be used to approximate the distribution function of  $S_t$ , and can be used to estimate probabilities, including tail probabilities.

Being finite-sample estimates, these probabilities should have a measure of uncertainty attached. This is obviously an issue for regulation, where the requirement is often to demonstrate that

$$\Pr(S_1 \ge s_0) \le \kappa_0$$

for some  $s_0$  which reflects the insurer's available capital, and some  $\kappa_0$  specified by the regulator. For Solvency II,  $\kappa_0 = 0.005$  for one-year total losses. A Monte Carlo point estimate of  $p_0 := \Pr(S_1 \geq s_0)$  which was less than  $\kappa_0$  would be much more reassuring if the whole of the 95% confidence interval for  $p_0$  were less than  $\kappa_0$ , than if the 95% confidence interval contained  $\kappa_0$ .

A similar problem is faced in ecotoxicology, where one recommendation would be equivalent in this context to requiring that the upper bound of a 95% confidence interval for  $p_0$  is no greater than  $\kappa_0$ ; see Hickey and Hart (2013). If we adopt this approach, though, it is incorrect simply to monitor the upper bound and stop sampling when it drops below  $\kappa_0$ , because the confidence interval in this case ought to account for the stochastic stopping rule, rather than being based on a fixed sample size. But it is possible to do a design calculation to suggest an appropriate value for n, the sample size, that will ensure that the upper bound will be larger than  $\kappa_0$  with specified probability, a priori, as we now discuss.

Let  $u_{1-\alpha}(x;n)$  be the upper limit of a level  $(1-\alpha)$  confidence interval for  $p_0$ , where x is the number of sample members that are at least  $s_0$ , and n