf{1}_{\{V_T \geq L\}}$$

and

$$\mathbf{1}_{\{\tau > T\}} = \mathbf{1}_{\{V_T \geq L\}}.$$

It correspond to a digital option in a Black-Scholes model, so

$$D(t, T) = \Phi(d_-) e^{-r(T-t)},$$

where

$$d_- = \frac{\log(\frac{V_t}{L}) + (r - \kappa - \frac{1}{2}\sigma_V^2)(T - t)}{\sigma_V \sqrt{(T - t)}}.$$

In his approach Merton also considers the possibility of a *partial recovery* in case of default in such a way that set of bond holders receive V_T if $V_T < L$. So, the price of this bond with this recovery rule will be

$$D(t, T) + \frac{1}{L} \mathbb{E}_{\mathbb{P}^*} \left(\frac{V_T \mathbf{1}_{\{V_T < L\}}}{e^{r(T-t)}} \middle| \mathcal{G}_t \right) = e^{-r(T-t)} - \frac{1}{L} P_t,$$

where P_t is the price of a put option with strike L sold by the bond holders to the owners of the firm. To see that note that in that case the payoff for a holder of a zero-coupon bond is

$$\begin{aligned} \mathbf{1}_{\{\tau > T\}} + \frac{V_T}{L} \mathbf{1}_{\{\tau \leq T\}} &= \mathbf{1}_{\{V_T \geq L\}} + \frac{V_T}{L} \mathbf{1}_{\{V_T < L\}} \\ &= 1 - \mathbf{1}_{\{V_T < L\}} + \frac{V_T}{L} \mathbf{1}_{\{V_T < L\}} \\ &= 1 - \left(1 - \frac{V_T}{L} \right) \mathbf{1}_{\{V_T < L\}} \\ &= 1 - \left(1 - \frac{V_T}{L} \right)_+ \\ &= 1 - \frac{1}{L} (L - V_T)_+. \end{aligned}$$

Hazard process approach

In this approach the total information available for the investors is given by a filtration $(\mathcal{G}_t)_{0 \leq t \leq T}$ and $\mathcal{G}_t = \sigma(\mathbf{1}_{\{\tau \leq s\}}, 0 \leq s \leq t) \vee \mathcal{F}_t$, where \mathcal{F}_t is the information of the *default free market* that includes the short rate process. On the other hand τ is not necessarily an (\mathcal{F}_t) -stopping time, but it is assumed that there exists a non negative (\mathcal{F}_t) -adapted process such that

$$\mathbb{P}^*(\tau > t | \mathcal{F}_t) = e^{-\int_0^t \lambda_s ds} > 0 \text{ for all } t \geq 0.$$

where \mathbb{P}^* is the risk neutral probability. Under this framework we have the following proposition.

Proposition

$$D(t, T) = \mathbf{1}_{\{t < \tau\}} \mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_t^T (\lambda_s + r_s) ds} \middle| \mathcal{F}_t \right).$$

The proof of this proposition is based in the following lemma.

Lemma

For any random variable X , integrable w.r.t. \mathbb{P}^* ,

$$\mathbf{1}_{\{t < \tau\}} \mathbb{E}_{\mathbb{P}^*} (X | \mathcal{G}_t) = \mathbf{1}_{\{t < \tau\}} \frac{\mathbb{E}_{\mathbb{P}^*} (X \mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)}{\mathbb{E}_{\mathbb{P}^*} (\mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)}.$$

Proof.

We have to prove that

$$\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{t < \tau\}} X \mathbf{1}_A) = \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t < \tau\}} \frac{\mathbb{E}_{\mathbb{P}^*}(X \mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)}{\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)} \mathbf{1}_A \right), \quad (1)$$

for all $A \in \mathcal{G}_t$. Then it is enough to consider sets of the form $A = \{\tau \leq s\} \cap B$, $B \in \mathcal{F}_t$ and $0 \leq s \leq t$ or $A \in \mathcal{F}_t$. If $A \in \mathcal{F}_t$

$$\begin{aligned} & \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t < \tau\}} \frac{\mathbb{E}_{\mathbb{P}^*}(X \mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)}{\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)} \mathbf{1}_A \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left(\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t) \frac{\mathbb{E}_{\mathbb{P}^*}(X \mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)}{\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t)} \mathbf{1}_A \right) \\ &= \mathbb{E}_{\mathbb{P}^*}(\mathbb{E}_{\mathbb{P}^*}(X \mathbf{1}_{\{t < \tau\}} \mathbf{1}_A | \mathcal{F}_t)) = \mathbb{E}_{\mathbb{P}^*}(X \mathbf{1}_{\{t < \tau\}} \mathbf{1}_A). \end{aligned}$$

If $A = \{\tau \leq s\} \cap B$, $B \in \mathcal{F}_t$, $\mathbf{1}_{\{t < \tau\}} \mathbf{1}_A = 0$, so both sides of (1) vanish. □

Proof.

Notice that we also need to prove that $\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{t < \tau\}} | \mathcal{F}_t) > 0$ when $\{t < \tau\}$. But in fact, set $Y_t := \mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{\tau > t\}} | \mathcal{F}_t)$, we have

$$\begin{aligned}\mathbb{P}^*(\{\tau > t\} \cap \{Y_t > 0\}) &= \mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{\tau > t\}} \mathbf{1}_{\{Y_t > 0\}}) \\ &= \mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{Y_t > 0\}} \mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{\tau > t\}} | \mathcal{F}_t)) \\ &= \mathbb{E}_{\mathbb{P}^*}(Y_t \mathbf{1}_{\{Y_t > 0\}}) \\ &= \mathbb{E}_{\mathbb{P}^*}(Y_t),\end{aligned}$$

where the last equality is due to the fact that $Y_t \geq 0$. Finally

$$\begin{aligned}\mathbb{P}^*(\{\tau > t\} \cap \{Y_t > 0\}) &= \mathbb{E}_{\mathbb{P}^*}(Y_t) = \mathbb{E}_{\mathbb{P}^*}(\mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{\tau > t\}} | \mathcal{F}_t)) \\ &= \mathbb{E}_{\mathbb{P}^*}(\mathbf{1}_{\{\tau > t\}}) = \mathbb{P}^*(\{\tau > t\})\end{aligned}$$



Proof.

(of the theorem)

$$\begin{aligned} D(t, T) &= \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{T < \tau\}} e^{-\int_t^T r_s ds} \middle| \mathcal{G}_t \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t < \tau\}} \mathbf{1}_{\{T < \tau\}} e^{-\int_t^T r_s ds} \middle| \mathcal{G}_t \right) \\ &= \mathbf{1}_{\{t < \tau\}} \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{T < \tau\}} e^{-\int_t^T r_s ds} \middle| \mathcal{G}_t \right) \\ &= \mathbf{1}_{\{t < \tau\}} \frac{\mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t < \tau\}} \mathbf{1}_{\{T < \tau\}} e^{-\int_t^T r_s ds} \middle| \mathcal{F}_t \right)}{\mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t < \tau\}} \middle| \mathcal{F}_t \right)} \\ &= \mathbf{1}_{\{t < \tau\}} \frac{\mathbb{E}_{\mathbb{P}^*} \left(\mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{T < \tau\}} \middle| \mathcal{F}_T \right) e^{-\int_t^T r_s ds} \middle| \mathcal{F}_t \right)}{\mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t < \tau\}} \middle| \mathcal{F}_t \right)} \\ &= \mathbf{1}_{\{t < \tau\}} \frac{\mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^T \lambda_s ds} e^{-\int_t^T r_s ds} \middle| \mathcal{F}_t \right)}{e^{-\int_0^t \lambda_s ds}} \end{aligned}$$



Note that we do not need the Key Lemma to calculate the price at time zero, because in such a situation $\mathcal{G}_0 = \mathcal{F}_0 = \{\phi, \Omega\}$, then

$$\begin{aligned} D(0, T) &= \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{T < \tau\}} e^{-\int_0^T r_s ds} \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left(\mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{T < \tau\}} \mid \mathcal{F}_T \right) e^{-\int_0^T r_s ds} \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^T \lambda_s ds} e^{-\int_0^T r_s ds} \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^T (\lambda_s + r_s) ds} \right). \end{aligned}$$

Credit default swaps (CDS)

CDS is a credit derivative that offers protection against default of a bond. Assume that the nominal of the bond is N and the recovery rate is $R < 1$, in such a way that the owner of the bond receives only NR in case of default.

Then the buyer of the CDS, in order to ensure the total nominal N , pays at time T_i

$$sN(T_i - T_{i-1})$$

provided the default time $\tau > T_i$, $i = 1, \dots, n$ and receives

$$N(1 - R)$$

at time τ if $\tau \leq T_n$.

We shall omit the so-called *accrual premium*: $sN(\tau - T_{\lfloor \tau \rfloor})$. Where

$$\lfloor \tau \rfloor := \min\{T_j \in \{T_0, T_1, T_2, \dots\} : T_j \leq \tau < T_{j+1}\}, \quad \tau > T_0 = 0.$$

In this way the discounted price of the CDS at time zero will be

$$\begin{aligned}
 & \mathbb{E}_{\mathbb{P}^*} \left(N(1 - R) \mathbf{1}_{\{\tau \leq T_n\}} e^{-\int_0^\tau r_s ds} - \sum_{i=1}^n sN(T_i - T_{i-1}) \mathbf{1}_{\{\tau > T_i\}} e^{-\int_0^{T_i} r_s ds} \right) \\
 &= N(1 - R) \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{\tau \leq T_n\}} e^{-\int_0^\tau r_s ds} \right) \\
 &\quad - \sum_{i=1}^n sN(T_i - T_{i-1}) \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{\tau > T_i\}} e^{-\int_0^{T_i} r_s ds} \right) \\
 &= N(1 - R) \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{\tau \leq T_n\}} e^{-\int_0^\tau r_s ds} \right) \\
 &\quad - \sum_{i=1}^n sN(T_i - T_{i-1}) \mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^{T_i} (r_s + \lambda_s) ds} \right).
 \end{aligned}$$

We can write, for $t_i = i \frac{T_n}{k}$,

$$\mathbf{1}_{\{\tau \leq T_n\}} e^{-\int_0^\tau r_s ds} = \lim_{k \rightarrow \infty} \sum_{i=1}^k \mathbf{1}_{\{t_{i-1} < \tau \leq t_i\}} e^{-\int_0^{t_{i-1}} r_s ds},$$

and

$$\begin{aligned} & \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{\tau \leq T_n\}} e^{-\int_0^\tau r_s ds} \right) \\ &= \lim_{k \rightarrow \infty} \sum_{i=1}^k \mathbb{E}_{\mathbb{P}^*} \left(\mathbf{1}_{\{t_{i-1} < \tau \leq t_i\}} e^{-\int_0^{t_{i-1}} r_s ds} \right) \\ &= \lim_{k \rightarrow \infty} \sum_{i=1}^k \mathbb{E}_{\mathbb{P}^*} \left(\left(\mathbf{1}_{\{\tau > t_{i-1}\}} - \mathbf{1}_{\{\tau > t_i\}} \right) e^{-\int_0^{t_{i-1}} r_s ds} \right) \\ &= \lim_{k \rightarrow \infty} \sum_{i=1}^k \mathbb{E}_{\mathbb{P}^*} \left(\left(e^{-\int_0^{t_{i-1}} \lambda_s ds} - e^{-\int_0^{t_i} \lambda_s ds} \right) e^{-\int_0^{t_{i-1}} r_s ds} \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left(\int_0^{T_n} e^{-\int_0^s (r_u + \lambda_u) du} \lambda_s ds \right). \end{aligned}$$

Since

$$\begin{aligned} & \lim_{k \rightarrow \infty} \sum_{i=1}^k \left(e^{-\int_0^{t_{i-1}} \lambda_s ds} - e^{-\int_0^{t_i} \lambda_s ds} \right) e^{-\int_0^{t_{i-1}} ds} \\ = & \lim_{k \rightarrow \infty} \sum_{i=1}^k \left(1 - e^{-\int_{t_{i-1}}^{t_i} \lambda_s ds} \right) e^{-\int_0^{t_{i-1}} (\lambda_s + r_s) ds} \\ = & \lim_{k \rightarrow \infty} \sum_{i=1}^k \lambda_{t_{i-1}} (t_i - t_{i-1}) e^{-\int_0^{t_{i-1}} (\lambda_s + r_s) ds} \end{aligned}$$

Therefore the price of the CDS is

$$N(1 - R)\mathbb{E}_{\mathbb{P}^*} \left(\int_0^{T_n} e^{-\int_0^s (r_u + \lambda_u) du} \lambda_s ds \right) - \sum_{i=1}^n sN(T_i - T_{i-1})\mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^{T_i} (r_s + \lambda_s) ds} \right).$$

s is chosen in such a way that the price of this contract is zero:

$$s = \frac{(1 - R)\mathbb{E}_{\mathbb{P}^*} \left(\int_0^{T_n} e^{-\int_0^s (r_u + \lambda_u) du} \lambda_s ds \right)}{\sum_{i=1}^n (T_i - T_{i-1})\mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^{T_i} (r_s + \lambda_s) ds} \right)}.$$

If the intensity λ is deterministic

$$\begin{aligned}\mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^{T_i} (r_s + \lambda_s) ds} \right) &= e^{-\int_0^{T_i} \lambda_s ds} \mathbb{E}_{\mathbb{P}^*} \left(e^{-\int_0^{T_i} r_s ds} \right) \\ &= \mathbb{P}^* (\tau > T_i) P(0, T_i)\end{aligned}$$

and

$$\mathbb{E}_{\mathbb{P}^*} \left(\int_0^{T_n} e^{-\int_0^s (r_u + \lambda_u) du} \lambda_s ds \right) = \int_0^{T_n} e^{-\int_0^s \lambda_u du} \lambda_s P(0, s) ds$$

then

$$s = \frac{(1 - R) \int_0^{T_n} e^{-\int_0^s \lambda_u du} \lambda_s P(0, s) ds}{\sum_{i=1}^n (T_i - T_{i-1}) \mathbb{P}^*(\tau > T_i) P(0, T_i)}.$$

and

$$\begin{aligned} s_j &= \frac{(1 - R) \int_0^{T_j} e^{-\int_0^s \lambda_u du} \lambda_s P(0, s) ds}{\sum_{i=1}^j (T_i - T_{i-1}) \mathbb{P}^*(\tau > T_i) P(0, T_i)} \\ &\approx \frac{(1 - R) \sum_{i=1}^j (\mathbb{P}^*(\tau > T_{i-1}) - \mathbb{P}^*(\tau > T_i)) P(0, T_i)}{\sum_{i=1}^j (T_i - T_{i-1}) \mathbb{P}^*(\tau > T_i) P(0, T_i)}. \end{aligned}$$

These values s_j , $j = 1, \dots, n$ are observed in the market and can be used to deduce $\mathbb{P}^*(\tau > T_i)$, $i = 1, \dots, n$ and obtain a piecewise constant hazard rate.

Market models. A market model for Swaptions

Consider a *payer swaption* with maturity $T < T_0$, *tenor structure* T_1, T_2, \dots, T_n , and *swap rate* R . Its payoff is

$$(S(T) - Z(T))_+$$

con

$$S(T) = P(T, T_0) - P(T, T_n)$$

that is the value of the floating payments minus the last fixed payment and

$$Z(T) = R\delta \sum_{i=1}^n P(T, T_i)$$

the value of payments with fixed rate minus the last payment.

We can take $Z(t)$ as *numeraire* and the price will be

$$Z(t)\mathbb{E}_{\mathbb{P}^{(Z)}}\left(\frac{(S(T)-Z(T))_+}{Z(T)}\middle|\mathcal{F}_t\right)=Z(t)\mathbb{E}_{\mathbb{P}^{(Z)}}\left(\left(\frac{S(T)}{Z(T)}-1\right)_+\middle|\mathcal{F}_t\right).$$

Then, if we assume that under \mathbb{P} , or \mathbb{P}^* we have an evolution

$$d\left(\frac{S(t)}{Z(t)}\right)=\frac{S(t)}{Z(t)}(\mu dt+\sigma dW_t),$$

with σ constant, it turns out that, under $\mathbb{P}^{(Z)}$

$$d\left(\frac{S(t)}{Z(t)}\right)=\frac{S(t)}{Z(t)}\sigma dW_t^Z,$$

so

$$\frac{S(T)}{Z(T)}=\frac{S(t)}{Z(t)}\exp\left\{\int_t^T\sigma dW_s^Z-\frac{1}{2}\int_t^T\sigma^2 ds\right\},$$

and we obtain the Black-Scholes formula of a call with strike 1 and $r = 0$, multiplied by $Z(t)$:

$$Z(t) \left(\frac{S(t)}{Z(t)} \Phi(d_+) - \Phi(d_-) \right) = S(t) \Phi(d_+) - Z(t) \Phi(d_-),$$

with

$$\Phi(d_{\pm}) = \frac{\log \frac{S(t)}{Z(t)} \pm \frac{1}{2} \sigma^2 (T - t)}{\sigma \sqrt{T - t}}.$$

This formula is known as *the Margrabe formula*. Remember that the *swap rate* is given by

$$R(t) = \frac{P(t, T_0) - P(t, T_n)}{\delta \sum_{i=1}^n P(t, T_i)},$$

so

$$\frac{S(t)}{Z(t)} = \frac{P(t, T_0) - P(t, T_n)}{R \delta \sum_{i=1}^n P(t, T_i)} = \frac{R(t)}{R}.$$

Therefore the volatility σ corresponds to the volatility of $R(t)$. The previous formula can be written more explicitly as

$$\text{Swaption}_t = (P(t, T_0) - P(t, T_n)) \Phi(d_+) - \left(R\delta \sum_{i=1}^n P(t, T_i) \right) \Phi(d_-),$$

where

$$\Phi(d_{\pm}) = \frac{\log(P(t, T_0) - P(t, T_n)) - \log(R\delta \sum_{i=1}^n P(t, T_i)) \pm \frac{1}{2}\sigma^2(T-t)}{\sigma\sqrt{T-t}}.$$

Definition

Let X be a payoff at T . A forward contract on X with delivering time T is a contract established at $t < T$ that specifies a forward price $f(t; T)$ **that will be paid at T** for receiving X . The price $f(t; T)$ is fixed in such a way that the contract price at t is zero.

Theorem

$$\begin{aligned} f(t; T) &= \frac{1}{P(t, T)} \mathbb{E}_{\mathbb{P}^*} \left(X \exp \left\{ - \int_t^T r_s ds \right\} \middle| \mathcal{F}_t \right) \\ &= \mathbb{E}_{\mathbb{P}^T}(X | \mathcal{F}_t). \end{aligned}$$

Proof.

The total payoff of this contract at T is

$$X - f(t; T),$$

so if \mathbb{P}^* is the risk neutral probability the price of this contract will be zero at t if and only if

$$\mathbb{E}_{\mathbb{P}^*} \left((X - f(t; T)) \exp \left\{ - \int_0^T r_s ds \right\} \middle| \mathcal{F}_t \right) = 0,$$

therefore

$$\mathbb{E}_{\mathbb{P}^*} \left(X \exp \left\{ - \int_0^T r_s ds \right\} \middle| \mathcal{F}_t \right) = f(t; T) \mathbb{E}_{\mathbb{P}^*} \left(\exp \left\{ - \int_0^T r_s ds \right\} \middle| \mathcal{F}_t \right).$$



Proof.

If we use \mathbb{P}^T

$$\mathbb{E}_{\mathbb{P}^T} \left(\frac{X - f(t; T)}{P(T, T)} \middle| \mathcal{F}_t \right) = 0,$$

Therefore

$$f(t; T) = \mathbb{E}_{\mathbb{P}^T} (X | \mathcal{F}_t).$$



Definition

Let X a payoff at T . A contract of futures on X and delivering time T is a financial asset with the following properties

- There exist a *future price* $F(t; T)$ on X at each time t .
- At T the owner of the contract pays $F(T; T)$ and receives X .
- For any arbitrary interval $(s, t]$ the owner receives $F(t; T) - F(s; T)$.
- At each time the price of the contract is zero.

Theorem

Let \mathbb{P}^ be a risk neutral probability measure such that the discounted value of a self-financing portfolio with one future contract is a \mathbb{P}^* -martingale, then*

$$F(t; T) = \mathbb{E}_{\mathbb{P}^*}(X|\mathcal{F}_t).$$

Proof.

Let V_t be the value of a self-financing portfolio formed by a bank account and one future contract

$$\begin{aligned} V_t &= \phi_t^0 e^{\int_0^t r_s ds} + \phi_t^1 \times 0 \\ &= \phi_t^0 e^{\int_0^t r_s ds} \end{aligned}$$

with $\phi_t^1 = 1$, but

$$\begin{aligned} dV_t &= r_t \phi_t^0 e^{\int_0^t r_s ds} dt + \phi_t^1 dF(t; T) \\ &= r_t V_t dt + \phi_t^1 dF(t; T), \end{aligned}$$

so

$$d\tilde{V}_t = e^{-\int_0^t r_s ds} \phi_t^1 dF(t; T),$$

with $F(T; T) = X$ and since \tilde{V} is a martingale (a Brownian one), with respect to \mathbb{P}^* , it turns out that $F(\cdot; T)$ is a martingale and therefore

$$F(t; T) = \mathbb{E}_{\mathbb{P}^*}(F(T; T) | \mathcal{F}_t) = \mathbb{E}_{\mathbb{P}^*}(X | \mathcal{F}_t).$$

Corollary

Future prices and forward prices coincide if and only if interest rates are deterministic.

Proof.

$f(t, T) = F(t; T)$ if and only if

$$\mathbb{P}^* = \mathbb{P}^T.$$

We know that

$$\frac{d\mathbb{P}^T}{d\mathbb{P}^*} = \frac{e^{-\int_0^T r_s ds}}{P(0, T)},$$

then $\frac{d\mathbb{P}^T}{d\mathbb{P}^*} = 1$ if and only if

$$P(0, T) = e^{-\int_0^T r_s ds}.$$

