

Reconstructing Within-Farm Transmission of a PRRSV-1 Outbreak Using Structured and Unstructured Phylogenetic Models

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1 **Abstract**

2 Porcine Reproductive and Respiratory Syndrome Virus (PRRSV) causes significant economic
3 losses and animal welfare concerns globally. While its rapid evolution makes it ideal for
4 phylodynamic inference, the accuracy of these models at the fine-scale, non-random contact
5 environments of individual farms remains unvalidated. By integrating genetic and
6 epidemiological data, phylodynamics offer a distinct advantage over traditional methods,
7 enabling the estimation of transmission parameters during unsampled periods, the
8 quantification of unreported cases, and the reconstruction of the outbreak's origin, independent
9 of clinical observation. This study investigates whether Bayesian birth-death models can
10 accurately reconstruct within-farm transmission dynamics and distinguish spread across
11 spatially distinct units. Building on the original study by Clilverd and colleagues (2023), which
12 used longitudinal incidence monitoring and phylogenetic contact tracing to characterize the
13 epidemic, we applied an unstructured Birth-Death Skyline (BDSKY) model to infer temporal
14 trajectories and a structured Birth-Death with Migration (BDMM) model to resolve pen- and
15 room-level spread.

16
17 The BDSKY analysis successfully recovered the established three-phase epidemic trajectory
18 (outbreak, dormancy, resurgence) by the original study and independently validated the timing of
19 the resurgence event. At finer resolution, the BDMM analysis indicated that within-pen
20 transmission was efficient ($Re=2.61$, 95% Highest Posterior Density (HPD): 0.52–5.21), while
21 spread between pens was significantly constrained within rooms ($Re=0.46$, 95% HPD: 0.09–0.94)
22 and across rooms ($Re=0.06$, 95% HPD: 0.02–0.12), reflecting effective physical
23 compartmentalization. Sensitivity analyses revealed that high-density Open Reading Frame 5
24 (ORF5) sampling enabled detection of an intense superspreading event, whereas whole-genome
25 sequencing (WGS) provided improved precision for molecular clock calibration.

26
27 These results demonstrate that integrating genetic data into birth–death models enables precise
28 reconstruction of epidemic trajectories and robust estimation of compartment-specific transmission
29 rates. This insight supports precision animal health management, complementing traditional
30 epidemiological approaches and potentially reducing the economic impact of PRRSV while
31 improving herd welfare.

32 **Introduction**

33 Porcine Reproductive and Respiratory Syndrome Virus (PRRSV) is globally recognized as one of
34 the most economically significant pathogens in the swine industry, causing severe respiratory and
35 reproductive disorders that greatly compromise the health and welfare of affected pigs (Chandra et
36 al., 2025; Clilverd et al., 2023; Sha et al., 2025). The virus is classified into two distinct species
37 based on their geographic origins: PRRSV-1, first identified in Europe (Wensvoort et al., 1991)
38 and PRRSV-2, first identified in North America (Collins et al., 1992). However, both species share
39 a conserved ~15 kb positive-sense RNA genome with at least 10 open reading frames (ORFs), with
40 ORF1a/b encoding non-structural proteins and ORF2–7 encoding structural proteins (Clilverd et
41 al., 2023). ORF5 sequencing, encoding the highly variable glycoprotein 5, is still commonly used
42 for viral classification (Murtaugh et al., 2010; Weng et al., 2025) and offers a key advantage in its
43 ease of amplification and Sanger sequencing from routine clinical samples (Zhang et al., 2017),
44 despite representing only ~4% of the genome and thus limiting capture of genome-wide diversity
45 (Murtaugh et al., 2010). In contrast, whole-genome sequencing (WGS) provides the advantage of
46 detecting variation across both structural and non-structural regions for comprehensive
47 epidemiological insight (Li et al., 2022).

48 It is known that PRRSV infection dynamics are highly age-dependent in both genetic variants:
49 neonatally infected pigs act as long-term reservoirs capable of shedding virus for up to 250 days
50 (Butler et al., 2014; Pileri & Mateu, 2016; Rowland et al., 2003; Wills et al., 2003; You et al.,
51 2022), whereas pigs infected post-birth typically clear infection within 42 days (Charpin et al.,
52 2012). Therefore, while vertical transmission often initiates circulation, prolonged shedding by
53 neonatally infected pigs bridges production stages and sustains endemicity through subsequent
54 horizontal spread (Pileri & Mateu, 2016). Furthermore, both species exhibit substantial genetic
55 variability driven by high mutation and recombination rates typical of positive-sense RNA viruses
56 (Franzo et al., 2021; Kappes & Faaberg, 2015; Parisio et al., 2024; Wu et al., 2024; Zhou et al.,
57 2024). Given this exceptional genetic diversity and rate heterogeneity, universal substitution
58 models are often unsuitable for characterizing PRRSV evolution (Balka et al., 2018; Gong et al.,
59 2024). While this fast evolution complicates disease control in PRRSV (Caserta et al., 2023), it
60 renders it particularly suitable for phylodynamic modelling (Duchene et al., 2020; Featherstone et
61 al., 2023). Unlike phylogenetics, which primarily focuses on evolutionary relationships and viral
62 classification (Frias-De-Diego et al., 2021), phylodynamics combines genetic and epidemiological
63 data to quantify pathogen transmission dynamics (Stadler et al., 2024). Central to these dynamics
64 is the effective reproductive number (R_e), defined as the expected number of secondary infections
65 caused by a single infected individual at a specific point in time (Stadler et al., 2024). By
66 estimating the R_e through time, it becomes possible to assess the epidemic trajectory relative to the
67 epidemiological threshold of 1, distinguishing between periods of growth ($R_e > 1$) and decline ($R_e < 1$)
68 (Andrade & Duggan, 2022).

69
70 In the field of animal health, phylodynamics is particularly valuable for addressing epidemiological
71 questions that are difficult to answer with traditional surveillance data alone, which relies primarily
72 on observational methods like case reporting and contact tracing (Guinat et al., 2021). In particular,
73 phylodynamics can estimate cross-species transmission rates between wild and domestic animals
74 and retrospectively assess the impact of control strategies, such as vaccination, by modeling
75 changes in transmission dynamics over time (Guinat et al., 2021)

76
77 Extensive research has successfully characterized transmission dynamics within large-scale
78 farming networks (Sequeira et al., 2025). For PRRSV-2, phylodynamic models have been utilized
79 to infer farm-to-farm transmission pathways and quantify farm-level transmissibility
80 (Pamornchainavakul et al., 2023), as well as to trace the historical migration of viral lineages across
81 different production companies (Alkhamis et al., 2016). Similarly, for PRRSV-1, migration of

83 lineages has been reconstructed across Europe to identify determinants of spread (Franzo et al.,
84 2022), and the role of pig movements in shaping viral diversity in Central Eastern Europe has been
85 investigated (Balka et al., 2018). Furthermore, Franzo et al. (2021) showed how phylodynamics
86 can track transmission within integrated pig companies, where farrowing, nursery, and finishing
87 units are managed by a single organization.

88 Complementing these studies, research at the within-farm resolution has largely relied on
89 phylogenetic and evolutionary approaches. For instance, phylogenetic analysis has been used to
90 investigate specific introduction and circulation events within a single farm (Li et al., 2022), while
91 the evolutionary dynamics and selection pressures facilitating persistence in vaccinated breeding
92 herds have also been characterized (Clilverd et al., 2024).

93 Despite these advances, quantitative phylodynamic estimates for transmission parameters at the
94 within-farm level remain limited (Meester et al., 2025). Applying these models to a single herd
95 presents distinct challenges, primarily because farm populations do not mix randomly. Instead, they
96 are highly structured environments where contact is constrained by physical barriers, such as
97 separation into specific pens and rooms. Consequently, the suitability of phylodynamic models to
98 resolve transmission dynamics within such a compartmentalized system, especially where viral
99 genetic diversity is constrained and sparse sampling can obscure rapid transmission chains, remains
100 to be fully validated (Alkhamis et al., 2016; Guinat et al., 2023; Meester et al., 2025).

101 To address these knowledge gaps, this study leverages the dataset from Clilverd and colleagues
102 (2023) to evaluate whether phylodynamic models can reliably infer within-farm transmission
103 dynamics. Unlike the original study, which characterized the epidemic utilizing longitudinal
104 incidence monitoring and phylogenetic contact tracing, this analysis validates the utility of
105 phylodynamic models by investigating their ability to reconstruct the established epidemiological
106 trajectory by Clilverd and colleagues (2023), differentiate within-pen transmission from
107 transmission between pens sharing the same room versus those in separate rooms, detect
108 differences in infectious periods between production stages (farrowing vs. nursery), and assess the
109 impact of genomic resolution (WGS vs. ORF5) on inference reliability. Following validation of the
110 temporal signal, absence of recombination and excluding viral introductions from sources outside
111 the farm during the study period, unstructured Birth-Death Skyline (BDSKY) and structured Birth-
112 Death with Migration (BDMM) models were applied to quantify these long-term and fine-scale
113 transmission parameters, respectively.

117 **Methods**

118 **Study Design and Sequence Acquisition**

119 The virological and epidemiological data used for this phylodynamic analysis were obtained from
120 a longitudinal study by Clilverd et al. (2023), which monitored a PRRSV-1 outbreak on a Spanish
121 300-sow farrow-to-wean farm. The original study followed three distinct batches of piglets from
122 birth to nine weeks of age at 1.5 months (Batch 1), 8 months (Batch 2), and 12 months (Batch 3)
123 after the clinical detection of the outbreak, which was characterized by the sudden appearance of
124 abortions, stillbirths, and weak-born animals. For all batches, umbilical cords were collected at
125 birth, and blood samples were drawn from individually ear-tagged piglets at 2, 4, 6, and 9 weeks
126 of age. Viral RNA was extracted from collected samples and screened for PRRSV-1 using RT-
127 qPCR. For samples with sufficient viral load ($C_t \leq 32.0$), viral sequencing was performed. This
128 generated ORF5 sequences via Sanger sequencing for animals in Batch 1 and Batch 3. Additionally,
129 WGS were generated from viral isolates using an Illumina MiSeq platform. No sequences could be
130 obtained from Batch 2. This phase was characterized by an absence of clinical symptoms and only
131 sporadic PCR-positive results (8/74 animals) with high C_t values (>31.5), which prevented the
132 recovery of sufficient genetic material for sequencing. All consensus sequences corresponding to
133 the WGS and ORF5, regions were deposited by Clilverd and colleagues (2023) in GenBank under
134 accession numbers OP688189–OP688223 and OP688224–OP688357, respectively.

135
136 In the current study, two primary datasets were compiled: a WGS dataset and an ORF5 dataset
137 (figure 1). Depending on the analysis performed, these were utilized either as restricted 'cohort'
138 datasets or comprehensive 'extended' datasets. The cohort WGS and ORF5 datasets consisted
139 exclusively of the consensus sequences derived directly from the outbreak study described by
140 Clilverd and colleagues (2023). These cohort-specific datasets were used for the phylodynamic
141 analyses. The extended WGS dataset combined the farm consensus sequences with ten regional
142 strains, five Spanish commercial PRRSV vaccines, and all GenBank WGS (taxid 1965066), while
143 the extended ORF5 dataset included farm sequences and all GenBank ORF5 entries. These datasets
144 were used to screen for recombination, and to rule out viral introductions from an outside source
145 during the study period.

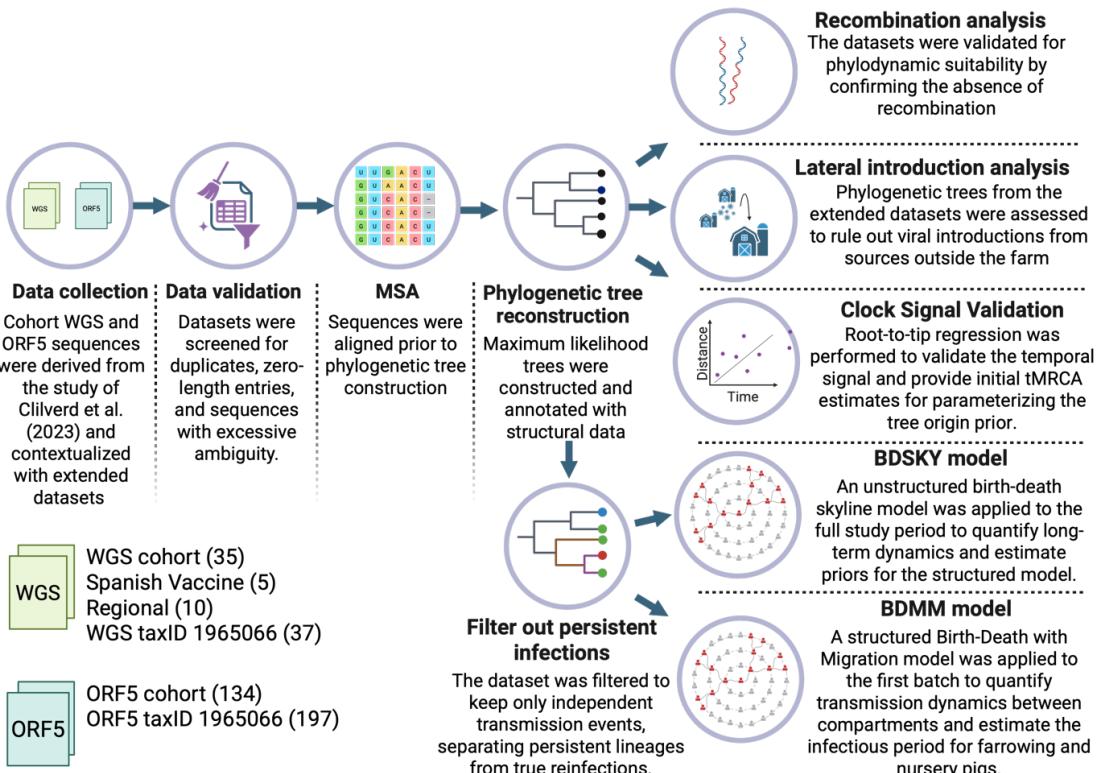


Figure 1. Schematic overview of the study workflow. The diagram shows computational pipeline used to analyze the PRRSV-1 outbreak. The process moves from Data Collection of Whole Genome Sequences (WGS) and ORF5 sequences (with sequence counts shown in brackets) to Data Validation and Multiple Sequence Alignment (MSA). Following Phylogenetic Tree Reconstruction, the dataset underwent validation steps, including recombination analysis, lateral introduction checks, and clock signal validation. Finally, the data were subjected to Birth-Death Skyline (BDSKY) and Birth-Death with Migration (BDMM) models to quantify long-term and short-term structured transmission dynamics, respectively. tMRCA=Time to Most Recent Common Ancestor. Created with BioRender.com.

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Data processing and validation

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All datasets were screened for duplicate or zero-length entries, invalid characters, and International Union of Pure and Applied Chemistry (IUPAC) ambiguity codes (Johnson, 2010). Sequences displaying excessive ambiguity, indicative of poor sequencing quality, were excluded from the datasets. However, sporadic ambiguity codes were retained, as these may represent biological signals of intra-host viral diversity rather than sequencing errors. Additionally, a pairwise identity analysis was performed on the cohort WGS and ORF5 datasets to identify and cluster 100% identical sequences.

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To enable further downstream analyses, multiple sequence alignments (MSAs) were constructed using the DECIPHER v3.4.0 package (Wright, 2016) on the extended WGS and ORF5 datasets for recombination and phylogenetic analyses.

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Classification of Independent Transmission Events

To ensure the phylodynamic model was informed by independent transmission events, a clade-based analysis was performed to differentiate between sequences representing prolonged viral persistence (within-host evolution) and those from new reinfection events in pigs sampled multiple times. This step was essential as standard birth-death models assume that every branching event in the phylogenetic tree represents a transmission event between hosts (Stadler et al., 2024). Including these sequences, which resemble intra-host evolutionary variants results in over-sampling and false

166 clusters that bias growth rate estimates upwards (Dearlove et al., 2017) and can lead to poor model
167 convergence (Alkhamis et al., 2022)

168
169 First, a patristic distance matrix was calculated from the phylogenetic tree using the ape v5.8-1
170 package (Paradis & Schliep, 2019), and hierarchical clustering (Kaufman & Rousseeuw, 1990) was
171 performed on this matrix. The resulting dendrogram was initially partitioned into five distinct
172 clades, aligning with the classification described by Clilverd and colleagues (2023), who identified
173 four co-circulating clades in Batch 1 and a single clade in Batch 3. For each pig with multiple
174 samples, the clade assignment of its virus was tracked chronologically. If a pig's subsequent sample
175 was assigned to the same clade as a previous sample, it was classified as a persistent infection; if it
176 was assigned to a different clade, it was classified as a reinfection. Since hierarchical clustering
177 requires a pre-determined number of clusters, a robustness test was performed by varying the
178 defined clade sizes to ensure that the classification of infection status for individuals remained
179 consistent across different thresholds. Based on this classification, we retained only the first
180 samples of persistent lineages, defined as initial infections, as well as samples from reinfections for
181 the main phylodynamic analyses

182 183 **Phylogenetic Reconstruction and Topology**

184 To assess viral diversity circulating on the farm, phylogenetic trees were constructed for the cohort
185 WGS and ORF5 datasets using IQ-TREE v3 (Wong et al., 2025). The optimal nucleotide
186 substitution model was inferred using the ModelFinder module (Kalyaanamoorthy et al., 2017),
187 and nodal support was assessed via Ultrafast Bootstrap Approximation with 1,000 replicates (Ecker
188 et al., 2024). These trees were rooted using phytools v2.5 (Revell, 2024) and subsequently
189 annotated with spatial metadata. The structural data used for this annotation was inferred directly
190 from the supplementary video provided by Clilverd and colleagues (2023). However, this spatial
191 annotation was restricted to Batch 1 sequences, as structural data for Batch 3 was not available. For
192 the Batch 1 sequences, the farm structure was defined as a farrowing unit comprising a single
193 compartment with 18 pens, and a nursery unit consisting of four distinct compartments, each
194 containing four pens.

195 196 **Assessment of Recombination and Lateral Introduction**

197 To rule out viral introductions from sources outside the farm during the study period, maximum
198 likelihood trees were generated for both the extended WGS and ORF5 datasets. These trees were
199 evaluated to determine whether farm-derived sequences clustered into a single, strongly supported
200 monophyletic clade, consistent with a single introduction event.. Following this, recombination
201 was assessed in both datasets, as it violates the bifurcation assumptions of standard phylogenetic
202 models (Driebe et al., 2015). This was done by visual inspection using Neighbor-Net in SplitsTree
203 (Huson & Bryant, 2024) and statistical validation via the Pairwise Homoplasy Index (Phi) test
204 (Bruen et al., 2006).

205 206 **Molecular Clock Validation and Temporal Signal**

207 To assess the temporal signal over the 12 month study period, root-to-tip regression analyses were
208 performed using Clockor2 (Featherstone et al., 2024). The analyses were also used to provide an
209 initial estimate of the time to the Most Recent Common Ancestor (tMRCA) required to
210 parameterize the tree origin prior for the subsequent BEAST analyses. In addition, the cohort ORF5
211 data were partitioned to assess whether Batch 1 and Batch 3 exhibited valid temporal signals under
212 separate local clocks.

213 214 **Phylodynamic Analysis**

215 Phylodynamic analysis was performed using Bayesian inference in BEAST v2.7.7 (Bouckaert et
216 al., 2019), allowing for the joint estimation of the phylogenetic tree, evolutionary parameters, and

epidemiological rates by integrating prior knowledge with the likelihood of the sequence data. Within this Bayesian framework, transmission dynamics were reconstructed using birth-death phylodynamic models, which simulate the epidemic as a forward-in-time stochastic process driven by rates of viral transmission (λ), recovery (μ), and sampling (ψ) (Stadler et al., 2024). To facilitate biological interpretation, the bdmm-prime v2.6.3 package (Kühnert et al., 2016; Scire et al., 2022; Vaughan & Stadler, 2025) was used, which re-parameterizes the raw stochastic rates (λ , μ and ψ) into derived epidemiological quantities: The Re ($Re = \lambda/\delta$), the sampling proportion ($p = \psi/\delta$) and the rate of becoming uninfected ($\delta = \mu + r\psi$) which can be converted to the average infectious period by taking its reciprocal ($1/\delta$) (Stadler et al., 2024).

Since the resulting joint posterior distribution is analytically intractable (Stadler et al., 2024), Markov chain Monte Carlo (MCMC) sampling was employed to numerically approximate the posterior estimates for all parameters. MCMC chains were run until convergence was achieved; unless otherwise stated, sufficient mixing and the quality of state-space exploration were confirmed using Tracer v1.7.2 (Rambaut et al., 2018), ensuring that all continuous parameters reached an Effective Sample Size (ESS) > 200 . The final log files were processed and summarized using the coda package v0.19.4.1 (Plummer et al., 2006) in R, with burn-in periods configured individually for each run to optimize posterior sampling.

Crucially, the Sampled Ancestors package (Gavryushkina et al., 2014) was used, with the removal probability upon sampling (r) set to 0 as animals remained in the herd after sampling. Unlike standard phylogenetic models, which typically treat all samples as terminal tips in the phylogenetic tree (Stadler et al., 2024), this approach allows sampled individuals to be direct ancestors of subsequent cases. This configuration is essential to capture within-farm dynamics, where high contact rates within small subpopulations create a high probability that sampled individuals are the direct ancestors of subsequently sampled pen-mates.

Long-term Temporal Dynamics (BDSKY)

To reconstruct the epidemic trajectory and estimate the Re over the 12-month study period, while evaluating the impact of genomic resolution (WGS vs. ORF5) on parameter estimates, an unstructured phylodynamic analysis was conducted using the BDSKY model (Stadler et al., 2013). Unlike constant-rate birth–death models, BDSKY allows epidemiological parameters to vary in a piecewise-constant manner across discrete time intervals, defined as “epochs”. This enables the reconstruction of the full trajectory of the outbreak by estimating the Re across distinct phases. Consequently, it allows for the precise identification of epidemiological shifts that may be missed by standard phylogenetic analyses, such as those previously applied by Clilverd and colleagues (2023).

To ensure robust inference, prior specifications were designed to address the parameter redundancy inherent in birth-death-sampling models, where diverse combinations of transmission, recovery, and sampling rates can produce identical likelihoods (Gavryushkina et al., 2014; Scire et al., 2022). Narrow priors were applied to the rate of becoming uninfected and origin, informed by established PRRSV-1 duration of infectivity (Charpin et al., 2012) and the confirmed outbreak timeline (Clilverd et al., 2023) (Table 1). Priors for the Re , sampling proportion, clock rate, and epoch transition times were kept uninformative to allow these parameters to be estimated directly from the genomic data rather than prior assumptions. By avoiding overly informative priors, this method reduces over-smoothing and preserves the ability to detect genuine changes in epidemiological patterns, including abrupt shifts in transmission dynamics (Parag et al., 2022). To validate these specifications, we performed a ‘sample-from-prior’ analysis (Drummond & Bouckaert, 2015) in BEAST, running the MCMC chain without sequence data to verify that the posterior distributions of inferred parameters were driven by the data rather than artifacts of the joint prior constraints.

Table 1: Prior distributions and rationale for the Bayesian phylodynamic models. Summary of the prior distributions applied to the Birth-Death Skyline (BDSKY) and Birth-Death with Migration (BDMM) models, including reference sources. HPD = Highest Posterior Density, tMRCA = time to Most Recent Common Ancestor

Parameter	Model	Prior Distribution	Rationale	Reference
Effective Reproduction Number (Re)	Both	Lognormal(0, 1.0)	Loose prior centered at 1, reflecting the epidemiological threshold. Its broad range allows genomic data to drive the estimates while preventing biologically implausible rates.	Andrade & Duggan (2022); Boskova et al. (2018)
Rate of Becoming Uninfectious	BDSKY	Lognormal(8.0, 0.1)	Tight prior (mean infectious period of ~45 days, 95% CI ~37–65 days) based on experimental data for nursery-age infected pigs which made up the majority of the study population	Charpin et al. (2012)
	BDMM	Lognormal(8.0, 0.5)	Broader prior (mean infectious period of ~45 days, 95% CI ~20–140 days) to evaluate differences between infections in utero and after birth.	Charpin et al. (2012); Butler et al., (2010; Rowland et al., (2003); R. W. Wills et al., (2003); You et al., (2022)
Sampling Proportion	BDSKY	Beta(1.0, 1.0)	Non-informed (uniform) prior allowing free estimation across the range 0 to 1.	-
	BDMM	Normal(0.33, 0.09)	Confined to the 95% HPD interval derived from the Batch 1 posterior of the BDSKY analysis.	This study
Clock Rate (Mean)	BDSKY	Lognormal(0.001, 1.25)	Broad distribution capturing the range for observed rates for PRRSV-1 centered around 0.001 substitutions per site per year.	Boskova et al. (2018); Li et al. (2022); Parisio et al. (2024); Shin et al. (2022)
	BDMM	Lognormal(0.0136, 0.1)	Confined by the posterior estimates from the BDSKY analysis due to weak local temporal signal in the Batch 1 only.	This study
Clock Rate (Std. Dev.)	Both	Gamma(0.5396, 0.3819)	Designed to complement the mean clock rate prior centered at 0.001, inducing an effective prior that provides a biologically plausible baseline while retaining the flexibility to estimate rate heterogeneity	Boskova et al. (2018)
Tree Origin (tMRCA)	BDSKY	Lognormal(1.20, 0.05)	Parameterized relative to the last Batch 3 sample (mean ~438 days, 95% CI 398–482 days) based on root-to-tip regression and the expected 1-to-8-week subclinical phase.	This study; Pedro Mil-Homens et al. (2024)
	BDMM	Lognormal(0.45, 0.1)	Parameterized relative to the last Batch 1 sample (mean ~164 days, 95% CI 134–199 days) based on root-to-tip regression and the expected 1-to-8-week subclinical phase.	This study; Pedro Mil-Homens et al. (2024)
Change Times (Epochs)	BDSKY	Uniform(0, 1.2)	Uninformative prior allowing free estimation of epoch configurations between the tMRCA and the end of Batch 3.	This study
Migration Rate	BDMM	Exponential(1.0)	Uninformative prior that accommodates a broad range of biologically plausible rates and ensures robustness in the presence of limited signal for population structure	Seidel et al. (2024)
Substitution Model	Both	bModelTest	Co-estimated with default priors for invariable sites, rate heterogeneity, and for base frequencies.	Bouckaert & Drummond (2017)

269 To infer the optimal nucleotide substitution model directly from the data instead of relying on
 270 predefined universal substitution models, the bModelTest package (Bouckaert & Drummond,
 271 2017) was used. Additionally, an uncorrelated relaxed clock model (Drummond et al., 2006) was
 272 assessed to evaluate whether a single mean molecular clock rate could be applied across the entire
 273 study period, or whether substitution rates varied among branches of the phylogeny.
 274

275 Distinct epoch strategies were used for the R_e and sampling proportion. The R_e was modeled using
 276 flexible 3-, 4-, and 5-epoch configurations with uniformly estimated transition times over the full
 277 study period (Figure 2). The 3-epoch configuration was applied to test whether the inferred
 278 dynamics could recover the established three-phase timeline (outbreak, dormancy, and resurgence)
 279 defined by Clilverd and colleagues (2023). Additionally, the 4- and 5-epoch configurations were
 280 used to investigate whether a higher temporal resolution could reveal further, previously
 281 unobserved epidemiological phases.

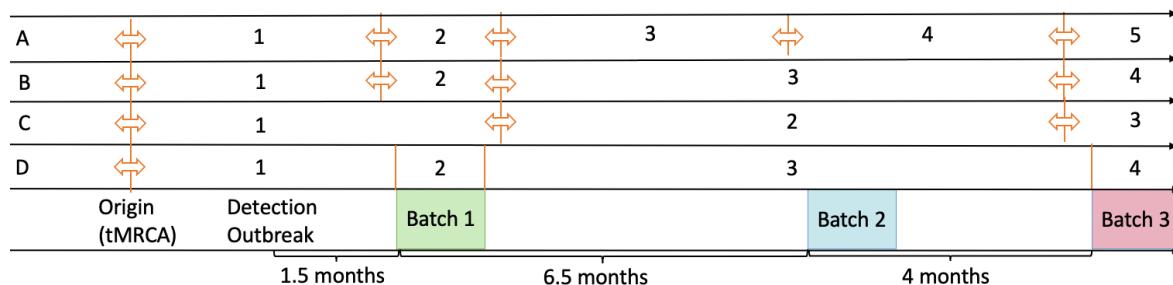


Figure 2: Schematic of temporal discretization strategies for the Birth-Death Skyline (BDSKY) model. The diagram illustrates the flexible production stage configurations (A–C) applied to the effective reproductive number (R_e), which utilized variable 3-, 4-, and 5-epoch intervals to evaluate if the model could reconstruct the outbreak phases. Orange arrows indicate estimated parameters, denoting both the flexible epoch transition times and the time of Most Recent Common Ancestor (tMRCA), reflecting the origin of the outbreak. This is contrasted with the sampling proportion configuration (D), where epoch intervals were fixed to strictly mirror the study schedule and sampling gaps. Here the sampling proportion of epoch 1 and 3 were fixed to 0 while this value was estimated for epoch 2 and 4.

282 The sampling proportion (p) was modeled using a fixed 4-epoch structure to reflect the sampling
 283 schedule of the study. In the general birth-death framework, p is defined as the ratio of the sampling
 284 rate to the total rate of becoming uninfected ($p=\psi/(\mu+r\psi)$) (Stadler et al., 2024). However,
 285 because the removal probability was set to zero ($r=0$) in the Sampled Ancestors model, this
 286 simplifies to the ratio of the sampling rate to the recovery rate ($p=\psi/\mu$). Furthermore, since
 287 longitudinal duplicates were excluded while distinct reinfection events were retained, p is
 288 functionally interpreted here as the proportion of distinct infection events that were successfully
 289 sequenced. Based on this definition, the parameter was constrained according to data availability:
 290 the first and third epochs, which contained no samples, were fixed to 0. Conversely, for the second
 291 and fourth epochs (covering Batch 1 and Batch 3), p was estimated using an uninformative prior,
 292 as the total number of infection events in the complete population could not be reliably inferred
 293 from the available metadata.
 294

295 Fine-Scale Spatial Transmission Dynamics (BDMM)

296 To quantify Batch 1 dynamics, specifically estimating compartment-specific transmission rates and
 297 infectious periods of farrowing and nursery stage pigs, a structured analysis was conducted using
 298 the BDMM model (Kühnert et al., 2016). This framework extends the temporal flexibility of the
 299 BDSKY model to structured populations by allowing parameters to vary not only across time but
 300 also across compartments. Given the insufficient WGS coverage for Batch 1, this structured

301 analysis was performed on the ORF5 dataset by partitioning the outbreak into three epochs that
 302 reflected distinct production stages: the initial sow herd phase (tMRCA to farrowing), the farrowing
 303 phase (first 4 weeks), and the nursery phase (weeks 4–9) (figure 3).

304
 305 The first epoch included a single unobserved (“ghost”) deme with a Re representing the sow
 306 reservoir which aimed to capture transmission via unsampled hosts or reservoirs and avoids bias in
 307 structured phylodynamic models (Ewing & Rodrigo, 2006; Müller et al., 2025). The root location
 308 probability was fixed at 100% for this deme, reflecting that the cohort was unborn at viral
 309 introduction to the farm and that all ancestral lineages originated there before migrating into
 310 observed populations.

311
 312 The spatial configuration of the demes was modeled directly on the pen-level topology described
 313 in the "Phylogenetic Reconstruction and Topology" section. However, because the BDMM model
 314 requires a consistent number of demes across all epochs, the overall structure was constrained to
 315 17 demes, comprising 16 pens and one ghost deme. While this configuration excluded two specific
 316 farrowing pens, the impact on transmission inference was considered negligible due to data
 317 sparsity; the farrowing phase yielded only six sequences, all of which originated from just two pens
 318 already included in the model.

319
 320 In this analysis, epoch change times were fixed but defined under two configurations: a production
 321 stage configuration and a migration configuration. The production stage configuration estimated
 322 the Re and rate of becoming uninfected across three biological stages: reservoir, farrowing, and
 323 nursery. The migration configuration, based on metadata from Clilverd et al. (2023), defined
 324 epochs of per-lineage, forward-in-time migration rates, assuming discrete movements between
 325 compartment at each sampling period. The analysis focused on four key one-day transitions: from
 326 sow to farrowing, from farrowing to nursery at week 4, and within the nursery at weeks 6 and 9.
 327 Movement within the farrowing unit at week 2 was excluded due to insufficient sequence data, as
 328 only six sequences were available for week 0 and none for week 2, precluding reliable estimation
 329 of migration rates for this interval.

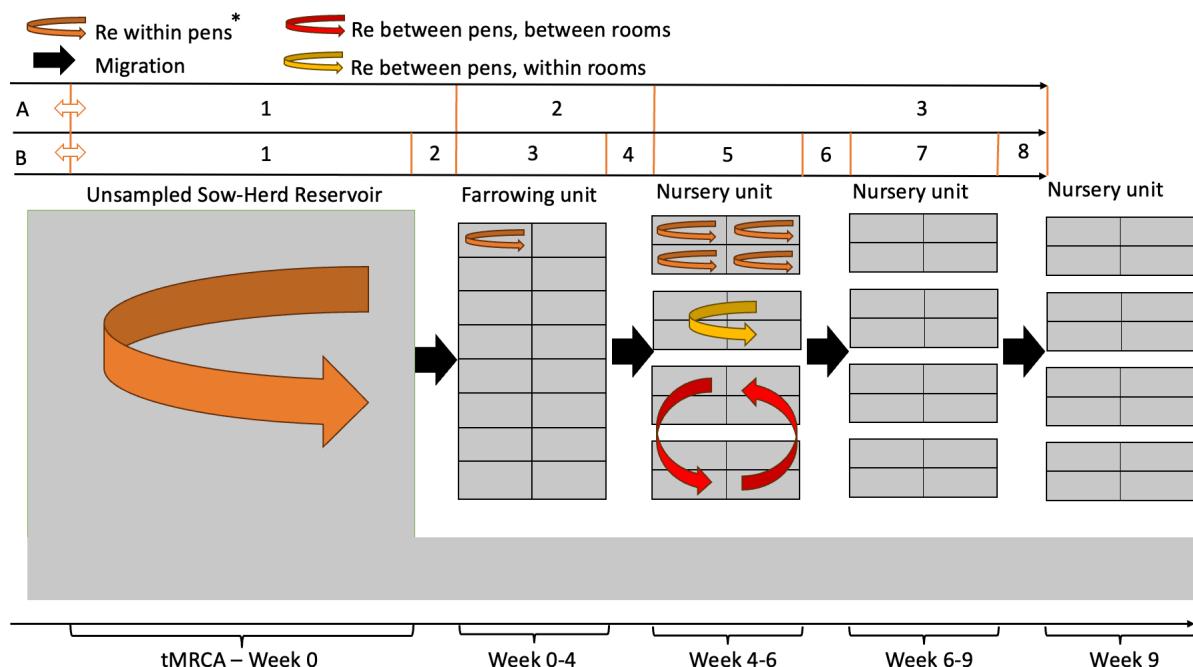


Figure 3: Schematic representation of the Birth-Death with Migration Model (BDMM) configuration. The diagram outlines the spatial structure and migration pathways modeled for Batch 1, including the unsampled sow-herd reservoir, farrowing unit, and nursery unit

compartments. Unlike the flexible unstructured model, the epoch intervals for the BDMM analysis were fixed based on defined production stages (A) and migration events (B).

*: For the Sow-Herd Reservoir, which is modeled as a single "ghost" deme, this parameter represents transmission within the reservoir population as a whole rather than within a specific physical pen.

To characterize transmission dynamics at different spatial levels, a parameter linking strategy was used to estimate three distinct R_e values. For the nursery phase, three linking strategies were applied: first, the R_e was estimated within pens and linked across all pens; second, the R_e between pens within rooms (e.g., pens 1–4, 5–8) were linked; and third, the R_e between pens across rooms were linked. For the farrowing phase, only within-pen R_e values were estimated and linked because horizontal transmission between pens could be excluded: all six sequences from this phase (five from one pen, one from another) were identified as vertical transmission (Clilverd et al., 2023), with no genetic evidence for horizontal transmission between pens. Furthermore, it was assumed that the farrowing and nursery units were isolated from the external farm environment after initial introduction, so the R_e between pens and the ghost deme was set to zero, treating the ghost deme as an ancestral lineage source, not a site of ongoing transmission.

The molecular clock rate for Batch 1 could not be reliably inferred due to insufficient temporal signal (see results: "Molecular Clock Validation and Temporal Signal"). To address this, the clock rate prior was specified using the posterior estimates from the preceding BDSKY analysis fitted to Batch 1 and Batch 3 (table 1). This approach aligns with recent HEV phylodynamic protocols where informative priors from larger datasets are applied to short-term subsets (Meester et al., 2025), and is supported by evidence that evolutionary rates converge and remain broadly stable across longer epidemic timescales (Bryant et al., 2007).

Furthermore, to address the parameter redundancy inherent to birth–death–sampling models, both the sampling proportion prior (centered on the BDSKY posterior estimate for batch 1, with the sow herd phase fixed at 0) and the origin time prior were assigned narrow distributions, ensuring reliable inference of the remaining epidemiological parameters. Finally, the bModelTest package was used using the same configuration as was done in the BDSKY model to investigate whether the inferred substitution model changed by excluding batch 3 from the analysis.

A robustness test was performed to challenge the assumption of discrete one-day animal movements by testing a continuous migration model, where migration was permitted continuously throughout the nursery phase. Finally, to check whether the posterior estimates were driven by genetic data rather than prior assumptions, the BDMM model was run with the sample-from-prior setting for both the continuous and discrete migration configurations.

363 Data and Code Availability

To ensure transparency and facilitate the reproducibility of the analyses presented here, all computational code, processing scripts, and final phylodynamic model configurations have been deposited in a public GitHub repository: **leon1603/prrsv-withinfarm-phylodynamics** (<https://github.com/leon1603/prrsv-withinfarm-phylodynamics>). Within this repository, a dedicated directory named "thesis" contains the digital version of this document alongside all referenced supplementary files. This includes an extensive metadata file (provided as Supplementary File 1 in the "thesis" folder), which details the specific structural-level data, infection classifications, clade assignment and identical sequence clusters for all cohort sequences in addition to pigs that were not sequenced but had metadata available which could be extracted from the supplementary files of Clilverd et al. (2023)

374

Results

375

Study Population and Sequence Validation

376 The final Cohort datasets were assembled, comprising a Cohort WGS dataset of 35 sequences
 377 (partitioned into n=24 for Batch 1 and n=11 for Batch 3) and a Cohort ORF5 dataset of 134
 378 sequences (partitioned into n=86 for Batch 1 and n=48 for Batch 3) (Table 2). Furthermore, the
 379 Extended datasets, consisted of 87 and 331 sequences for WGS and ORF5, respectively.
 380

381 Sequence validation revealed minor and explainable inconsistencies in the cohort data; a 'Y'
 382 ambiguity code (C/T) was identified in two WGS, consistent with previously reported intra-host
 383 viral diversity in this dataset by Clilverd and colleagues (2023) (see Supplementary File 2 for
 384 detailed per-sequence ambiguity information). Furthermore, negligible ambiguity (<0.9%) was
 385 observed in all external reference strains which were retained to avoid compromising phylogenetic
 386 signal (Lozano-Fernandez, 2022).

Table 2: Summary of PRRSV-1 sequence datasets, quality metrics, and batch partitioning.
*Overview of the Cohort and Extended datasets for Whole Genome Sequences (WGS) and ORF5,
 detailing sequence counts, length ranges and ambiguity codes.*

Dataset	Sequence Count	Min Length	Max Length	Total Ambiguities	Ambiguity Characters
Cohort WGS Consensus	35	15098	15098	2	Y
Batch 1	24				
Batch 3	11				
Cohort ORF5 Consensus	134	606	606	0	None
Batch 1	86				
Batch 3	48				
Regional Sequences	10	14443	14910	132	K, M, N, R, S, W, Y
Licensed Vaccines in Spain	5	14758	15120	0	None
WGS (taxID 1965066)	37	14932	15428	225	K, N, R, S, Y
ORF5 (taxID 1965066)	197	606	606	97	K, M, R, S, W, Y

387 Assessment of genetic redundancy revealed that all sequences in the WGS cohort were unique. In
 388 contrast, the ORF5 dataset comprised only 47 unique sequences, corresponding to a redundancy of
 389 64.93% and forming 20 groups of identical sequences (see Supplementary File 3 for detailed
 390 tables).
 391

392 Additionally, an inconsistency in the sampling dates for the cohort ORF5 sequences was also
 393 identified and corrected. The recorded metadata for Batch 1 listed identical sampling dates for week
 394 6 and week 9 at 13/07/2017, except for a single week 9 sequence dated 25/07/2017. This was
 395 determined to be a likely annotation error. Consequently, all Batch 1, week 9 sequences were
 396 corrected to the 25/07/2017 date to accurately reflect the sampling schedule.
 397

398

Classification of Independent Transmission Events

399 To differentiate persistent infections from reinfection events, a clade-based classification was
 400 performed. The final classification for the WGS dataset identified 24 initial infections, 10 persistent
 401 infections, and 1 reinfection. For the ORF5 dataset, the analysis identified 80 initial infections, 53
 402 persistent infections, and 1 reinfection. Table 3 summarizes sequence inclusion after the exclusion
 403 of persistent infections. This analysis proved to be robust for both datasets: the number of persistent

404 and new infections remained stable when the number of clades was set to any value between three
 405 and seven for WGS, and between four and seven for the ORF5 dataset (Supplementary file 4).
 406 Additionally, both ORF5 and WGS analyses identified a reinfection event in animal 515. This event
 407 was recorded at week 6 in the ORF5 dataset (where the week 9 sample was unavailable) and at
 408 week 9 in the WGS dataset (where the week 6 sample was unavailable).

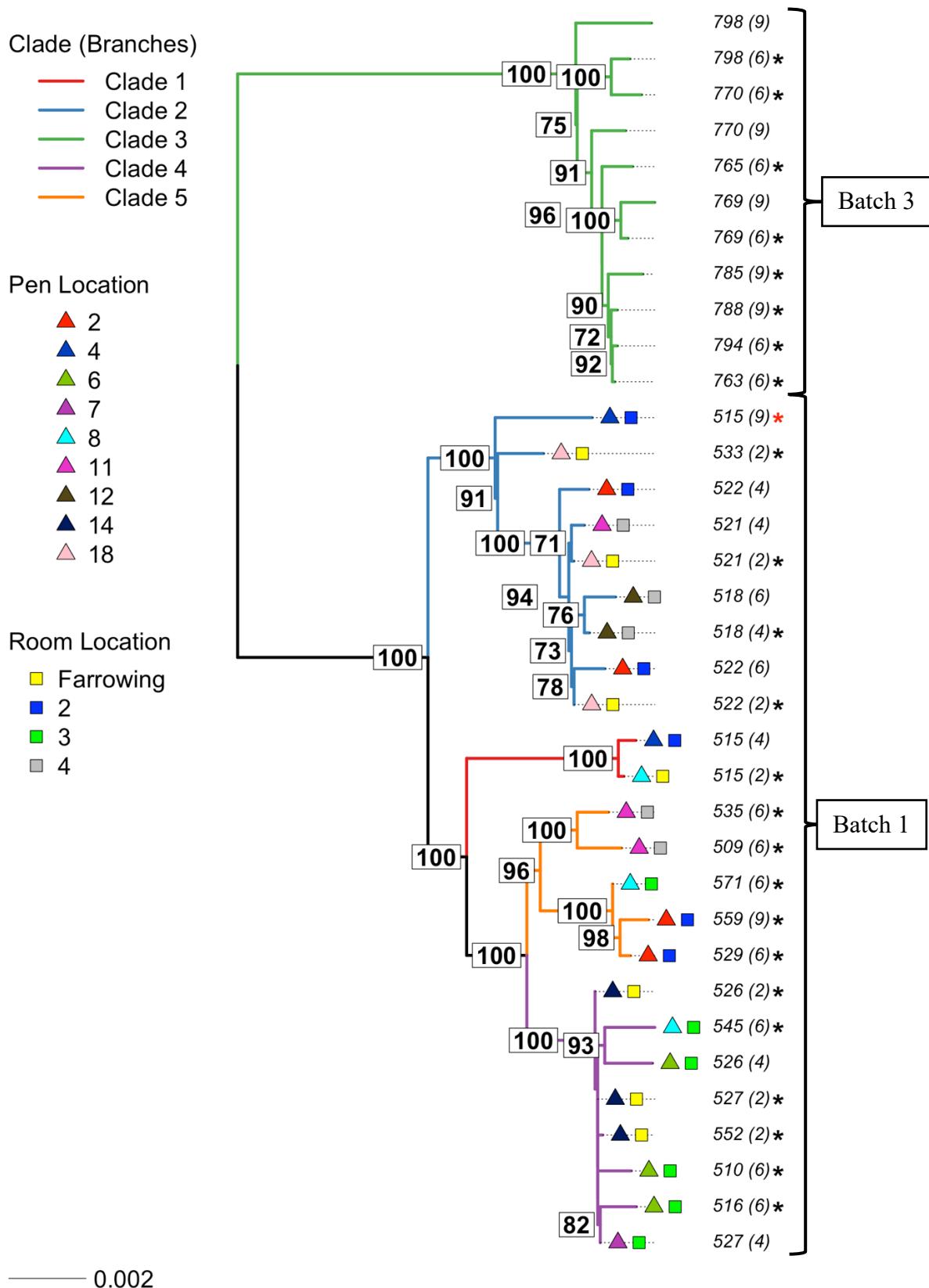
Table 3: Metadata summary of ORF5 sequences before and after filtering for persistent infections. Data are presented as Total Sequences / Included Sequences.

General Cohorts			Spatial Distribution (Batch 1)				
Week	Batch 1	Batch 3	Farrowing	Room 1	Room 2	Room 3	Room 4
Week 0	6 / 6	0 / 0	6 / 6	0 / 0	0 / 0	0 / 0	0 / 0
Week 2	2 / 0	5 / 5	2 / 0	0 / 0	0 / 0	0 / 0	0 / 0
Week 4	8 / 8	0 / 0	0 / 0	0 / 0	3 / 3	5 / 5	0 / 0
Week 6	34 / 25	22 / 18	0 / 0	9 / 9	8 / 5	17 / 11	0 / 0
Week 9	36 / 14	21 / 5	0 / 0	12 / 5	4 / 1	11 / 2	9 / 6
Total	86 / 53	48 / 28	8 / 6	21 / 14	15 / 9	33 / 18	9 / 6

409 After all persistent infections were excluded from the ORF5 datasets, the number of total and
 410 unique sequences dropped from 134 to 81 and 47 to 35 sequences, respectively. The redundancy
 411 percentage improved from 64.93% to 56.59% and 11 identical sequence groups were observed
 412 compared to the 20 identical groups in the ORF5 dataset without exclusion. After removal of
 413 persistent infections from the WGS dataset, the total number of total sequences dropped from 35
 414 to 25 sequences.
 415

Phylogenetic Reconstruction and Topology

416 The phylogenetic reconstruction of the consensus WGS and ORF5 datasets yielded topologies
 417 characterized by high statistical confidence, as indicated by the high nodal bootstrap values
 418 observed across the backbone of both rooted trees (Figure 4). The clustering algorithm used for the
 419 reinfection analysis identified five monophyletic groups in the WGS dataset, yielding a cluster
 420 composition identical to the clades described by Clilverd and colleagues (2023). Farrowing unit
 421 sequences consistently occupied ancestral nodes relative to nursery unit sequences, which appeared
 422 as descendant taxa within the topology. In contrast, the ORF5 phylogeny failed to correctly recover
 423 this structure. Specifically, the clustering algorithm could not resolve one clade as a monophyletic
 424 group; instead, one lineage was split and interspersed among other clades. Additionally, the clear
 425 ancestral relationship observed in the genomic data, where farrowing sequences precede nursery
 426 sequences, was not evident in the ORF5 tree. Finally, a strong spatial signal is evident in the
 427 terminal branches of the ORF5 phylogeny. It can be observed that the clusters of 100% identical
 428 sequences consist largely of samples originating from the same room. While these identical
 429 sequence groups also show a clear tendency to cluster within the same physical pens, this signal
 430 is most pronounced at the room level.
 431

A

B

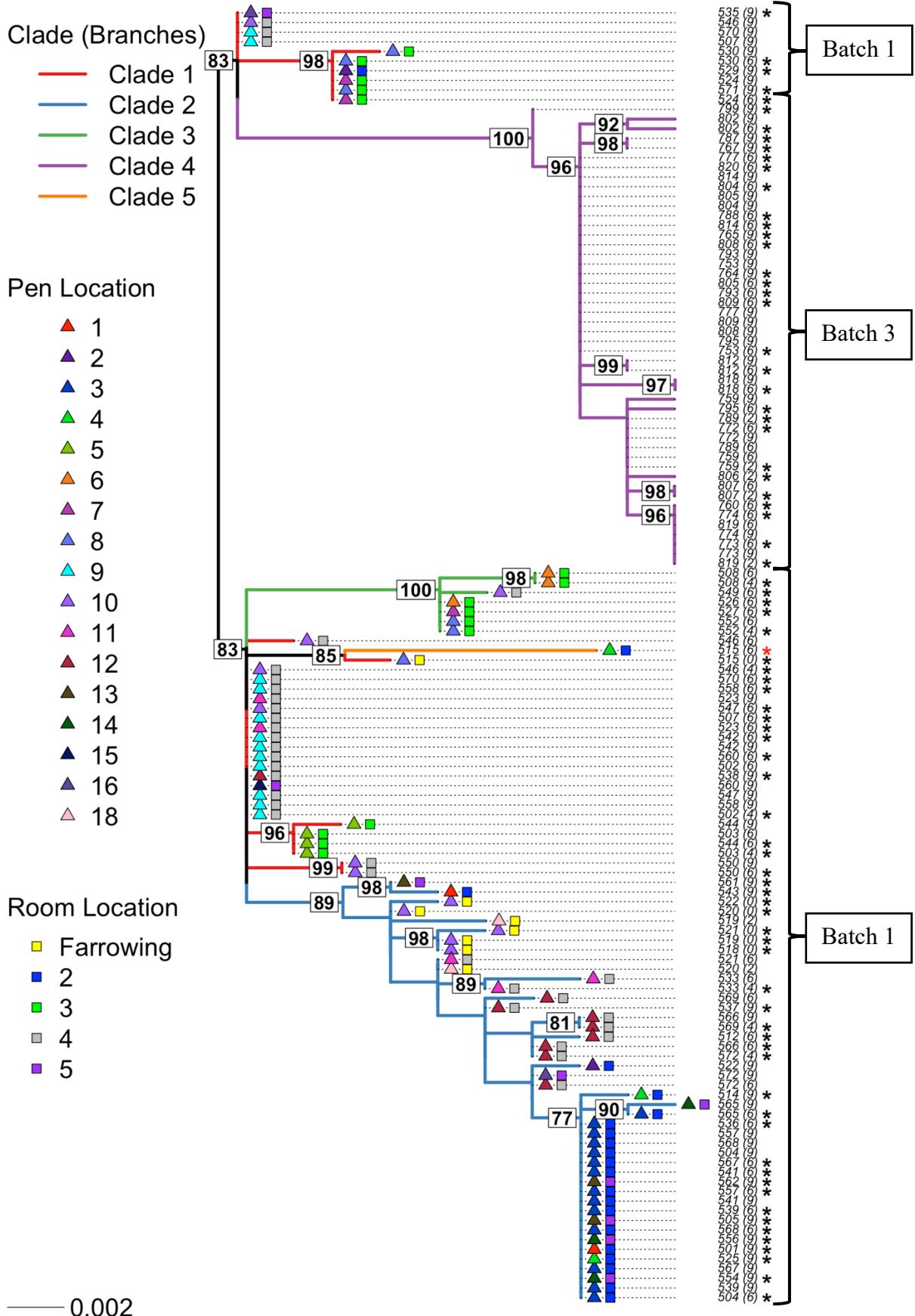


Figure 4: Maximum likelihood phylogenetic reconstruction of the 35 cohort Whole Genome Sequences (WGS) (A) and 134 consensus ORF5 sequences (B). The tree displays the evolutionary relationships of the cohort sequences with branches colored according to the five-clade classification. The scale bar represents genetic distance expressed as nucleotide substitutions per site. Colored triangles and squares denote pen- and room-level locations, respectively. Nodal support values are indicated on the tree nodes (values below 70 are removed for clarity). An asterisk (*) following the sequence annotation denotes an initial infection, whereas a red asterisk indicates a reinfection event.

432 Assessment of Recombination and Lateral Introduction

433 Phylogenetic reconstruction of the extended WGS and ORF5 datasets revealed strongly supported
434 monophyletic clades of farm sequences in both datasets (Supplementary File 5), confirming a single
435 introduction and validating that the outbreak was constrained to within-farm dynamics for the
436 duration of the 12-month study period. Subsequent screening for recombination revealed no
437 evidence of reticulation in Neighbor-Net networks (Supplementary file 6) nor significant signals in
438 the Phi test (WGS: $p=1.0$; ORF5: $p=0.84$), validating its suitability for phylodynamic inference.
439

440 Molecular Clock Validation and Temporal Signal

441 Root-to-tip regression analysis, performed using Clockor2 on the maximum likelihood
442 phylogenetic trees to obtain priors for the final BEAST models, yielded strong positive correlation
443 rates for both the cohort WGS ($R^2 = 0.992$) and ORF5 ($R^2 = 0.836$) datasets. The estimated
444 evolutionary rates were 1.354×10^{-2} and 1.698×10^{-2} substitutions/site/year for WGS and ORF5,
445 respectively. Furthermore, the tMRCA was estimated to be 2017/01/23 for the ORF5 dataset and
446 2017/02/24 for the WGS dataset, placing the origin of the outbreak approximately 8 weeks before
447 clinical detection at 2017/04/08. This confirmed the combined datasets' suitability for temporal
448 analysis.
449

450 However, the partitioned analysis (local clock model for ORF5 data, partitioned by Batch 1 and
451 Batch 3), revealed that Batch 1 exhibited a poor temporal signal on its own ($R^2 = 0.144$) with an
452 evolutionary rate of 2.693×10^{-2} substitutions/site/year, while Batch 3 displayed no temporal signal
453 (negative evolutionary rate). This finding indicated that a reliable clock rate could not be estimated
454 from the Batch 1 and 3 data separately.
455

456 Long-term Temporal Dynamics (BDSKY)

457 To reconstruct the viral spread and estimate evolutionary parameters over the 12-month study
458 period, the BDSKY model was applied to both the cohort ORF5 and WGS datasets. A direct
459 comparison of the inferred epidemiological trajectories showed that both datasets successfully
460 resolved the outbreak into the three distinct phases (initial outbreak, dormancy and resurgence)
461 previously described by Clilverd and colleagues (2023). Under the 3-epoch configuration, the
462 ORF5 and WGS datasets yielded consistent epidemiological patterns, though with slight variations
463 in magnitude (Figure 5 A and B). The ORF5 analysis estimated an initial outbreak phase with a Re
464 of 2.48 (95% Highest Posterior Density (HPD): 1.81–3.15), followed by a significant reduction in
465 transmission during the dormant phase ($Re = 0.20$; 95% HPD: 0.03–0.40), and a final resurgence
466 phase $Re = 3.17$; 95% HPD: 2.02–4.40). The WGS analysis mirrored this trajectory, estimating an
467 initial outbreak Re of 1.80 (95% HPD: 1.20–2.44), a dormant phase Re of 0.35 (95% HPD: 0.05–
468 0.71), and a resurgence Re of 3.85 (95% HPD: 1.62–6.42). Furthermore, the estimated transition
469 times between these phases aligned with the study's metadata for both datasets, placing Batch 1
470 within the outbreak phase, Batch 2 in the dormant phase, and Batch 3 in the resurgence phase.

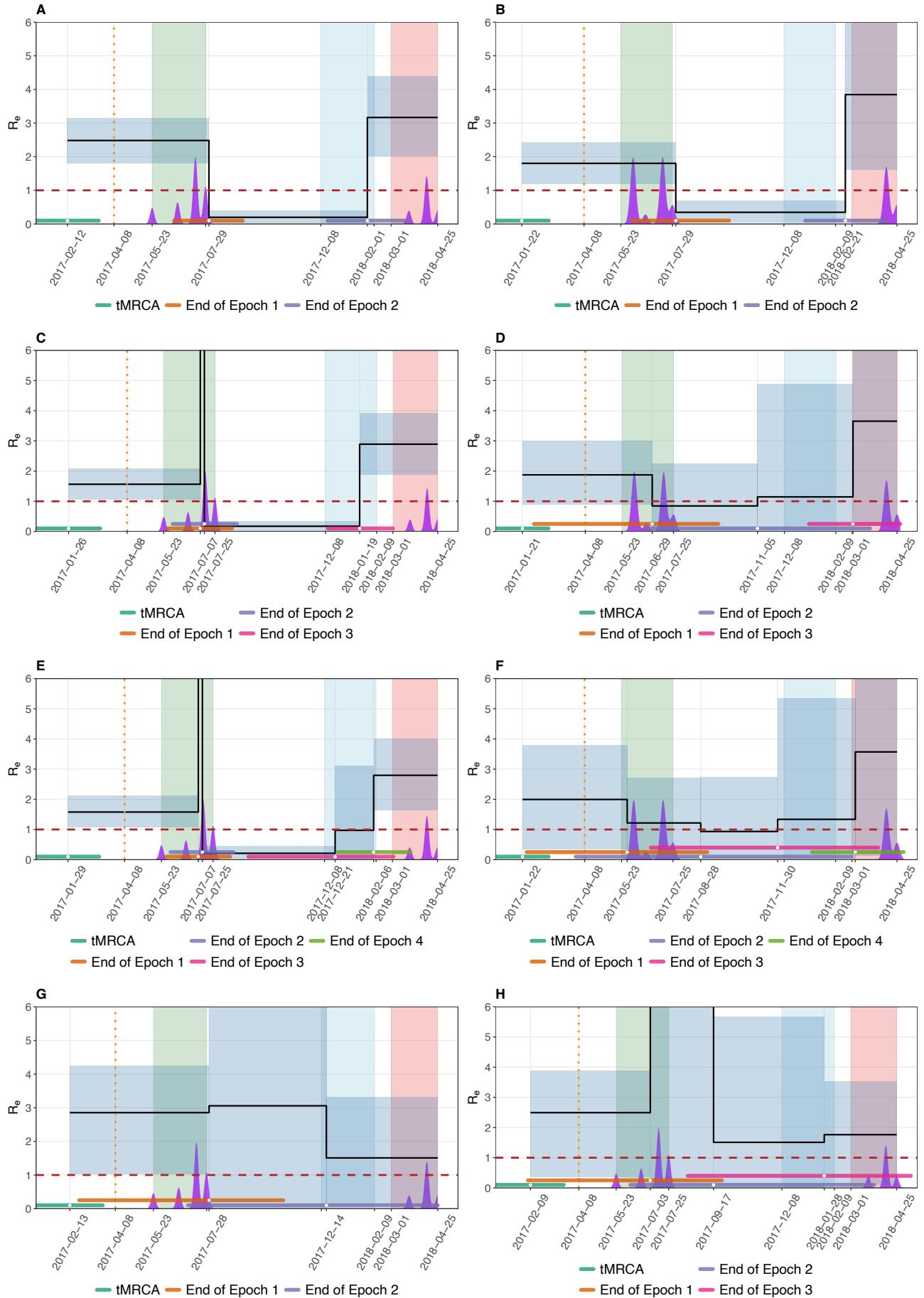


Figure 5: Comparison of the temporal trajectory of the Effective Reproductive Number (R_e) inferred from ORF5 and Whole Genome Sequence (WGS) datasets, alongside prior sensitivity analysis. The upper panels compare the epidemiological reconstruction for the ORF5 dataset (Left Column: A, C, E) and the WGS dataset (Right Column: B, D, F) across three temporal configurations: 3-epoch (A, B), 4-epoch (C, D), and 5-epoch (E, F). The bottom panels display the corresponding sample-from-prior analysis using the 3-epoch (G) and 4-epoch (H) configurations. The R_e is depicted by the solid black line (mean) and the shaded background regions indicate the 95% Highest Posterior Density (HPD) intervals. The horizontal-colored bars positioned above the x-axis denote the 95% HPD intervals for the epoch transition times and the time to Most Recent Common Ancestor (tMRCA), while the purple density plots along the timeline indicate the sampling density. The red dashed line represents the epidemiological threshold of $R_e = 1$ and the orange dotted vertical line indicates the detection of the outbreak on the farm. The colored vertical bars represent the sampling period for batch 1 (green), batch 2 (blue) and batch 3 (red). R_e estimates exceeding 6 are omitted from the visualization for clarity but are available in supplementary file 7.

471 While the 3-epoch model yielded comparable parameter estimates across both datasets, increasing
 472 the temporal resolution to 4- and 5-epoch models resulted in differing model behavior between the
 473 datasets. In the ORF5 analysis, higher-resolution models resolved a short-duration, high-magnitude
 474 transmission peak during the early outbreak period, with estimated R_e values reaching
 475 approximately 20 (95% HPD: ~7–40) (Figure 5 C and E). In contrast, application of 4- and 5-epoch
 476 models to the WGS dataset resulted in increased uncertainty in parameter estimates. For these
 477 models, estimations for epoch transition times and R_e values exhibited overlapping 95% HPD
 478 intervals, and distinct temporal phases were not clearly resolved (Figure 5D and F). Finally, the
 479 sampling proportion was estimated consistently across models for both Batches 1 and 3. Focusing
 480 on Batch 1, the WGS dataset yielded posterior estimates centered around 0.25 (95% HPD: 0.10–
 481 0.40), compared to approximately 0.30 (95% HPD: 0.15–0.45) for the ORF5 dataset (Figure 6).
 482

483 Conversely, the comparison of evolutionary parameters showed that the WGS dataset offered
 484 higher precision than ORF5. The mean molecular clock rate for WGS was estimated at 1.12×10^{-2}
 485 substitutions/site/year (95% HPD: $0.97 \times 10^{-2} – 1.33 \times 10^{-2}$), compared to the 1.41×10^{-2} (95%
 486 HPD: $0.85 \times 10^{-2} – 2.00 \times 10^{-2}$) estimated for ORF5. Despite this difference in precision, the 95%
 487 HPD intervals for the Coefficient of Variation (CoV) consistently exceeded 0.1 across all models,
 488 supporting the use of an uncorrelated relaxed clock (Drummond & Bouckaert, 2015). Additionally,
 489 the analysis identified distinct substitution models: the K81 model for ORF5 and Model 24 for
 490 WGS, which couples A→T and G→T rates while distinguishing A→C from G→T transversion
 491 rates (Bouckaert & Drummond, 2017).

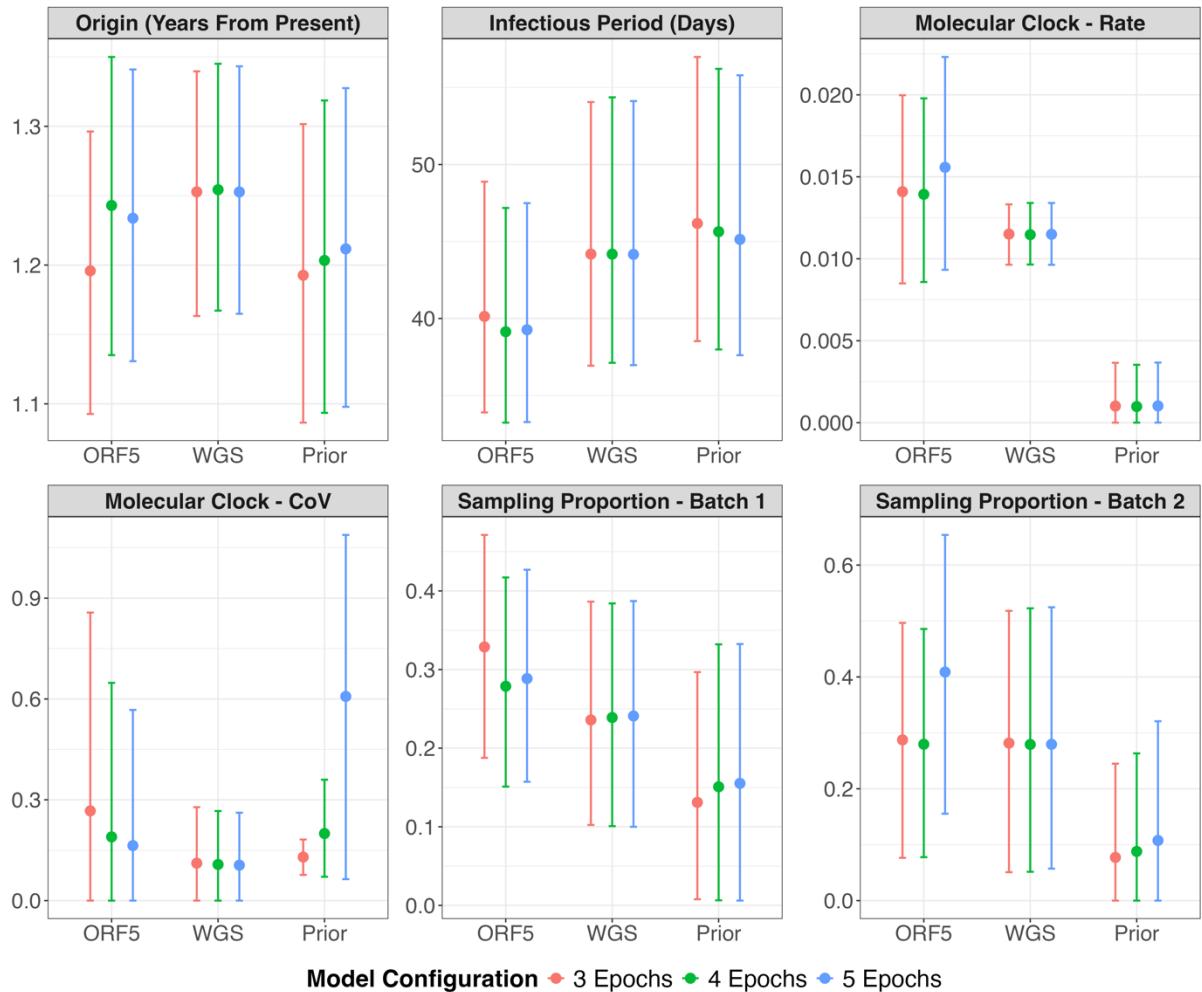


Figure 6: Sensitivity analysis of evolutionary and epidemiological parameter estimates across the Birth-Death Skyline (BDSKY) model configurations. The plots compare the posterior mean estimates and 95% Highest Posterior Density (HPD) intervals for parameters inferred from the ORF5 and Whole Genome Sequence (WGS) datasets, alongside estimates resulting from the sample-from-prior analysis (indicated on the x-axis). All estimates are represented with colors differentiating the temporal resolutions applied: 3-epoch (orange), 4-epoch (green), and 5-epoch (blue) configurations. The comparison highlights the increased precision of the WGS dataset in estimating evolutionary parameters (Clock Rate and Coefficient of Variation (CoV)) relative to ORF5. Note that the posterior estimate of the rate of becoming uninfected, as output by BEAST, was converted into the infectious period per animal, expressed in days. Additionally, the the 4- and 5-epoch sample-from-prior models failed to achieve an Effective Sample Size (ESS) of 200 for the CoV,

492 The comparison between the data-driven and sample-from-prior analysis revealed that the prior-
 493 only analyses for both the 3- and 4-epoch configurations could not resolve distinct phases or
 494 differences in the Re (Figure 5G and H). Specifically, regarding the 4-epoch configuration, the
 495 prior-only analysis showed no evidence of the explosive outbreak peak observed in the data-driven
 496 model, exhibiting very wide 95% HPD intervals for both the Re and epoch transition estimates.
 497 Notably, the 4- and 5-epoch sample-from-prior models failed to achieve an ESS of 200 for the
 498 estimated epoch transition times, Re values, and the CoV (see Supplementary File 7). Tracer
 499 diagnostics indicated poor mixing and an inability to distinguish between multiple local optima;
 500 this was considered a stable limitation of the prior landscape that would not likely improve with
 501 additional sampling. Regarding evolutionary parameters, significant deviations were observed

502 between the prior and data-driven results, particularly for the molecular clock rate and CoV,
503 confirming that the sequence data contained a strong phylogenetic signal that drove the inference.
504

505 Fine-Scale Spatial Transmission Dynamics (BDMM)

506 To quantify transmission dynamics at the pen and room levels and to determine whether the
507 infectious period differed between the farrowing and nursery phases, the BDMM analysis was
508 performed using the discrete migration configuration.
509

510 For the farrowing unit, the estimated R_e value was 1.67 (95% HPD: 0.24–3.26). This estimate did
511 not deviate substantially from the sample-from-prior analysis (Figure 7), indicating that these
512 results were largely influenced by prior assumptions rather than the genomic data. In contrast,
513 transmission dynamics within the nursery unit revealed a biological signal that diverged from the
514 prior distributions. While within-pen R_e was estimated to be 2.61 (95% HPD: 0.52–5.21), the R_e
515 between pens showed to significantly below the epidemiological threshold of 1. Crucially, the 95%
516 HPD intervals for between-pen transmission were considerably narrower than the analysis run only
517 on prior configurations. Specifically, the R_e within the same room was estimated at 0.46 (95%
518 HPD: 0.09–0.94), while transmission between rooms was even lower at 0.06 (95% HPD: 0.02–
519 0.12). This stands in contrast to the sample-from-prior analysis, where the within-room
520 transmission interval included 1 (mean 0.80; 95% HPD: 0.10–1.73), confirming that the discrete
521 model successfully resolved fine-scale constraints on viral spread that were not present in the prior
522 assumptions.
523

524 Furthermore, the model indicated no significant difference in the infectious period between the
525 farrowing units, with a mean of 44 days (95% HPD: 26–103 days), and nursery units, with a mean
526 of 37 days (95% HPD: 21–99 days). While the sample-from-prior analysis yielded a substantially
527 wider 95% HPD interval for the infectious period in the nursery phase compared to the data-driven
528 analysis, the interval for the farrowing phase remained comparable to the prior distribution.
529 Regarding movement dynamics, the migration rates estimated in the data-driven analysis closely
530 matched those obtained from the corresponding sample-from-prior analysis.
531

532 The posterior estimates for the sampling proportion, tree origin, and mean molecular clock rate
533 closely matched their respective prior distributions, consistent with the narrow priors applied to
534 these parameters. Additionally, consistent with the BDSKY analysis, the K81 model was identified
535 as the optimal nucleotide substitution model.
536

537 Finally, to challenge the assumption of discrete one-day animal movements, a robustness test was
538 performed using a continuous migration model. This analysis yielded highly consistent parameter
539 estimates with the discrete configuration described above, showing no major deviations in
540 epidemiological or evolutionary parameters (Supplementary File 8).

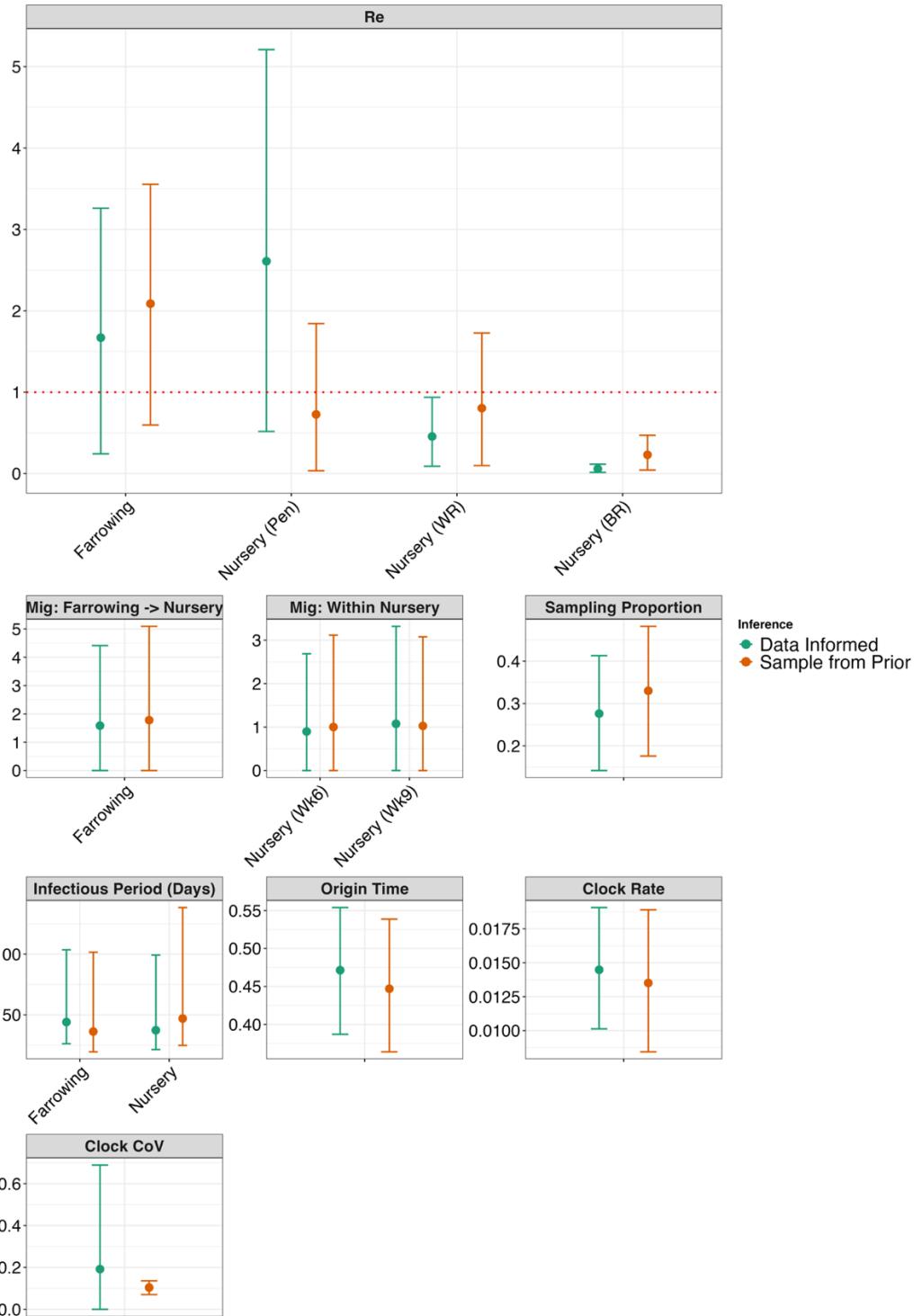


Figure 7: Comparison of posterior parameter estimates for the Birth-Death with Migration (BDMM) model using the discrete migration configuration against the sample-from-prior analysis. The plots display the mean posterior estimates and 95% Highest Posterior Density (HPD) intervals for epidemiological and evolutionary parameters. The analysis compares the Data Informed analysis (blue) with the sample-from-prior analysis (orange) to validate the influence of the genetic data on the posterior estimates. The red dotted horizontal line represents the epidemiological threshold of $Re = 1$. Note that the posterior estimate of the rate of becoming uninfected, as output by BEAST, was converted into the infectious period per animal, expressed in days. , BR=Between Room, CoV=Coefficient of Variation, Mig=Migration rate, Re = Effective reproduction number, WK = Week, WR = Within Room

541 **Discussion**

542 The primary aim of this study was to evaluate the capacity of phylodynamic models to reliably
543 infer previously unquantified epidemiological parameters and distinguish transmission potential
544 across spatially distinct farm units. By leveraging genetic data to reconstruct the transmission
545 dynamics of a PRRSV-1 outbreak, this analysis recovered the established three-phase
546 epidemiological trajectory of outbreak, dormancy, and resurgence described by Clilverd et al.
547 (2023). Crucially, the BDSKY model dated the resurgence event to shortly before Batch 3,
548 supporting the hypothesis proposed by Clilverd and colleagues (2023) that this resurgence was
549 driven by a selection process immediately preceding Batch 3 sampling, as evidenced by the limited
550 genetic diversity of the emerging escape variant.

551 At the fine-scale resolution, the BDMM successfully distinguished between within- and between-
552 pen transmission in the nursery unit. While the within-pen R_e were estimated to be 2.61 (95% HPD:
553 0.52–5.21), estimates for transmission between pens were significantly below the epidemiological
554 threshold of 1, dropping to 0.46 (95% HPD: 0.09–0.94) for pens within the same room and 0.06
555 (95% HPD: 0.02–0.12) for pens across different rooms. This drop in transmission efficiency is
556 validated by our observed phylogenetic structure which displayed a strong spatial signal where
557 clusters of 100% identical sequences consist of samples originating from the same pen and room.
558 The 100% sequence similarity suggests that these individuals likely infected one another, forming
559 localized transmission chains that are effectively constrained by the farm's physical
560 compartmentalization. Additionally, experimental data for related swine viruses also support our
561 findings. Research on Classical Swine Fever Virus (CSFV) shows that transmission within a pen
562 is nearly 20 times faster than between pens separated by open barriers (Weesendorp et al.,
563 2014). Similarly, the negligible R_e for between-room spread aligns with PRRSV aerosol studies
564 where the virus, though detectable in barn air, often fails to infect pigs in separate compartments
565 despite exposure to contaminated air current (Otake et al., 2002; Trincado et al., 2004). This finding
566 demonstrates that phylodynamic models can successfully quantify the structured transmission
567 dynamics characteristic of farm environments. Specifically, the models reveal that while direct
568 contact drives efficient transmission among pen-mates, horizontal spread between spatially
569 separated units remains a rare event.

570
571 However, the R_e estimated for the farrowing unit requires a distinct interpretation. Unlike the
572 nursery estimates, this value does not reflect horizontal transmission between pen mates, as genetic
573 analysis confirmed these infections were acquired in-utero (Clilverd et al., 2023). Instead, similar
574 to the "ghost deme" sow reservoir, the farrowing R_e serves a structural function: it bridges the
575 temporal gap between introduction and the nursery outbreak, allowing the lineage to persist in the
576 model without quantifying local horizontal spread.

577
578 Notably, the estimated infectious period did not differ between the farrowing and nursery phases.
579 While this may reflect a genuine similarity between phases, the ability to detect a difference was
580 likely constrained due to sampling sparsity: in the reconstructed tree, farrowing-phase sequences
581 did not appear as ancestors of nursery-derived lineages, indicating that nursery infections likely
582 arose from unsampled farrowing or nursery stage pigs or the sow reservoir. Consequently, the
583 model could not interpret farrowing infections as long-term reservoirs bridging the nursery stage.
584 This highlights a significant limitation of the model's "isolated" assumption, which treated the
585 farrowing and nursery units as closed systems after the initial introduction from the sow herd.
586 Moreover, the large 95% HPD intervals associated with the infectious period estimates in both
587 phases obscured the detection of a statistically meaningful difference. This uncertainty can be
588 explained by a parameter non-identifiability issue, in which the model cannot distinguish between
589 "fast turnover" (high transmission with rapid recovery) and "slow turnover" (low transmission with
590 prolonged infectiousness) scenarios because both yield indistinguishable likelihoods (Zarebski et
591

592 al., 2022). The impact of this limitation is apparent in the nursery unit results, where both within-
593 pen Re and the infectious period were associated with wide 95% HPD intervals. A commonly
594 adopted approach to alleviate such identifiability issues is to fix the rate of becoming uninfected
595 using a well-supported prior (Weber et al., 2023).

596
597 A critical outcome of this study is the evaluation of genomic resolution on model performance,
598 revealing complementary strengths between targeted gene sequencing and whole-genome
599 approaches. The larger ORF5 dataset produced more confident epidemiological estimates for
600 the Re and transition times than the sparse WGS dataset, supporting the argument that high taxon
601 density is required to distinguish epidemiological rate shifts from background noise in BDSKY
602 models (Culshaw et al., 2019). Notably, the higher resolution of the ORF5 dataset resolved a brief,
603 intense transmission peak during the early outbreak that the WGS dataset failed to capture. We
604 hypothesize that this peak reflects the superspreading event reported by Clilverd and colleagues
605 (2023), in which a single pig accounted for 28.6% of all traced horizontal transmission
606 events. Conversely, the WGS dataset provided superior resolution for evolutionary parameters,
607 yielding narrower 95% HPD intervals for the clock rate and CoV, and a stronger temporal signal
608 ($R^2=0.992$) compared to ORF5 ($R^2=0.836$). This aligns with the 'temporal horizon' concept (Dudas
609 & Bedford, 2019), where longer sequences provide the mutational density necessary for precise
610 clock calibration. This increased information density likely also drove the inference of a more
611 complex substitution model for the WGS dataset (Model 24) compared to ORF5 (K81), suggesting
612 that the exceptional genetic diversity of PRRSV requires data-driven models rather than reliance
613 on default assumptions. Consequently, researchers should select genomic resolution based on
614 specific objectives: high-density, targeted sequencing is recommended for monitoring fine-scale
615 epidemiological dynamics and detecting superspreading events. Conversely, longer sequences
616 should be prioritized for evolutionary precision, as increased mutational density enables accurate
617 molecular clock calibration. While not the primary focus of this study, such precision is essential
618 for analyses aiming to determine the origin of an outbreak by estimating the tMRCA.
619

620 A key strength of the BDSKY analysis lies in the strict parameterization of the sampling proportion
621 to mirror the sampling schedule. This fixed configuration was necessary to distinguish sampling
622 gaps from low viral prevalence and prevent poor statistical identifiability (Featherstone et al., 2021;
623 Parag et al., 2022). By preventing the model from inherently interpreting the lack of samples during
624 the dormant phase as a halt of transmission, this strategy ensured that parameter estimates remained
625 robust and driven by genomic signals rather than prior configurations. This was further validated
626 by sample-from-prior analyses, which confirmed that the prior configuration of the model alone
627 could not resolve the epidemiological phases. However, a key limitation of the BDSKY model is
628 its assumption of an unstructured population which is violated according to our BDMM analysis
629 that confirmed that transmission was dependent on spatial units. Consequently, ignoring this
630 structure can lead to biased transmission rate estimates, as the model incorrectly assumes equal
631 contact probabilities across the entire herd (Seidel et al., 2024). Specifically, the model may
632 misinterpret localized viral persistence within a single pen as broader, herd-wide transmission,
633 potentially inflating the Re estimate (Seidel et al., 2024; Volz & Frost, 2014). Accordingly, it
634 should be noted that the Re values reported in the BDSKY analysis may represent overestimates
635 of the true transmission intensity.

636
637 Furthermore, the study faced several further strengths and limitations. The clade-based analysis
638 effectively differentiated persistent infections from reinfections, supported by the recovery of the
639 single reinfection event (pig 515) reported by Clilverd et al. (2023) and robustness checks
640 confirming that infection classifications remained consistent across varying numbers of estimated
641 clades. However, the ORF5 dataset failed to maintain monophyly for one lineage, probably due to
642 a high proportion of identical sequences. This aligns with the finding that high sequence identity
643 increases "monophyly-violating interclass coalescence," increasing the likelihood that lineages mix

before coalescing (Mehta et al., 2016). To overcome the limitation of low unique genetic variation, future studies could leverage intra-host diversity as demonstrated by Ortiz et al. (2023). By tracking shared minor variants that are typically discarded in consensus sequencing, this method resolves transmission links between hosts with identical consensus sequences, effectively distinguishing specific transmission pathways within otherwise unresolved clusters. Furthermore, the finding that transmission is predominantly confined to within-pen dynamics introduces a challenge regarding the "burnout" phenomenon (Parsons et al., 2024). In these small groups, the virus quickly exhausts the limited pool of susceptible pigs, causing outbreaks to terminate abruptly. This premature truncation of transmission chains can bias R_e estimates, as standard models assume ongoing spread in a larger susceptible population rather than extinction due to host depletion.

To overcome the limitations associated with data sparsity and structure, future study designs targeting within-farm transmission should be optimized for phylodynamic inference. We recommend a shift from fixed longitudinal sampling of the same individuals to a rotational strategy. By replacing repeated sampling of already infected individuals at subsequent time points with sampling of previously unsampled pen-mates or animals that have not yet tested positive, researchers can avoid redundancy arising from persistent infections and instead maximize the recovery of distinct transmission events. Furthermore, sampling must extend beyond the farrowing and nursery units to include the sow herd to replace the "ghost deme" assumption with empirical data, enabling the direct quantification of vertical transmission and spillover events. Finally, future phylodynamic studies that aim to characterize transmission between isolated compartments should incorporate metadata on personnel movement and ventilation to better distinguish transmission pathways. Arruda and colleagues (2019) highlight that distinguishing between aerosol transmission and local spread via fomites is often confounded by a lack of data on internal biosecurity measures and contact networks. Integrating these factors is methodologically feasible in phylodynamic models, as Meester et al. (2025) recently used BDMM to successfully distinguish environmental from direct transmission routes. Therefore, detailed metadata would allow models to explicitly parameterize the physical and procedural barriers defining subpopulation interactions.

In conclusion, this study demonstrates that phylodynamic models can successfully reconstruct within-farm PRRSV-1 outbreak dynamics and quantify fine-scale transmission parameters using standard surveillance data. By confirming the three-phase epidemic trajectory and distinguishing within-pen spread from between-pen transmission, our findings validate the utility of phylodynamics for resolving epidemiological events that traditional methods may miss. The analysis also highlights critical trade-offs: while targeted ORF5 sequencing offers the high sampling density necessary to resolve rapid transmission shifts, whole-genome sequencing provides the mutational resolution required for robust evolutionary clock calibration. Ultimately, moving from observational surveillance to quantitative, genetics-driven modeling allows for a more nuanced understanding of outbreak mechanics, providing the foundation needed to refine internal biosecurity and targeted intervention strategies in the swine industry

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