

Applications of Econometric & Data Science Methods

Homework 2

Due Date: Monday, April 11, 18:30 CST

Note: all vectors are column vectors.

1. [35 pts] Frisch-Waugh Theorem with a Realized Sample

Suppose we have n observations of an outcome with p covariates: $\{(y_i, x_i)\}_{i=1}^n$. In matrix form, we observe y and X . We divide our covariates in two groups, so that $x_i = (x_{i1}, x_{i2})$ for all i , or $X = (X_1, X_2)$. Suppose the first p_1 covariates are in group 1. Let $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2)$ be the best linear predictor of y using X and consider the following matrices:

$$P_1 = X_1(X_1'X_1)^{-1}X_1'; \quad M_1 = I - P_1.$$

Let $\tilde{X}_2 = M_1X_2$ and $\tilde{y} = M_1y$.

a. [3 pts]

What are the dimensions of y ? Of X_2 ? Of $\hat{\beta}_1$?

b. [5 pts]

Show that the *normal equations* – the ones that give rise to $\hat{\beta}$ – are:

$$\begin{aligned} X_1'X_1\hat{\beta}_1 + X_1'X_2\hat{\beta}_2 &= X_1'y \\ X_2'X_1\hat{\beta}_1 + X_2'X_2\hat{\beta}_2 &= X_2'y \end{aligned}$$

c. [5 pts]

Show that

$$X_1\hat{\beta}_1 + P_1X_2\hat{\beta}_2 = P_1y.$$

d. [5 pts]

Show that

$$\tilde{y} = \tilde{X}_2 \hat{\beta}_2 + \hat{\epsilon},$$

where $\hat{\epsilon}$ is the error from the projection of y on X .

e. [5 pts]

Show that

$$\hat{\beta}_2 = \left(\tilde{X}_2' \tilde{X}_2 \right)^{-1} \tilde{X}_2' \tilde{y}.$$

f. [5 pts]

Explain your result in words.

g. [7 pts]

Observation i , (y_i, x_i) , is a realization of random vector (Y_i, \mathbf{X}_i) where $\mathbf{X}_i = (\mathbf{X}_i^1, \mathbf{X}_i^p)$, so that our data as a whole can be seen as a realization of random vector $((Y_1, \mathbf{X}_1), (Y_2, \mathbf{X}_2), \dots, (Y_n, \mathbf{X}_n))$.

Answer Yes or No.

(i) [2 pts] Does your result require that (Y_i, \mathbf{X}_i) and (Y_j, \mathbf{X}_j) be identically distributed for all $i, j \in \{1, \dots, n\}$?

(ii) [2 pts] Does your result require that $\{(Y_1, \mathbf{X}_1), (Y_2, \mathbf{X}_2), \dots, (Y_n, \mathbf{X}_n)\}$ be mutually independent?

(iii) [1 pts] Does your result require that $((Y_1, \mathbf{X}_1), (Y_2, \mathbf{X}_2), \dots, (Y_n, \mathbf{X}_n))$ be i.i.d.?

(iv) [2 pts] Suppose there are random variables (W_1, \dots, W_n) whose realizations you do not observe and that, for some $i \in \{1, \dots, n\}$,

$$\text{Cov}(\mathbf{X}_i, W_i) \equiv \begin{pmatrix} \text{Cov}(\mathbf{X}_i^1, W_i) \\ \vdots \\ \text{Cov}(\mathbf{X}_i^p, W_i) \end{pmatrix} > \begin{pmatrix} 0 \\ \vdots \\ 0 \end{pmatrix}.$$

Does your result require that $\text{Cov}(Y_i, W_i) = 0$?

2. [35 pts] Measurement Error in OLS

Consider an outcome y and covariates x_1, \dots, x_K . These are all random variables. Now fix $\beta = (\beta_0, \dots, \beta_K) \in \mathbb{R}^{K+1}$ such that

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_K x_K + \epsilon$$

where $\mathbb{E}[\epsilon] = 0$ and $\mathbb{E}[\epsilon x_k] = 0$ for each $k \in \{1, \dots, K\}$.

Suppose we have an i.i.d. sample $\{(y_i, 1, x_{1i}, \dots, x_{Ki})\}_{i=1}^N$. Let $x_i = (1, x_{1i}, \dots, x_{Ki})$, $X = (x_1, \dots, x_N)$ and $\mathbf{y} = (y_1, \dots, y_N)$. X has dimensions $N \times (K+1)$. Assume away multicollinearity in population and in sample. Under these conditions, recall that

$$\sqrt{N}(\hat{\beta} - \beta) \xrightarrow{d} N(0, \mathbb{E}[x_i x_i']^{-1} \mathbb{E}[\epsilon_i^2 x_i x_i'] \mathbb{E}[x_i x_i']^{-1}),$$

where \xrightarrow{d} denotes convergence in distribution as N grows to infinity, and $\hat{\beta} = (X'X)^{-1}X'\mathbf{y}$.

a. [5 pts] Homoskedasticity

Suppose in addition that

- (i) $\text{Var}(y_i | x_i) = \sigma^2$
- (ii) $\mathbb{E}[y_i | x_{1i}, \dots, x_{Ki}] = \beta_0 + \beta_1 x_{1i} + \dots + \beta_K x_{Ki}$.

Show that

$$\sqrt{N}(\hat{\beta} - \beta) \xrightarrow{d} N(0, \sigma^2 \mathbb{E}[x_i x_i']^{-1}).$$

b. [15 pts] Measurement Error in the Outcome Variable

Suppose that we can only work with an imperfect measure of y : \tilde{y} , where

$$\tilde{y} = y + u_y.$$

(i) [6 pts] What assumption(s) is (are) required to consistently estimate β by simply working with $(\tilde{y}_1, \dots, \tilde{y}_N)$ instead of the (unobservable) y_i 's?

(ii) [6 pts] Suppose your assumption(s) in (i) hold. For simplicity, assume homoskedasticity and suppose that $\text{Var}(u_y | x_1, \dots, x_K) = \sigma_u^2$. Moreover, suppose that u_y and ϵ are uncorrelated. How does the asymptotic variance of $\tilde{\beta} \equiv (X'X)^{-1}X'\tilde{\mathbf{y}}$ compare with that of $\hat{\beta}$?

(iii) [3 pts] Under the assumptions in (ii), if you had access to both y and \tilde{y} , would it be a good idea to work with \tilde{y} ? Justify your answer.

c. [15 pts] Measurement Error in the Covariates

Suppose instead that we observe y, x_1, \dots, x_{K-1} perfectly, but we can only work with an imperfect measure of x_K : \tilde{x}_K , where

$$\tilde{x}_K = x_K + u_K.$$

(i) [5 pts] What assumption(s) is (are) required to consistently estimate β by simply working with $(\tilde{x}_{K1}, \dots, \tilde{x}_{KN})$ instead of the (unobservable) x_{Ki} 's?

(ii) [1 pts] If your assumption(s) in (i) hold and u_K and ϵ are uncorrelated. Is it a good idea to work with \tilde{x}_K if you had access to \tilde{x}_K and x_K ? Why, or why not?

(iii) [4 pts] Clear your assumptions and suppose that $\text{Cov}(x_K, u_K) = 0$. Can β be consistently estimated by working with \tilde{x}_K instead of x_K ? Why, or why not?

(iv) [5 pts] Clear all assumptions. Consider now the simpler case where $K = 1$. Show that the probability limit of $\tilde{\beta}_1$ when using \tilde{x}_1 instead of x_1 is

$$\text{plim}_{N \rightarrow \infty} \tilde{\beta}_1 = \frac{\beta_1 [\text{Var}(x_{1i}) + \text{Cov}(x_{1i}, u_{1i})] + \text{Cov}(\epsilon_i, u_{1i})}{\text{Var}(x_{1i}) + \text{Var}(u_{1i}) + 2\text{Cov}(x_{1i}, u_{1i})}.$$

What happens when the measurement error is “random”, i.e. independent of x_1 and ϵ ? In metrics slang, this is called the *attenuation bias*.

3. [30 pts] Checking the Exclusion Restriction

Consider an outcome y and K covariates, x_1, \dots, x_K . These are all random variables. You know that, holding everything else constant (*caeteris paribus*), an increase in the realization of x_k of one unit increases the realized outcome by β_k units, for all $k \in \{1, \dots, K\}$.

Let $\epsilon = y - (\beta_0 + \beta_1 x_1 + \dots + \beta_K x_K)$, where $\mathbb{E}[\epsilon] = 0$ (or $\beta_0 \equiv \mathbb{E}[y - (\beta_1 x_1 + \dots + \beta_K x_K)]$). While x_1, \dots, x_{K-1} are all exogenous, you suspect that x_K is endogenous, i.e. correlated with ϵ . To measure β_0, \dots, β_K you come up with a plausible instrument z_K . It is a strong instrument, and you would like to test whether it is also excluded, i.e. uncorrelated with ϵ .

a. [10 pts] An ideal case

Suppose someone gave you β_0, \dots, β_K .

(i) [3 pts] Do you then observe the realizations of ϵ in any given dataset containing realizations of y and the K covariates? Answer Yes or No (one word).

(ii) [7 pts] For a given sample with n observations, you can form the statistic $\widehat{\text{Cov}}(z_K, y - [\beta_0 + \beta_1 x_1 + \dots + \beta_K x_K])$, a sample covariance. Derive the asymptotic distribution of this statistic. How would you test for the exclusion restriction of z_K ?

Hint: Use the central limit theorem and the delta method. You DO NOT need to give an explicit form for the variance-covariance matrix.

b. [20 pts] Real life

Now you do not know what β_0, \dots, β_K are. Your goal is the same: to check whether z_K satisfies the exclusion restriction.

Let $x = (1, x_1, \dots, x_K)$ and $z \equiv (1, x_1, \dots, x_{K-1}, z_K)$.

(i) [7 pts] Show that the residual from the IV regression that uses z as an instrument for x is:

$$\epsilon^{IV} = y - x' \mathbb{E}[zx']^{-1} \mathbb{E}[zy]$$

(ii) [3 pts] What is $\mathbb{E}[z\epsilon^{IV}]$?

(iii) [10 pts] Reflect and comment: Is your goal met, i.e. have you found a way to test the exclusion restriction of z ?