Phase transitions in Generalized Linear Models

Cargèse summer school

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Generalized Linear Models

Definition

Statistical model

▶ One observes for $1 \le \mu \le m$

$$Y_{\mu} \sim P_{\text{out}} \left(\cdot \mid \langle \mathbf{\Phi}_{\mu}, \mathbf{X}^* \rangle \right)$$

- ▶ $\mathbf{X}^* \in \mathbb{R}^n$: signal vector of dimension n.
- $lackbox{\Phi}_1,\ldots,lackbox{\Phi}_m\in\mathbb{R}^n$: measurement vectors.
- $ightharpoonup P_{\text{out}}$: transition kernel.

Goal: recover X^* from Y (and Φ).

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- ▶ When is it information-theoretically possible?
- ► When is it computationally tractable?

Examples

Some interesting particular cases

Linear model:

$$\mathbf{Y} = \mathbf{\Phi} \mathbf{X}^*$$

► Phase retrieval:

$$\mathbf{Y}=\left|\mathbf{\Phi}\mathbf{X}^{*}\right|$$

► 1-bit CS ("Planted" perceptron):

$$\mathbf{Y} = \mathsf{sign} ig(\mathbf{\Phi} \mathbf{X}^* ig)$$

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$$\mathbf{Y} = \mathbf{\Phi} \mathbf{X}^* + \mathsf{Noise}$$

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► 1-bit CS ("Planted" perceptron):

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$$\mathbf{Y} = \left| \mathbf{\Phi} \mathbf{X}^* \right| + \mathsf{Noise}$$

► Logistic model:

$$Y_{\mu} = \begin{cases} +1 & \text{with probability} \quad \frac{1}{1 + \exp(-\lambda \langle \mathbf{\Phi}_{\mu}, \mathbf{X}^{*} \rangle)} \\ \\ -1 & \text{with probability} \quad \frac{1}{1 + \exp(\lambda \langle \mathbf{\Phi}_{\mu}, \mathbf{X}^{*} \rangle)} \end{cases}$$

$$\mathbf{Y} \sim P_{\mathrm{out}} \Big(\cdot \ \Big| \mathbf{\Phi} \mathbf{X}^* \Big)$$

- $\qquad \text{Asymptotic regime: } n \to \infty, \quad m/n \to \alpha > 0.$
- $\mathbf{X}^* = (X_1^*, \dots, X_n^*) \overset{\text{i.i.d.}}{\sim} P_0, \qquad \mathbb{E}_{P_0} X^2 = \rho.$

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- ▶ P_{out} has to be "regularized" by some (small) Gaussian noise: $\forall x \in \mathbb{R}$, " $P_{\mathrm{out}}(\cdot \mid x) = \widetilde{P}_{\mathrm{out}}(\cdot \mid x) + \mathcal{N}(0, \sigma^2)$ ", where $\sigma > 0$.
- ▶ If P_{out} takes values in \mathbb{N} , no need for regularization ($\sigma = 0$).

Assumptions

$$\mathbf{Y} \sim P_{\mathrm{out}} \Big(\cdot \ \Big| \ \mathbf{\Phi} \mathbf{X}^* \Big)$$

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- $\mathbf{X}^* = (X_1^*, \dots, X_n^*) \stackrel{\text{i.i.d.}}{\sim} P_0, \qquad \mathbb{E}_{P_0} X^2 = \rho.$
- $\bullet \ \, \left(\Phi_{i,j}\right) \text{ are independent, } \begin{cases} \mathbb{E}\Phi_{i,j} = 0 \\ \mathbb{E}\Phi_{i,j}^2 = 1/n \\ \sup_{i,j} \mathbb{E}|\Phi_{i,j}|^3 \text{ remains bounded.} \end{cases}$
- $ightharpoonup \mathbb{E}[|Y_{\mu}|^{2+\epsilon}]$ remains bounded, for some $\epsilon > 0$.
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The statistician knows the model, i.e. P_0 and P_{out} .

Information-theoretic study

The mutual information

Posterior distribution $P(\mathbf{X}^*|\mathbf{Y}, \mathbf{\Phi})$:

$$P(\mathbf{x}|\mathbf{Y}, \mathbf{\Phi}) = \frac{1}{\mathcal{Z}_n} P_0^{\otimes n}(\mathbf{x}) \prod_{\mu=1}^m P_{\text{out}}(Y_\mu | \langle \mathbf{\Phi}_{\mu}, \mathbf{x} \rangle)$$

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$$f_n = -\frac{1}{n} \mathbb{E} \log \mathcal{Z}_n = -\frac{1}{n} \mathbb{E} \left[\log \int_{\mathbf{x} \in \mathbb{R}^n} dP_0^{\otimes n}(\mathbf{x}) \prod_{\mu=1}^m P_{\text{out}}(Y_\mu | \langle \mathbf{\Phi}_{\mu}, \mathbf{x} \rangle) \right]$$

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Equivalently, we are going to study the mutual information:

$$\frac{1}{n}I(\mathbf{X}^*;\mathbf{Y}|\mathbf{\Phi}) = f_n + \mathsf{Constant} + o_n(1) \,.$$

"Replica Symmetric" formula

Theorem

$$\frac{1}{n}I(\mathbf{X}^*;\mathbf{Y}|\mathbf{\Phi}) \xrightarrow[n\to\infty]{} \inf_{q\in[0,\rho]} \sup_{r\geq0} \left\{ I_{P_0}(r) + \alpha \mathcal{I}_{P_{\text{out}}}(q) - \frac{r}{2}(\rho-q) \right\}$$

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Example: Linear regression

- $Y = \Phi X^* + \sigma Z.$
- ► "Tanaka formula", proved by Barbier et al., 2016 and Reeves and Pfister, 2016.

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Example: 'planted' perceptron.

- $\mathbf{Y} = \operatorname{sign}(\mathbf{\Phi}\mathbf{X}^*), \text{ where } X_1^*, \dots, X_n^* \overset{\text{i.i.d.}}{\sim} \mathcal{U}(+1, -1).$
- $\blacktriangleright \ S_n = \left\{ \mathbf{x} \left| \forall \mu, \, \mathsf{sign}(\mathbf{\Phi}_{\mu} \mathbf{x}) = Y_{\mu} \right. \right\}$
- $\blacktriangleright \frac{1}{n} I(\mathbf{X}^*; \mathbf{Y} | \mathbf{\Phi}) = \log 2 \frac{1}{n} \mathbb{E} \left[\log \# S_n \right].$
- ► Formula obtained by Gardner and Derrida, 1989.

"Replica Symmetric" formula

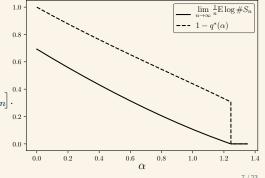
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Two scalar inference channels

Explanation of the formula

$$\begin{split} \text{Recall: } \mathbf{Y} &\sim P_{\text{out}}\big(\cdot \mid \mathbf{\Phi} \mathbf{X}^*\big) \\ & \frac{1}{n} I(\mathbf{X}^*; \mathbf{Y} | \mathbf{\Phi}) \xrightarrow[n \to \infty]{} \inf_{q \in [0, \rho]} \sup_{r > 0} \Big\{ I_{P_0}(r) + \alpha \, \mathcal{I}_{P_{\text{out}}}(q) - \frac{r}{2} (\rho - q) \Big\} \end{split}$$

Additive Gaussian channel

$$I_{P_0}(r) = I(X^*; \sqrt{r}X^* + Z)$$

where $X^* \sim P_0$ and $Z \sim \mathcal{N}(0, 1)$.

Two scalar inference channels

Explanation of the formula

Recall:
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Non-linear Gaussian retrieval

$$\mathcal{I}_{P_{\text{out}}}(q) = I(W^*; Y^{(q)}|V)$$

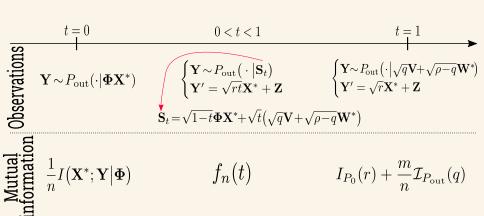
where $V, W^* \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0,1)$ and

$$Y^{(q)} \sim P_{\text{out}} \left(\cdot \mid \sqrt{q}V + \sqrt{\rho - q}W^* \right)$$

Proof technique

The interpolation method

In the spirit of Talagrand's interpolation scheme for the perceptron.



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$$\begin{array}{c|c} & t = 0 & 0 < t < 1 & t = 1 \\ \hline \text{Substituting of States} & \mathbf{Y} \sim P_{\mathrm{out}}(\cdot | \mathbf{\Phi} \mathbf{X}^*) & \begin{cases} \mathbf{Y} \sim P_{\mathrm{out}}(\cdot | \mathbf{S}_t) \\ \mathbf{Y}' = \sqrt{rt} \mathbf{X}^* + \mathbf{Z} \end{cases} & \begin{cases} \mathbf{Y} \sim P_{\mathrm{out}}(\cdot | \sqrt{q} \mathbf{V} + \sqrt{\rho - q} \mathbf{W}^*) \\ \mathbf{Y}' = \sqrt{r} \mathbf{X}^* + \mathbf{Z} \end{cases} \\ \mathbf{S}_t = \sqrt{1 - t} \mathbf{\Phi} \mathbf{X}^* + \sqrt{t} \left(\sqrt{q} \mathbf{V} + \sqrt{\rho - q} \mathbf{W}^* \right) \\ \hline \\ \mathbf{I}_{P_0}(r) + \frac{m}{n} \mathcal{I}_{P_{\mathrm{out}}}(q) \end{aligned}$$

Goal: show that
$$f'_n(t) \simeq \frac{r}{2}(\rho - q)$$
.

Interpolation method

Derivative of the interpolating mutual information

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► We have to show that the overlap

$$\frac{1}{n} \sum_{i=1}^{n} x_i^{(t)} X_i^*$$

concentrates around some value, and then choose $\it q$ to be equal to this value.

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Derivative of the interpolating mutual information

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- ▶ In the case of Bayes-optimal inference problems this is true under mild assumptions: Montanari, 2008, Korada and Macris, 2010.
- More details about the techniques in Jean Barbier's talk on Saturday.

Limit of the overlap

Minimal Mean Squared Error

Theorem

For almost all $\alpha>0$, the infimum of the "Mutual Information formula" admits a unique minimizer $q_*(\alpha)$ and

$$\left|\frac{1}{n}\sum_{i=1}^n x_i X_i^*\right| \xrightarrow[n\to\infty]{} q_*(\alpha), \quad \text{in probability,}$$

where $\mathbf{x} \sim P(\mathbf{X}^* = \cdot \mid \mathbf{\Phi}, \mathbf{Y})$ independently of everything else.

One deduces:

$$\mathrm{MMSE}_n(\alpha) := \frac{1}{n^2} \mathbb{E} \left\| \mathbf{X}^* \mathbf{X}^{*\intercal} - \mathbb{E} \left[\mathbf{X}^* \mathbf{X}^{*\intercal} | \mathbf{\Phi}, \mathbf{Y} \right] \right\|^2 \xrightarrow[n \to \infty]{} \rho^2 - q_*(\alpha)^2$$

Algorithmic analysis

Generalized Approximate Message Passing (GAMP)

- Precursors in physics: Mezard, 1989, Kabashima, 2008.
- ▶ Generalization of AMP (Donoho et al., 2009) introduced by Rangan, 2011. Iterative algorithm: produces estimates $\hat{\mathbf{x}}^0, \dots, \hat{\mathbf{x}}^t$.

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- ► Its performance can be rigorously tracked:

State evolution, Javanmard and Montanari, 2013

$$\frac{1}{n^2} \mathbb{E} \left\| \mathbf{X}^* \mathbf{X}^{*\intercal} - \widehat{\mathbf{x}}^t \widehat{\mathbf{x}}^{t\intercal} \right\|^2 \xrightarrow[n \to \infty]{} \rho^2 - (q^t)^2$$

where q^t is given by the recursion $(q^0 = 0)$:

$$\begin{cases} q^{t+1} = \rho - 2I'_{P_0}(r^t) \\ r^t = -2\alpha \mathcal{I}'_{P_{\text{out}}}(q^t) \end{cases}$$

▶ GAMP converges to a stationary point $(q^{\rm alg}, r^{\rm alg})$ of the MI formula and if $q^{\rm alg} = q_*(\alpha)$, then GAMP achieves the MMSE!

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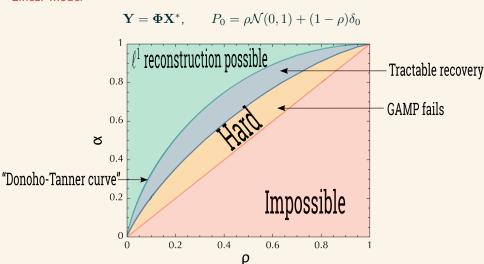
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- ▶ GAMP converges to a stationary point $(q^{\rm alg}, r^{\rm alg})$ of the MI formula and if $q^{\rm alg} = q_*(\alpha)$, then GAMP achieves the MMSE!
- ▶ Main belief: GAMP is optimal among all polynomial-time algorithms.

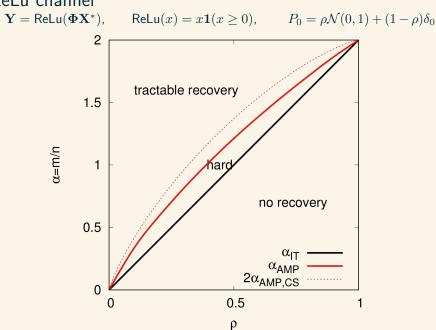
Phase diagrams: warm-up

Linear model

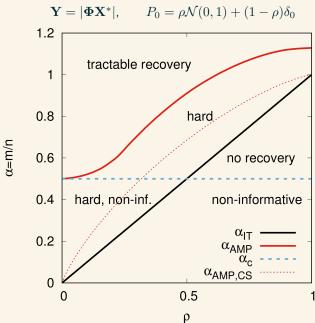


Phase diagram from Krzakala et al., 2012

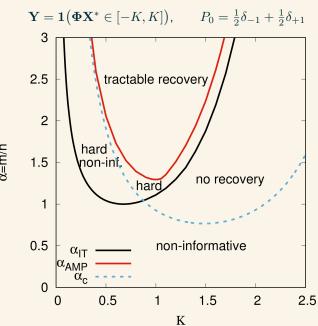
ReLu channel



Absolute value channel



Symmetric door channel



A learning problem

A different point of view

- ▶ The points $\{(\Phi_1, Y_1), \dots, (\Phi_m, Y_m)\}$ can be seen as data generated by some relation $\mathbf{Y} \sim P_{\mathrm{out}}(\cdot | \Phi \mathbf{X}^*)$.
- ▶ **Question**: How difficult is it to learn this relation?

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- ▶ Question: How difficult is it to learn this relation?
- ► What is the optimal generalization error

$$\mathcal{E}_n^{\text{gen}} = \min_{\widehat{\theta}} \mathbb{E} \Big[\big(Y^{(\text{new})} - \widehat{\theta} (\mathbf{\Phi}^{(\text{new})}; \mathbf{Y}, \mathbf{\Phi}) \big)^2 \Big]$$

where $Y^{(\mathrm{new})} \sim P_{\mathrm{out}} \big(\cdot \mid \langle \mathbf{\Phi}^{(\mathrm{new})}, \mathbf{X}^* \rangle \big)$ is a new sample.

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Theorem

$$\mathcal{E}_n^{\mathrm{gen}} \xrightarrow[n \to \infty]{} E(q_*(\alpha))$$

where

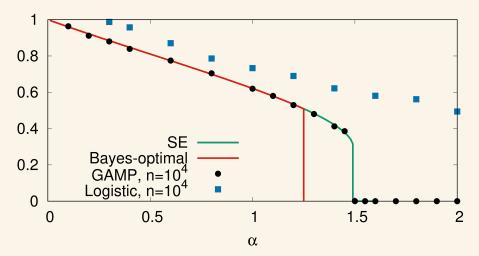
$$E(q) = \mathbb{E}\left[\left(Y^{(q)} - \mathbb{E}[Y^{(q)}|V]\right)^2\right]$$

Recall the second scalar channel: $Y^{(q)} \sim P_{\mathrm{out}}(\cdot \mid \sqrt{q}V + \sqrt{\rho - q}W)$, $V, W \overset{\mathrm{i.i.d.}}{\sim} \mathcal{N}(0,1)$.

Classification: the perceptron

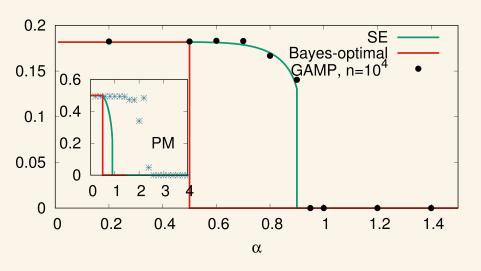
$$\mathbf{Y} = \text{sign}(\mathbf{\Phi}\mathbf{X}^*), \qquad P_0 = \frac{1}{2}\delta_{-1} + \frac{1}{2}\delta_{+1}$$

Computed by Györgyi, 1990 and also Seung et al., 1992:

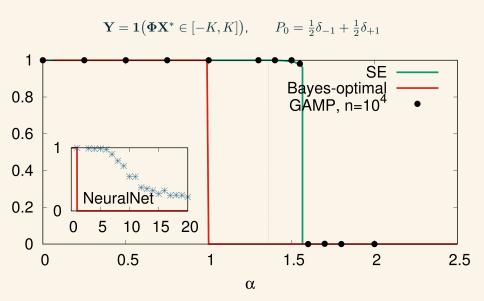


Regression: phase retrieval

$$\mathbf{Y} = |\mathbf{\Phi} \mathbf{X}^*|, \qquad P_0 = \frac{1}{2}\delta_{-1} + \frac{1}{2}\delta_{+1}$$



Classification: the symmetric door



Thank you for your attention.

Any questions?

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