```
In [1]: # Time Series Modeling for 30-Year U.S. Mortgage Rate
        import pandas as pd
        import numpy as np
        import matplotlib.pyplot as plt
        from statsmodels.tsa.arima.model import ARIMA
        from statsmodels.graphics.tsaplots import plot acf, plot pacf
        from statsmodels.tsa.stattools import adfuller
        from statsmodels.stats.diagnostic import acorr_ljungbox
        import warnings
        warnings.filterwarnings("ignore")
        # Step 1: Load and preprocess data
        df = pd.read_csv("m-mortg.txt", sep="\s+", header=None, names=["Year", "M
        df["Date"] = pd.to_datetime(df[["Year", "Month", "Day"]])
        df.set_index("Date", inplace=True)
        df = df[["Rate"]]
        df["log rate"] = np.log(df["Rate"])
        # Step 2: Visualize original and log-transformed data
        df[["Rate", "log_rate"]].plot(title="Mortgage Rate vs Log-Rate", subplots
        plt.tight_layout()
        plt.show()
        # Step 3: ADF test for stationarity
        result = adfuller(df['log_rate'])
        print("ADF Test p-value:", result[1])
        # Step 4: First difference the series
        df['log_rate_diff'] = df['log_rate'].diff()
        # Step 5: ACF and PACF plots
        plot_acf(df['log_rate_diff'].dropna(), lags=30)
        plt.title("ACF of Differenced Series")
        plt.show()
        plot_pacf(df['log_rate_diff'].dropna(), lags=30)
        plt.title("PACF of Differenced Series")
        plt.show()
        # Step 6: Auto ARIMA grid search
        def auto_arima_grid_search(series, d, p_range=range(0, 4), q_range=range(
            best_aic = np.inf
            best_order = None
            best_model = None
            for p in p_range:
                for q in q_range:
                    try:
                        model = ARIMA(series, order=(p, d, q))
                         result = model.fit()
                        if result.aic < best_aic:</pre>
                            best_aic = result.aic
                            best_order = (p, d, q)
                            best_model = result
                    except:
                        continue
            return best_order, best_model
        best_order, best_model = auto_arima_grid_search(df["log_rate"], d=1)
```

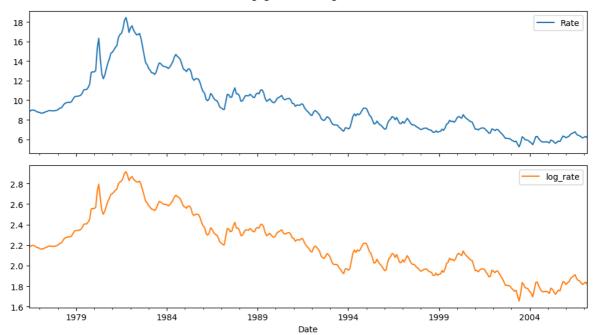
```
print("Best ARIMA model:", best_order)
print(best_model.summary())

# Step 7: Model diagnostics
resid = best_model.resid
lb_test = acorr_ljungbox(resid, lags=[12], return_df=True)
print("\nLjung-Box Q(12) Test:")
print(lb_test)

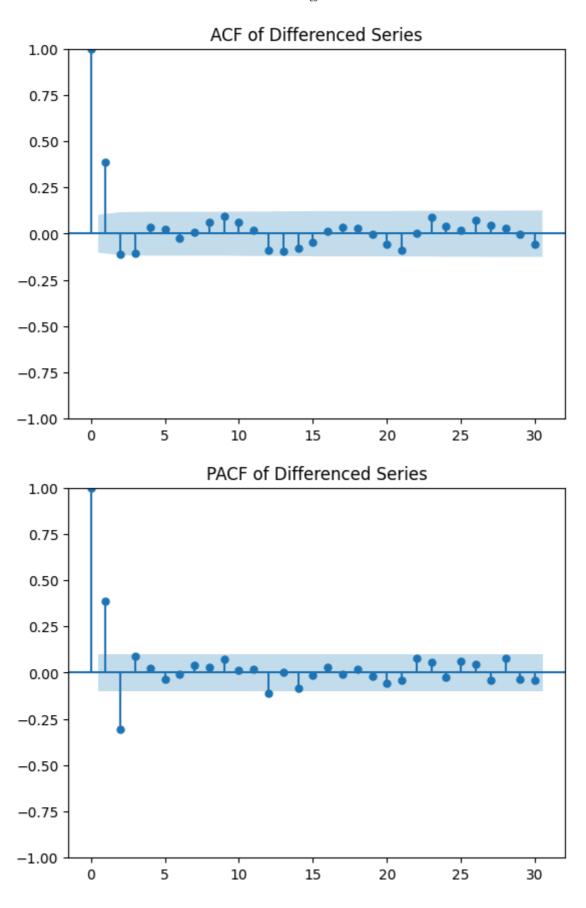
# Step 8: Forecast 1-4 steps ahead
forecast = best_model.get_forecast(steps=4)
log_forecast = forecast.predicted_mean
rate_forecast = np.exp(log_forecast)

# Step 9: Display forecast results
print("\n1 to 4-step forecast (original rate):")
for i in range(4):
    print(f"Step {i+1}: {rate_forecast.iloc[i]:.4f}")
```

## Mortgage Rate vs Log-Rate



ADF Test p-value: 0.8196642809000326



Best ARIMA model: (3, 1, 0)

## SARIMAX Results

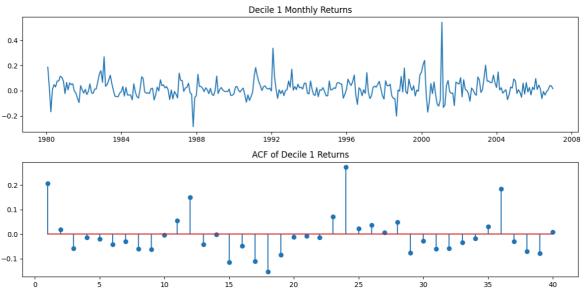
Time:			SAR. 	LMAX Resul 	.ts 		
370 Model: ARIMA(3, 1, 0) Log Likelihood 82 0.726 Date: Sat, 05 Apr 2025 AIC -163 3.453 Time: 13:23:24 BIC -161 7.810 Sample: 06-01-1976 HQIC -162 7.239 -03-01-2007 Covariance Type: opg							
0.726  Date: Sat, 05 Apr 2025 AIC -163 3.453 Time: 13:23:24 BIC -161 7.810 Sample: 06-01-1976 HQIC -162 7.239 - 03-01-2007  Covariance Type: opg	•		log_ra	ate No.	Observations:		
Date: Sat, 05 Apr 2025 AIC -163 3.453 Time: 13:23:24 BIC -161 7.810 Sample: 06-01-1976 HQIC -162 7.239 - 03-01-2007 Covariance Type: opg	Model:		ARIMA(3, 1,	0) Log	Likelihood		82
3.453 Time: 13:23:24 BIC -161 7.810 Sample: 06-01-1976 HQIC -162 7.239 - 03-01-2007 Covariance Type: opg	0.726						
Time: 13:23:24 BIC161 7.810	Date:	Sa	at, 05 Apr 20	025 AIC			-163
7.810 Sample: 06-01-1976 HQIC -162 7.239 - 03-01-2007 Covariance Type: opg	3.453						
Sample: 06-01-1976 HQIC -162 7.239 - 03-01-2007  Covariance Type: opg	Time:		13:23:	:24 BIC			-161
7.239  - 03-01-2007  Covariance Type: opg							
- 03-01-2007  Covariance Type: opg	•		06-01-19	976 HQIC	•		-162
Covariance Type: opg	7.239						
coef std err z P> z  [0.025 0 975]			- 03-01-20	007			
coef std err z P> z  [0.025 0 975]					:========	========	=====
975] ar.L1	====						
ar.L1	_	coef	std err	Z	P> z	[0.025	0
ar.L1	975]						
0.629 ar.L2							
ar.L2	ar.L1	0.5323	0.049	10.758	0.000	0.435	
0.269 ar.L3	0.629						
ar.L3		-0.3491	0.041	-8 <b>.</b> 560	0.000	-0.429	_
0.195 sigma2							
sigma2		0.0894	0.054	1.654	0.098	-0.017	
0.001 ==================================							
======================================	-	0.0007	3.12e-05	21.949	0.000	0.001	
Ljung-Box (L1) (Q): 0.00 Jarque-Bera (JB): 281.52  Prob(Q): 0.96 Prob(JB): 0.00  Heteroskedasticity (H): 0.92 Skew: 0.67  Prob(H) (two-sided): 0.64 Kurtosis: 7.07  ===========  Warnings: [1] Covariance matrix calculated using the outer product of gradients (coplex-step).  Ljung-Box Q(12) Test: b_stat lb_pvalue	==========	======	=========	=======		========	
281.52 Prob(Q):	======================================	(0):		0 00	larque_Rera	(1R)•	
Prob(Q): 0.96 Prob(JB): 0.00  Heteroskedasticity (H): 0.92 Skew: 0.67  Prob(H) (two-sided): 0.64 Kurtosis: 7.07  ==========  Warnings: [1] Covariance matrix calculated using the outer product of gradients (coplex-step).  Ljung-Box Q(12) Test:		(4).		0.00	Jai que Bei a	(30).	
<pre>0.00 Heteroskedasticity (H):</pre>				0 06	Proh(1R):		
Heteroskedasticity (H): 0.92 Skew: 0.67 Prob(H) (two-sided): 0.64 Kurtosis: 7.07 ==================================				0.90	1100(30).		
0.67 Prob(H) (two-sided):		city (H)		0 92	Skew:		
Prob(H) (two-sided): 0.64 Kurtosis: 7.07		CICY (11/1	ı	0132	Sicw.		
7.07 ===================================		ided):		0.64	Kurtosis:		
Warnings: [1] Covariance matrix calculated using the outer product of gradients (coplex-step).  Ljung-Box Q(12) Test:     lb_stat lb_pvalue 12 0.05655     1.0  1 to 4-step forecast (original rate): Step 1: 6.0752 Step 2: 6.0807 Step 3: 6.1018 Step 4: 6.1035	7.07	iucu, i		0101	1101 105151		
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<pre>[1] Covariance matrix calculated using the outer product of gradients (copplex-step).  Ljung-Box Q(12) Test:</pre>	Warnings:						
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Ljung-Box Q(12) Test:     lb_stat lb_pvalue 12    0.05655		IIIati IX (	accutateu us	sing the o	dicer product	or gradient	13 (00
lb_stat lb_pvalue 12 0.05655	prox stop,:						
12 0.05655							
1 to 4-step forecast (original rate): Step 1: 6.0752 Step 2: 6.0807 Step 3: 6.1018 Step 4: 6.1035							
Step 1: 6.0752 Step 2: 6.0807 Step 3: 6.1018 Step 4: 6.1035	12 0:05055	1.0					
Step 2: 6.0807 Step 3: 6.1018 Step 4: 6.1035			original rate	e):			
Step 3: 6.1018 Step 4: 6.1035	•						
Step 4: 6.1035	•						
·	•						
# Time Series Modeling for Decile 1 Monthly Returns using SARIMA	Step 4: 6.1035						
	# Time Series	Modeling	g for Decile	1 Monthly	/ Returns usi	ng SARIMA	

In [2]: |# Time Series Modeling for Decile 1 Monthly Returns using SARIMA

import pandas as pd

```
import numpy as np
from statsmodels.tsa.statespace.sarimax import SARIMAX
import matplotlib.pyplot as plt
from statsmodels.tsa.stattools import acf
from statsmodels.stats.diagnostic import acorr_ljungbox
from scipy import stats
# Step 1: Load data
dec1 = pd.read_csv('m-dec1-8006.txt', sep='\s+', header=None)
dec1.columns = ['date', 'return']
# Step 2: Create datetime index
dates = pd.date_range(start='1980-01-01', periods=len(dec1), freq='M')
dec1_ts = pd.Series(dec1['return'].values, index=dates)
# Step 3: Plot time series and ACF
plt.figure(figsize=(12, 6))
plt.subplot(211)
plt.plot(dec1 ts)
plt.title('Decile 1 Monthly Returns')
plt.subplot(212)
acf_values = acf(dec1_ts.values, nlags=40)[1:]
plt.stem(range(1, len(acf_values) + 1), acf_values)
plt.title('ACF of Decile 1 Returns')
plt.tight_layout()
plt.show()
# Step 4: Fit SARIMA(0,0,1)(1,0,1)[12] model
model = SARIMAX(dec1_ts, order=(0, 0, 1), seasonal\_order=(1, 0, 1, 12))
results = model.fit()
print(results.summary())
# Step 5: Residual diagnostics
residuals = results.resid
plt.figure(figsize=(12, 8))
plt.subplot(221)
plt.plot(residuals)
plt.title('Residuals Over Time')
plt.subplot(222)
plt.hist(residuals, bins=30)
plt.title('Histogram of Residuals')
plt.subplot(223)
acf_values = acf(residuals.values, nlags=40)[1:]
plt.stem(range(1, len(acf_values) + 1), acf_values)
plt.title('ACF of Residuals')
plt.subplot(224)
stats.probplot(residuals, dist="norm", plot=plt)
plt.title('Q-Q Plot of Residuals')
plt.tight_layout()
plt.show()
# Step 6: Ljung-Box test (Q(24))
lb_result = acorr_ljungbox(residuals, lags=[24], return_df=True)
q_stat = lb_result['lb_stat'].iloc[0]
p_val = lb_result['lb_pvalue'].iloc[0]
print(f"Ljung-Box Q(24) Statistic: {q_stat:.4f}")
print(f"p-value: {p_val:.4f}")
# Step 7: Conclusion
alpha = 0.05
```

```
if p_val > alpha:
    print(f"Residuals show no significant autocorrelation (p > {alpha}).
else:
    print(f"Residuals show significant autocorrelation (p ≤ {alpha}). Mod
```



RUNNING THE L-BFGS-B CODE

\* \* \*

Machine precision = 2.220D-16N = 4 M = 10

At X0 0 variables are exactly at the bounds

At iterate f = -1.22381D + 00|proj g| = 6.68059D-01At iterate 5 f = -1.22788D + 00|proj g|= 1.03522D+00 At iterate 10 f = -1.23114D+00|proj g| = 8.61734D-02At iterate 15 f = -1.23382D + 001.13082D+00 |proj g|= At iterate 20 f = -1.25902D + 00|proj g| = 2.80640D-01

This problem is unconstrained.

```
At iterate 25 f=-1.28620D+00 |proj g|=1.53820D-01 At iterate 30 f=-1.29205D+00 |proj g|=2.60927D-01 At iterate 35 f=-1.29438D+00 |proj g|=1.65326D-01 At iterate 40 f=-1.29491D+00 |proj g|=2.71945D-02 At iterate 45 f=-1.29502D+00 |proj g|=1.48042D-03
```

\* \* \*

Tit = total number of iterations

Tnf = total number of function evaluations

Tnint = total number of segments explored during Cauchy searches

Skip = number of BFGS updates skipped

Nact = number of active bounds at final generalized Cauchy point

Projg = norm of the final projected gradient

F = final function value

\* \* \*

N Tit Tnf Tnint Skip Nact Projg F 4 47 59 1 0 0 1.412D-03 -1.295D+00 F = -1.2950159735225941

CONVERGENCE: REL\_REDUCTION\_OF\_F\_<=\_FACTR\*EPSMCH SARIMAX Results

\_\_\_\_\_

=======================================		
Dep. Variable:	У	No. Observations:
Model: 419.585	SARIMAX(0, 0, 1)×(1, 0, 1, 12)	Log Likelihood
Date:	Sat, 05 Apr 2025	AIC
-831.170 Time: -816.047	13:23:26	BIC
Sample: -825.134	01-31-1980	HQIC
0231134	- 12-31-2006	
Covariance Type:	opg	:======================================

====	coef	std err	z	P> z	[0.025	0.
975] 						
ma.L1 0.312	0.2455	0.034	7.196	0.000	0.179	
ar.S.L12 1.023	0.9997	0.012	85.199	0.000	0.977	
ma.S.L12 0.412	-0.9862	0.293	-3.366	0.001	-1.561	_
sigma2 0.006	0.0041	0.001	4.008	0.000	0.002	

0.02 Jarque-Bera (JB):

=======

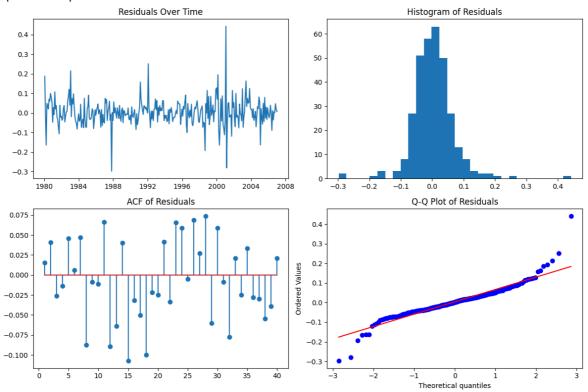
Ljung-Box (L1) (Q):

1059.98

=======

## Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).



Ljung-Box Q(24) Statistic: 23.7876

p-value: 0.4738

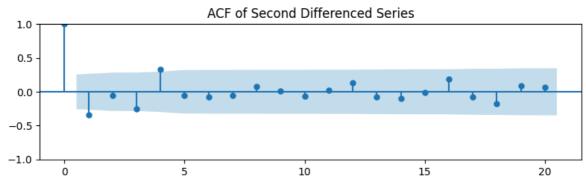
Residuals show no significant autocorrelation (p > 0.05). Model is adequat e.

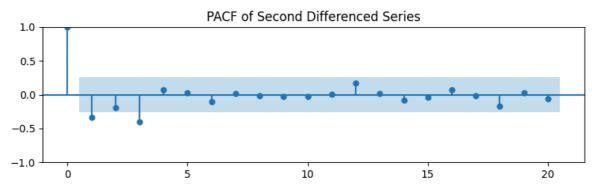
```
In [3]: # Alcoa Quarterly EPS Time Series Modeling and Forecasting
        import pandas as pd
        import numpy as np
        import matplotlib.pyplot as plt
        import itertools
        import warnings
        from statsmodels.tsa.arima.model import ARIMA
        from statsmodels.graphics.tsaplots import plot_acf, plot_pacf
        from statsmodels.tsa.stattools import adfuller
        from statsmodels.stats.diagnostic import acorr_ljungbox
        warnings.filterwarnings("ignore")
        # Step 1: Load and preprocess data
        df = pd.read_csv("q-aa-earn.txt", sep="\s+", header=None, names=["Day", "
        df["Date"] = pd.to_datetime(df[["Year", "Month", "Day"]])
        df.set_index("Date", inplace=True)
        df = df.sort_index()
        eps = df["EPS"]
```

```
# Step 2: Stationarity check (ADF Test)
result = adfuller(eps)
print("ADF p-value (original series):", result[1])
# Step 3: Apply second-order differencing (since series is non-stationary
diff_eps = eps.diff().dropna().diff().dropna()
result diff2 = adfuller(diff eps)
print("ADF p-value (second difference):", result_diff2[1])
# Step 4: Plot ACF and PACF for second-differenced series
fig, axes = plt.subplots(2, 1, figsize=(8, 5))
plot_acf(diff_eps, lags=20, ax=axes[0])
axes[0].set_title('ACF of Second Differenced Series')
plot_pacf(diff_eps, lags=20, ax=axes[1])
axes[1].set_title('PACF of Second Differenced Series')
plt.tight_layout()
plt.show()
# Step 5: Try multiple ARIMA models and select based on AIC
candidate_orders = [(0, 2, 1), (1, 2, 0), (1, 2, 1), (2, 2, 0), (0, 2, 2)]
results = []
for order in candidate_orders:
    model = ARIMA(eps, order=order)
    fitted model = model.fit()
    results.append((order, fitted_model))
    print(f"ARIMA{order} - AIC: {fitted_model.aic:.2f}")
best_order, best_model = min(results, key=lambda x: x[1].aic)
print(f"\nBest ARIMA model: {best order} (AIC = {best model.aic:.2f})")
print(best_model.summary())
# Step 6: Residual diagnostics
resid = best_model.resid
lb_result = acorr_ljungbox(resid, lags=[12], return_df=True)
print("\nLjung-Box Q(12) Test:")
print(lb_result)
# Step 7: Forecast next 4 quarters
forecast = best_model.get_forecast(steps=4)
mean_forecast = forecast.predicted_mean
conf_int = forecast.conf_int()
# Step 8: Plot forecast
last_date = df.index[-1]
future_dates = pd.date_range(start=last_date + pd.DateOffset(months=3), p
plt.figure(figsize=(10, 6))
plt.plot(df.index, df['EPS'], label='Historical EPS')
plt.plot(future_dates, mean_forecast, color='red', label='Forecast')
plt.fill_between(future_dates, conf_int.iloc[:, 0], conf_int.iloc[:, 1],
plt.title('Alcoa Quarterly EPS Forecast')
plt.xlabel('Date')
plt.ylabel('EPS')
plt.legend()
plt.grid(True)
plt.tight_layout()
plt.show()
# Step 9: Print forecast values
```

```
print("\nForecast Summary:")
for i in range(4):
    print(f"Quarter {i+1}: {mean_forecast.iloc[i]:.4f}, 95% CI: [{conf_in}]
```

ADF p-value (original series): 0.6181931025603755 ADF p-value (second difference): 4.894164446594406e-22





```
ARIMA(0, 2, 1) - AIC: -70.09

ARIMA(1, 2, 0) - AIC: -42.49

ARIMA(1, 2, 1) - AIC: -72.07

ARIMA(2, 2, 0) - AIC: -47.25

ARIMA(0, 2, 2) - AIC: -75.51

ARIMA(2, 2, 2) - AIC: -73.68
```

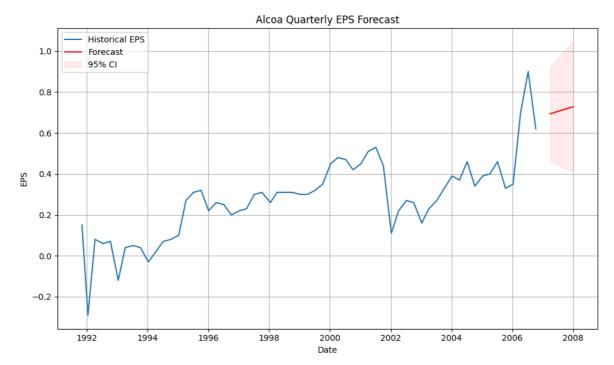
Best ARIMA model: (0, 2, 2) (AIC = -75.51) SARIMAX Results

========	=========	========		========	========	=====
==== Dep. Varia	able:	E	EPS No.	Observations	:	
61 Model:	,	ARIMA(0, 2,	2) Log	Log Likelihood		
0.753 Date:	Sa <sup>-</sup>	t, 05 Apr 20	)25 AIC			-7
5.506 Time:		12.22.	. 27 DIC			6
9.274		15:25:	27 BIC			-6
Sample: 3.073			0 HQI	С		-7
3.073		_	61			
Covariance	e Type:	C	pg			
====						=====
975]	coef	std err	Z	P> z	[0.025	0.
ma.L1 0.549	-1.4639	0.467	-3.137	0.002	-2.378	_
ma.L2	0.4688	0.175	2.686	0.007	0.127	
0.811 sigma2 0.026	0.0135	0.007	2.045	0.041	0.001	
========	:=======	========	======	=========	========	=====
Ljung-Box 63.12	(L1) (Q):		3.04	Jarque-Bera	(JB):	
Prob(Q):			0.08	Prob(JB):		
	dasticity (H):		1.44	Skew:		
0.76 Prob(H) (t 7.83	two-sided):		0.42	Kurtosis:		
========	=========	========	======	:========	========	=====

## Warnings:

[1] Covariance matrix calculated using the outer product of gradients (com plex-step).

Ljung-Box Q(12) Test: lb\_stat lb\_pvalue 12 7.956544 0.788517



Forecast Summary:

Quarter 1: 0.6937, 95% CI: [0.4650, 0.9224] Quarter 2: 0.7051, 95% CI: [0.4447, 0.9656] Quarter 3: 0.7166, 95% CI: [0.4269, 1.0063] Quarter 4: 0.7281, 95% CI: [0.4110, 1.0452]

In []:

In [ ]: