

## 270A - Homework 3

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**Problem 1.** Show that if  $X$  and  $Y$  are independent, integer-valued random variables, then for all  $n \in \mathbb{Z}$ ,

$$\mathbb{P}[X + Y = n] = \sum_{m \in \mathbb{Z}} \mathbb{P}[X = m] \cdot \mathbb{P}[Y = n - m].$$

*Proof.*

$$\begin{aligned} \mathbb{P}[X + Y = n] &= \mathbb{P}[Y = n - X] \\ &= \sum_{m \in \mathbb{Z}} \mathbb{P}[Y = n - m, X = m] \\ &= \sum_{m \in \mathbb{Z}} \mathbb{P}[Y = n - m] \cdot \mathbb{P}[X = m], \end{aligned}$$

where the last line follows from the independence of  $X$  and  $Y$ . □

**Problem 3.** Let  $X_1, X_2, \dots$  be independent random variables that satisfy

$$\frac{\text{Var}[X_i]}{i} \rightarrow 0$$

as  $i \rightarrow \infty$ . Let  $S_n = X_1 + \dots + X_n$ . Prove that

$$\frac{S_n - \mathbb{E}[S_n]}{n} \rightarrow 0 \quad \text{in probability.}$$

*Proof.* By Chebyshev's inequality, we have that

$$\begin{aligned} \mathbb{P} \left[ \left| \frac{S_n - \mathbb{E}[S_n]}{n} \right| > \epsilon \right] &\leq \frac{\text{Var}[(S_n - \mathbb{E}[S_n])/n]}{\epsilon^2} \\ &= \frac{1}{n^2 \epsilon^2} \sum_{i=1}^n \text{Var}[X_i] \\ &\leq \frac{1}{n \epsilon^2} \cdot \max_{i \leq n} \text{Var}[X_i] \\ &\rightarrow 0. \end{aligned}$$

□

**Problem 4.**

(a) Show that

$$d(X, Y) = \mathbb{E} \left[ \frac{|X - Y|}{1 + |X - Y|} \right]$$

defines a metric on the set of random variables.

*Proof.* It's clear that  $d$  is symmetric. Since  $\frac{|X-Y|}{1+|X-Y|}$  is nonnegative, its expectation is zero if and only if  $|X - Y|$  is zero almost surely, which happens if and only if  $X = Y$  almost surely. This just leaves the triangle inequality.

Consider the function  $f(t) = \frac{t}{1+t} = 1 - \frac{1}{1+t}$ . Since  $\frac{1}{1+t}$  is clearly decreasing, we have that  $f$  is increasing for  $t \geq 0$ . For any three random variables  $X$ ,  $Y$ , and  $Z$ , we then have

$$\begin{aligned} d(X, Z) &= f(|X - Z|) \\ &\leq f(|X - Y| + |Y - Z|) \\ &= \frac{|X - Y|}{1 + |X - Y| + |Y - Z|} + \frac{|Y - Z|}{1 + |X - Y| + |Y - Z|} \\ &\leq \frac{|X - Y|}{1 + |X - Y|} + \frac{|Y - Z|}{1 + |Y - Z|} \\ &= d(X, Y) + d(Y, Z). \end{aligned}$$

□

(b) Show that  $d(X_n, X) \rightarrow 0$  if and only if  $X_n \rightarrow X$  in probability.

*Proof.* Suppose that  $X_n \rightarrow X$  in probability. Fix  $\epsilon > 0$  and let  $E_n = \{|X_n - X| > \epsilon\}$ . For sufficiently large  $n$ , we have that  $\mathbb{P}[E_n] < \epsilon$ . Let's bound the expectation, using the fact that  $\frac{t}{1+t} \leq 1$  for all  $t \geq 0$ .

$$\begin{aligned} d(X_n, X) &= \int_{E_n} \frac{|X_n - X|}{1 + |X_n - X|} d\mathbb{P} + \int_{E_n^c} \frac{|X_n - X|}{1 + |X_n - X|} d\mathbb{P} \\ &\leq \mathbb{P}[E_n] \cdot 1 + \mathbb{P}[E_n^c] \cdot \epsilon \\ &\leq 2\epsilon. \end{aligned}$$

Suppose that  $X_n$  doesn't converge to  $X$  in probability. Then for some  $\epsilon$ , there are infinitely many  $n$  such that  $\mathbb{P}[|X_n - X| > \epsilon] > \epsilon$ . Since  $t \mapsto \frac{t}{1+t}$  is increasing, for such an  $n$  we have (with  $E_n$  defined as above)

$$\begin{aligned} d(X_n, X) &\geq \int_{E_n} \frac{|X_n - X|}{1 + |X_n - X|} d\mathbb{P} \\ &\geq \mathbb{P}[E_n] \cdot \frac{\epsilon}{1 + \epsilon} \\ &\geq \frac{\epsilon^2}{1 + \epsilon}, \end{aligned}$$

so  $d(X_n, X)$  doesn't go to zero.

□

**Problem 5.** Let  $X_1, X_2, \dots$  be independent  $\text{Ber}(p_n)$  random variables.

(a) Show that  $X_n \rightarrow 0$  in probability if and only if  $p_n \rightarrow 0$ .

*Proof.* Fix  $\epsilon > 0$ . We then have

$$\mathbb{P}[|X_n| > \epsilon] = \mathbb{P}[X_n = 1] = p_n.$$

We then have that  $\mathbb{P}[|X_n| > \epsilon] \rightarrow 0$  if and only if  $p_n \rightarrow 0$ . □

(b) Show that  $X_n \rightarrow 0$  a.s. if and only if  $\sum p_n < \infty$ .

*Proof.* Fix  $\epsilon > 0$  and let  $E_n = \{|X_n| > \epsilon\} = \{X_n = 1\}$ . We then have

$$\sum \mathbb{P}[E_n] = \sum p_n < \infty.$$

By Borel-Cantelli, we have

$$\mathbb{P}[\limsup E_n] = \mathbb{P}[X_n = 1 \text{ infinitely often}] = 0.$$

Taking complements gives

$$1 = \mathbb{P}[\liminf E_n^c] = \mathbb{P}[X_n = 0 \text{ eventually}],$$

so  $X_n \rightarrow 0$  almost surely. On the other hand, if  $\sum p_n = \infty$ , then since the  $X_j$ 's are independent, Borel-Cantelli says that

$$\mathbb{P}[\limsup E_n] = \mathbb{P}[X_n = 1 \text{ infinitely often}] = 1.$$

Since  $X_n = 1$  infinitely often,  $X_n$  doesn't converge to zero almost surely. □

**Problem 6.** Let  $X_1, X_2, \dots$  be a sequence of random variables on  $(\Omega, \mathcal{F}, \mathbb{P})$  where  $\Omega$  is a countable set and  $\mathcal{F} = 2^\Omega$ . Show that  $X_n \rightarrow X$  in probability implies  $X_n \rightarrow X$  a.s.

*Proof.* Suppose  $X_n$  didn't converge to  $X$  almost surely. Then since  $\Omega$  is a discrete space, there is some singleton  $\{\omega\} \in \mathcal{F}$  with  $\mathbb{P}[\{\omega\}] > 0$  (we'll abuse notation and write  $\mathbb{P}[\omega] > 0$ ) and  $X_n(\omega)$  doesn't converge to  $X(\omega)$ . We can then find  $\epsilon > 0$  so that for infinitely many  $n$ ,  $|X_n(\omega) - X(\omega)| > \epsilon$ . But then  $X_n$  can't converge to  $X$  in probability since

$$\mathbb{P}[|X_n - X| > \epsilon] \geq \mathbb{P}[\omega] > 0,$$

for infinitely many  $n$ . □

**Problem 7.** Show that for any sequence of random variables  $X_1, X_2, \dots$  there exists a sequence of positive real numbers  $c_1, c_2, \dots$  such that  $c_n X_n \rightarrow 0$  a.s.

*Proof.* Since  $X_n$  takes values in the real numbers, we have that  $\mathbb{P}[|X_n| > t] \rightarrow 0$  as  $t \rightarrow \infty$ . There is then some  $b_n$  so that  $\mathbb{P}[|X_n| > b_n] < 2^{-n}$ . Now let  $c_n = \frac{1}{n \cdot b_n}$ . For any  $\epsilon > 0$  we have

$$\sum \mathbb{P}[|c_n X_n| > \epsilon] = \sum \mathbb{P}[|X_n| > \epsilon \cdot n b_n].$$

For  $n$  sufficiently large,  $\epsilon n \geq 1$  and  $\mathbb{P}[|X_n| > \epsilon n b_n] \leq \mathbb{P}[|X_n| > b_n] < 2^{-n}$ . The tail of the above series is then summable, and we have

$$\mathbb{P}[\limsup\{|c_n X_n| > \epsilon\}] = 0$$

for all  $\epsilon > 0$ , so  $c_n X_n \rightarrow 0$  a.s. □

**Problem 8.** Let  $X_1, X_2, \dots$  be independent random variables (from  $(\Omega, \mathcal{F}, \mathbb{P})$  to  $\mathbb{R}$ ). Show that  $\sup_n X_n < \infty$  a.s. if and only if there exists  $M \in \mathbb{R}$  such that

$$\sum_n \mathbb{P}[X_n > M] < \infty.$$

*Proof.* Suppose there exists a real  $M$  so that the above sum is finite. Then by Borel-Cantelli we have that  $\mathbb{P}[X_n > M \text{ infinitely often}] = 0$ , or  $\mathbb{P}[X_n \leq M \text{ eventually}] = 1$ . For any  $\omega \in \Omega$ , there is some  $N_\omega$  so that  $X_n(\omega) \leq M$  for all  $n > N_\omega$ . We then have that

$$X_n(\omega) \leq \max\{X_1(\omega), X_2(\omega), \dots, X_{N_\omega}(\omega), M\}.$$

Since this holds for a.e.  $\omega$ , we have that  $\sup_n X_n < \infty$  a.s.

Conversely, suppose that for every  $M$  we have  $\sum \mathbb{P}[X_n > M] = +\infty$ . Since the  $X_n$ 's are independent, Borel-Cantelli tells us that  $\mathbb{P}[X_n > M \text{ infinitely often}] = 1$ . For almost any  $\omega \in \Omega$  and for any real number  $M$ , there are infinitely many  $n$  such that  $X_n(\omega) > M$ . Consequently,  $\sup_n X_n(\omega)$  is infinite for almost every  $\omega$ . □

**Problem 9.** Let  $X_0 = 1$  and define  $X_n$  inductively by choosing  $X_{n+1}$  uniformly at random from the interval  $[0, X_n]$ . Prove that

$$\frac{\log X_n}{n} \rightarrow c$$

a.s. and find the value of  $c$ .

*Solution.* Let  $U_1, U_2, \dots$  be iid  $\text{Unif}[0, 1]$  random variables. We claim that  $X_n$  and  $U_1 \cdot U_2 \cdots U_n$  have the same distribution. This is clearly true for  $X_1$ , so suppose that  $X_n$  and  $U_1 \cdots U_n$  have the same distribution and consider  $X_{n+1}$ . We have that

$$\mathbb{P}[X_{n+1} \leq t] = \frac{t}{X_n}, \quad \text{for all } t \leq X_n.$$

On the other hand, we have

$$\mathbb{P}[U_{n+1} \cdot X_n \leq t] = \mathbb{P}[U_{n+1} \leq t/X_n] = \frac{t}{X_n}.$$

We then have

$$\frac{\log X_n}{n} = \frac{1}{n} \sum_{k=1}^n \log U_k.$$

Let's compute an expectation

$$\begin{aligned} \mathbb{E}[\log U_k] &= \int_0^1 \log x \, dx \\ &= -1. \end{aligned}$$

By the strong law of large numbers, we have that

$$\frac{\log X_n}{n} = \frac{1}{n} \sum_{k=1}^n \log U_k \rightarrow -1 \text{ a.s.}$$

□

**Problem 10.** Let  $X_1, X_2, \dots$  be independent random variables such that  $X_n$  takes value  $n$  with probability  $1/(2n \log n)$  and value  $-n$  with the same probability, and the value 0 with the remaining probability  $1 - 1/(n \log n)$ . Show that this sequence obeys the weak law, but not the strong law in the sense that

$$\frac{1}{n} \sum_{i=1}^n X_i \rightarrow 0$$

in probability but not a.s.

*Proof.* Let  $S_n = \sum_{i=2}^{n+1} X_i$ . Since the  $X_i$  are centered, so is  $S_n$  and  $\mathbb{E}[S_n/n] = 0$ . Let's compute the variance of  $X_n$ .

$$\begin{aligned} \text{Var}[X_n] &= \frac{n^2}{2n \log n} + \frac{(-n)^2}{2n \log n} \\ &= \frac{n}{\log n}. \end{aligned}$$

Since the  $X_n$ 's are independent, the variances add and we have

$$\text{Var}[S_n/n] = \frac{1}{n^2} \sum_{k=2}^{n+1} \frac{k}{\log k}.$$

Now let's fix  $\epsilon > 0$  and use Chebyshev

$$\mathbb{P} \left[ \left| \frac{S_n}{n} \right| > \epsilon \right] \leq \frac{1}{n^2 \epsilon^2} \sum_{k=2}^{n+1} \frac{k}{\log k} \tag{1}$$

Now  $\frac{k}{\log k}$  is increasing for  $k \geq 3$ , so we can bound the above probability

$$\begin{aligned} \mathbb{P} \left[ \left| \frac{S_n}{n} \right| > \epsilon \right] &\leq \frac{1}{n^2 \epsilon^2} \cdot \frac{n(n+1)}{\log(n+1)} \\ &\rightarrow 0, \end{aligned}$$

so  $S_n/n \rightarrow 0$  in probability.

Now consider the sum

$$\sum_{n=2}^{\infty} \mathbb{P}[X_n = n] = \sum_{n=2}^{\infty} \frac{1}{2n \log n} = +\infty.$$

Since the  $X_n$ 's are independent, Borel-Cantelli tells us that  $\mathbb{P}[X_n = n \text{ infinitely often}] = 1$ . Consequently,  $X_n/n = 1$  infinitely often, so the sum  $\frac{1}{n} \sum_{k=2}^{n+1} X_k$  cannot converge to zero.  $\square$