

STDA Homework 1

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Note:

You can get full, run-able code files at my github page: Visit <https://github.com/letsjdosth/SpaTempoDA>. Because I provide the full code file separately, in this report, I will show key-code block only instead of bringing the full, verbose code.

1 Problem 1

Suppose we want to simulate a random vector $Y \sim N(\mu, \Sigma)$. If Σ (Matern) is symmetric and positive definite, it can be represented using the Cholesky decomposition $\Sigma = LL'$, where L is a lower triangular matrix. Consider the following algorithm for simulating Y :

- Calculate the matrix L .
- Sample $Z \sim N(0, I)$, where I is the $n \times n$ identity matrix.
- Let $Y = \mu + LZ$

1.1 Problem 1-(a)

Show that Y generated in this way has the correct distribution. You may use the fact that a linear function of a multivariate normal random variable is again multivariate normal; just show the mean and variance are correct.

From the elementary observation from mathematical statistics, we have for $a, b \in \mathcal{R}$ and for a random variable X ,

$$E[aX + b] = aE[X] + b$$

$$Var[aX + b] = a^2 Var[X]$$

And its multivariate version is, for proper $A \in \mathcal{R}^{n \times n}$, $b \in \mathcal{R}^n$, and for a random vector X ,

$$E[AX + b] = AE[X] + b$$

$$Var[AX + b] = AVar[X]A'$$

Then, using this observation and the fact at this problem, we can get

$$Y \sim N(\mu + 0, LIL')$$

Since $LL' = \Sigma$ by our setting, we get $Y \sim N(\mu, \Sigma)$, so we can confirm that the algorithm is valid.

1.2 Problem 1-(b)

Write a function or a few lines of code in R to implement this method for arguments `mu` and `Sigma`. You may use the built-in function `chol` for the Cholesky decomposition and `rnorm` to generate Z .

Here is implementation. Also you may see the 'HW1sol_problem.R' file.

```
#R code

rnorm_chol_algorithm = function(mu, Sigma) {
  # mu is vector
  # Sigma should be symmetric, positive definite matrix
  dimension = dim(Sigma)
  vecZ = rnorm(dimension, 0, 1)
  chol_upperL = chol(Sigma)
  return(mu + t(chol_upperL) %*% vecZ)
}
```

1.3 Problem 1-(c)

For a mean and covariance function of your choosing, use your code from (b) and make a few plots illustrating realizations of a Gaussian process on $[0;1]$, but changing the different parameters in the model. These differences will be easier to see if you keep the same Z sample but just change μ and Σ .

For convenience, I use `fields$Matern` function to calculate covariance matrix from parameters of Matern. Here are 3 test cases. (In full code, there are one more case for ordinary(not Matern) GP.)

- $\nu = 0.1$ (in code, test 3)
- $\nu = 0.5$ (in code, test 2)
- $\nu = 1$ (in code, test 4)

Note that $\nu = 0.5$ case is the exponential covariance.

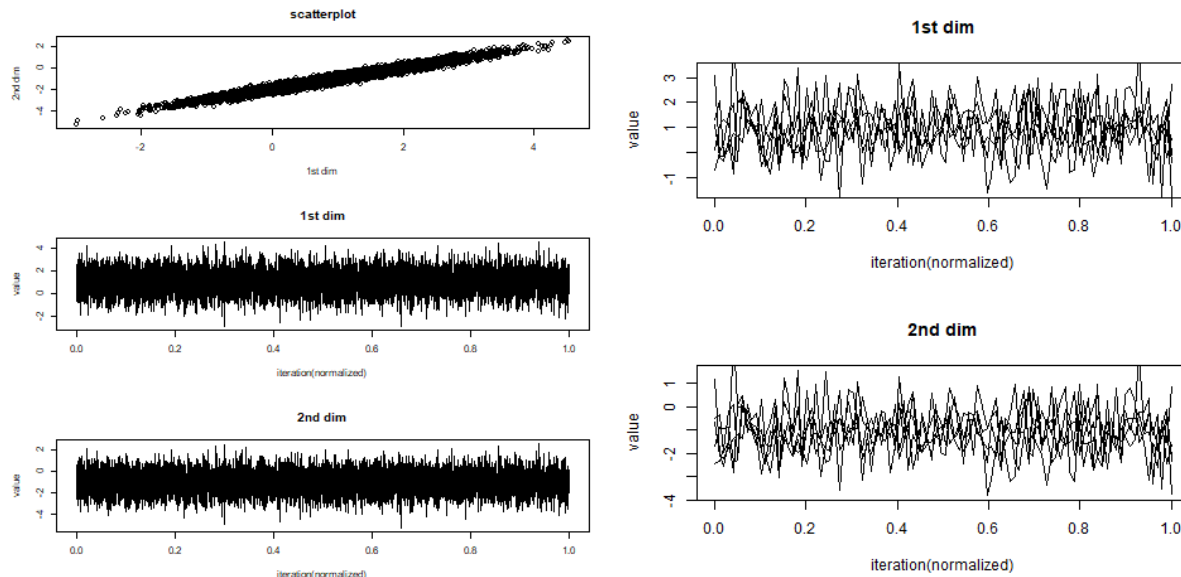


Figure 1: $\nu = 0.1$

2 Problem 2

The file `CAtemps.RData` contains two R objects of class `SpatialPointsDataFrame`, called `CAtemp` and `CAgrid`. `CAtemp` contains a average temperatures from 1961-1990 at 200 locations (latitude and longitude) in

California in degrees Fahrenheit, along with their elevation in meters. CAgrid contains elevations in meters over a grid of locations. I've given you some code to get started with this data in HW1.R

Consider the following model for the temperature data.

$$Y_i = \mu(s_i; \beta) + e(s_i; \sigma^2, \rho, \tau)$$

where $\mu(s_i; \beta) = \beta_0 + \beta_1 \text{Longitude}(s) + \beta_2 \text{Latitude}(s) + \beta_3 \text{Elevation}(s)$ and $e(s_i; \sigma^2, \rho, \tau)$ is a zero mean stationary Gaussian process with exponential covariance function.

Another way of writing this is as

$$Y_i = \mu(s_i; \beta) + Z(s_i; \sigma^2, \rho) + \epsilon_i$$

where now Z is mean zero Gaussian process like e but without the nugget term, and the ϵ are iid $N(0, \tau^2)$, independent of Z . This is important because we want to predict $\mu(s_i; \beta) + Z(s_i; \sigma^2, \rho)$ without the measurement error.

2.1 Problem 2-(a)

Using the CAtemp data, form a preliminary estimate of β using ordinary least squares and make a color plot of the residuals. Include your estimates and plot.

2.2 Problem 2-(b)

Estimate the variogram non-parametrically and then fit the exponential variogram to it using weighted least squares. Make and include a plot of the nonparametric and parametric variogram functions. Also store your parameter estimates and report them.

2.3 Problem 2-(c)

We will now form the GLS estimate of β by hand, rather than using the gls function.

- Use the `rdist` function in `fields` to create a matrix of distance (in miles) between pairs of locations in CAtemp.
- Create the covariance matrix, plugging in your estimates from the fitted variogram. (Hint: Sum tow matrix, one without a nugget and one using the `diag` function to create the matrix $\tau^2 I$)
- Invert the covariance matrix and store it for later reference.
- Create the X matrix (Hint: Use `cbind`.)
- Put all the pieces together to form $\hat{\beta}_{GLS}$.

2.4 Problem 2-(d)

Calculate and plot the EBLUP of $\mu + Z$ at the location in CAgrid, plugging in your estimates from (b) and (c). Calculate and plot the (estimated) standard error of Z at each prediction location

Since $Z \sim N(0, C)$, by definition of functional distribution in weak sense, $\langle Z, x \rangle \sim N(\langle 0, x \rangle, \langle C(x), x \rangle)$ for all $x \in \mathcal{H} = \mathcal{L}^2([0, 1])$. So, with eigenfunctions $\{v_i\}$, we get multivariate normal distribution for

$$\{\langle Z, v_i \rangle\}_{i=1,2,\dots,m} \sim \text{Normal}_m \left(\begin{bmatrix} \langle 0, v_1 \rangle \\ \langle 0, v_2 \rangle \\ \dots \\ \langle 0, v_m \rangle \end{bmatrix}, \begin{bmatrix} \langle C(v_1), v_1 \rangle & \langle C(v_1), v_2 \rangle & \dots & \langle C(v_1), v_m \rangle \\ \langle C(v_2), v_1 \rangle & \langle C(v_2), v_2 \rangle & \dots & \langle C(v_2), v_m \rangle \\ \dots & \dots & \dots & \dots \\ \langle C(v_m), v_1 \rangle & \langle C(v_m), v_2 \rangle & \dots & \langle C(v_m), v_m \rangle \end{bmatrix} \right)$$

For simplicity, denote above expression's covariance matrix as Σ_m . Then we simply write above as

$$\{\langle Z, v_i \rangle\}_{i=1,2,\dots,m} \sim Normal_m(0, \Sigma_m)$$

Since the pdf of multivariate normal is well known, I write it without additional explanation. If denote the vector $[\langle Z, v_i \rangle]_i, i = 1, 2, \dots, m$ as z_m , then

$$f_m(z_m) = \frac{1}{(\sqrt{2\pi})^m \det(\Sigma_m)} \exp\left(-\frac{1}{2}(z_m - 0)^T \Sigma_m^{-1} (z_m - 0)\right)$$

Likewise, since $X := \mu + Z$,

$$\{\langle X, v_i \rangle\}_{i=1,2,\dots,m} \sim Normal_m\left(\begin{bmatrix} \langle \mu, v_1 \rangle \\ \langle \mu, v_2 \rangle \\ \dots \\ \langle \mu, v_m \rangle \end{bmatrix}, \begin{bmatrix} \langle C(v_1), v_1 \rangle & \langle C(v_1), v_2 \rangle & \dots & \langle C(v_1), v_m \rangle \\ \langle C(v_2), v_1 \rangle & \langle C(v_2), v_2 \rangle & \dots & \langle C(v_2), v_m \rangle \\ \dots & \dots & \dots & \dots \\ \langle C(v_m), v_1 \rangle & \langle C(v_m), v_2 \rangle & \dots & \langle C(v_m), v_m \rangle \end{bmatrix}\right)$$

with denoting the mean vector as μ_m covariance matrix as Σ_m . Then we simply write above as

$$\{\langle X, v_i \rangle\}_{i=1,2,\dots,m} \sim Normal_m(\mu_m, \Sigma_m)$$

and pdf is, where $x_m = [\langle X, v_i \rangle]_i, i = 1, 2, \dots, m$,

$$f_m(x_m) = \frac{1}{(\sqrt{2\pi})^m \det(\Sigma_m)} \exp\left(-\frac{1}{2}(x_m - \mu_m)^T \Sigma_m^{-1} (x_m - \mu_m)\right)$$