Chapter 05: Resampling Methods

Solutions to Exercises

February 01, 2023

CONCEPTUAL

EXERCISE 1:

$$\begin{split} Var(\alpha X + (1-\alpha)Y) &= Var(\alpha X) + Var((1-\alpha)Y) + 2Cov(\alpha X, (1-\alpha)Y) \\ &= \alpha^2 \sigma_X^2 + (1-\alpha)^2 \sigma_Y^2 + 2\alpha(1-\alpha)\sigma_{XY} \\ &= \alpha^2 \sigma_X^2 + (1+\alpha^2-2\alpha)\sigma_Y^2 + (2\alpha-2\alpha^2)\sigma_{XY} \\ &= \alpha^2 \sigma_X^2 + \sigma_Y^2 + \alpha^2 \sigma_Y^2 - 2\alpha\sigma_Y^2 + 2\alpha\sigma_{XY} - 2\alpha^2\sigma_{XY} \\ &\frac{\partial}{\partial \alpha} : 2\alpha\sigma_X^2 + 0 + 2\alpha\sigma_Y^2 - 2\sigma_Y^2 + 2\sigma_{XY} - 4\alpha\sigma_{XY} = 0 \\ &(2\sigma_X^2 + 2\sigma_Y^2 - 4\sigma_{XY})\alpha = 2\sigma_Y^2 - 2\sigma_{XY} \\ &\alpha = \frac{\sigma_Y^2 - \sigma_{XY}}{\sigma_X^2 + \sigma_Y^2 - 2\sigma_{XY}} \end{split}$$

EXERCISE 2:

Part a)

Probability is equal to not selecting that one observation out of all observations: $\frac{n-1}{n}$

Part b)

Because bootstrap uses replacement, the probability is the same as Part a: $\frac{n-1}{n}$

Part c)

Probability of not selecting the jth observation is the same for each selection. After n selections, the probability of never selecting the jth observation is: $(\frac{n-1}{n})^n = (1-\frac{1}{n})^n$

Part d)

1-(1-1/5)^5

[1] 0.67232

Part e)

```
1-(1-1/100)^100
```

[1] 0.6339677

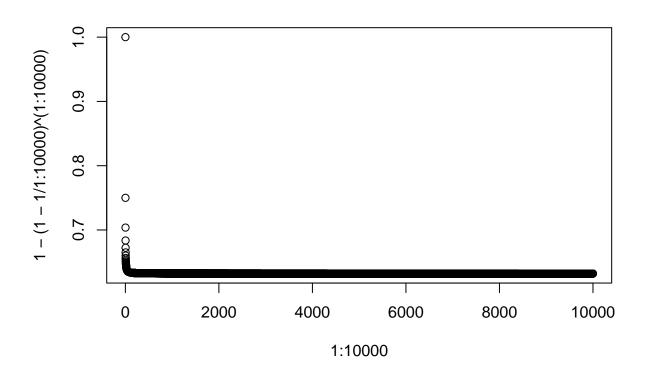
Part f)

```
1-(1-1/10000)^10000
```

[1] 0.632139

Part g)

```
plot(1:10000, 1-(1-1/1:10000)^(1:10000))
```



Probability pretty quickly reaches mid-60%

Part h)

```
store <- rep(NA, 10000)
for (i in 1:10000)
  store[i] <- sum(sample(1:100, rep=TRUE)==4) > 0
mean(store)
```

[1] 0.636

The resulting fraction of 10,000 bootstrap samples that have the 4th observation is close to our predicted probability of $1-(1-1/100)^2100 = 63.4\%$

EXERCISE 3:

Part a)

From page 181 in the text, the k-fold CV approach "involves randomly dividing the set of observations into k groups, or folds, of approximately equal size. The first fold is treated as a validation set, and the method is fit on the remaining k-1 folds. The mean squared error, MSE, is then computed on the observations in the held-out fold. This procedure is repeated k times."

Part b)

- Compared to the validation set approach, k-fold CV has less variance but more bias
- Compared to LOOCV approach, k-fold CV has more variance but less bias

EXERCISE 4:

We can use the bootstrap method to sample with replacement from our dataset and estimate Y's from each sample. With the results of different predicted Y values, we can then estimate the standard deviation of our prediction.

APPLIED

EXERCISE 5:

Part a)

```
require(ISLR2)
data(Default)
set.seed(1)
fit1 <- glm(default ~ income + balance, data=Default, family=binomial)
summary(fit1)</pre>
```

```
##
## Call:
## glm(formula = default ~ income + balance, family = binomial,
       data = Default)
## Deviance Residuals:
                     Median
                10
                                   30
                                           Max
## -2.4725 -0.1444 -0.0574 -0.0211
                                        3.7245
##
## Coefficients:
                 Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.154e+01 4.348e-01 -26.545 < 2e-16 ***
               2.081e-05 4.985e-06 4.174 2.99e-05 ***
## balance
               5.647e-03 2.274e-04 24.836 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for binomial family taken to be 1)
##
##
      Null deviance: 2920.6 on 9999 degrees of freedom
## Residual deviance: 1579.0 on 9997 degrees of freedom
## AIC: 1585
##
## Number of Fisher Scoring iterations: 8
Part b)
set.seed(1)
train <- sample(nrow(Default), nrow(Default)*0.5)</pre>
fit2 <- glm(default ~ income + balance, data=Default, family=binomial, subset=train)
prob2 <- predict(fit2, Default[-train,], type="response")</pre>
pred2 <- ifelse(prob2 > 0.5, "Yes", "No")
table(pred2, Default[-train,]$default)
##
## pred2
          No
              Yes
##
    No 4824
               108
##
     Yes
           19
                49
mean(Default[-train,]$default != pred2) # test error
## [1] 0.0254
Part c)
set.seed(2) # Repeat 1
train <- sample(nrow(Default), nrow(Default)*0.5)
fit2 <- glm(default ~ income + balance, data=Default, family=binomial, subset=train)
prob2 <- predict(fit2, Default[-train,], type="response")</pre>
pred2 <- ifelse(prob2 > 0.5, "Yes", "No")
mean(Default[-train,]$default != pred2) # test error
```

```
## [1] 0.0238
```

Call:

```
set.seed(3) # Repeat 2
train <- sample(nrow(Default), nrow(Default)*0.5)</pre>
fit2 <- glm(default ~ income + balance, data=Default, family=binomial, subset=train)</pre>
prob2 <- predict(fit2, Default[-train,], type="response")</pre>
pred2 <- ifelse(prob2 > 0.5, "Yes", "No")
mean(Default[-train,]$default != pred2) # test error
## [1] 0.0264
set.seed(4) # Repeat 3
train <- sample(nrow(Default), nrow(Default)*0.5)</pre>
fit2 <- glm(default ~ income + balance, data=Default, family=binomial, subset=train)</pre>
prob2 <- predict(fit2, Default[-train,], type="response")</pre>
pred2 <- ifelse(prob2 > 0.5, "Yes", "No")
mean(Default[-train,]$default != pred2) # test error
## [1] 0.0256
The test error seems consistent around 2.5% (variance is not large)
Part d)
set.seed(1)
train <- sample(nrow(Default), nrow(Default)*0.5)</pre>
fit3 <- glm(default ~ income + balance + student, data=Default, family=binomial, subset=train)</pre>
prob3 <- predict(fit3, Default[-train,], type="response")</pre>
pred3 <- ifelse(prob3 > 0.5, "Yes", "No")
mean(Default[-train,]$default != pred3) # test error
## [1] 0.026
Test error with the student feature included is similar to without including student (no significant reduc-
tion)
EXERCISE 6:
Part a)
require(ISLR2)
data(Default)
set.seed(1)
fit1 <- glm(default ~ income + balance, data=Default, family=binomial)</pre>
summary(fit1)
```

glm(formula = default ~ income + balance, family = binomial,

```
##
      data = Default)
##
## Deviance Residuals:
      Min 1Q Median
##
                                3Q
                                          Max
## -2.4725 -0.1444 -0.0574 -0.0211
##
## Coefficients:
##
                Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.154e+01 4.348e-01 -26.545 < 2e-16 ***
## income
               2.081e-05 4.985e-06 4.174 2.99e-05 ***
## balance
               5.647e-03 2.274e-04 24.836 < 2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for binomial family taken to be 1)
##
##
      Null deviance: 2920.6 on 9999 degrees of freedom
## Residual deviance: 1579.0 on 9997 degrees of freedom
## AIC: 1585
##
## Number of Fisher Scoring iterations: 8
Estimated standard error is 0.000004985 for income and 0.0002274 for balance
Part b)
set.seed(1)
boot.fn <- function(df, trainid) {</pre>
  return(coef(glm(default ~ income + balance, data=df, family=binomial, subset=trainid)))
boot.fn(Default, 1:nrow(Default)) # check match with summary
     (Intercept)
                       income
                                    balance
## -1.154047e+01 2.080898e-05 5.647103e-03
Part c)
require(boot)
##
            : boot
boot(Default, boot.fn, R=100)
##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
##
## Call:
## boot(data = Default, statistic = boot.fn, R = 100)
##
##
## Bootstrap Statistics :
```

```
## original bias std.error
## t1* -1.154047e+01 8.556378e-03 4.122015e-01
## t2* 2.080898e-05 -3.993598e-07 4.186088e-06
## t3* 5.647103e-03 -4.116657e-06 2.226242e-04
```

Part d)

Standard error estimates are pretty close using glm summary function versus bootstrap with R=100

- income: 4.985e-06 with glm summary, 4.128e-06 using bootstrap
- balance: 2.274e-04 with glm summary, 2.106e-04 using bootstrap

EXERCISE 7:

Part a)

```
require(ISLR2)
data(Weekly)
set.seed(1)
fit1 <- glm(Direction ~ Lag1 + Lag2, data=Weekly, family=binomial)
summary(fit1)</pre>
```

```
##
## Call:
## glm(formula = Direction ~ Lag1 + Lag2, family = binomial, data = Weekly)
## Deviance Residuals:
##
     Min
              1Q Median
                              3Q
                                     Max
## -1.623 -1.261 1.001
                           1.083
                                   1.506
##
## Coefficients:
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) 0.22122
                          0.06147
                                    3.599 0.000319 ***
              -0.03872
                          0.02622 -1.477 0.139672
## Lag1
               0.06025
                                    2.270 0.023232 *
## Lag2
                          0.02655
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for binomial family taken to be 1)
##
      Null deviance: 1496.2 on 1088 degrees of freedom
## Residual deviance: 1488.2 on 1086 degrees of freedom
## AIC: 1494.2
## Number of Fisher Scoring iterations: 4
```

Part b)

```
set.seed(1)
fit2 <- glm(Direction ~ Lag1 + Lag2, data=Weekly, family=binomial, subset=2:nrow(Weekly))
summary(fit2)</pre>
```

```
##
## Call:
## glm(formula = Direction ~ Lag1 + Lag2, family = binomial, data = Weekly,
       subset = 2:nrow(Weekly))
##
## Deviance Residuals:
                     Median
       Min
            10
                                   30
                                           Max
## -1.6258 -1.2617 0.9999
                                         1.5071
                               1.0819
##
## Coefficients:
               Estimate Std. Error z value Pr(>|z|)
                                     3.630 0.000283 ***
## (Intercept) 0.22324
                           0.06150
                           0.02622 -1.466 0.142683
## Lag1
               -0.03843
               0.06085
                           0.02656
                                    2.291 0.021971 *
## Lag2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##
       Null deviance: 1494.6 on 1087 degrees of freedom
## Residual deviance: 1486.5 on 1085 degrees of freedom
## AIC: 1492.5
##
## Number of Fisher Scoring iterations: 4
Part c)
ifelse(predict(fit2, Weekly[1,], type="response")>0.5, "Up", "Down")
##
## "Up"
Weekly[1,]$Direction
## [1] Down
## Levels: Down Up
The first observation was incorrectly classified (predicted Up, actually Down)
Part d)
set.seed(1)
loocv.err <- rep(0,nrow(Weekly))</pre>
for (i in 1:nrow(Weekly)) {
 myfit <- glm(Direction ~ Lag1 + Lag2, data=Weekly[-i,], family=binomial)</pre>
 mypred <- ifelse(predict(myfit, Weekly[1,], type="response")>0.5, "Up", "Down")
 loocv.err[i] <- ifelse(Weekly[i,]$Direction==mypred, 0, 1)</pre>
str(loocv.err)
## num [1:1089] 1 1 0 0 0 1 0 0 0 1 ...
```

Part e)

mean(loocv.err)

[1] 0.444444

Estimated test error with LOOCV is 44.4%

EXERCISE 8:

Part a)

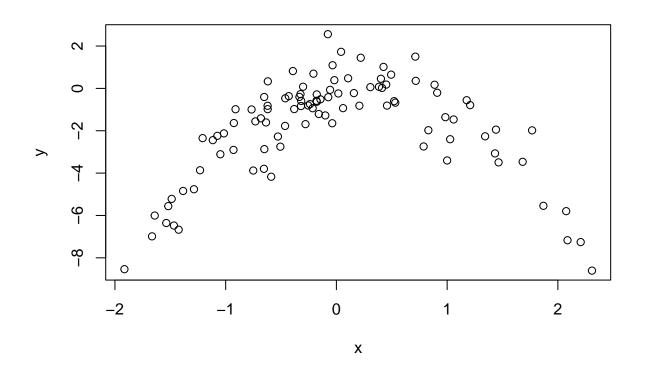
```
set.seed(1)
y <- rnorm(100) # why is this needed?
x <- rnorm(100)
y <- x - 2*x^2 + rnorm(100)</pre>
```

 $Y = X - 2X^2 + \epsilon$

n = 100 observations

p=2 features

Part b)



Relationship between X and Y is quadratic

```
Part c)
```

```
set.seed(1)
df <- data.frame(y, x, x2=x^2, x3=x^3, x4=x^4)
fit1 \leftarrow glm(y \sim x, data=df)
cv.err1 <- cv.glm(df, fit1)</pre>
cv.err1$delta
## [1] 5.890979 5.888812
fit2 \leftarrow glm(y \sim x + x2, data=df)
cv.err2 <- cv.glm(df, fit2)</pre>
cv.err2$delta
## [1] 1.086596 1.086326
fit3 <- glm(y \sim x + x2 + x3, data=df)
cv.err3 <- cv.glm(df, fit3)</pre>
cv.err3$delta
## [1] 1.102585 1.102227
fit4 \leftarrow glm(y \sim x + x2 + x3 + x4, data=df)
cv.err4 <- cv.glm(df, fit4)</pre>
cv.err4$delta
## [1] 1.114772 1.114334
Part d)
set.seed(2)
df \leftarrow data.frame(y, x, x2=x^2, x3=x^3, x4=x^4)
fit1 <- glm(y ~ x, data=df)</pre>
cv.err1 <- cv.glm(df, fit1)</pre>
cv.err1$delta
## [1] 5.890979 5.888812
fit2 <- glm(y \sim x + x2, data=df)
cv.err2 <- cv.glm(df, fit2)</pre>
cv.err2$delta
## [1] 1.086596 1.086326
fit3 \leftarrow glm(y \sim x + x2 + x3, data=df)
cv.err3 <- cv.glm(df, fit3)</pre>
cv.err3$delta
```

[1] 1.102585 1.102227

```
fit4 <- glm(y ~ x + x2 + x3 + x4, data=df)
cv.err4 <- cv.glm(df, fit4)
cv.err4$delta</pre>
```

[1] 1.114772 1.114334

Results are exactly the same because LOOCV predicts every observation using the all of the rest (no randomness involved)

Part e)

The quadratic model using X and X^2 had the lowest error. This makes sense because the true model was generated using a quadratic formula

Part f)

```
fit0 <- lm(y ~ poly(x,4))
summary(fit0)</pre>
```

```
##
## Call:
## lm(formula = y \sim poly(x, 4))
##
## Residuals:
##
      Min
               1Q Median
                               3Q
                                      Max
## -2.8914 -0.5244 0.0749 0.5932
                                   2.7796
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.8277
                           0.1041 -17.549
                                            <2e-16 ***
## poly(x, 4)1
                2.3164
                           1.0415
                                    2.224
                                            0.0285 *
## poly(x, 4)2 -21.0586
                           1.0415 -20.220
                                            <2e-16 ***
## poly(x, 4)3 -0.3048
                           1.0415 -0.293
                                            0.7704
                                            0.6373
## poly(x, 4)4 -0.4926
                           1.0415 -0.473
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.041 on 95 degrees of freedom
## Multiple R-squared: 0.8134, Adjusted R-squared: 0.8055
## F-statistic: 103.5 on 4 and 95 DF, p-value: < 2.2e-16
```

Summary shows that only X and X^2 are statistically significant predictors. This agrees with the LOOCV results that indicate using only X and X^2 produces the best model.

EXERCISE 9:

Part a)

```
require(MASS)
require(boot)
data(Boston)
(medv.mu <- mean(Boston$medv))</pre>
```

```
## [1] 22.53281
Part b)
(medv.sd <- sd(Boston$medv)/sqrt(nrow(Boston)))</pre>
## [1] 0.4088611
Part c)
set.seed(1)
mean.fn <- function(var, id) {</pre>
  return(mean(var[id]))
(boot.res <- boot(Boston$medv, mean.fn, R=100))
##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
##
## Call:
## boot(data = Boston$medv, statistic = mean.fn, R = 100)
##
## Bootstrap Statistics :
##
       original
                     bias
                              std. error
## t1* 22.53281 0.009027668 0.3482331
Estimation from bootstrap with R=100 is 0.38, reasonably close to 0.41
Part d)
boot.res$t0 - 2*sd(boot.res$t) # lower bound
## [1] 21.83634
boot.res$t0 + 2*sd(boot.res$t) # upper bound
## [1] 23.22927
t.test(Boston$medv)
##
   One Sample t-test
##
##
## data: Boston$medv
## t = 55.111, df = 505, p-value < 2.2e-16
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## 21.72953 23.33608
## sample estimates:
## mean of x
## 22.53281
```

```
Part e)
(medv.median <- median(Boston$medv))</pre>
## [1] 21.2
Part f)
set.seed(1)
median.fn <- function(var, id) {</pre>
 return(median(var[id]))
}
(boot.res <- boot(Boston$medv, median.fn, R=100))</pre>
##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
## Call:
## boot(data = Boston$medv, statistic = median.fn, R = 100)
##
##
## Bootstrap Statistics :
                       std. error
      original bias
           21.2 -0.029 0.3461316
## t1*
Estimated standard error is 0.368
Part g)
(medv.mu10 <- quantile(Boston$medv, 0.1))</pre>
## 10%
## 12.75
Part h)
set.seed(1)
quantile10.fn <- function(var, id) {
 return(quantile(var[id], 0.1))
(boot.res <- boot(Boston$medv, quantile10.fn, R=100))</pre>
##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
##
## Call:
## boot(data = Boston$medv, statistic = quantile10.fn, R = 100)
##
##
## Bootstrap Statistics :
      original bias std. error
         12.75 0.008 0.5370477
## t1*
```

Estimated standard error is 0.499