# CS189/CS289A Introduction to Machine Learning Lecture 9: Regression

Peter Bartlett

February 17, 2015

• Review: Decision theory.

- Review: Decision theory.
- Empirical risk minimization.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.
  - Maximum likelihood is least squares.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.
  - Maximum likelihood is least squares.
  - Distributions of parameter estimates.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.
  - Maximum likelihood is least squares.
  - Distributions of parameter estimates.

### Regression with quadratic loss

Outcomes are in  $\mathcal{Y} = \mathbb{R}$ .

We consider the quadratic loss function,  $\ell(\hat{y}, y) = (\hat{y} - y)^2$ .

Risk is expected squared error:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}.$$

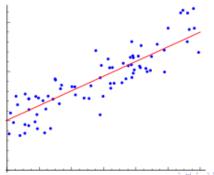
### Regression with quadratic loss

Outcomes are in  $\mathcal{Y} = \mathbb{R}$ .

We consider the quadratic loss function,  $\ell(\hat{y}, y) = (\hat{y} - y)^2$ .

Risk is expected squared error:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}.$$



Risk:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}$$

Risk:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^2 = \mathbb{E}\mathbb{E}[(f(X) - Y)^2|X].$$

Risk:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^2 = \mathbb{E}\mathbb{E}[(f(X) - Y)^2|X].$$

For each X, we minimize the conditional expectation of the loss,

$$\mathbb{E}\left[(f(X)-Y)^2|X\right].$$

Risk:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^2 = \mathbb{E}\mathbb{E}[(f(X) - Y)^2|X].$$

For each X, we minimize the conditional expectation of the loss,

$$\mathbb{E}\left[(f(X)-Y)^2|X\right].$$

#### Bias-variance decomposition

$$R(f) = \mathbb{E}\left[\underbrace{\left[\left(f(X) - \mathbb{E}[Y|X]\right)^{2}\right]}_{\text{bias}^{2}} + \mathbb{E}\left[\underbrace{\left[\left(\mathbb{E}[Y|X] - Y\right)^{2}\right]}_{\text{variance}}\right]$$

#### Bias-variance decomposition

$$R(f) = \mathbb{E}\underbrace{\left[ (f(X) - \mathbb{E}[Y|X])^2 \right]}_{\text{bias}^2} + \mathbb{E}\underbrace{\left[ (\mathbb{E}[Y|X] - Y)^2 \right]}_{\text{variance}}$$
$$= \mathbb{E}\left[ (f(X) - f^*(X))^2 \right] + \mathbb{E}\left[ (f^*(X) - Y)^2 \right]$$

#### Bias-variance decomposition

$$R(f) = \mathbb{E}\underbrace{\left[ (f(X) - \mathbb{E}[Y|X])^2 \right]}_{\text{bias}^2} + \mathbb{E}\underbrace{\left[ (\mathbb{E}[Y|X] - Y)^2 \right]}_{\text{variance}}$$

$$= \mathbb{E}\left[ (f(X) - f^*(X))^2 \right] + \mathbb{E}\left[ (f^*(X) - Y)^2 \right]$$

$$= \mathbb{E}\left[ (f(X) - f^*(X))^2 \right] + R(f^*).$$

#### Bias-variance decomposition

$$R(f) = \mathbb{E}\underbrace{\left[ (f(X) - \mathbb{E}[Y|X])^2 \right]}_{\text{bias}^2} + \mathbb{E}\underbrace{\left[ (\mathbb{E}[Y|X] - Y)^2 \right]}_{\text{variance}}$$

$$= \mathbb{E}\left[ (f(X) - f^*(X))^2 \right] + \mathbb{E}\left[ (f^*(X) - Y)^2 \right]$$

$$= \mathbb{E}\left[ (f(X) - f^*(X))^2 \right] + R(f^*).$$

$$R(f) - R^* = \mathbb{E}\left[ (f(X) - f^*(X))^2 \right].$$

Consider  $X \in \mathbb{R}^p$ ,  $Y \in \mathbb{R}$ , and consider linear (affine) prediction rules

Consider  $X \in \mathbb{R}^p$ ,  $Y \in \mathbb{R}$ , and consider linear (affine) prediction rules,

$$F_{lin} := \{x \mapsto x'\beta + \beta_0 : \beta \in \mathbb{R}^p, \beta_0 \in \mathbb{R}\}.$$

Consider  $X \in \mathbb{R}^p$ ,  $Y \in \mathbb{R}$ , and consider linear (affine) prediction rules,

$$F_{lin} := \left\{ x \mapsto x'\beta + \beta_0 : \beta \in \mathbb{R}^p, \beta_0 \in \mathbb{R} \right\}.$$

Two ways to motivate least squares:

Consider  $X \in \mathbb{R}^p$ ,  $Y \in \mathbb{R}$ , and consider linear (affine) prediction rules,

$$F_{lin} := \left\{ x \mapsto x'\beta + \beta_0 : \beta \in \mathbb{R}^p, \beta_0 \in \mathbb{R} \right\}.$$

Two ways to motivate least squares:

Onsider the class of linear prediction rules. Minimize empirical risk over the class of linear prediction rules.

Consider  $X \in \mathbb{R}^p$ ,  $Y \in \mathbb{R}$ , and consider linear (affine) prediction rules,

$$F_{lin} := \left\{ x \mapsto x'\beta + \beta_0 : \beta \in \mathbb{R}^p, \beta_0 \in \mathbb{R} \right\}.$$

Two ways to motivate least squares:

- Onsider the class of linear prediction rules. Minimize empirical risk over the class of linear prediction rules.
- ② Model the process generating the  $Y_i$ s as a linear function of the  $X_i$ s, plus additive Gaussian noise.
  - Compute the maximum likelihood estimate for the linear coefficients.

Consider  $X \in \mathbb{R}^p$ ,  $Y \in \mathbb{R}$ , and consider linear (affine) prediction rules,

$$F_{lin} := \left\{ x \mapsto x'\beta + \beta_0 : \beta \in \mathbb{R}^p, \beta_0 \in \mathbb{R} \right\}.$$

Two ways to motivate least squares:

- Consider the class of linear prediction rules. Minimize *empirical risk* over the class of linear prediction rules.
- ② Model the process generating the  $Y_i$ s as a linear function of the  $X_i$ s, plus additive Gaussian noise.

Compute the maximum likelihood estimate for the linear coefficients.

In both cases, we arrive at the *normal equations*: the choice of  $\beta$  corresponds to a projection on to a linear sub-space.

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.
  - Maximum likelihood is least squares.
  - Distributions of parameter estimates.

#### Risk and empirical risk

Risk is the expected squared error:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}.$$

#### Risk and empirical risk

Risk is the expected squared error:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}.$$

Empirical risk is the sample average of squared error:

#### Risk and empirical risk

Risk is the expected squared error:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}.$$

Empirical risk is the sample average of squared error:

$$\hat{R}(f) = \hat{\mathbb{E}}_n \ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2.$$

#### Risk and empirical risk

Risk is the expected squared error:

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2}.$$

Empirical risk is the sample average of squared error:

$$\hat{R}(f) = \hat{\mathbb{E}}_n \ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2.$$

Here,  $\hat{\mathbb{E}}_n$  means expectation under the *empirical distribution*, which puts mass 1/n at each  $(X_i, Y_i)$  pair in the sample.

We want to choose a linear prediction rule  $f \in F_{lin}$  to minimize risk.

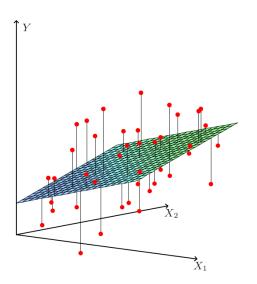
We want to choose a linear prediction rule  $f \in F_{lin}$  to minimize risk. One approach is to choose the linear prediction rule that minimizes empirical risk:

$$\hat{f} := \arg\min_{f \in F_{lin}} \hat{\mathbb{E}}_n \ell(f(X), Y)$$

We want to choose a linear prediction rule  $f \in F_{lin}$  to minimize risk. One approach is to choose the linear prediction rule that minimizes empirical risk:

$$\hat{f} := \arg\min_{f \in F_{lin}} \hat{\mathbb{E}}_n \ell(f(X), Y)$$

$$= \arg\min_{f \in F_{lin}} \sum_{i=1}^n (f(X_i) - Y_i)^2.$$



**FIGURE 3.1.** Linear least squares fitting with  $X \in \mathbb{R}^2$ . We seek the linear function of X that minimizes the sum of squared residuals from Y.

Just as we did when we were considering linear classifiers, we'll simplify notation by bundling the offset term  $(\beta_0)$  into the parameter vector  $\beta$  and assuming that the covariates  $X_i$  include a constant 1 component.

Just as we did when we were considering linear classifiers, we'll simplify notation by bundling the offset term  $(\beta_0)$  into the parameter vector  $\beta$  and assuming that the covariates  $X_i$  include a constant 1 component.

Then  $f \in F_{lin}$  is of the form  $f(x) = x'\beta$ .

We wish to find  $\hat{f}: x \mapsto x'\hat{\beta}$ 

We wish to find  $\hat{f}: x \mapsto x'\hat{\beta}$ , where

$$\hat{\beta} = \arg\min_{eta \in \mathbb{R}^p} \sum_{i=1}^n (X_i' \beta - Y_i)^2$$

We wish to find  $\hat{f}: x \mapsto x'\hat{\beta}$ , where

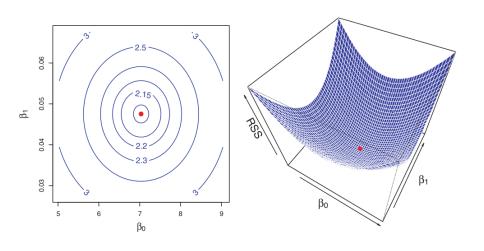
$$\hat{\beta} = \arg\min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (X_i'\beta - Y_i)^2$$

$$= \arg\min_{\beta \in \mathbb{R}^p} ||X\beta - y||^2,$$
RSS

where the design matrix  $X \in \mathbb{R}^{n \times p}$  and response vector  $y \in \mathbb{R}^n$  are

$$X = \begin{pmatrix} X_1' \\ X_2' \\ \vdots \\ X_n' \end{pmatrix}, \qquad y = \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix}.$$

(Think of  $n \gg p$ , so X is tall.)



Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$

Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$
$$= \frac{1}{2} (X\beta - y)' (X\beta - y)$$

### Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$
$$= \frac{1}{2} (X\beta - y)' (X\beta - y)$$
$$= \frac{1}{2} \beta' X' X \beta - y' X \beta + \frac{1}{2} y' y,$$

Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$
$$= \frac{1}{2} (X\beta - y)' (X\beta - y)$$
$$= \frac{1}{2} \beta' X' X \beta - y' X \beta + \frac{1}{2} y' y,$$

we can differentiate wrt  $\beta$ :

$$\nabla_{\beta} RSS(\beta) = X'X\beta - X'y, \qquad \nabla^{2}_{\beta} RSS(\beta) = X'X.$$

Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$
$$= \frac{1}{2} (X\beta - y)' (X\beta - y)$$
$$= \frac{1}{2} \beta' X' X \beta - y' X \beta + \frac{1}{2} y' y,$$

we can differentiate wrt  $\beta$ :

$$\nabla_{\beta} RSS(\beta) = X'X\beta - X'y, \qquad \nabla^{2}_{\beta} RSS(\beta) = X'X.$$

Now,  $X'X \succeq 0$ 

Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$
$$= \frac{1}{2} (X\beta - y)' (X\beta - y)$$
$$= \frac{1}{2} \beta' X' X \beta - y' X \beta + \frac{1}{2} y' y,$$

we can differentiate wrt  $\beta$ :

$$\nabla_{\beta} RSS(\beta) = X'X\beta - X'y, \qquad \nabla^{2}_{\beta} RSS(\beta) = X'X.$$

Now,  $X'X \succeq 0$ , so setting  $\nabla_{\beta}RSS(\beta) = 0$  gives a minimum of RSS

Defining

$$RSS(\beta) = \frac{1}{2} \|X\beta - y\|^2$$
$$= \frac{1}{2} (X\beta - y)' (X\beta - y)$$
$$= \frac{1}{2} \beta' X' X \beta - y' X \beta + \frac{1}{2} y' y,$$

we can differentiate wrt  $\beta$ :

$$\nabla_{\beta} RSS(\beta) = X'X\beta - X'y, \qquad \nabla^{2}_{\beta} RSS(\beta) = X'X.$$

Now,  $X'X\succeq 0$ , so setting  $\nabla_{\beta}RSS(\beta)=0$  gives a minimum of RSS when

$$X'X\beta = X'y$$
.

### Normal equations

$$X'X\beta=X'y.$$

### Normal equations

$$X'X\beta = X'y$$
.

$$\hat{\beta} = (X'X)^{-1}X'y.$$

#### A projection viewpoint

We are aiming to find  $\beta$  to minimize  $||y - X\beta||$ .

#### A projection viewpoint

We are aiming to find  $\beta$  to minimize  $||y - X\beta||$ .

Writing

$$X=\begin{pmatrix}x_1&x_2&\cdots&x_p\end{pmatrix},$$

#### A projection viewpoint

We are aiming to find  $\beta$  to minimize  $||y - X\beta||$ .

Writing

$$X = \begin{pmatrix} x_1 & x_2 & \cdots & x_p \end{pmatrix},$$

we have

$$y - X\beta = y - \sum_{j=1}^{p} \beta_j x_j.$$

#### A projection viewpoint

We are aiming to find  $\beta$  to minimize  $||y - X\beta||$ .

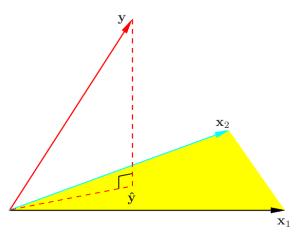
Writing

$$X = \begin{pmatrix} x_1 & x_2 & \cdots & x_p \end{pmatrix},$$

we have

$$y - X\beta = y - \sum_{j=1}^{p} \beta_j x_j.$$

That is, we want to find a linear combination of the columns  $x_j \in \mathbb{R}^n$  of X that minimizes Euclidean distance to  $y \in \mathbb{R}^n$ .



**FIGURE 3.2.** The N-dimensional geometry of least squares regression with two predictors. The outcome vector  $\mathbf{y}$  is orthogonally projected onto the hyperplane spanned by the input vectors  $\mathbf{x}_1$  and  $\mathbf{x}_2$ . The projection  $\hat{\mathbf{y}}$  represents the vector of the least squares predictions

### Projection Theorem

The optimal approximation  $\hat{y}$  in the space spanned by the columns  $x_j$  of X has an error  $y - \hat{y}$  that is orthogonal to that column space.

#### Projection Theorem

The optimal approximation  $\hat{y}$  in the space spanned by the columns  $x_i$  of X has an error  $y - \hat{y}$  that is orthogonal to that column space.

$$(y - \hat{y})'X = 0$$
  $\Leftrightarrow$   $X'(y - X\beta) = 0$   $\Leftrightarrow$   $X'y = X'X\beta$ .

$$X'(y-X\beta)=0$$

$$\Leftrightarrow$$

$$X'y = X'X\beta.$$

#### Projection Theorem

The optimal approximation  $\hat{y}$  in the space spanned by the columns  $x_i$  of Xhas an error  $y - \hat{y}$  that is orthogonal to that column space.

$$\zeta = 0$$
  $\Leftrightarrow$ 

$$(y-\hat{y})'X=0$$
  $\Leftrightarrow$   $X'(y-X\beta)=0$   $\Leftrightarrow$   $X'y=X'X\beta.$ 

$$\Leftrightarrow$$

$$X'y = X'X\beta.$$

#### Normal equations

$$X'X\beta = X'y$$
.

#### Projection Theorem

The optimal approximation  $\hat{y}$  in the space spanned by the columns  $x_i$  of X has an error  $y - \hat{y}$  that is orthogonal to that column space. That is,

$$\Leftrightarrow$$

$$(y - \hat{y})'X = 0$$
  $\Leftrightarrow$   $X'(y - X\beta) = 0$   $\Leftrightarrow$   $X'y = X'X\beta$ .

$$\Leftrightarrow$$

$$X'y=X'X\beta.$$

#### Normal equations

$$X'X\beta = X'y$$
.

$$\hat{\beta} = (X'X)^{-1}X'y.$$

#### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

#### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

$$\hat{R}(f) = \hat{\mathbb{E}}_n \ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2.$$

### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$
  
$$\hat{R}(f) = \hat{\mathbb{E}}_{n}\ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^{n} (f(X_{i}) - Y_{i})^{2}.$$

When is the risk of the empirical risk minimizer  $\hat{f}$  close to the minimal risk?

#### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

$$\hat{R}(f) = \hat{\mathbb{E}}_{n}\ell(f(X), Y) = \frac{1}{n}\sum_{i=1}^{n}(f(X_{i}) - Y_{i})^{2}.$$

When is the risk of the empirical risk minimizer  $\hat{f}$  close to the minimal risk?

### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

$$\hat{R}(f) = \hat{\mathbb{E}}_{n}\ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^{n} (f(X_{i}) - Y_{i})^{2}.$$

When is the risk of the empirical risk minimizer  $\hat{f}$  close to the minimal risk?

It suffices if

the X come from a compact set,

#### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

$$\hat{R}(f) = \hat{\mathbb{E}}_{n}\ell(f(X), Y) = \frac{1}{n}\sum_{i=1}^{n}(f(X_{i}) - Y_{i})^{2}.$$

When is the risk of the empirical risk minimizer  $\hat{f}$  close to the minimal risk?

- the X come from a compact set,
- ② the  $Y_i$ s have tails that are not too heavy (e.g., sub-Gaussian),

### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

$$\hat{R}(f) = \hat{\mathbb{E}}_{n}\ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^{n} (f(X_{i}) - Y_{i})^{2}.$$

When is the risk of the empirical risk minimizer  $\hat{f}$  close to the minimal risk?

- 1 the X come from a compact set,
- ② the  $Y_i$ s have tails that are not too heavy (e.g., sub-Gaussian),
- $||\hat{\theta}||$  is not too large, and

### Risk versus empirical risk

$$R(f) = \mathbb{E}\ell(f(X), Y) = \mathbb{E}(f(X) - Y)^{2},$$

$$\hat{R}(f) = \hat{\mathbb{E}}_{n}\ell(f(X), Y) = \frac{1}{n} \sum_{i=1}^{n} (f(X_{i}) - Y_{i})^{2}.$$

When is the risk of the empirical risk minimizer  $\hat{f}$  close to the minimal risk?

- the X come from a compact set,
- ② the  $Y_i$ s have tails that are not too heavy (e.g., sub-Gaussian),
- $0 n \gg p$ .

#### Outline

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.
  - Maximum likelihood is least squares.
  - Distributions of parameter estimates.

Linear model

#### Linear model

Model the conditional distribution of Y given X = x as

$$P(Y|X=x) = \mathcal{N}(x'\beta, \sigma^2).$$

#### Linear model

Model the conditional distribution of Y given X = x as

$$P(Y|X=x) = \mathcal{N}(x'\beta, \sigma^2).$$

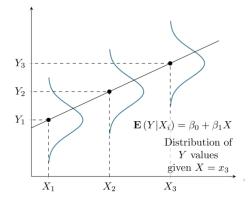
Equivalently:  $Y = x'\beta + \epsilon$ , where  $\epsilon \sim \mathcal{N}(0, \sigma^2)$ .

#### Linear model

Model the conditional distribution of Y given X = x as

$$P(Y|X = x) = \mathcal{N}(x'\beta, \sigma^2).$$

Equivalently:  $Y = x'\beta + \epsilon$ , where  $\epsilon \sim \mathcal{N}(0, \sigma^2)$ .



How to estimate  $\beta$ ?

How to estimate  $\beta$ ?

```
Maximum likelihood
```

How to estimate  $\beta$ ?

#### Maximum likelihood

Conditional likelihood:

$$L(\beta) = \prod_{i=1}^{n} p(Y_i|X_i,\beta) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - X_i'\beta)^2\right).$$

How to estimate  $\beta$ ?

#### Maximum likelihood

Conditional likelihood:

$$L(\beta) = \prod_{i=1}^{n} p(Y_i|X_i,\beta) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - X_i'\beta)^2\right).$$

Log likelihood:

$$\ell(\beta) = (\text{function of } \sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - X_i'\beta)^2.$$

How to estimate  $\beta$ ?

#### Maximum likelihood

Conditional likelihood:

$$L(\beta) = \prod_{i=1}^{n} p(Y_i|X_i, \beta) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - X_i'\beta)^2\right).$$

Log likelihood:

$$\ell(\beta) = (\text{function of } \sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - X_i'\beta)^2.$$

Maximum likelihood is least squares.

Bias and variance of  $\hat{eta}$ 

Fix X.

# Bias and variance of $\hat{\beta}$

Fix X. Provided  $\mathbb{E}y = X\beta$  and  $Cov(y) = \sigma^2 I$ ,

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta}$$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$
$$= (X'X)^{-1}X'\mathbb{E}y$$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$
$$= (X'X)^{-1}X'\mathbb{E}y$$
$$= (X'X)^{-1}X'X\beta$$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$
$$= (X'X)^{-1}X'\mathbb{E}y$$
$$= (X'X)^{-1}X'X\beta$$
$$= \beta.$$

## Bias and variance of $\hat{\beta}$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$
$$= (X'X)^{-1}X'\mathbb{E}y$$
$$= (X'X)^{-1}X'X\beta$$
$$= \beta.$$

 $Cov(\hat{\beta})$ 

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$

$$= (X'X)^{-1}X'\mathbb{E}y$$

$$= (X'X)^{-1}X'X\beta$$

$$= \beta.$$

$$\operatorname{Cov}(\hat{\beta}) = \mathbb{E}\left[ (\hat{\beta} - \beta)(\hat{\beta} - \beta)' \right]$$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$

$$= (X'X)^{-1}X'\mathbb{E}y$$

$$= (X'X)^{-1}X'X\beta$$

$$= \beta.$$

$$\operatorname{Cov}(\hat{\beta}) = \mathbb{E}\left[ (\hat{\beta} - \beta)(\hat{\beta} - \beta)' \right]$$

$$= \mathbb{E}\left[ ((X'X)^{-1}X'y - \beta) ((X'X)^{-1}X'y - \beta)' \right]$$

Fix X. Provided 
$$\mathbb{E}y = X\beta$$
 and  $Cov(y) = \sigma^2 I$ ,

$$\mathbb{E}\hat{\beta} = \mathbb{E}\left[ (X'X)^{-1}X'y \right]$$

$$= (X'X)^{-1}X'\mathbb{E}y$$

$$= (X'X)^{-1}X'X\beta$$

$$= \beta.$$

$$\operatorname{Cov}(\hat{\beta}) = \mathbb{E}\left[ (\hat{\beta} - \beta)(\hat{\beta} - \beta)' \right]$$

$$= \mathbb{E}\left[ ((X'X)^{-1}X'y - \beta) ((X'X)^{-1}X'y - \beta)' \right]$$

$$\vdots$$

$$= \sigma^{2}(X'X)^{-1}.$$



#### Statistical tests

If the data is generated by a linear model with  $\mathcal{N}(0,\sigma^2)$  noise,

#### Statistical tests

If the data is generated by a linear model with  $\mathcal{N}(0, \sigma^2)$  noise, then:

#### Statistical tests

If the data is generated by a linear model with  $\mathcal{N}(0, \sigma^2)$  noise, then:

• We can compute distributions of parameter estimates:

$$\hat{\beta} \sim \mathcal{N}(\beta, (X'X)^{-1}\sigma^2)$$

#### Statistical tests

If the data is generated by a linear model with  $\mathcal{N}(0, \sigma^2)$  noise, then:

• We can compute distributions of parameter estimates:

$$\hat{\beta} \sim \mathcal{N}(\beta, (X'X)^{-1}\sigma^2)$$

• We can calculate approximate confidence sets for the parameters: the standardized coefficient is

$$z_j = \frac{\hat{\beta}_j}{\sigma \sqrt{\mathsf{v}_j}},$$

which is normal (here,  $v_i$  is the jth diagonal entry of  $(X'X)^{-1}$ ).

#### Statistical tests

If the data is generated by a linear model with  $\mathcal{N}(0, \sigma^2)$  noise, then:

We can compute distributions of parameter estimates:

$$\hat{\beta} \sim \mathcal{N}(\beta, (X'X)^{-1}\sigma^2)$$

 We can calculate approximate confidence sets for the parameters: the standardized coefficient is

$$z_j = \frac{\hat{\beta}_j}{\sigma \sqrt{\mathbf{v}_j}},$$

which is normal (here,  $v_j$  is the *j*th diagonal entry of  $(X'X)^{-1}$ ).

• In particular, we can design tests for non-zero values of parameters.

#### Statistical tests

If the data is generated by a linear model with  $\mathcal{N}(0, \sigma^2)$  noise, then:

We can compute distributions of parameter estimates:

$$\hat{\beta} \sim \mathcal{N}(\beta, (X'X)^{-1}\sigma^2)$$

• We can calculate approximate confidence sets for the parameters: the standardized coefficient is

$$z_j = \frac{\hat{\beta}_j}{\sigma \sqrt{\mathsf{v}_j}},$$

which is normal (here,  $v_j$  is the jth diagonal entry of  $(X'X)^{-1}$ ).

• In particular, we can design tests for non-zero values of parameters.

#### Outline

- Review: Decision theory.
- Empirical risk minimization.
  - Least squares.
  - Normal equations.
- Linear model with additive Gaussian noise.
  - Maximum likelihood is least squares.
  - Distributions of parameter estimates.