

Lecture 8: Uncertainty Quantification (I)

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- Frequentist approach: confidence interval for $g(\hat{\theta})$.
- Bayesian approach: conditional distribution for $g(\hat{\theta})$.
- Also known as uncertainty quantification
- Source of uncertainty: estimation error of $\hat{\theta}$
- One popular way: find the variance of $g(\hat{\theta})$.

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- Question: how do we find σ ?
- We will focus on the case $g(\theta) = f_{\theta}(x)$.

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- In linear regression: $f(x) = x^T \hat{\beta}$
- Nonlinear problem: need Monte Carlo.

Bootstrap method

- Draw n_B samples from S randomly with replacement, denote as S^b .
- Repeat this for $b = 1, \dots, B$.
- For each bootstrap dataset S^b , we apply ML procedure to find \hat{f}_{S^b}
- We now have $\hat{f}_{S^1}(x), \dots, \hat{f}_{S^B}(x)$
- They “simulate” the scenario where data are randomly obtained.
- Their variance is approximately the same as the $\hat{f}_S(x)$

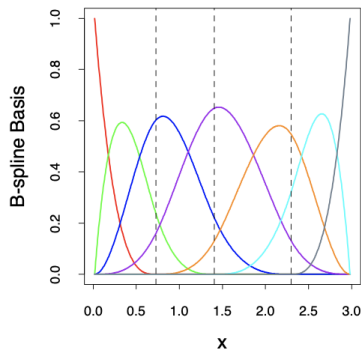
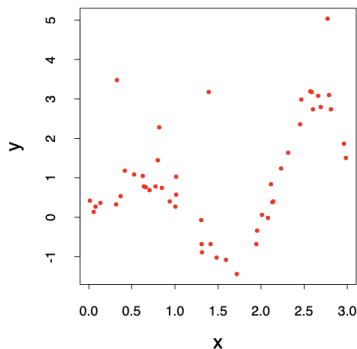
True variance: the case where the only source of randomness is y

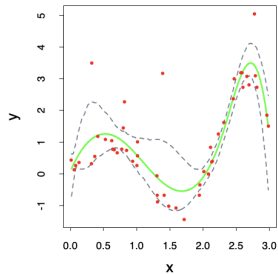
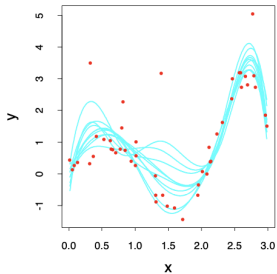
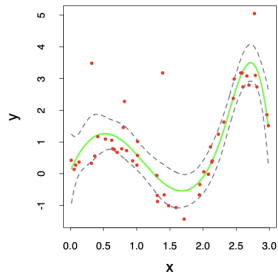
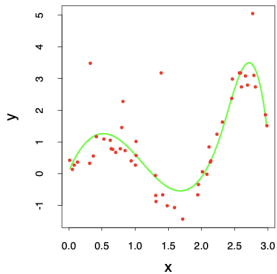
- Recall that $\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} = \beta^* + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \epsilon$
- Given \mathbf{X} , the estimated covariance matrix is $\sigma_\epsilon^2 (\mathbf{X}^T \mathbf{X})^{-1}$
- At x_0 , the variance of $\hat{f}(x)$ is $\sigma_\epsilon^2 x_0^T (\mathbf{X}^T \mathbf{X})^{-1} x_0$

Bootstrap

- $\hat{\beta}_b = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}_b = \beta^* + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \epsilon_b$
- Estimation for x_0 : $x_0^T \hat{\beta}_b$
- The bootstrap covariance comes from randomness in ϵ_b
- The variance is $\sigma_\epsilon^2 (\mathbf{X}^T \mathbf{X})^{-1}$.

- A standard way to deal with functional data
- We want to fit $\mu(x) = \beta_1 h_1(x) + \cdots \beta_7 h_7(x) + \epsilon$
- $h_j(x)$: spline functions
- Data fitting: $y_i = \mu(x_i) + \epsilon_i$
- Linear basis regression $y_i = \beta_1 h_1(x_i) + \cdots + \beta_7 h_7(x_i) + \epsilon_i$





Using Fisher information

In many applications we model the data from a parameteric family

$$z_i \sim p(\theta, z)$$

The overall likelihood is given by

$$p(\theta, S) = \prod_{i=1}^n p(\theta, z_i)$$

We can try to maximize its log $l(\theta, S) = \log(p(\theta, S))$

$$l(\theta, S) = \sum_{i=1}^n l(\theta, z_i) = \sum_{i=1}^n \log p(\theta, z_i)$$

The Maximum Likelihood Estimator (MLE) is obtained as

$$\hat{\theta} = \arg \max l(\theta, S).$$

The score function is defined as

$$\nabla_{\theta} l(\theta, S) = \sum_{i=1}^n \nabla_{\theta} l(\theta, z_i)$$

The Fisher information matrix is its expectation w.r.t z

$$I(\theta) = \mathbb{E}_S[\nabla l(\theta, S) \nabla l(\theta, S)^T] = n \mathbb{E}_z[\nabla l(\theta, z) \nabla l(\theta, z)^T]$$

The MLE estimator converges

$$\hat{\theta} \rightarrow \mathcal{N}(\theta^*, I(\theta^*)^{-1})$$

Using the approximation $\hat{\theta} \approx \theta^*$, we have

$$\hat{\theta} - \theta^* \approx \mathcal{N}(0, I(\hat{\theta})^{-1})$$

If we want to have a confidence interval of θ_j^* , we can use

$$\hat{\theta}_j - q^{(1-\alpha)} \sqrt{(I(\hat{\theta})^{-1})_{j,j}}, \quad \hat{\theta}_j + q^{(1-\alpha)} \sqrt{(I(\hat{\theta})^{-1})_{j,j}},$$

where $q^{1-\alpha}$ is the Gaussian quantile of order $1 - \alpha$.

- Linear model $y_i = \beta_*^\top x_i + \epsilon_i$, $\epsilon_i \sim \mathcal{N}(0, \sigma_\epsilon^2)$.
- For general β , the loglikelihood is given by

$$l(\beta) = -\frac{n}{2} \log(2\pi\sigma_\epsilon^2) - \frac{1}{2\sigma_\epsilon^2} \sum_{i=1}^n (y_i - x_i^\top \beta)^2$$

- The Fisher information matrix is given by

$$I(\beta_*) = \frac{1}{\sigma_\epsilon^2} X^\top X$$

- This agrees with earlier estimates.

Example

Suppose using linear regression, we have $X^T X = nI_d$, $\hat{\sigma}^2 = 1$, $\hat{\beta} = [1, 0]$ What will be the confidence interval for $f([1, 1])$?

Answer: $[1 - 1.96 \times \frac{\sqrt{2}}{\sqrt{n}}, 1 + 1.96 \times \frac{\sqrt{2}}{\sqrt{n}}]$