

Applied Bayesian Modeling module 4:

Deriving a posterior, role of prior information

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Recap of module 3:

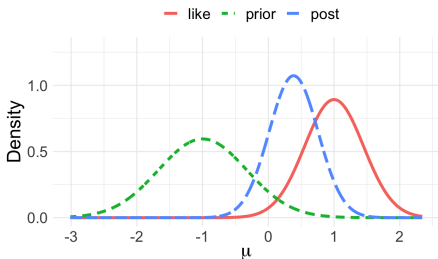
- ▶ Bayesian inference about a parameter μ , using data \mathbf{y} :
 - (1) relate data \mathbf{y} to μ through a likelihood function $p(\mathbf{y}|\mu)$
 - (2) set a prior distribution for μ , $p(\mu)$
 - (3) use Bayes' rule to get posterior distribution $p(\mu|\mathbf{y}) = \frac{p(\mu)p(\mathbf{y}|\mu)}{p(\mathbf{y})}$,
 - (4) use the posterior $p(\mu|\mathbf{y})$ to provide summaries of interest, e.g. point estimates and uncertainty intervals, called credible intervals (CIs).
- ▶ Doing this for a normal-normal set-up:

$$y_i|\mu, \sigma^2 \sim N(\mu, \sigma^2)(\text{independent}); \mu \sim N(m_0, s_{\mu 0}^2);$$
$$\Rightarrow \mu|\mathbf{y}, \sigma^2 \sim N\left(\frac{m_0/s_{\mu 0}^2 + n \cdot \bar{y}/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2}, \frac{1}{1/s_{\mu 0}^2 + n/\sigma^2}\right).$$

- ▶ This module: derive the posterior, discuss the role of prior info (the module with the most equations)
- ▶ Note: conditioning on σ is left out from equations to make it easier to focus on main points, σ^2 assumed known throughout.

How did the normal posterior come about?

- ▶ Bayes' rule: $p(\mu|\mathbf{y}) = \frac{p(\mu)p(\mathbf{y}|\mu)}{p(\mathbf{y})}$
- ▶ Good news: To recognize the posterior as a normal distribution (or more generally, a specific parametric form for the density), we only have to consider the terms that include μ : $p(\mu|\mathbf{y}) \propto p(\mu)p(\mathbf{y}|\mu)$



- ▶ $p(\mu|\mathbf{y})$ is a probability density function with $\int p(\mu|\mathbf{y})d\mu = 1$,
- ▶ what matters is how $p(\mu|\mathbf{y})$ depends on μ
- ▶ $p(\mathbf{y})$ does NOT vary with μ and thus ends up being just a scaling factor in Bayes' rule.

How did the normal posterior come about? (ctd)

- ▶ We work towards finding m and v in

$$\mu|\mathbf{y}, \sigma^2 \sim N(m, v),$$

so towards

$$p(\mu|\mathbf{y}) \propto \exp\left(-\frac{1}{2}f(\mu)\right),$$

where

$$f(\mu) = \frac{1}{v}(\mu - m)^2 \propto \frac{1}{v}(\mu^2 - 2m\mu).$$

- ▶ Re-arranging terms in the prior and likelihood will get you to that form, and thus m and v
- ▶ We will use that summation in log-likelihood can be rewritten as follows:

$$\sum_{i=1}^n (y_i - \mu)^2 \propto \sum y_i \cdot \mu + n\mu^2 \propto n(\bar{y} - \mu)^2.$$

Let's do it!

- Starting with Bayes' rule:

$$\begin{aligned} p(\mu|\mathbf{y}) &\propto p(\mu)p(\mathbf{y}|\mu) \\ &\propto \exp\left(\frac{-1}{2s_{\mu 0}^2}(\mu - m_0)^2\right) \cdot \exp\left(\frac{-1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right), \\ &\propto \exp\left(\frac{-1}{2s_{\mu 0}^2}(\mu - m_0)^2\right) \cdot \exp\left(\frac{-n}{2\sigma^2}(\bar{y} - \mu)^2\right) \\ &\propto \exp\left(-\frac{1}{2}f(\mu)\right), \end{aligned}$$

where

$$f(\mu) = 1/s_{\mu 0}^2(\mu^2 - 2 \cdot m_0 \cdot \mu) + n/\sigma^2(\mu^2 - 2 \cdot \bar{y} \cdot \mu).$$

Re-arranging terms, toward $f(\mu) \propto \frac{1}{v}(\mu^2 - 2m\mu)$

► We have

$$\begin{aligned} f(\mu) &= 1/s_{\mu 0}^2(\mu^2 - 2 \cdot m_0 \cdot \mu) + n/\sigma^2(\mu^2 - 2 \cdot \bar{y} \cdot \mu) \\ &\propto (1/s_{\mu 0}^2 + n/\sigma^2)\mu^2 - 2 \cdot (1/s_{\mu 0}^2 m_0 + n/\sigma^2 \bar{y}) \cdot \mu, \end{aligned}$$

► Let $v = (1/s_{\mu 0}^2 + n/\sigma^2)^{-1}$, then

$$f(\mu) \propto 1/v(\mu^2 - v \cdot 2 \cdot (1/s_{\mu 0}^2 m_0 + n/\sigma^2 \bar{y})\mu).$$

► Comparing this to $f(\mu) \propto \frac{1}{v}(\mu^2 - 2m\mu)$, we recognize that

$$\mu|\mathbf{y} \sim N\left(\frac{m_0/s_{\mu 0}^2 + n \cdot \bar{y}/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2}, \frac{1}{1/s_{\mu 0}^2 + n/\sigma^2}\right).$$

Summary so far

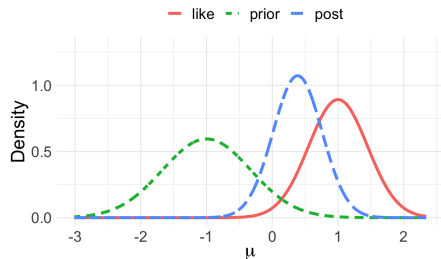
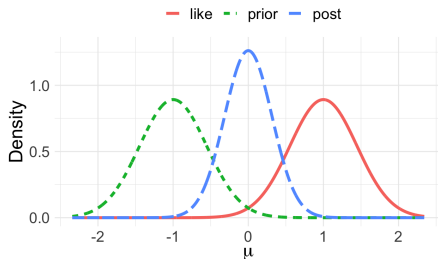
- ▶ Bayes' rule simplifies to: $p(\mu|\mathbf{y}) \propto p(\mu)p(\mathbf{y}|\mu)$
- ▶ We derived the posterior in the normal-normal set-up:

$$\begin{aligned}y_i|\mu, \sigma^2 &\sim N(\mu, \sigma^2)(\text{independent}); \\ \mu &\sim N(m_0, s_{\mu 0}^2); \\ \Rightarrow \mu|\mathbf{y}, \sigma^2 &\sim N\left(\frac{m_0/s_{\mu 0}^2 + n \cdot \bar{y}/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2}, \frac{1}{1/s_{\mu 0}^2 + n/\sigma^2}\right).\end{aligned}$$

- ▶ Now let's examine the expressions for the posterior mean and variance

Posterior variance

- Interpretation is easiest in terms of precision = $1/\text{variance}$
- Posterior precision = $1/s_{\mu 0}^2 + n/\sigma^2$;
This is the sum of the prior precision and the precision of \bar{y}
- Examples: same likelihood with $\bar{y} = 1$; prior differs



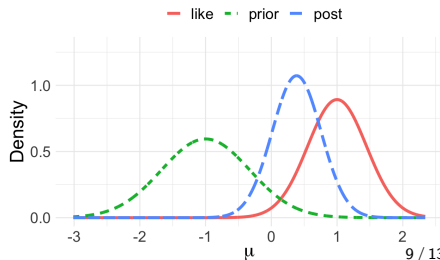
Posterior mean

- Rewrite as weighted combination of prior mean and \bar{y} :

$$E(\mu|\mathbf{y}) = \frac{m_0/s_{\mu 0}^2 + n \cdot \bar{y}/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2} = \frac{1/s_{\mu 0}^2}{1/s_{\mu 0}^2 + n/\sigma^2} m_0 + \frac{n/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2} \bar{y},$$

with the weight assigned to the prior mean (sample mean) given by the share of prior (sample mean) precision out of the overall precision.

- Examples: same likelihood with $\bar{y} = 1$; prior differs



If helpful, an informal way to see how we arrive at the posterior

- ▶ Posterior \propto prior \cdot likelihood
- ▶ What shows up in the posterior after the rewrite of the likelihood?

$$p(\mu|\mathbf{y}) \propto \exp\left(\frac{-1}{2s_{\mu 0}^2}(\mu - m_0)^2\right) \cdot \exp\left(\frac{-n}{2\sigma^2}(\mu - \bar{y})^2\right).$$

- ▶ The likelihood contribution to the posterior is equivalent to a contribution of “ $\mu \sim N(\bar{y}, \sigma^2/n)$ ”
- ▶ So posterior can be considered as arising from product of $\mu \sim N(m_0, s_{\mu 0}^2)$ and $\mu \sim N(\bar{y}, \sigma^2/n)$
- ▶ Posterior precision = $1/s_{\mu 0}^2 + n/\sigma^2$, sum of the prior precision and precision coming from contribution of the likelihood
- ▶ Posterior mean = precision-weighted means

Summary (ctd)

- ▶ Bayes' rule simplifies to: $p(\mu|\mathbf{y}) \propto p(\mu)p(\mathbf{y}|\mu)$
- ▶ We derived the posterior in the normal-normal set-up:

$$y_i|\mu, \sigma^2 \sim N(\mu, \sigma^2)(\text{independent}); \quad (1)$$

$$\mu \sim N(m_0, s_{\mu 0}^2); \quad (2)$$

$$\Rightarrow \mu|\mathbf{y}, \sigma^2 \sim N\left(\frac{m_0/s_{\mu 0}^2 + n \cdot \bar{y}/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2}, \frac{1}{1/s_{\mu 0}^2 + n/\sigma^2}\right). \quad (3)$$

- ▶ Posterior mean and variance determined by prior and sample mean and variance

Sampling properties of Bayesian estimators

- ▶ A Bayesian point estimate for μ is given by $E(\mu|\mathbf{y})$.
 - ▶ What are the sampling properties of this estimator, and how do they compare to those of the maximum likelihood estimator \bar{y} ?
 - ▶ Estimator = $E(\mu|\mathbf{y})$ as a function of the yet-to-be-observed data
 - ▶ Sampling properties = behavior under hypothetically repeatable surveys or experiments
 - ▶ Consider bias and mean squared error (MSE):
 - ▶ Bias = expected value of estimator - (unknown) true value
 - ▶ MSE = squared difference between expected value and the (unknown) true value; $\text{MSE} = \text{bias}^2 + \text{variance of estimator}$
 - ▶ In a HW (optional) extra credit question, we show that in a normal-normal setting
 - ▶ Bayes estimator is biased if prior mean \neq true value, \bar{y} is unbiased
 - ▶ But prior information can result in smaller MSE for Bayes estimator, compared to \bar{y}
- ⇒ Illustration of bias-variance trade off of using prior information

Summary and outlook

- ▶ Bayes' rule simplifies to: $p(\mu|\mathbf{y}) \propto p(\mu)p(\mathbf{y}|\mu)$
- ▶ We derived the posterior in the normal-normal set-up:

$$\begin{aligned}y_i|\mu, \sigma^2 &\sim N(\mu, \sigma^2)(\text{independent}); \\ \mu &\sim N(m_0, s_{\mu 0}^2); \\ \Rightarrow \mu|\mathbf{y}, \sigma^2 &\sim N\left(\frac{m_0/s_{\mu 0}^2 + n \cdot \bar{y}/\sigma^2}{1/s_{\mu 0}^2 + n/\sigma^2}, \frac{1}{1/s_{\mu 0}^2 + n/\sigma^2}\right),\end{aligned}$$

and discussed how the posterior depends on prior information and data.

- ▶ Next: sampling-based approach to do inference in settings where posterior is available in closed form